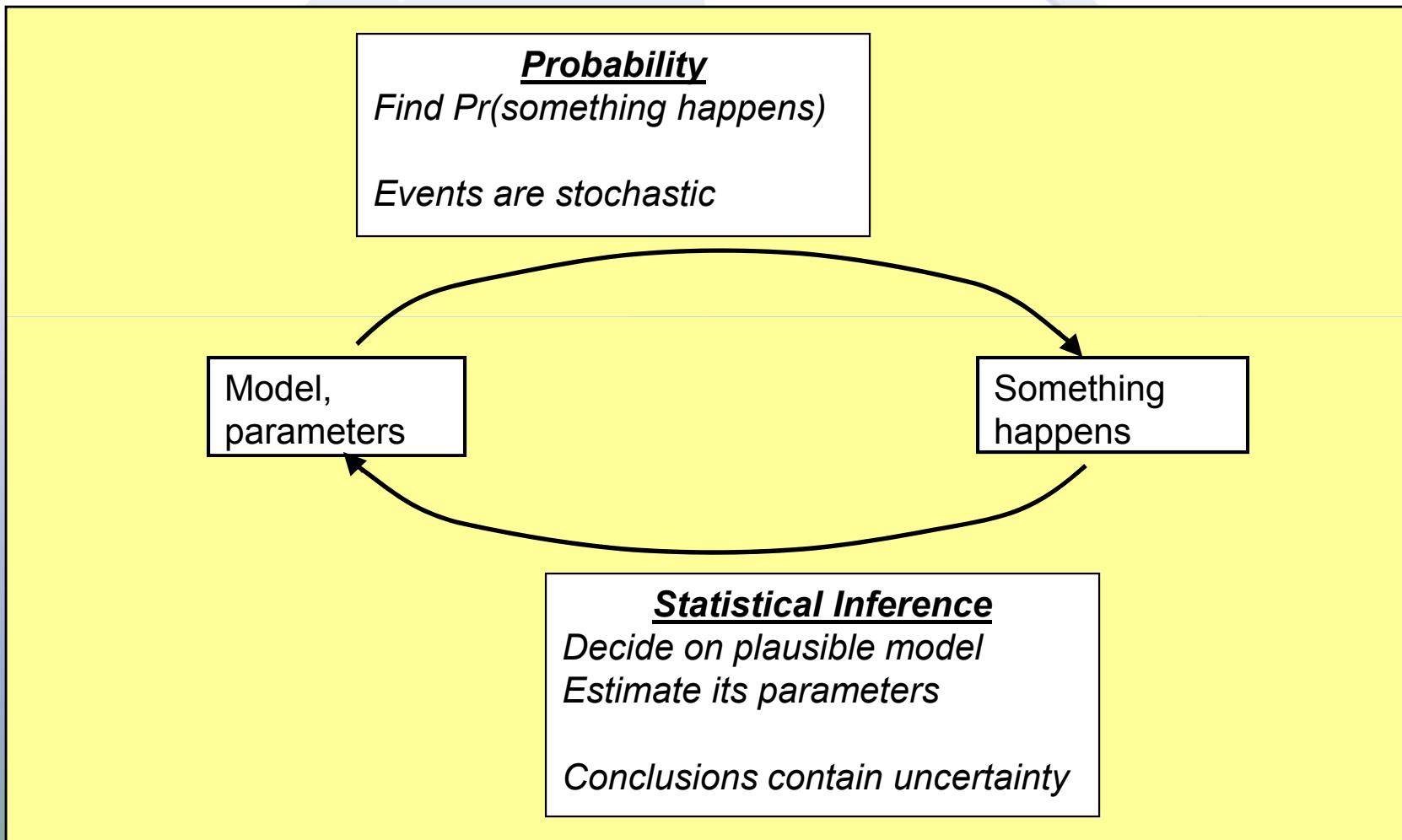


Section 3: Frequentist Statistical Inference

- *Purpose*
 - Students will learn about maximum likelihood estimators, confidence intervals, and about methods of model validation
- *Objectives*
 - Students will learn
 - Definition of maximum likelihood estimator (MLE) and confidence interval
 - Application of these estimators to Poisson, binomial, and exponential data
 - Graphical tools for model validation
 - Intro to hypothesis-testing for model validation, with example applications

Difference between Probability and Statistical Inference



Frequentist Statistical Inference

- *Estimation*
 - *Point Estimates*
 - *Interval Estimates (Confidence Intervals)*
- *Model Validation*
 - *Graphical Methods*
 - *Tests of Hypotheses*

LOSP Example Data

- *In the LOSP example, suppose that we have collected data related to this type of event*
- *The rate of experiencing LOSP*
 - 1 initiating event in 9.2 operating years
- *The probability of not starting a diesel generator (DG)*
 - 1 failure to start in 75 demands
- *The rate of a DG not operating*
 - 0 failures to run in 146 running hours

Tips for Solving Many Statistics Problems

- Answer the following questions, in order:
 - What are the data?
 - Which part(s) of data considered to be observations of a random variable?
 - Which model (distribution) generated the observed data?
 - State which parameter(s) of distribution are known/which are unknown
 - For Bayesian analysis, what is prior distribution of unknown parameter(s)?
 - What is to be found?
 - Point estimate, confidence interval, Bayes posterior distribution, Bayes credible interval, ...
 - Find the formula for it in HOPE, course slides, or notes
- Replace symbols in the formula by known quantities from the data. Look up any needed values from tables, and write out the answer!



Point Estimation

- **Point estimator**
 - A function of random data that estimates unknown parameter
- **Point estimate**
 - Value of estimator for actual observed data
- **That is**
 - Estimator is a random variable (upper case, Λ)
 - Estimate is a number, the value the random variable takes (lower case, λ)



Likelihood Function

- *For discrete random variable*
 - **Likelihood** = $\Pr(\text{data})$, considered as a function of the unknown parameter(s)
- *For continuous random variable*
 - **Likelihood** = density of the data, considered as a function of the unknown parameter(s)
- *In other words, the **likelihood** is the pdf written as a function of the unknown parameter(s)*

Likelihood — Discrete Example

- *Recall for Poisson and binomial distributions (both discrete)*

$$\Pr(X = x) = e^{-\lambda t} (\lambda t)^x / x!$$

$$\Pr(X = x) = \binom{n}{x} p^x (1-p)^{n-x}$$

- *Binomial likelihood example*

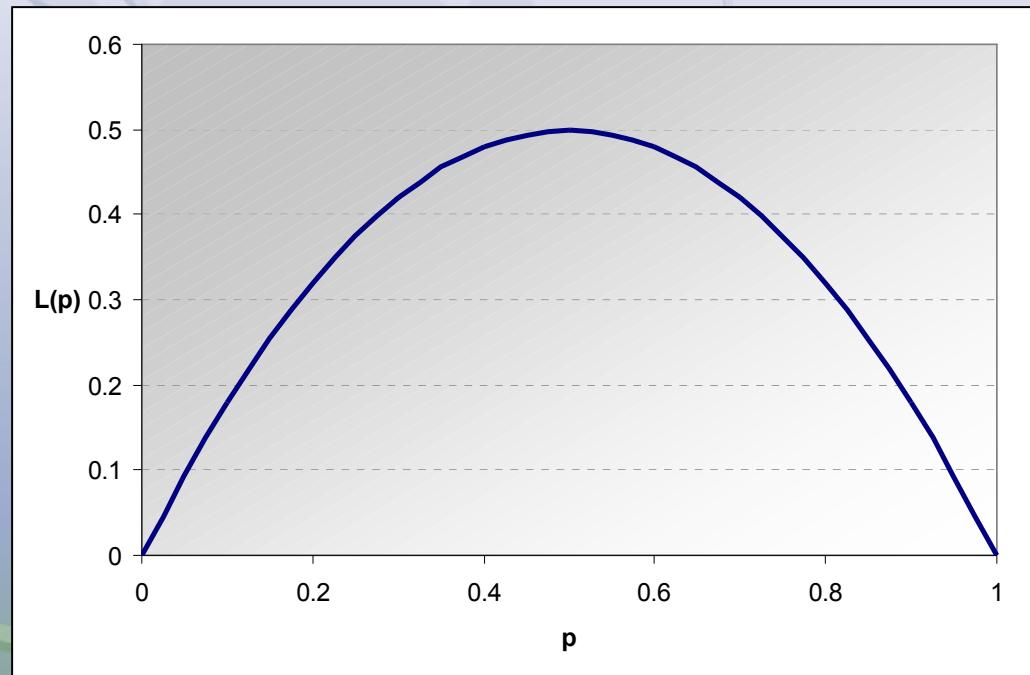
- One failure out of two trials, $x = 1, n = 2$

$$L(p) = \Pr(X = 1 | p) = \binom{2}{1} p^1 (1-p)^{2-1}$$

$$= \frac{2!}{1!(2-1)!} p(1-p) = 2p(1-p)$$

Likelihood — Discrete Example (cont.)

- *What does this binomial likelihood look like?*
 - One failure out of two trials, $x = 1, n = 2$
 - $L(p) = 2p(1 - p) = 2p - 2p^2$
- *The question then is*
 - *What value of p maximizes $L(p)$?*
 - *In other words, what value of p most likely yields data of one failure out two trials?*



Likelihood — Continuous Example

- Recall for exponential distribution (continuous)

$$f(t) = \lambda e^{-\lambda t}$$

- Observe n **independent** event times, each with same exponential distribution.

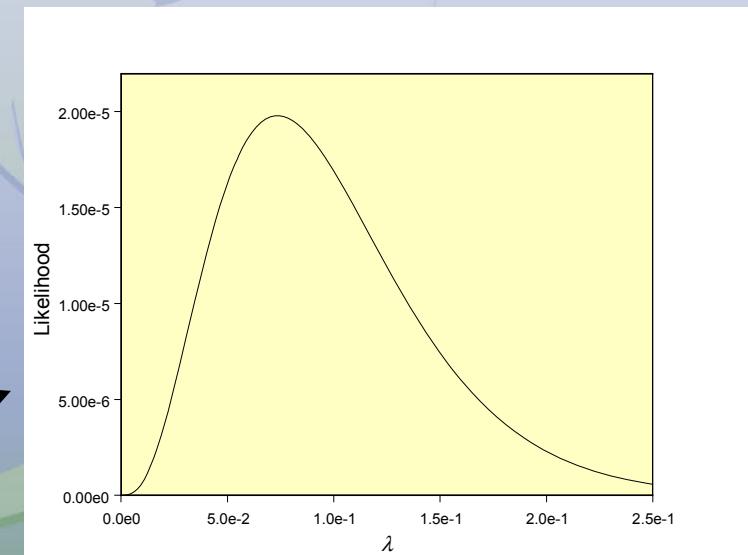
- The observations are t_1, t_2, \dots, t_n

- The likelihood is

$$L(\lambda) = \prod f(t_i | \lambda)$$

$$= \prod \lambda e^{-\lambda t_i} = \lambda^n e^{-\lambda \sum t_i}$$

- Example with $n=3, \sum t_i=40.8$



Maximum Likelihood Estimate (MLE)

- Given the data, the MLE is the parameter value that maximizes the likelihood function
 - This is what we saw on the likelihood example
- Formulas for common cases
 - If $X \sim \text{Poisson}(\lambda t)$, and x is observed, the MLE is
$$\hat{\lambda} = x / t$$
 - If $X \sim \text{binomial}(n, p)$, and x is observed, the MLE is
$$\hat{p} = x / n$$
 - If $T_i \sim \text{exponential}(\lambda)$, independently distributed, $i = 1, \dots, n$, and t_1, \dots, t_n are observed, the MLE is
$$\hat{\lambda} = n / \sum t_i$$

MLE (cont.)



- *Maximum likelihood estimates for LOSP example*
 - If $X \sim \text{Poisson}(\lambda t)$, and $x=1$ event (in 9.2 years), the MLE is

$$\hat{\lambda} = 1 / 9.2 \text{ years} = 0.11 \text{ per year}$$

- If $X \sim \text{binomial}(n, p)$, and $x=1$ failure (in 75 demands), MLE is

$$\hat{p} = 1 / 75 = 0.013$$

- If $X \sim \text{Poisson}(\lambda t)$, and $x=0$ failures (in 146 hours), the MLE is

$$\hat{\lambda} = 0 / 146 \text{ hours} = 0 \text{ per hour}$$

Moment Estimates

- *Estimate moments by the corresponding sample moments*
 - *Sample mean*
$$\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$$
 - *Sample variance*
$$s^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2$$
- *In this course, these are useful for estimating moments of distributions produced by Monte Carlo simulations, when n is large*

Confidence Intervals

- A $100(1 - \alpha)\%$ **confidence interval** has the form (L, U) , where L and U are functions of the data
- The interval satisfies:
 - $\Pr(L \leq \text{parameter} \leq U) \approx 100(1 - \alpha)\%$
 - We would like exact equality, but with discrete data we have to settle for
$$\Pr(L \leq \text{parameter} \leq U) \geq 100(1 - \alpha)\%$$
- **IMPORTANT**
 - In this equation, the parameter is fixed, and L and U are considered random.



Pages 6-6, 6-7, 6-32, 6-33, 6-51, B-5, B-6

Formulas for 90% confidence intervals (with analogous formulas for other confidence levels)

- If $X \sim \text{Poisson}(\lambda t)$,

$$(L, U) = (\chi^2_{0.05}(2X)/(2t), \chi^2_{0.95}(2X + 2)/(2t))$$

- If $X \sim \text{binomial}(n, p)$,

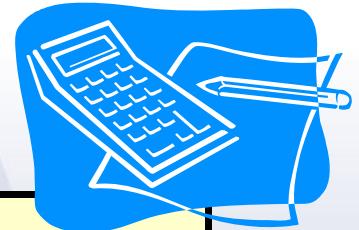
$$(L, U) = (\text{beta}_{0.05}(X, n - X + 1), \text{beta}_{0.95}(X + 1, n - X))$$

- If $T_i \sim \text{exponential}(\lambda)$, independently distributed, $i = 1, \dots, n$, and $T = \sum T_i$,

$$(L, U) = (\chi^2_{0.05}(2n)/(2T), \chi^2_{0.95}(2n)/(2T))$$

- One-sided intervals are also possible (L, ∞) or $(-\infty, U)$.

Summary So Far of Estimates for LOSP Example



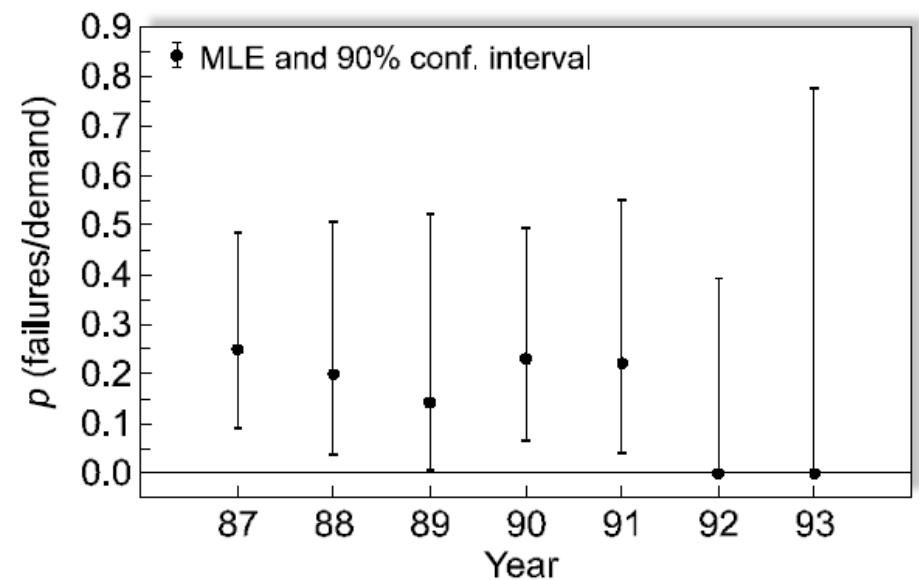
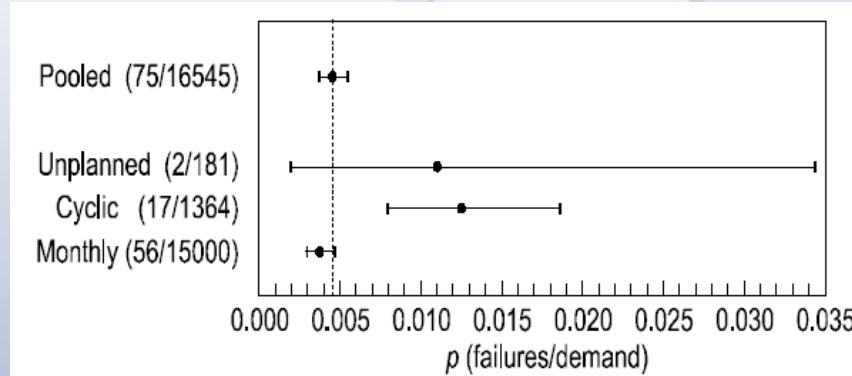
Parameter	Data	MLE	90% Conf. Int.
λ_{LOSP}	1 event in 9.2 op. yrs.	$1.1E-1 \text{ yr}^{-1}$	$(5.6E-3, 5.2E-1) \text{ yr}^{-1}$
p_{FTS}	1 failure in 75 demands	$1.3E-2$	$(6.8E-4, 6.2E-2)$
λ_{FTR}	0 failures in 146 hrs	0 hr^{-1}	$(0, 2.1E-2) \text{ hr}^{-1}$

MODEL VALIDATION

- *It is important to check the model assumptions, by seeing if the data and model are consistent — you can make big errors if you forget to check model assumptions*
- *Graphical methods provide insights about possible violations of assumptions*
- *Hypothesis tests quantify the strength of the evidence against the assumptions*

Two simple graphs are useful

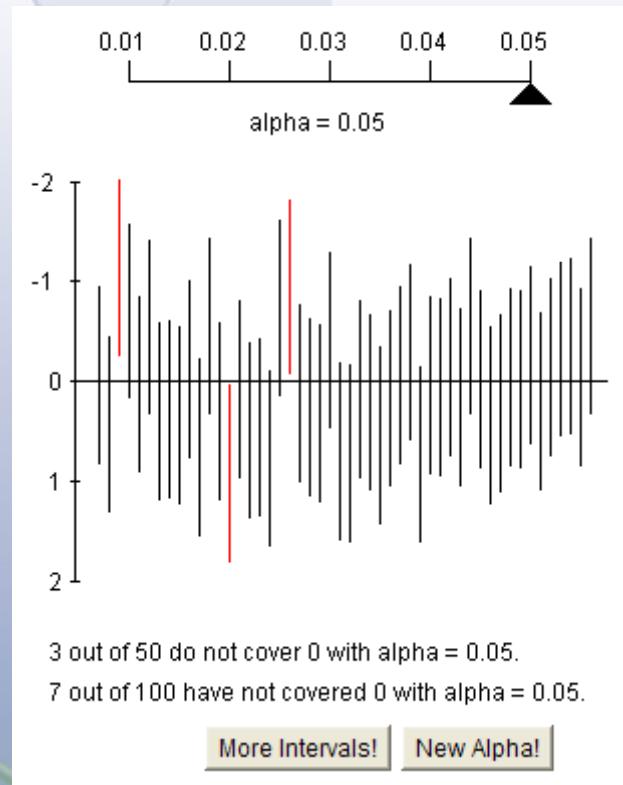
- Side-by-side confidence intervals*



Pages 6-21, 6-25, 6-48

Two simple graphs are useful (cont.)

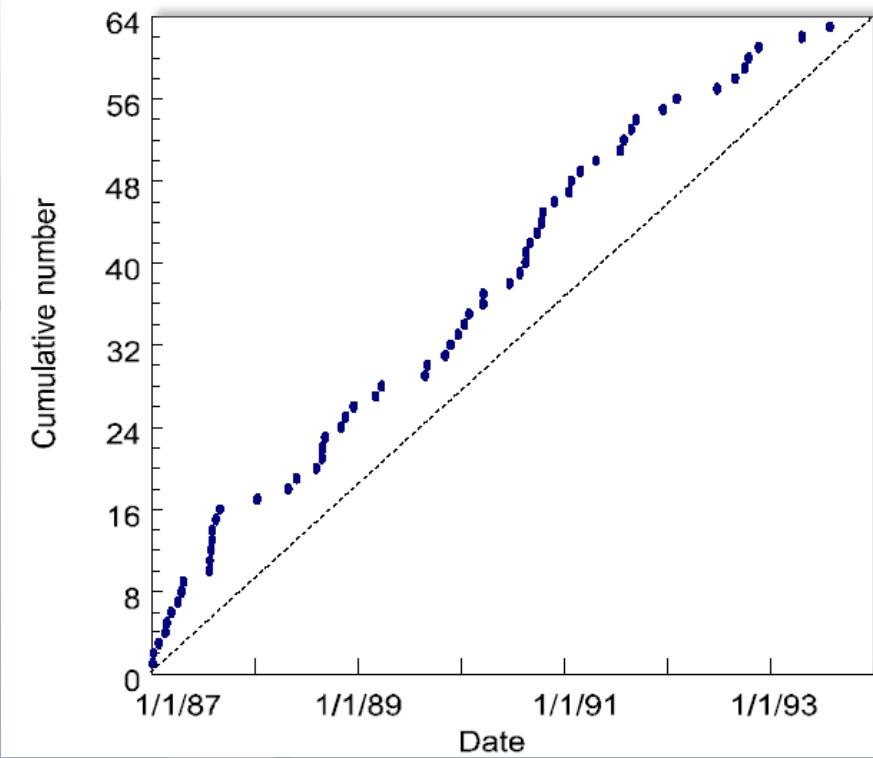
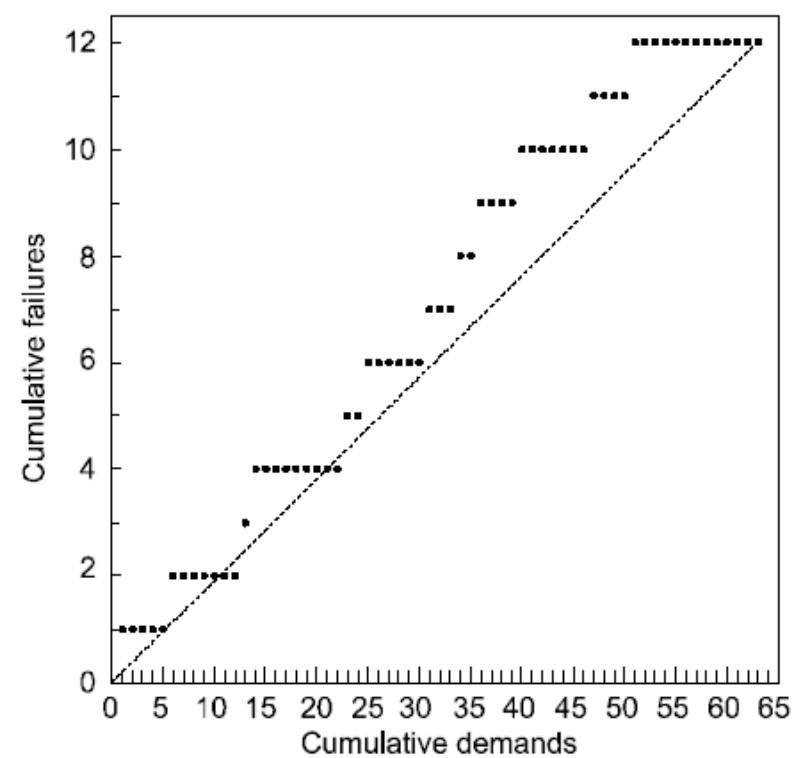
- *Side-by-side confidence intervals*



<http://www.stat.sc.edu/%7Ewest/javahtml/ConfidenceInterval.html>

Two simple graphs are useful (cont.)

- Cumulative plots*



Pages 6-25, 6-48

Test of a Hypothesis – Framework

- *Define two possibilities, or hypotheses*
 - H_0 *the null hypothesis*
 - A simple assumption that is used unless the data give good reason not to believe hypothesis
 - H_1 *the alternative hypothesis*
 - A more complex assumption that will be used only if the null hypothesis is rejected
- *But, two kinds of error are possible*
 - Type I error — reject H_0 when H_0 is true
 - Type II error — accept H_0 when H_0 is false



Test of a Hypothesis – Decision Rule

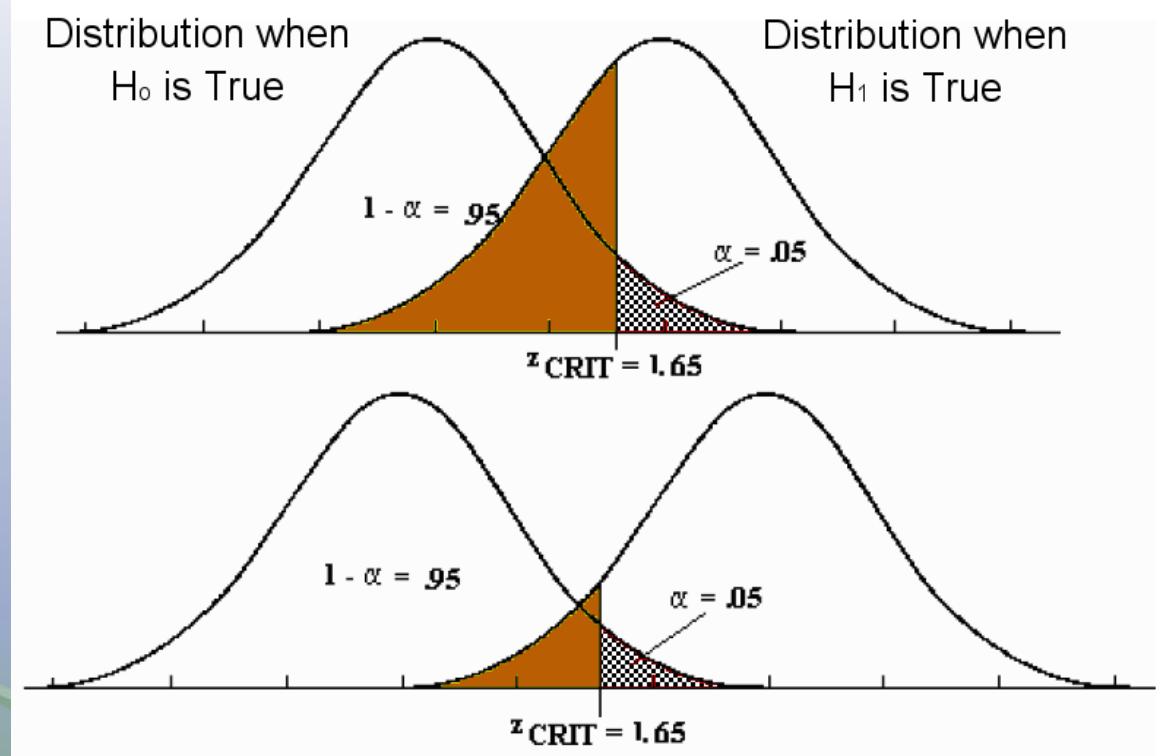
- *The “decision rule”*
 - Construct test statistic (Y), which is a function of the data
- Decide to limit $\Pr(\text{Type I error})$
 - Set α = desired $\Pr(\text{reject } H_0 \mid H_0 \text{ true})$
 - A typical value of α is 0.05 (other values are possible)
 - The more important the decision, in general the smaller one would want to set α
 - This probability of false alarm is the chance you make wrong inference when original hypothesis was correct

Test of a Hypothesis – Decision Rule

- *Construct some “critical region”, with*
$$\Pr(Y \text{ in critical region} \mid H_0 \text{ true}) = \alpha.$$
 - *If α is small, we do not expect to see Y fall in critical region, unless H_0 is false.*
- *Collect data, calculate the value of Y*
 - *If calculated value of Y is in critical region, **reject** H_0 in favor of H_1 ,*
 - *Otherwise, give H_0 benefit of doubt, “**accept**” H_0 (do not reject the null hypothesis)*

Test of a Hypothesis – Graphical Interpretation

- *What we are really “testing” are two sampling distributions, one for H_0 and one for H_1*
- *When these two distributions change, then decision may change*
 - *Example when the mean value varies (two data sets)*

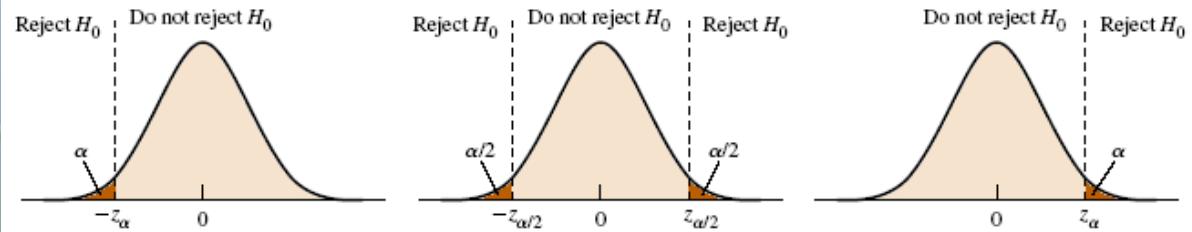


Test of a Hypothesis – Summary

- *From the
“Pocket
Dictionary
of Statistics”*

hypothesis and an **alternative hypothesis**. A null hypothesis is usually tested and either rejected in favor of an alternative hypothesis or not rejected, in which case the alternative hypothesis cannot be sustained.

1. State a null hypothesis (H_0) based on the specific question or phenomenon to be investigated.
2. State an alternative hypothesis. This may be one-sided or two-sided depending on the problem being investigated as defined in the null hypothesis.
3. Specify the **level of significance (α)**. This is commonly taken as 0.05 and represents the maximum acceptable probability of incorrectly rejecting the null hypothesis.
4. Determine an appropriate **sampling distribution** of the **sample statistic** of interest. Select a one-tailed or two-tailed test, depending on the alternative hypothesis.
5. Evaluate the **standard error** or, more generally, an estimate of the standard error of the sample statistic; the formula for the standard error depends on the sample statistic in question.
6. Compute the true value of the **test statistic** and locate its value on the sampling distribution.
7. Reject or do not reject H_0 , depending on whether or not the sample statistic is located on the sampling distribution at or beyond the value of the test statistic at a given α .



Test of a Hypothesis – Another Way of Stating Conclusions

- ***The p-value***
 - Value of α at which H_0 would just **barely** be rejected
 - In other words, “shift” α left or right until H_0 is not accepted
- *It is common to collect data, calculate y , and then report the p-value*
 - This approach measures strength of evidence against H_0
 - A small p-value corresponds to strong evidence against H_0
- A p-value < 0.05 corresponds to “statistical significance”

Commonly Used Tests

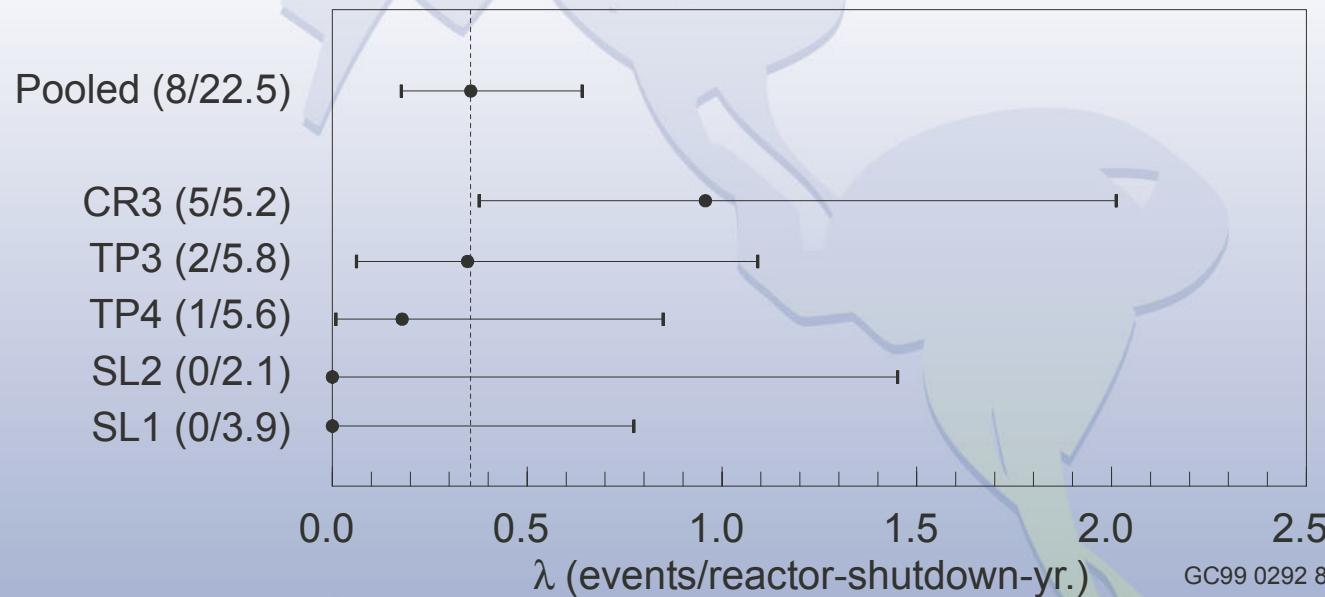
- *To test H_0 : $\lambda = \lambda_0$*
 - *Reject H_0 if confidence interval for λ does not contain λ_0*
- *The “chi-squared test” is one test of more complicated hypotheses.*
- *Many other tests exist for various situations*



Pages 6-22 and 6-23

Chi-Squared Test — Example

- LOSP was considered for 5 plants while in shutdown



- We want to test the hypothesis
 - H_0 : λ is same at all five plants
 - H_1 : λ is not same at all five plants

Chi-Squared Test — Example (cont.)

- The data, and start of calculations for test

Plant	Events, x_i	Plant shut- down yrs, t_i	Expected count, $e_i = \hat{\lambda}t_i$
CR3	5	5.224	1.857
SL1	0	3.871	1.376
SL2	0	2.064	0.734
TP3	2	5.763	2.048
TP4	1	5.586	1.985
<i>Totals</i>	8	22.508	8.000

$$\hat{\lambda} = \sum x_i / \sum t_i = 0.355$$



Chi-Squared Test — Example (cont.)

- e_i is expected count assuming H_0 true
- Compare observed, x_i , with expected, e_i
- Combine the results for each cell by

$$X^2 = \sum \frac{(x_i - e_i)^2}{e_i}$$

- If H_0 true, distribution of X^2 is approximately chi-squared, if e_i values are not too small
 - Degrees of freedom = no. of cells - 1, i.e. 4 in example
 - For “not too small”, see HOPE, p. 6-24
- If H_0 false, X^2 tends to be larger than chi-squared random variable. So reject H_0 if X^2 in tail of chi-squared distr.

Chi-Squared Test — Example (cont.)

- *In this example, we can calculate $\chi^2 = 7.92$*
- *Close to 90th percentile of $\chi^2(4)$ distribution*
- *Conclusion: evidence is borderline, but not quite strong enough to justify rejecting H_0 .*
- *This conclusion was also suggested by the picture*
 - *But interpretation of the picture is somewhat subjective*
 - *Hypothesis test objectively quantifies the strength of the evidence against H_0*

Hypothesis Testing Example – Maintenance Rule Performance Criteria

- Component has assumed probability of failure (PRA point estimate)
- Component tests or demands can reveal failure
- How many failures are too many for assumed probability of failure to be accepted?
- Null hypothesis – $H_0: p = p_o$
- Alternative hypothesis – $H_1: p > p_o$
- Choose $\alpha = 0.05$

Hypothesis Testing Example – Maintenance Rule Performance Criteria

- Assume 24 tests will be performed over next evaluation period
- Assume number of failures follows binomial distribution with $n = 24$ and $p = p_o$ (under H_o). Assume $p_o = 0.06$.
- Need to find x_{crit} , such that $\Pr(\text{rejecting } H_o | H_o \text{ true}) = 0.05$.
- That is, $\Pr(X > x_{crit} | n=24, p_o = 0.06) = 0.05$
- Easier to use $\Pr(X \leq x_{crit} | n=24, p_o = 0.06) = 0.95$
- Because X is discrete, don't strive for exact equality in the above

Hypothesis Testing Example – Maintenance Rule Performance Criteria

- *Using binomial distribution, find $Pr(X \leq 3 | n=24, p_o = 0.06) = 0.947$, so performance criterion is set at 3 or fewer failures in 24 tests.*
- *What is Type II error probability?*
 - Depends on what p actually is
 - Assume $p_{act} = 3 p_o = 0.18$
 - $Pr(\text{Type II error}) = Pr(\text{accepting } H_0 | H_1 \text{ true}) = Pr(X \leq 3 | n=24, p_{act} = 0.18) = 0.35$

Summary of Frequentist Estimates for LOSP Example



Parameter	Point Est. (MLE)	90% Interval
λ_{LOSP}	$1.1E-1 \text{ yr}^{-1}$	$(5.6E-3, 5.2E-1) \text{ yr}^{-1}$
p_{FTS}	$1.3E-2$	$(6.8E-4, 6.2E-2)$
λ_{FTR}	0 hr^{-1}	$(0, 2.1E-2) \text{ hr}^{-1}$