# Information Bias and Lifetime Mortality Risks of Radiation-Induced Cancer 

## Low LET Radiation

## Prepared by

L. E. Peterson, X. J. Schull. B. R. Davis, P. A. Buffler

School of Public Health
University of Texas, Houston

Prepared for
U.S. Nuclear Regulatory Commission

## NOTICE

This document was prepared with the support of the U.S. Nuclear Regulatory Commission (NRC) Grant Program. The purpose of the NRC Grant Program is to support basic, advanced, and developmental scientific research for a public purpose in areas relating to nuclear safety. The nature of NRC's Grant Program is such that the grantee bears prime responsibility for the conduct of the research and exercises judgement and original thought toward attaining the scientific goals. The opinions, findings, conclusions and recommendations expressed herein are therefore those of the author(s) and do not necessarily reflect the views of the NRC.

## Available from

Superintendent of Documents
U.S. Government Printing Office

Mail Stop SSOP
Washington, DC 20402-9328
and
National Technical Information Service
Springfield, VA 22161

## Information Bias and Lifetime Mortality Risks of Radiation-Induced Cancer

Low LET Radiation

Manuscript Completed: February 1994
Date Published: April 1994

Prepared by
L. E. Peterson, W. J. Schull, B. R. Davis, P. A. Buffler*

School of Public Health
University of Texas,
Health Science Center
Houston, TX 77225

Prepared for
Division of Regulatory Applications
Office of Nuclear Regulatory Research
U.S. Nuclear Regulatory Commission

Washington, DC 20555-0001
FIN G1992

[^0]
#### Abstract

Additive and multiplicative models of relative risk were used to measure the effect of cancer misclassification and DS86 random errors on lifetime risk projections in the Life Span Study (LSS) of Hiroshima and Nagasaki atomic bomb survivors. The true number of cancer deaths in each stratum of the cancer mort ${ }^{14}+\mathrm{y}$ crossclassification was estimated using sufficient statistics from the EM algorithm. Average survivor doses in the strata were corrected for DS86 random error ( $\sigma=0.45$ ) by use of reduction factors. Poisson regression was used to model the corrected and uncorrected mortality rates with covariates for age at-time-of-bombing, age at-time-of-death and gender. Excess risks were in good agreement with risks in RERF Report 11 (Part 2) and the BEIR-V Report. Bias due to DS86 random error typically ranged from $-15 \%$ to $-30 \%$ for both sexes, and all sites and models. The total bias, including diagnostic misclassification, of excess risk of nonleukemia for exposure to 1 Sv from age 18 to 65 under the non-constant relative projection model was $-37.1 \%$ for males and $-23.3 \%$ for females. Total excess risks of leukemia under the relative projection model were biased $-27.1 \%$ for males and $-43.4 \%$ for females. Thus, nonleukemia risks for 1 Sv from ages 18 to 65 (DR$\mathrm{REF}=2$ ) increased from $1.91 \% / \mathrm{Sv}$ to $2.68 \% / \mathrm{Sv}$ among males and from $3.23 \% / \mathrm{Sv}$ to $4.02 \% / \mathrm{Sv}$ among females. Leukemia excess risks increased from $0.87 \% / \mathrm{Sv}$ to $1.10 \% / \mathrm{Sv}$ among males and from $0.73 \% / \mathrm{Sv}$ to $1.04 \% / \mathrm{Sv}$ among females. Bias was dependent on the gender, site, correction method, exposure profile and projection model considered. Future studies that use LSS data for U.S. nuclear workers may be downwardly biased if lifetime risk projections are not adjusted for random and systematic errors. (Supported by U.S. NRC Grant NRC-04-091-92).


This report makes use of data obtained from the Radiation Effects Research Foundation (RERF) in Hiroshima, Japan. RERF is a private foundation funded equally by the Japanese Ministry of Health and Welfare and the U.S. Department of Energy through the U.S. National Academy of Sciences. The conclusions in this report are those of the authors and do not necessarily reflect the scientific judgement of RERF or its funding agencies.

## Contents

1 EXECUTIV! SUMMARY ..... 1-1
2 INTRODUCTION ..... 2-1
2.1 Precision, Validity, Generalizability and Bias ..... 2-1
2.2 Effects of Random Error on Dose-Response ..... ,-2
2.3 Effects of Systematic Error on Dose-Response ..... L-8
2.4 Studies of Death Certificate Misclassification ..... 2-9
2.5 Studies of Lifetime Risk ..... 2-11
2.6 Reseazch Objectives ..... 2-12
3 MATERIALS AND METHODS ..... 3-1
3.1 Sources of Data ..... 3-1
3.1.1 RERF Autopsy Program ..... 3-1
3.1.2 RERF Cancer Mortality Data ..... 3-1
3.1.3 RERF Average Body Transmission Factors ..... 3-1
3.1.4 Reduction Factors for DS86 Random Error ..... 3-4
3.2 Use of Confirmation Rates to Adjust Cancer Deaths ..... 3-4
3.2.1 Diagnostic Screening ..... 3-4
3.2.2 Estimation of True Cancer Deaths ..... 3-5
3.3 Dose-Response Analysis ..... 3-6
3.3.1 Excess Relative Risks ..... 3-6
3.3.2 Absolute Risks ..... 3-8
3.3.3 Goodness-of-Fit (GOF) Statistics ..... 3-8
3.4 Projection of Lifetime Mortality Risks ..... 3-9
3.4.1 Risk Coefficients and Projection Models ..... 3-9
3.4.2 Baseline Rates and Life Tables ..... 3-11
3.4.3 Death Certificate Correction Factors (DCCF) ..... 3-11
4 RESULTS ..... 4-1
4.1 Per Cent Distribution of True Cancer Deaths ..... 4-1
4.2 Poisson Regression ..... 4-1
4.2.1 Models with Age ATB, Age ATD and Gender ..... 4-1
4.2.2 BEIR-V Models ..... 4-2
4.3 Lifetime Risk Projection ..... 4-3
4.3.1 Lifetime Risks Without Adjustments ..... 4-3
4.3.2 Bias in Absolute Projection Models ..... 4-5
4.3.3 Bias in Transported Relative Projection Models ..... 4-5
4.3.4 Bias in Relative Projection Models ..... 4-7
4.3.5 Bias in BEIR-V Relative Projection Models ..... 4-9
5 DISCUSSION ..... 5-1
5.1 Confirmation Rates ..... 5-1
5.2 Regression Methods ..... 5-1
5.3 DS86 Random Error ..... 5-3
5.4 Autopsy Program and Diagnostic Misclassiciation ..... 5-4
5.5 Bias in Lifetime Murtality Risks ..... 5-4
6 SUMMARY ..... 6-1
7 REFERENCES ..... 7-1
8 NOTATION ..... 8-1
9 ABBREVIATIONS ..... 9-1
10 APPENDLX A. Dose-Respouse Modeling ..... A-1
10.1 General Approach ..... A-1
10.2 Correction of Shielded Kerma for Random Unctrtainty ..... A-1
10.3 Neutron Relative Biological Effectiveness and Estimation of Organ Dose Equivalents from Corrected Shielded Kerma ..... A-1
10.4 RERF Models ..... A-2
10.4.1 Relative and Excess Relative Risks ..... A-2
10.4.2 Non-constant Excess Relative Risk Models for Leukemia, Nonleukemia, Stomach and Breast Sites ..... A-2
10.4.3 Constant Excess Relative Risk Models for Lung, Bladder, Liver, Colon, and Ovary Sites ..... A-3
10.4.4 Determining Excess Relative Risk from Regression Coefficients ..... A-3
10.4.5 Determining Absolute Risks from Regression Coefficients ..... A-4
10.5 BEIR-V Models ..... A-5
10.5.1 BEIR-V Leukemia model ..... A-5
10.5.2 BEIR-V Breast Model ..... A-5
10.5.3 BEIR-V Respiratory Model ..... A-6
10.5.4 BEIR-V Digestive Model ..... A-6
10.5.5 BEIR-V Other Cancers Model ..... A-6
10.6 Regression Diagnostics and Goodness-of-Fit (GOF) ..... A-7
11 APPENDIX B. Lifetime Risk Projection ..... B-1
11.1 Introduction ..... B-1
11.2 Hazard functions for Radiation-Induced Cancer ..... B-1
11.3 Double-Decrement Life Table (Radiation-Induced Cancers) ..... B-2
11.4 Single-Decrement Life Table (Baseline cancers) ..... B-2
11.5 Lifetime Risks Based on Method of Elandt-Johnson and Johnson ..... B-3
11.6 Years of Life Lost Per Premature Radiation-Induced Cancer Death ..... B-6
11.7 Probability of Causation ..... B-6
11.8 Error Propagation ..... B-6
11.8.1 Constant and Non-constant Absolute and Relative Projection ModelsB-6
11.8.2 BEIR-V Relative Projection Model ..... B-8
11.8.3 Probability of Causation ..... B-9
11.9 Credibility Intervals of Lifetime Risk ..... B-9
12 APPENDIX C. Relative and Absolute Risk Coefficients. ..... C-1
13 APPENDIX D. Lifetime Mortality Risks (\%/Sv). ..... D-1

## List of Tables

1 Cross-classification of LSS cancer mortality data. ..... 3-4
2 Data arrangement of screening results in the RERF ..... 3-5
3 Probabilities of misclassification of disease. ..... 3-5
4 Crude confirmation rates for cancer, $\theta_{0}$, and non-cancer, $\theta_{n c}$, in the Life Span Study Pathology Report 4 ..... 3-6
5 Sex-specific confirmation rates for cancer, $\theta_{c}$, and non-cancer, $\theta_{n c}$, in the Life Span Study Pathology Report 4. ..... 3-8
6 Sex- and city-specific confirmation rates for cancer, $\theta_{c}$, and non-cancer, $\theta_{n c}$, in the Life Span Study Pathology Report 4. ..... 3-9
7 Age ATD-specific confirmation rates for cancer, $\theta_{c}$, and non-cancer, $\theta_{n c}$, in the Life Span Study Pathology Report 4. ..... 3-9
8 T65DR-specific confirmation rates for cancer, $\theta_{c}$, and non-cancer, $\theta_{n c}$, in the Life Span Study Pathology Report 4. ..... 3-9
9 Site- and sex-specific excess risks (\%/Sv) for the absolute projection model without adjustment for diagnostic misclassification and DS86 random error (DRREF=2) ..... 4-4
10 Site- and sex-specific excess risks (\%/Sv) for the transported relative pro- jection model without adjustment for diagnostic misclassification and DS86 random error (DRREF $=2$ ) ..... 4-4
11 Site- and sex-specific excess risks (\%/Sv) for the relative projection model without adjustment for diagnostic misclassification and DS86 random error (DRREF = 2) ..... 4-5
12 Site- and sex-specific excess risks (\%/Sv) for the BEIR-V relative projection model without adjustment for diagnostic misclassification and DS86 random error (DRREF=2) ..... 4-6
13 Negative bias of excess risk (\%/Sv) among ma es for the absolute projection model (DRREF $=2$ ) ..... 4-6
14 Negative bias of excess risk (\%/Sv) among females for the absolute projec- tion model (DRREF=2). ..... 4-7
15 Negative bias of excess risk (\%/Sv) among males for the transported relative projection model (DRREF=2). ..... 4-8
16 Negative bias of excess risk (\%/Sv) among females for the transported rel- ative projection model (DRREF=2). ..... 4-8
17 Negative bias of excess risk (\%/Sv) among males for the relative projection model (DRREF $=2$ ). ..... 4-12
18 Negative bias of excess risk (\%/Sv) among females for the relative projection model (DRREF=2). ..... 4-13
19 Negative bias of excess risk (\%/Sv) among males for the BEIR-V projection model (DRREF $=2$ ) ..... 4-14
20 Negative bias of excess risk (\%/Sv) among females for the BEIR-1' projec- tion model (DRREF=2). ..... 4-14
21 Weighted per cent distribution of non-cancer deaths among both sexes in both cities in the Life Span Study (1950-85). ..... 5-1
22 Weighted per cent distribution of non-cancer deaths among males in both cities in the Life Span Study (1950-85). ..... 5-1
23 Weighted per cent distribution of non-cancer deaths among females in both cities in the Life Span Study (1950-85). ..... 5-2
24 Weighted per cent distribution of non-cancer deaths among Hiroshima males in the Life Span Study $(1950-85)$. ..... 5-2
25 Weighted per cent distribution of non-cancer deaths among Nagasaki males in the Life Span Study (1950-85). ..... $5-3$
26 Weighted per cent distribution of non-cancer deatias among Hiroshima fe- males in the Life Span Study (1950-85). ..... 5-3
27 Weighted per cent distribution of non-cancer deaths among Nagasaki fe- males in the Life Span Study (1950-85). ..... 5-3
28 Error components of lifetime risk. ..... B-9

## List of Figures

1 Random and systematic error in epidemiologic research. (Reproduced with permission, Kleinbaum, D.G., Kupper, L.L., Morgenstern, H. Epidemiologic Research: Principles and Quantitative Methods. (c1982 Van Nostrand Reinhold)
Hierarchy of populations in epidemiologic research. (Reproduced with per-
mission, Kleinbaum, D.G., Kupper, L.L., Morgenstern, H. Epidemiologic
Research: Principles and Quantitative Methods. (c)1982 Van Nostrand
Reinhold)
3 Hierarchy of populations in the Life Span Study. . . . . . . . . . . . . . . . 2-5
4 Chromosome aberration (CA) dose-response within epilation groups. Doses adjusted for $45 \%$ random error. Adapted from Sposto et al. (1991). . . . 2-6

6 Typical methodology for estimating lifetime mortality risk of radiation- $\begin{aligned} & \text { induced cancer. . . . . . . . . . . . . . . . . . . . . . } 3-2\end{aligned}$
7 Methodology for estimating lifetime mortality risk of radiation-induced $\begin{aligned} & \text { cancer in this investigation. }\end{aligned}$
8 Conditional probabilities of radiation-induced nonleukemia among males in this study and results based on the Sposto et al. regression coefficients. 4-10
) Conditional probabilities of radiation-induced nonleukemia among females in this study and results based on the Sposto et al. regression coefficients. 4-11

## ACKNOWLEDGEMENTS

The successful completion of this project ${ }^{1}$ would not have been possible without the help, suggestions and scientific advice of several individuals. The following individuals contributed, in part, to this study and are gratefully acknowledged:

Ms. Barbara Brooks<br>Dr. Sharon Cooper<br>Dr. Tom Downs<br>Dr. Mark Durst<br>Dr. Edward Frome<br>Dr. Ethel Gilbert<br>Dr. Charles Land<br>Dr. Mary McConney<br>Dr. Colin Muirhead<br>Dr. Donald Pierce<br>Dr. Dale Preston<br>Dr. Yukiko Shimizu<br>Dr. Mary Ann Smith<br>Ms. Terry Smith<br>Dr. Shlumo Yaniv<br>Department of Energy<br>University of Texas<br>University of Texas<br>Lawrence Berkeley Laboratory<br>Oak Ridge National Laboratory<br>Battelle Pacific Northwest Laboratory<br>National Cancer Institute<br>Hirosoft International Corporation<br>National Radiological Protection Borrd(UK)<br>Radiation Effects Research Foundation<br>Radiation Effects Research Foundation<br>Radiation Effects Research Foundation<br>University of Texas<br>UT M.D. Anderson Cancer Center<br>Nuclear Regulatory Commission

[^1]
## FOREWORD

The Nuclear Regulatory Commission, through its Grant Program has supported several educational institutions. One of the criteria considered in awarding a grant is the benefit to the graduate research program of the institution, e.g., graduate student training.

NUREG/GR-0011 "Information Bias and Lifetime Mortality Risks of Radiation Induced Cancer" is based on the doctoral dissertation of the senior author. The research was supported by an NRC grant.

NUREG/GR-0011 is published for information only. Publication of this report does not necessarily constitute NRC approval or agreement with the information contained herein.


Donald A. Cool, Chief
Radiation Protection and Health Effects Branch
Division of Regulatory Applications Office of Nuclear Regulatory Research

## 1 EXECUTIVE SUMMARY

The purpose of the present study was to measure the effect of random and systematic errors in the measurement of radiation exposures and cancer-specific mortality misclassification in the Life Span Study (LSS) of Hiroshima and Nagasaki atomic bomb survivors and the Surveillance, Epidemiology, and End Results (SEER) program on lifetime mortality risks of radiation-induced cancer for U.S. nuclear workers. The LSS is a radiation effects cohort study that has been conducted by the Radiation Effects Research Foundation (RERF), formerly known as the Atomic Bomb Casualty Commission (ABCC), since 1947. The RERF is a private non-profit Japanese Foundation, supported equally by the Government of Japan through the Ministry of Health and Welfare, and the U.S. Government through the National Academy of Sciences under contract with the Department of Energy. The SEER program is a nation-wide cancer reporting system run by the National Cancer Institute of the National Institutes of Health.

Although every imaginable aspect of the effect of information bias ${ }^{1}$ on radiationinduced cancer in the LSS bas been subject to scrutiny over the last decade, the one part that has managed to elude systematic investigation has been the joint analysis of information bias and lifetime risk projections for U.S. nuclear workers. The primary intent of this study was to measure the bias in lifetime mortality risks of radiation-induced cancer that have been generated with and without adjustment for Dosimetry System - 1986 (DS86) random error and diagnostic misclassification of mortality rates in the LSS and SEER program. Adjustments for DS86 random errors and diagnostic misclassification of LSS cancer deaths were made during dose-response analysis with Poisson regression using the AMFIT ${ }^{2}$ computer program. Adjustments for diagnostic misclassification of sancer deaths in the SEER data were made during lifetime risk projection using the SURVRAD ${ }^{3}$ computer program. Death certificate and confirmation and detection rates for the LSS in the years 1950-1975 were based on results of the RERF Pathology Studies. Confirmation and detection rates for the SEER program were obtained from reports published in the open literature.

The major findings of this investigation were:
(1). As age at death increased a greater proportion of true cancer deaths were attributable to non-cancer deaths because the true number of cancer deaths is equal to the sum of the product of the observed cancer deaths and the probability that the observed cancer deaths are correctly classified and the product of the observed non-cancer deaths and one minus the the probability that the observed non-cancer deaths were correctly classified (see Eq. 5 in §3.2.2).
(2). Poisson regression resulted in fitted maximum likelihood models that were in concordance with the observed data. When the goodness-of-fit of regression models containing time-dependent covariates is reasonable, non-constant lifetime risk projection should be used.
(3). Excess relative risk coefficients for the RERF and BEIR-V models were in good agreement with those published in RERF Report 11 (Part 2) and the BEIR-V report. Small differences existed between regression results for RERF models that contained parameters for age at-time-of-bombing (ATB), age at-time-of-death (ATD), and gender because organ dose estimates were used rather than shielded kerma. Thus, the lifetime

[^2]based on these models were slightly higher than those that would obtain from the use of coefficients in RERF Report 11.
(4). Statistical modeling with the BEIR-V models provided regression coefficients that were almost exactly identical to those in the BEIR-V report. For leukemia, the linearquadratic contribution of dose to excess mortality was slightly lower than that in the BEIRV report. Lifetime risks based on the BEIR-V models were similar to those published in the BEIR-V Report (NRC, 1990). Bias due to DS86 randoh error for the digestive site was smaller than bias in the RERF non-constant nonleukemia projection models, which was most likely due to truncation of dose equivalent to 4 Sv . The correction of diagnostic misclassification in excess risks for the BEIR-V digestive cancer site had little effect on bias $(-2 \%)$ because records with an age at death beyond 75 , when cancer misclassification rises markedly, were excluded.
(5). Using a Dose-Rate Reduction Effectiveness Factor (DRREF) of two and no correction for DS86 raudom error or diagnostic misclassification in the non-constant relative projection model, lifetime risks (\%/Sv) of nonleukemia among males exposed acutely to 1 Sv at age $25,45,65$ or continuously from ages 18 to 65 were $2.10 \%, 2.78 \%, 1.20 \%$ and $1.91 \%$. For females, nonleukemia lifetime risks for the same exposure profiles were $3.49 \%$, $4.32 \%, 1.97 \%$ and $3.23 \%$. Excess leukemia risks for 1 Sv at 25,45 and 65 and over the years 18 to 65 were $0.35 \%, 0.46 \%, 2.46 \%$ and $0.87 \%$ for males and $0.26 \%, 0.41 \%, 1.96 \%$ and $0.73 \%$ for femaies. These data were in good agreement with the results of Land and Sinclair (1991). By way of comparison, for exposure from ages 18 to 65 , excess nonleukemia risks based on the constant relative projection model were $2.84 \%$ for males and $4.75 \%$ for fernales. The risks of leukemia among males was $0.75 \%$ and among females was $0.64 \%$. Therefore, lifetime risk estimates based on constant models did not underestimate risks projected by non-constant models.
(6). The correction of differential diagnostic misclassification with leukemia and nonleukemia (and non-cancer) confirmation rates that were stratified on T65DR dose (DS86 shielded kerma was converted to T65DR shielded in order to select T65DR-specific confirmation rates) resulted in bias that was negative. Confirmation rates for leukemia and nonleukemia that were stratified on age ATD did not provide bias that was more negative than that obtained with DS86-specific confirmation rates. Correction of diagnostic misclassification using confirmation rates that were crude or stratified on either gender or city and gender resulted in bias that was negative or positive. The bias of excess risk of nonleukemia due to diagnostic misclassification for 1 Sv at age $25,45,65$ or continuously from ages 18 to 65 under the non-constant relative projection model was $-5.0 \%(2.13 \% / \mathrm{Sv}$ vs. $2.24 \% / \mathrm{Sv}),-7.3 \%(2.78 \% / \mathrm{Sv}$ vs. $2.99 \% / \mathrm{Sv}),-38.9 \%(1.20 \% / \mathrm{Sv}$ vs. $1.67 \% / \mathrm{Sv})$ and $-11.3 \%$ $(1.91 \% / \mathrm{Sv}$ vs. $2.13 \% / \mathrm{Sv})$ for males and $-1.5 \%(3.49 \% / \mathrm{Sv}$ vs. $3.54 \% / \mathrm{Sv}),-3.9 \%(4.32 \% / \mathrm{Sv}$ vs. $4.49 \% / \mathrm{Sv}),-26.2 \%(1.97 \% / \mathrm{Sv}$ vs. $2.48 \% / \mathrm{Sv})$ and $-6.3 \%(3.23 \% / \mathrm{Sv}$ vs. $3.43 \% / \mathrm{Sv})$ for females. For leukemia excess risks under the same dose profiles, the bias due to diagnostic misclassification was $-6.0 \%(0.36 \% / \mathrm{Sv}$ vs. $0.37 \% / \mathrm{Sv}),-69.7 \%(0.46 \% / \mathrm{Sv}$ vs. $0.77 \% / \mathrm{Sv})$, $-23.3 \%(2.46 \% / \mathrm{Sv}$ vs. $3.04 \% / \mathrm{Sv})$ and $-23.9 \%(0.87 \% / \mathrm{Sv}$ vs. $1.09 \% / \mathrm{Sv})$ for males and $-12.1 \%(0.26 \% / \mathrm{Sv}$ vs. $0.30 \% / \mathrm{Sv}),-83.4 \%(0.41 \% / \mathrm{Sv}$ vs. $0.75 \% / \mathrm{Sv}),-40.9 \%(1.96 \% / \mathrm{Sv}$ vs. $2.77 \% / \mathrm{Sv})$, and $-42.8 \%(0.73 \% / \mathrm{Sv}$ vs. $1.05 \% / \mathrm{Sv})$ for females. When the nonleukemia Poisson regression corfficients from Sposto et al. (1992) were used to project lifetime risks under the non-constant relative model, the bias due to diagnostic misclassification for 1 Sv acute at 25,45 , or 65 and over a career ( 18 to 65 ) was $-1.0 \%$ ( $2.61 \% / \mathrm{Sv}$ vs. $2.64 \% / \mathrm{Sv}$ ), $-4.0 \%(5.72 \% / \mathrm{Sv}$ vs. $5.95 \% / \mathrm{Sv}), 1.3 \%(2.39 \% / \mathrm{Sv}$ vs. $2.36 \% / \mathrm{Sv})$, and $-10.0 \%(4.90 \% / \mathrm{Sv}$ vs. $5.39 \% / \mathrm{Sv}$ ) for males and $13.3 \%(3.02 \% / \mathrm{Sv}$ vs. $2.62 \% / \mathrm{Sv}), 3.2 \%(6.49 \% / \mathrm{Sv}$ vs. $6.28 \% / \mathrm{Sv})$, $5.2 \%(2.46 \% / \mathrm{Sv}$ vs. $2.32 \% / \mathrm{Sv})$ and $2.5 \%(5.92 \% / \mathrm{Sv}$ vs. $5.77 \% / \mathrm{Sv})$ for females.
(7). The use of reduction factors to correct for DS86 random error in survivor doses indicated that lifetime risks were negatively biased $15 \%-30 \%$. Bias of excess risk (nonconstant relative projection and correction for diagnostic misclassification) of nonleukemia due to DS86 random errors for 1 Sv at age $25,45,65$ or continuously from ages 18 to 65 was $-27.1 \%(2.24 \% / \mathrm{Sv}$ vs. $2.28 \% / \mathrm{Sv}),-23.5 \%(2.99 \% / \mathrm{Sv}$ vs. $3.69 \% / \mathrm{Sv}),-24.6 \%(1.67 \% / \mathrm{Sv}$ vs. $2.08 \% / \mathrm{Sv})$ and $-23.7 \%(2.13 \% / \mathrm{Sv}$ vs. $2.63 \% / \mathrm{Sv})$ for males and $-19.6 \%(3.54 \% / \mathrm{Sv}$ vs. $4.24 \% / \mathrm{Sv}),-13.9 \%(4.49 \% / \mathrm{Sv}$ vs. $5.12 \% / \mathrm{Sv}),-15.9 \%(2.48 \% / \mathrm{Sv}$ vs. $2.88 \% / \mathrm{Sv})$ and $-14.9 \%(3.43 \% / \mathrm{Sv}$ vs. $3.94 \% / \mathrm{Sv})$ for females. For leukemia excess risks under the same dose profiles, the bias due to DS86 random error was $-17.4 \% ~(0.37 \% / \mathrm{Sv}$ vs. $0.44 \% / \mathrm{Sv})$, $-14.2 \%(0.77 \% / \mathrm{Sv}$ vs. $0.88 \% / \mathrm{Sv}),-13.3 \%(3.04 \% / \mathrm{Sv}$ vs. $3.44 \% / \mathrm{Sv})$ and $-14.0 \%(1.09 \% / \mathrm{Sv}$ vs. $1.24 \% / \mathrm{Sv})$ for males and $-15.1 \%(0.30 \% / \mathrm{Sv}$ vs. $0.34 \% / \mathrm{Sv}),-11.6 \%(0.75 \% / \mathrm{Sv}$ vs. $0.84 \% / \mathrm{Sv}),-10.9 \%(2.77 \% / \mathrm{Sv}$ vs. $3.07 \% / \mathrm{Sv})$, and $-11.4 \%(1.05 \% / \mathrm{Sv}$ vs. $1.17 \% / \mathrm{Sv})$ for females.
(8). The correction of mortality misclassification in SEER baseline rates used in lifetime risk projection (non-constant relative model) increased excess risks by $2.1 \%$ for nonleukemia and decreased risk by $10.8 \%$ for leukemia.
(9). The total bias of excess risk of nonleukemia for exposure from age 18 to 65 under the non-constant relative projection model was $-37.1 \%$ for males and $-23.3 \%$ for females. For leukemia excess risks under the relative projection model, the total bias was $-27.1 \%$ for males and $-43.4 \%$ for females. Thus, nonleukemia risks increased $37.1 \%$ for males ( $1.91 \% / \mathrm{Sv}$ to $2.68 \% / \mathrm{Sv}$ ) and $23.3 \%$ for females $(3.23 \% / \mathrm{Sv}$ to $4.02 \% / \mathrm{Sv})$ and leukemia risks increased $27.1 \%(0.87 \% / \mathrm{Sv}$ to $1.10 \% / \mathrm{Sv})$ for males and $43.4 \%(0.73 \% / \mathrm{Sv}$ to $1.04 \% / \mathrm{Sv}$ ).
(10). In most cases, bias due to diagnostic misclassification for lifetime risk projections using the relative model was more positive and less erratic than bias for the absolute and transported relative models. With regard to risk projection and future studies of information bias, we recommend the relative model because its use, when compared with other models, resulted in biases with lower variation across gender, sites and exposure profiles.

It is patently clear that the effects of diagnostic misclassification and DS86 random errors are dependent on gender, site, correction methods, exposure profiles and projection models. The effects of increased internal validity on the generalizability of Japanese radiation risk information to U.S. nuclear workers are only revealed when lifetime risks are projected after adjustments are made for random and systematic errors. Future studies in which LSS data are generalized to U.S. nuclear workers may be biased if lifetime risks are not adjusted for random and systematic errors.

Readers who favor our results should not let their enthusiasm overtake their knowledge of bias and regard our assumptions as fixed verities, rather than empirical hypotheses. The major purpose for undertaking this study was to confirm the impression that there are certain advantages of projecting lifetime risk after performing Poisson regression when studying information bias in the LSS. Since we did not employ logistic regression to estimate cancer misclassification probabilities and did not fully implement the EM algorithm to impute missing data where there was no autopsy information, this study should be regarded as an investigation into the most fundamental assumptions. As a result, new phenomena in the LSS should not force a reevaluation of this study's findings.

## 2 INTRODUCTION

Recent studies have conclusively demonstrated that diagnostic misclassification and random errors in the Dosimetry System-1986 (DS86) are major components of information bies that can affect lifetime risk projections based on the Life Span Study (LSS) of Hiroshima and Nagasaki atomic bomb survivors (Sposto et al., 1991; Sposto et al., 1992; Pierce and Vaeth, 1989; Pierce et al., 1990, Pierce and Vaeth, 1991; Pierce et al., 1991; Ron et al., 1991). On a simplistic level, the relationship between information bias and lifetime risk is evident to the epidemiologist who would adjust cancer mortality rates used in dose-response analyses with the ratio of the cancer confirmation ${ }^{4}$ rate to the cancer detection rate, or to the statistician who would suggest that a single excess relative risk coefficient adjusted for DS86 random error results in an increase in projected lifetime risk. A complete analysis of information bias or lifetime risk projection, however, requires a more thorough understanding of both topics. Yet, few scientists who write on these general areas actually study the effects of information bias on lifetime risk by conducting dose-response analysis and projecting lifetime risks with the results.

The purpose of this study, then, was to quantify changes in excess cancer mortality risks by correcting for 1) random error in individual DS86 dose equivalents; 2) diagnostic misclassification of cancer rates in the LSS used in dose-response analysis; and 3) dis そnostic misclassification of cancer mortality in U.S. vital statistics that are used in I fetime risk projection.

### 2.1 Precision, Validity, Generalizability and Bias

Random error is a precision issue in epidemiologic research that is related to statistical variation of estimates. It is based mainly on sampling variation and is not generally considered to be of the same importance as systematic error (validity). One should not sacrifice validity for the sake of precision, at least under the stultifying conditions that an increasing number of epidemiologists work in today. Figure 1 shows the relationship between precision (random error) and validity (systematic error) in epidemiologic research.

Kleinbaum et al. (1982) suggest that there are four populations (Figure 2) typically involved in an epidemiologic investigation of disease etiology. Under the present study, the external population is the group of U.S. nuclear workers for which inclusion into the LSS has been restricted but to which results are generalized. Because risk information from the LSS is generalized to U.S. working populations, one must ensure that the external validity is hinged on several criteria related to biologic plausibility, strength of association, doseresponse gradients, temporality, disease specificity and consistency with other findings which, collectively, act to discredit the null or biological hypothesis. The study population is comprised of LSS subjects from which the effect estimator, $\hat{\theta}$, is measured. The effect, estimate, $\theta^{\circ}$, for the actual population, is represented by $\hat{\theta}$ in the sampled study population. The true effect measure, $\theta$, is for the target population through which internal validity and to which statistical (random) and methodological (systematic) issues apply. Figure 3 shows schernatically the hierarchy of populations in the LSS.

The effect measure, $\hat{\theta}$, is an asymptotically unbiased estimator of $\theta$ if random and systematic errors are corrected and

$$
\begin{equation*}
\lim _{n \rightarrow \infty} E(\hat{\theta})=\theta \tag{1}
\end{equation*}
$$

[^3]where $E(\hat{\theta})$ is the expectation of $\theta$. The bias of $\theta$ relative to $\hat{\theta}$ when $\hat{\theta}$ is uncorrected for random and/or systematic error is functionally composed as
\[

$$
\begin{equation*}
B I A S(\hat{\theta}, \theta)=(\hat{\theta}-\theta) / \theta \tag{2}
\end{equation*}
$$

\]

and serves as the underlying construct for comparing lifetime risks that are estimated with and without adjustments for DS86 random error and diagnostic misclassification. When the effect estimate, $\hat{\theta}$, is greater than the true association, $\theta$, bias is positive, however, if the effect estimate, $\dot{\theta}$, is less than the true association, $\theta$, then bias is negative. If $\hat{\theta}$ and $\theta$ are both on either side of the null value and $\hat{\theta}$ is closer to the null value than $\theta$, then bias is defined as being toward the null. If, on the contrary, $\hat{\theta}$ is further away from the null than $\theta$, provided they are both on the same side of the null value, then the bias is said to be away from the null. As an example, if the lifetime mortality risk of radiation-induced cancer, $\hat{\theta}$, is $3 \% / \mathrm{Sv}$ when no correction for random and systematic error is made and the true estimate, $\theta$, is $5 \% / \mathrm{Sv}$ (after error is adjusted), the bias is negative and is toward the null. Likewise, if the uncorrected lifetime risk is $3 \% / \mathrm{Sv}$ and the corrected estimate is $2 \% / \mathrm{Sv}$, then the bias is positive and is away from the null. A simple point to remember is that if excess risks are increased after making an adjustment for random and systematic Giror, the bias is negative and towards the null.

### 2.2 Effects of Random Error on Dose-Response

Random error in, say, the DS86 systern is attributable to the methodology used for estimating DS86 doses and survivor response (Thiessen and Kaul, 1991). As a paradigm, Sposto et al. (1991) recently sampled 1028 subjects from the Adult Health Study (AHS) population (which is a sample of the LSS population) to estimate random error in the DS86 by modeling the dose-response of the combined effects of severe epilation and chromosome aberrations (CA). Figure 4 shows the proportion of cells with CA as a function of corrected and uncorrected DS86 dose for the 1028 survivors. While the two straight lines represent the fitted regression lines for the per cent CA of the no epilation and epilation groups, respectively, the two curvilinear lines represent the fitted dose-response functions (same groups) assuming a $45 \%$ coefficient of variation (CV) in random error. This finding by Sposto et al. should not be surprising because recent analyses in the LSS have shown that the average survivor true dose, $\operatorname{Avg}(x \mid z)$, is less than the estimated dose, $z$, at any level of $z$ because the deviations $(x-z)$ tend more toward the negative rather than the positive (Pierce et al., 1990; Pierce and Vaeth, 1991; Pierce et al., 1991). In addition, the ratio of $\operatorname{Avg}(x \mid z)$ to the average estimated dose, $\operatorname{Avg}(z \mid x)$, decreases from unity as $z$ increases bucause, for increasing $z$, there are fewer survivors. The effect of random error on the dose-response in the previous example suggests that a majority of individuals with severe epilation and notable CA have been assigned DS86 doses which are equal to AHS participants without severe epilation and CA. Thus, to correct for the random errors in DS86, one would most likely increase the slope of the no epilation group by a factor, which according to Pierce and Vaeth (1989) is called the Linear Extrapolation Overestimation Factor or Dose Rate Reduction Effectiveness Factor. Figure 5 illustrates the the relationship between $\operatorname{Avg}(z \mid x)$ and $\operatorname{Avg}(x \mid z)$ as a function of CV for the two cities.


Figure 1: Random and systematic error in epidemiologic research. (With permission, (c)1982 Van Nostrand Reinhold)

External population


Figure 2: Hierarchy of populations in epidemiologic research. (With permission, (c) 1982 Van Nostrand Reinhold)


Figure 3: Hierarchy of populations in the Life Span Study.


Figure 4: Chromosome aberration (CA) dose-response within epilation groups. Doses adjusted for $45 \%$ random error. Adapted from Sposto et al. (1991).


$$
\begin{aligned}
& -\quad 30 \% \mathrm{CV} \\
& -\cdots-35 \mathrm{CV} \\
& -\cdots-40 \% \mathrm{CV} \\
& -\cdots-\cdots 45 \% \mathrm{CV}
\end{aligned}
$$



| $-\sim$ | $30 \% \mathrm{CV}$ |
| :---: | :---: |
| $-\cdots$ | $35 \% \mathrm{CV}$ |
| $-\cdots$ | $40 \% \mathrm{CV}$ |
| $-\cdots-\cdots$ | $45 \% \mathrm{CV}$ |

Figure 5: Plot of survivor true dose, $\operatorname{Avg}(x \mid z)$, as a function of estimated dose, $\operatorname{Avg}(z \mid x)$, for four levels of random error. Adapted from Pierce and Vaeth (1991).

Thiessen and Kaul (1991) reported that random uncertainties in DS86 range from 15\% to $40 \%$ and are mainly attributable to survivor location and shielding parameters (DS86 input) and the choice of shielding factors and an appropriate model (DS88 methodology). Jablon (1971) has suggested that a survivor's reported location on the questionnaires used for dosimetry could have been affected by simple errors due to postconcussion amnesia or to deliberate mistatements hinged on beneficial welfare laws that were dependent on distance from the hypocenter. Random errors in T65DR doses were also analyzed by Gilbert (1982). Her results indicated that by truncating shielded kerma to 600 Gy , bias was reduced but at the expense of a substantial loss of power. Moreover, if the standard errors of the estimates are large, then moderate bias due to random error in doses may be moot.

The overall effect of DS86 random error on dose-response is a downward bias of the risk coefficients in either the linear ( L ) or linear-quadratic ( LQ ) models. If we model with cancer rates held constant and use doses that are not underestimated, then there will be a downward bias of the regression (risk) coefficients. On the contrary, if dose is underestimated, then the regression coefficients in the L and LQ models must make up for the difference between the logarithms of baseline and fitted excess rates, thus, risks become increased. It warrants noting that increases in neutron relative biological effectiveness (RBE) can increase dose equivalent and result in decreased risk coefficients. However, if the neutron component of dose equivalent is reduced, then there will be a increase of risk (regression) coefficients. This is what happened with the reduction of the neutron component in Hiroshima during DS86: the risk coefficients increased because shielded kerma (gamma and neutron) decreased. It should be pointed out that the risk coefficients for DS86 when considering kerma are about $40 \%$ higher than T65DR, but when organ doses and new transmission factors are considered, the coefficients are similar (Shimizu et al., 1988).

### 2.3 Effects of Systematic Error on Dose-Response

Loss of validity in the LSS is attributable to information bias caused by 1) systematic errors in analytic and numerical calculations in DS86 and subsequent misclassification of exposure; 2) misclassification of disease in LSS subjects and 3) selection bias caused by using the 1950 census for cohort construction.

For dosimetry, Thiessen and Kaul (1991) cite sources of systematic error in DS86 arising from the spectral yield, burst altitude, megaton yield and efficiency, and cross sections to determine when the devices went critical and how much the air and shielding materials attenuated and scattered the incident radiation. Systematic errors in DS86 range from $10 \%$ to $15 \%$.

Systematic error in the LSS is also attributable to a selection bias brought about by not sampling the Hiroshima and Nagasaki populations for subjects before the 1950 city censuses were available. The most recent study of systematic error and its effect on the dose-response curve in the LSS was carried out by Sposto et al. (1992). In their analysis, they estimated misclassification probabilities for cancer and non-cancer and discovered that, when adjusting for a $22 \%$ cancer misclassification probability, 839 non-cancer deaths needed to be reclassified as cancer deaths. In addition, after the correction was made for the $22 \%$ misclassification rate, they found that (for males at age ATB 25) the cancer excess relative risk increased from 0.494 to $0.553(12 \%)$ and the number of excess deaths increased from 274 to 317 (16\%). Their findings indicated that a downward bias of risk existed as a result of the underreporting of cancer as the underlying cause of death on
death certificates. More importantly, they demonstrated how to employ the results of the RERF Autopsy Program to increase the validity and generalizability of LSS results to other populations.

Another correction of a downward bias of excess relative and absolute risks occurred with the implementation of the new DS86 shielded kerma values, which resulted in an upward correction of a downward bias in shielded kerma and a subtle upward correction of a downward bias of organ doses. For the reader who is interested in comparisons of sex- and site-specific tress relative risks and absolute risks for the T65DR and DS86, see Report 11, Part. 1 'Shimizu et al., 1987).

### 2.4 Studies of Death Certificate Misclassification

The first extensive evaluation of death certificate validity in the Atomic Bomb Casualty Commission (ABCC) pathology studies was done by Stone and Anderson (1960) on 1165 Hiroshima autopsy cases obtained from 1949 through 1959. In their analysis, they tried to answer several questions generally related to death certificate validity: how representative of the target populat . the sample of cases?; what was the accuracy of autopsy in terms of specifying a single underlying cause of death?; how accurately was the coding performed?; and finally, how comparable were the autopsy diagnoses and underlying cause of death reported on death certificates?

First, they found that the underlying cause of death on the death certificate affected the likelihood of being autopsied and that there was a higher proportion of deaths due to malignancy that were autopsied, rather than non-neopiastic diseases, as indicated by the high correspondence between necropsy and death due to neoplasm. This led them to believe that the population for whom the cases represented was simply unknown. Second, the requirement for a single cause of death caused more difficulty in terms of assigning a correct cause. They also discovered that anatomical findings may be variously interpreted. For example, when clinical information was not available at the time of post mortem evaluation, it was difficult to discern renal insufficiency from diabetes mellitus, hypertension, atherosclerosis, or the combination thereof. Third, it was known that from a sample of 1000 deaths in the ABCC study, there was a $97 \%$ agreement between autopsy diagnoses and underlying cause of death on death certificates. Coding in the present analysis was done by two trained coders and was therefore believed to be very accurate. Finally, the correspondence of underlying cause of disease and autopsy diagnoses for all neoplasms and leukernia were $92 \%$ and $86 \%$, respectively. This showed that the International Statistical Classification (WHO,1959) worked quite well when comparing underlying cause of disease and autopsy diagnoses in this study.

In 1962, a joint pathology study of the A-bomb survivors was instituted among the ABCC, Japanese National Institutes of Health (JNIH), Hiroshima and Nagasaki City Medical Associations, Departments of Pathology of Hiroshima and Nagasaki University Medical Schools, the Hiroshima Red Cross Hospital, Atomic Bomb Hospitals of uovin cities, and Hiroshima University Research Institute for Nuclear Medicine and Biology (Zeldis and Matsumoto, 1962) ${ }^{5}$. This effort was largely due to the Unified Study Plan which called for the mutual support of well-controlled studies to combine clinical, pathologic and vital statistics investigations on 100,000 individuals who were either present in these cities and received large doses of radiation, present in these cities but suffered no radiation injury, or not in the cities at all; this sample of 100,000 persons was called the Life Span Study.

[^4]The plan also called for a new autopsy procurement plan in the LSS since previous work showed strong evidence for selection (Stone and Anderson, 1960). In the ensuing pathology studies, factors influencing autopsy selection were analyzed objectively to determine how the autopsy series might be used for epidemiologic investigations. Immediately, systematic coverage of both cities was begun to collect information on recent deaths. Screening was implemented to determine status within the LSS sample and permission to conduct autopsy was sought from families and others who were concerned. The results of this work were published in Reports 1 through 4 of the ABCC-JNIH Pathology Studies in Hiroshima and Nagasaki (Angevine et al., 1963; Beebe et al., 1967; Steer et al., 1973; Yamamoto et al., 1978).

The latest report of the LSS Pathology Studies, Report 4, suggested that a peak autopsy rate of $45 \%$ was reached in 1963 after which time the rate dropped to $15 \%$ in 1975 (Yanamoto et al., 1978). The rate averaged $19 \%$ from 1971 to 1975. An unusual finding in the report was that from 1961-75, there was a $25.5 \%$ autopsy rate on individuais dying at home; this was a direct result of implementing the autopsy procurement plan. Confirmation and detection rates for neoplasms were higher than those for cerebrovascular and cardiovascular disease, however, there was often disagreement between death certificate and autopsy diagnosis. Nonetheless, it must be kept in mind that the purpose of these analyses was to verify death certificate accuracy in the context of specifying radiation effects. The use of autopsy information alone is limited by the amount and selective nature of such data. In 1975, it was recommended that the autopsy program be terminated. Since then, approximately 8 autopsies have been performed each year, thus leaving Japanese vital statistics as the primary source of information concerning death certificate validity.

Jablon and colleagues (1966) conducted another study of death certificate validity in the LSS and stated that vital statistics for all malignancies were $14 \%$ too low. Specifically, mortality rates for malignant neoplasms of digestive organs were $13 \%$ too low (stomach cancer was $21 \%$ too low, cancer of other digestive organs was $3 \%$ too high); cancer of the respiratory system was $40 \%$ too low; and uterus $4 \%$ too high. It follows that in this setting the true mortality rates for malignant neoplasms were underestimated by Japanese vital statistics.

More recently, in an RERF study on cancer mortality among A-bomb survivors, it was recognized that a wide variation existed for confirmation and detection rates for various causes of death, however, the authors went on to say that there was no evidence to suggest that inaccuracies of death certificates were consistently related to A-bomb exposure (Preston et al., 1986). Two years later, in RERF Report 11 (Shimizu et al., 1988), the investigators recognized that risk projection was affected to some degree by death certificate inaccuracies and recommended the site-specific correction of these insufficiencies, however, they used a crude correction of 1.23 , which was identical to that used by the BEIR-III committee (NRC, 1980). In BEIR-V, although no correction for death certificate misclassification was made, the problem of diagnostic misclassification was circumvented by restricting analyses to survivors whose attained age was less than 75 , since it was known that misclassification increases dramatically after an attained age of 75 or thereabout.

Although much work has been done by the RERF in the way of providing insight about death certificate validity and the selective nature of autopsy in the LSS, little has been done to use these site-specific data for risk estimation (NRC, 1990).

A recent study of the LSS autopsy data revealed that, overall, cancer mortality is underestimated by about $18 \%$ (Ron et al., 1991). In addition, for Cancers of Interest (lymphoma, breast, brain, multiple myeloma and melanoma) they found a $40 \%$ increase
in mortality rates between 1962 and 1982. Their results, as reporte-d. do not really lend themselves well for use in this analysis because they did not providr accer misclassification probabilities that were stratified by site, sex, city, age ATB, age ATD and follow-up period, since the study only addressed cancer mortality trends. Sposto et al. (1992) recently performed a dose-response analysis using LSS nonleukemia data corrected for a $22 \%$ cancer misclassification probability and observed a $12 \%$ increase in excess relative risk and a $16 \%$ increase in absolute risk for Hiroshima males exposed at age 25 . Although they modeled and used nonleukemia misclassification probabilities as a function of city, sex, age ATB, age ATD and dose in the EM algorithm to impute missing data, they did not project lifetime risks to reveal the full effect of misclassification since the main focus was on the possibility that an apparent increase in the non-cancer death rate was attributable to cancer deaths being misclassified as non-cancer deaths.

In the United States, studies of death certificate misclassification for malignant disease have been conducted since 1941 (Dorn and Horn, 1941). Some involved a small number of cases and were limited in scope (Moriyama et al., 1958; James et al., 1955). Among the large-scale studies are Dorn and Horn's on the First National Cancer Survey, Dorn and Cutler's on the Second National Cancer Survey, and the Pan American Health Association's study (Dorn and Cutler, 1958; Puffer and Griffith, 1967). More recent studies by Percy et al. showed that according to the underlying cause death, $65 \%$ of the death certificates were accurate (Percy et al., 1981; Percy et al., 1990). Ron et al. (1991) report on a historical review of cancer mortality misclassification in the U.S.

### 2.5 Studies of Lifetime Risk

Several studies reflect the state-of-the-art in lifetime risk projection. With the exception of female breast cancer, the BEIR-V study relied solely on the LSS data to project lifetime risk of developing cancer in various sites (NRC,1990). While detailed descriptions on risk projection and respective uncertainties were well documented throughout the report, no effort was made to correct for site-specific death certificate misclassification. However, the BEIR-V analyses only included data for which survivor attained age was less than 75 an age at which misclassification starts to increase. The BEIR-III committee corrected for death certificate incompleteness, but instead of taking a site-specific approach, they used a crude correction factor of 1.23 (NRC,1980). The United Nations Scientific Committee on the Effects of Atomic Radiation (UNSCEAR) report "Sources, Effects and Risks of Ionizing Radiation" also relied to a large extent on the A-bomb data for the purpose of making lifetime risk estimates (UN, 1988). Here again, the authors recognized the uncertainties due to death certificate misclassification and underscored the need to account for such variation. Unfortunately, this comprehensive evaluation of lifetime risk from radiation exposure followed the already suffering method of providing risk estimates without correcting for random and systematic error.

Gilbert's classic health effects studies reported on radiation-induced late effects for an exposed working population (Gilbert, 1989a; Gilbert 1989b; Gilbert 1991). She used the life table approach and combined U.S. vital statistics and LSS data to obtain risk estimates constrained by lower, middle, and upper boundary conditions. However, a correction for site-specific misclassification was not made. Stather and his colleagues (Stather et al., 1988) conducted a health effects study using radiation risk data published in the 1988 UNSCEAR report. Their results indicated that risk estimates for human exposure to radiation are three times higher than risk estimates introduced by the International Commission on Radiological Protection in 1977 (ICRP, 1977). This is in good agreement with
the findings of the BEIR-V committee, who suggested a 4- to 5 -fold increase in risk since the BEIR-III committee published its findings in 1980. These apparent increases in risk caused much concern in the area of radiation risk assessment and warranted a reappraisal of current radiation protection guidelines by the International Commission on Radiological Protection (ICRP, 1991). In ICRP Report 60, however, there is no discussion about the effects of diagnostic misclassification on lifetime risk estimates. In another ICRP study, Land and Sinclair (1991) used risk coefficients from Tables 5A and 5B of RERF Report 11 (Shimizu et al., 1988) to project risk for a number of Western populations, but did not adjust their lifetime risk estimates for DS86 random error or diagnostic misclassification. Within this framework, it was propitious to pursue this investigation in view of such findings.

The NIH study to develop radioepidemiological tables also deserves mention (Rall et al., 1985). Multiplicative and additive risk data were used to determine age- and sexspecific risk at a point in time from a previous exposure. No mention was made for the correction of site-specific misclassification and its effect on radiation risk estimates.

What these and other studies lack is an analytic evaluation of the degree to which sitespecific diagnostic misclassification and DS86 random error jointly affect lifetime mortality risks. This investigation has the distinct advantage of complementing the above studies in order to increase internal validity in the LSS (by correcting estimates of lifetime risk) to therefore understand changes in the generalizability of results to U.S. nuclear workers.

### 2.6 Research Objectives

The following is a list of specific objectives for this investigation:
(1). Obtain for the years 1950-75, confirmation and detection rates for the leukemia and nonleukemia sites published in RERF Pathology Reports.
(2). Estimate cancer and non-cancer confirmation rates for the BEIR-V digestive cancer site by combining data for rubrics such as the stomach and colon.
(3). Estimate the true number of cancer deaths in each subpopulation of the LSS cancer mortality data by using sufficient statistics of the expectation-maximization (EM) algorithm ${ }^{6}$.
(4). Calculate organ radiation absorbed doses ${ }^{7}$ from shielded kerma ${ }^{8}$ using body selfshielding transmission factors for the marrow, stomach, and colon.
(5). Model the excess relative risk (ERR) of radiation-induced cancer mortality for the leukemia, nonleukemia and BEIR-V digestive cancer sites with and without use of sufficient statistics and adjustment for DS86 random error using non-linear Poisson regression. Variables to be used in the analysis are organ radiation dose equivalent and covariates (effect modifiers) such as age ATB, age ATD, sex, and city (Hiroshima or Nagasaki). The L and LQ dose-response models will be used and the Pearson chi-square, deviance and Freeman-Tukey goodness-of-fit residuals determined for each model.
(6). Determine a Death Certificate Correction Factor (DCCF) for baseline rates of leukemia, nonleukemia and BEIR-V digestive cancers by dividing each site's confirmation rate by its detection rate obtained from the SEER data.
(7). Use a life-table method to combine ERR and absolute risk (AR) coefficients with

[^5]SEER baseline rates to obtain lifetime risk coefficients for a working U.S. population with and without using SEER-based DCCFs. Generate $90 \%$ confidence intervals of lifetime risk coefficients based on "model" and "non-model" geometric standard deviations, DRREFs and linear or linear-quadratic models. Use a DRREF of two to generate sex-specific lifetime risks (excess deaths/Sv/ 100,000 population) for the following exposure profiles: 1 Sv at age $25,1 \mathrm{~Sv}$ at age $45,1 \mathrm{~Sv}$ at age 65 , and $0.02128 \mathrm{~Sv} / \mathrm{y}$ from age 18 to 65 ( 1 Sv total).
(8). Ascertain the effect of nondifferential and differential misclassification of cancer mortality on point estimates of lifetime risk.

## 3 MATERIALS AND METHODS

The following sections outline the various methods employed in the study. Figure 6 shows the typical methodology used for risk assessment in radioepidemiologic studies and Figure 7 illustrates the method used in the present study.

### 3.1 Sources of Data

### 3.1.1 RERF Autopsy Program

Between January 1961 and December 1975, the RERF performed 4,920 autopsies during the Autopsy Progam. Results of the Autopsy Program are reported in RERF Pathology Reports 1-4 and contain autopsy characteristics as a function of city (Hiroshima or Nagasaki), place (RERF or other), exposure (T65D shieided kerma), sex and age ATD (Angevine et al., 1963; Beebe et al., 1967; Steer et al., 1973; Yamamoto et al., 1978). Although Reports 1-4 list confirmation and detection rates (discussed below) for leukemia and nonleukemia, there were no data for the BEIR-V digestive cancer site. Section 3.2 describes confirmation and detection rates, estimation of confirmation rates for the BEIR-V leukemia and digestive models, and the use of cancer and non-cancer confirmation rates to determine the true number of cancer deaths in each subpopulation (stratum).

### 3.1.2 RERF Cancer Mortality Data

The RERF continually maintains a computer data base which contains the status of LSS subjects at the time of each 5 -year follow-up. The mortality status of each survivor at follow-up is determined by searching for LSS study subjects in the obligatory household registries (koseki) throughout Japan. Death certificate information, namely, underlying cause of death, for any survivor is obtained from the Vital Statistics Death Schedules and appended to the computer data base. At present, the LSS listing contains information on 5,936 cancer deaths for the years 1950-85.

As of 1985 , there were 120,128 survivors in the extended cohort of the Life Span Study (LSS-E85) of which 75,991 have been assigned radiation doses from the Dosimetry System 1986 (DS86) (Beebe and Usagawa, 1968; Shimizu et al., 1988). Survivors for which DS86 dose estimates do not exist include 26,517 who were not in the cities (NIC) at time of the bombings, 2,383 with insufficient shielding information and 15,237 who had doses from the Tentative Dosimetry System 1965 Revised (TD65R) (Milton and Shohoji, 1968) but for which DS86 doses could not be calculated. The LSS mortality data are cross-classified into several age-, sex-, age ATB-, age ATD-, and dose-specific categories as shown in Table 1. Since confirmation and detection rates are proportions (discussed later), site-specific sample sizes were based on the higher of the two sample size estimates for each proportion (Cochran, 1977). The site-specific precisions expected from using all of the data were less than 0.05 , except for the colon ( 0.074 ) and breast ( 0.066 ). In addition to the categorical covariates in Table 1, there are several person-year weighted continuous variables for the mean age ATB, mean age ATD, and the gamma and neutron components of shielded kerma (See Appendix A).

### 3.1.3 RERF Average Body Transmission Factors

The RERF has maintained dosimetry information for all of the study subjects in the LSS. These data include DS86 estimates of the shielded kerma from gamma rays and neutrons


Figure 6: Typical methodology for estimating lifetime mortality risk of radiationinduced cancer.

## Relative model



* Healthy worky cifect is a works-in-progress

Figure 7: Methodology for estimating lifetime mortality risk of radiation-induced cancer in this investigation.

Table 1: Cross-classification of LSS cancer mortality data.

| Category | Levels | Description |
| :--- | :--- | :--- |
| City | 2 | Hiroshima, Nagasaki |
| Sex | 2 | Males, Fernales |
| Age at exposure | 13 | $0-4,5-9,10-14$, 15-19, 20-24, |
|  |  | 25-29, 30-34, 35-39, 40-44, |
| Follow-up period | 7 | 45-49, 50-54, 55-59, 60+ |
|  |  | 1 Oct 1950 to 31 Dec 1955 |
|  | 1 Jan 1956 to 31 Dec 1960 |  |
|  |  | 1 Jan 1961 to 31 Dec 1965 |
|  |  | 1 Jan 1966 to 31 Dec 1970 |
|  | 1 Jan 1971 to 31 Dec 1975 |  |
|  |  | 1 Jan 1976 to 31 Dec 1980 |
|  |  | 1 Jan 1981 to 31 Dec 1985 |

in units of mGy, the location and radiation shielding at age ATB, sex, city, and distance in meters from the hypocenter of the blast. Information on the organ-, city-, age ATB-, and radiation-specific body transmission factors are also available to convert kerma to organ dose. The average body transmission factors are also cross-classified into several organ-, age ATB-, and city-specific categories and were used for converting whole-body shielded kerma into organ absorbed doses (see Appendix A). For the neutron portion of dose equivalent, an RBE factor of 10 was used for the RERF models. In the BEIR-V relative risk models, a neutron RBE of 20 was used.

### 3.1.4 Reduction Factors for DS86 Random Error

City-specific reduction factors used to adjust DS86 survivor doses were adapted from previous work on DS86 random errors (Pierce and Vaeth, 1991). Reduction factors were multiplied by organ dose equivalents which were used in the dose-response analysis described in the next section (see Appendix A).

### 3.2 Use of Confirmation Rates to Adjust Cancer Deaths

### 3.2.1 Diagnostio Screening

Comparisons if ortality between death certificates and autopsy records reported in the RERF Aute $\quad$. rogram (Yamamoto et al., 1978) are identical to the results of screening tests (Lilir and Lilienfeld, 1980; Fleiss, 1981; Kramer, 1988). Data arrangement for results of the LSS Autopsy Program are arranged in Table 2.

As one notices in Table 2, the sensitivity, which is equivalent to the cancer detection rate and the ratio $a /(a+c)$, is defined as the probability of correctly assigning cancer X as the underlying cause of death on a death certificate given that the principal autopsy finding was cancer X . The specificity, which is equivalent to the non-cancer detection rate and the ratio $d /(b+d)$, is defined as the probability of correctly assigning nou-cancer as the underlying cause of death on a death certificate given that the decedent's principal autopsy finding was non-cancer. The predictive value positive $\left(P V^{+}\right)$, which is equivalent to the cancer confirmation rate and the ratio $a /(a+b)$, is defined as the probability that

Table 2: Data arrangement of screening results in the RERF.

|  | Autopsy diagnosis |  |  |
| :--- | :---: | :---: | :--- |
| Death Certificate | Cancer $X$ | Non-cancer |  |
| Cancer $X$ | $a$-confirmed | $b$-false positives | $a+b=d_{c}$ |
| Non-cancer | $c$-false negatives | $d$-absence of Cancer $X$ | $c+d=d_{n c}$ |
|  | $a+c=D_{c}$ | $b+d=D_{n c}$ | $a+b+c+d=d_{T}$ |

Table 3: Probabilities of misclassification of disease.

|  | Autopsy diagnosis |  |  |
| :--- | :---: | :---: | :---: |
| Death Certificate | Cancer $X$ | Non-cancer |  |
| Cancer $X$ | $\phi$ | $(1-\psi)$ |  |
| Non-cancer | $(1-\phi)$ | $\psi$ |  |
|  | $D_{c}$ | $D_{n c}$ | $D_{T}$ |

an individual with cancer X as the underlying cause on their death certificate actually died of cancer. Lastly, the predictive value negative $\left(\mathrm{PV}^{-}\right)$, which is equivalent to the non-cancer confirmation rate and the ratio $d /(c+d)$, is defined as the probability that an individual with non-cancer as the underlying cause on their death certificate actually did not die of cancer X. The observed number of cancer deaths on death certificates of a sample of LSS survivors is $d_{c}$ and the observed number of non-cancer deaths is $d_{n c}$. The total number of deaths due to cancer and non-cancer is $d_{T}$. When sensitivity and specificity differ across exposure levels, misclassification is termed differential. However, when sensitivity and specificity are equal across exposure levels, the misclassification is called non-differential.

Confirmation rates for the BEIR-V digestive (ICD 150-159) cancer sites were estimated as the ratio of the total number confirmed (a) to the total number of death certificates sampled ( $a+b=d_{c}$ ) within each rubric.

### 3.2.2 Estimation of True Cancer Deaths

In order to adjust the observed number of cancer deaths in a given subpopulation, $d_{\mathrm{c}}$, for diagnostic misclassification, it was necessary to estimate the true number of cancer deaths, $D_{e}$ and the true number of non-cancer deaths, $D_{\text {ne }}$. If we denote the sensitivity as $\phi$, specificity as $\psi$, cancer confirmation rate as $\theta_{c}$, non-cancer confirmation rate as $\theta_{n c}$, true cancer rate, $\pi_{c}$, as $D_{c} / d_{T}$ and the true non-cancer rate, $\pi_{n c}$, as $D_{n c} / d_{T}$, then one can see that Tables 2 and 3 can be combined to determining the relationships between each of the above parameters.

Arithmetically, the cancer confirmation rate $\theta_{c}$ is related to $\phi$ and $\psi$ by the relationship

$$
\begin{equation*}
\theta_{\mathrm{e}}=\frac{\phi \pi_{\mathrm{c}}}{\phi \pi_{\mathrm{e}}+(1-\psi) \pi_{n c}} \tag{3}
\end{equation*}
$$

and the relationship between the non-cancer confirmation rate, $\theta_{n c}$, and $\phi$ and $\psi$ is

$$
\begin{equation*}
\theta_{n c}=\frac{\phi \pi_{n c}}{\phi \pi_{n c}+(1-\psi) \pi_{c}} \tag{4}
\end{equation*}
$$

The sufficient statistics for estimating $D_{c}$ and $D_{n c}$ in each cell of the cross-tajuiatad LSS person-year table are

$$
\begin{equation*}
D_{c}=\theta_{c} d_{c}+\left(1-\partial_{n c}\right) d_{n c} \tag{5}
\end{equation*}
$$

Table 4: Crude confirmation rates for cancer, $\theta_{c}$, and non-cancer, $\theta_{n c}$, in the Life Span Study Pathology Report 4.

| Cancer site | $\theta_{c}$ | $\theta_{n e}$ |
| :--- | :---: | :---: |
| Leukemia | 0.857 | 0.999 |
| Nonleukemia | 0.657 | 0.800 |
| Digestive | 0.782 | 0.914 |

and

$$
\begin{equation*}
D_{n c}=\theta_{n c} d_{n c}+\left(1-\theta_{c}\right) d_{c} \tag{6}
\end{equation*}
$$

Confirmation rates for cancer and non-cancer estimated from RERF Pathology Report 4 (Yamamoto et al., 1978) were used in Eq. 5 to estimate the true number of cancer deaths for each stratum of the LSS cancer mortality data before modeling dose-response. Minimum latency periods of 2 years for leukemia and 10 years for solid cancers were user! so that the estimation affected only deaths that were likely to be radiation-induced "ables 4-8 list the cancer and non-cancer confirmation rates from RERF Pathology \%eport 4 for all covariables jointly (crude), and as a function of gender, city, age ATD and ' f 65 DR shielded kerma that were used in Eq. 5 for adjusting mortality for diagnostic misclassification. In order to use the confirmation rates in Pathology Report 4 that were stratified on T65DR shielded kerma (Table 8) with cancer mortality data hinged on the DS86 doses (Table 1), it was necessary to convert DS88 kerma into T65DR kerma.

Using average house transmission factors from Table 1 of the Appendix of Shimizu et al. (1987), we estimated the city-specific T65DR neutron organ dose equivalents as

$$
\begin{equation*}
D_{i j, n, 65, \text { eity }}^{*}=\frac{D_{i j, n, 86, \text { eity }}^{*}}{\Omega_{n, 86, \text { eity }} / \Omega_{n, 65, \text { eity }}} \tag{7}
\end{equation*}
$$

where $D_{i j, n, 86, \text { city }}^{*}$ is the city-specific DS86 organ dose equivalent from neutrons corrected for random error and $\Omega_{n, 86 \text {, city }}$ and $\Omega_{n, 65 \text {, city }}$ are city-specific average house transmission factors for the DS86 and T65DR systems from Table 1 of the Appendix in Shimizu et al. (1987). The T65DR $\gamma$-ray organ dose equivalents were functionally composed as

$$
\begin{equation*}
D_{i j, \gamma, 65, \text { city }}^{*}=\frac{D_{i j, \gamma, 86, \text { eity }}^{*}}{\Omega_{\gamma, 86, \text { city }}^{*} / \Omega_{\gamma, 65, \text { city }}} \tag{8}
\end{equation*}
$$

where $D_{i j, \gamma, 86, \text { city }}^{*}$ is the city-specific DS86 organ dose equivalent for $\gamma$-rays corrected for random error and $\Omega_{\gamma, 86 \text {,city }}$ and $\Omega_{\gamma, 65, \text { city }}$ are city-specific average house transmission factors from Table 1 of the Appendix in Shimizu et al. (1987). The city-specific neutron, $D_{i j, n, 65, \text { city }}^{*}$, and $\gamma-\mathrm{r}$; organ doses, $D_{i j, \gamma, 65, \text { city }}^{*}$, were summed to provide the total organ dose equivalent for selecting a confirmation rate in Table 8 based on a given T65DR dose range. (Appendix A provides a thorough explanation of the methods used for estimating organ dose equivalents).

### 3.3 Dose-Response Analysis

### 3.3.1 Excess Relative Risks

Additive and multiplicative models of relative risk wet used to estimate cancer risk coefficients for each sex, age ATB and age ATD category (Brown and Chu, 1989; Kodell and

Gaylor, 1989; Kodell et al., 1991). The ERR risk model used in this investigation followed that used in RERF Report 11 (Shimizu et al., 1988). For the reader who is interested in further study, Muirhend and Darby (1987) provide an extensive evaluation of estimating radiation risks with additive and multiplicative maximum likelihood (ML) methods. Using the mortality data described in the previous section, we define the mortality rate, $\lambda_{i j}$, in the ith stratum of city, sex and age ATB categories and jth exposure category as

$$
\begin{equation*}
\lambda_{i j}=\lambda_{i 0} \Phi_{R R}(a) \tag{9}
\end{equation*}
$$

where $\lambda_{i 0}$ is the mortality rate ( $D_{c} /$ person-years $\times 10,000$ ) in the 0 dose category of the $i$ th stratum of city, sex and ATB cross-classification and $\Phi_{R R}(a)$ is the relative risk coefficient for exposure at age ATB a. Since the relative risk is related to the excess relative risk as

$$
\begin{equation*}
\Phi_{R R}(a)=\left[1+\Phi_{E R R}(a)\right] \tag{10}
\end{equation*}
$$

we can obtain maximum likelihood (ML) estimates of $\Phi_{E R R}(a)$ by first fitting a model of the form

$$
\begin{equation*}
\lambda_{i j}=\alpha_{i,} e^{\beta_{0}}\left[1+\left\{\beta_{1} D_{i j}^{*} e^{\left(\boldsymbol{\beta}^{T} ; \boldsymbol{z}\right)}\right\}\right] \tag{11}
\end{equation*}
$$

where $\alpha_{s}$ is an unknown nuisance parameter for the stratification of background rates ( $\lambda_{i 0}$ ) on sex ( 2 levels), city ( 2 levels), age ATB ( 13 levels) and follow-up period ( 7 leveis) resulting in $i=364$ strata, $\exp \left(\beta_{0}\right)$ is a constant term, $\beta_{1}$ is the contribution of dose equivalent to excess relative risk, $D_{i j}^{*}$ is the organ dose equivalent and $z$ is a row vector of covariatcs representing age ATB, age ATD or gender.

Once the model has been fit and ML estimates of nuisance parameters and regression coefficients are known, then the excess relative risk at the 1 Sv level for exposure at ag a under the constant model, $\Phi_{E R R}(a)$, can be determined as

$$
\begin{equation*}
\Phi_{E R R}(a)=\beta_{1} e^{\left(\beta^{T} ; \boldsymbol{z}\right)} \tag{12}
\end{equation*}
$$

where $\beta_{1}$ is a ML er. f the linear contribution of dose equivalent to the outcome effect, $\boldsymbol{\beta}^{T}$ is the transform of row vector $\beta$ of coeflicients and $\boldsymbol{z}$ is a row vector of covariates representing age ATB, age ATD and gender. When covariates for age ATD are included in the regression model of Eq. 11, we can obtain the excess relative risk, $\Phi_{E R R}(a, t)$, for the non-constant model, which changes with attained age $t$. To fit the model in Eq. 11, the computer program AMFIT was used for grouped Poisson regression with a stratified excess relative risk model (Fieston and Pierce, 1993). Appendix A provides a detailed description of model formulation, coding methods and matrix operations used for estimating sex-, age ATB- and age ATD-specific $\Phi_{E R R}(a)$ and $\Phi_{E R R}(a, t)$.

Age ATB-, age ATD-, and sex-specific non-constant excess relative risk coefficients, $\Phi_{E R R}(a, t)$, in units of $\% / \mathrm{Sv}$ were estimated for the leukemia and nonleukernia sites with neutron RBEs of $10 .{ }^{9}$

The BEIR-V ERR model for estimating $\Phi_{E R R}(a, t)$ in each LSS subpopulation exposed at age $a$ at $t$ years since exposure (NRC, 1990) was

$$
\begin{equation*}
\Phi_{E R R}(a, t)=f(d) g(\beta) \tag{13}
\end{equation*}
$$

[^6]Table 5: Sex-specific confirmation rates for cancer, $\theta_{c}$, and non-cancer, $\theta_{n c}$, in the Life Span Study Pathology Report 4.

|  | $\theta_{c}$ |  |  | $\theta_{n c}$ |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Cancer site | Males | Female |  | Males | Female |
|  | 0.850 | 0.863 |  | 1.000 | 0.998 |
| Leukemia | 0.638 |  | 0.792 | 0.810 |  |
| Nonleukemia | 0.688 | 0.630 |  |  |  |
| Digestive | 0.787 | 0.764 |  | 0.900 | 0.927 |

where $f(d)$ is a function of either the linear $\left(\alpha_{1} D_{i j}^{*}\right)$ or linear-quadratic $\left(\alpha_{1} D_{i j}^{*}+\alpha_{2} D_{i j}^{* 2}\right)$ contribution of radiation dose and $g(\beta)$ is a link function equal to $\exp \left(\boldsymbol{\beta}^{T} ; \boldsymbol{z}\right)$ dependent on sex, age ATB, and time since exposure (see Appendix A). Absolute risks (excess deaths $/ 10^{4} \mathrm{PYSv}$ ) were not estimated from regression coefficients of the BEIR-V models. When fitting the model for digestive cancer, records were dropped if the time since exposure was $\leq 10$ years, attained age exceeded 75 years or organ dose equivalents (neutron $\mathrm{RBE}=20$ ) exceeded 4 Sv . However, when fitting the leukemia mocel, records were dropped if the bone marrow dose equivelent (neutron $R B E=20$ ) exceeded 4 Sv or attained age exceeded 75 years. ${ }^{10}$

### 3.3.2 Absolute Risks

Absolute risks, or the number of excess deaths per $10^{4}$ person-years at the 1 Sv level were estimated by use of the formula

$$
\begin{equation*}
\Phi_{A R}(a)=\left(\sum_{i} \sum_{k}\left(P Y_{i j} \lambda_{i 0} \Phi_{E R R}(a) D_{i j}^{*}\right) / \sum_{i} \sum_{j}\left(P Y_{i j} D_{i j}^{*}\right)\right) \times 10^{4} \tag{14}
\end{equation*}
$$

where $P Y_{i j}$ is the person-years of follow-up in each subpopulation and the other parameters are defined above. Age ATB-, age ATD-, and sex-specific AR coefficients, $\Phi_{A R}(a, t)$, in units of deaths $/ 10^{4} \mathrm{PYSv}$ were estimated for the leukemia and nonleukemia sites with neutron RBEs of 10 when $\Phi_{E R R}(a, t)$ was used in the above equation. Absolute risks were not estimated from regression coefficients of the BEIR-V models.

### 3.3.3 Goodners-of-Fit (GOF) Statistics

Regression residuals, defined as the squared difference between the observed cancer deaths, $y_{i}$, and the predicted deaths, $\hat{\mu}_{i}$, were determined to ascertain how well each model fitted the observed data (Rayner and Best, 1989). Cressie and Read (1984) introduced the power divergence family of test statistics, which were used in the present study for assessing goodness-of-fit (GOF). When $\hat{\mu}_{i} \geq 5$ for all $i$ then Pearson $\chi^{2}$ residuals

$$
\begin{equation*}
r_{P}=\left(y_{i}-\hat{\mu}_{i}\right)^{2} / \hat{\mu}_{i} \tag{15}
\end{equation*}
$$

and GOF statistic, $\chi^{2}=\sum r_{p}^{2}$ are adequate measures of dispersion. If all $\hat{\mu}_{i} \leq 1$ or $\hat{\mu}_{i} \rightarrow 0$, then deviance residuals

$$
\begin{equation*}
r_{D}=2\left[y_{i} \log \frac{y_{i}}{\mu_{i}}\right]^{1 / 2} \tag{16}
\end{equation*}
$$

[^7]Table 6: Sex- and city-specific confirmation rates for cancer, $\theta_{0}$, and non-cancer, $\theta_{n c}$, in the Life Span Study Pathology Report 4.

|  | Hiroshima, $\theta_{c}$ |  | Nagasaki, $\theta_{c}$ |  | Hiroshima, $\theta_{n c}$ |  | Nagasaki, $\theta_{n c}$ |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Cancer site | Males | Females | Males | Females | Males | Females | Males | Females |
| Leukemia | 0.846 | 0.769 | 0.857 | 0.999 | 1.000 | 0.999 | 1.000 | 0.996 |
| Nonleukemia | 0.685 | 0.629 | 0.695 | 0.670 | 0.954 | 0.959 | 0.947 | 0.970 |
| Digestive | 0.775 | 0.779 | 0.833 | 0.702 | 0.964 | 0.974 | 0.990 | 0.994 |

Table 7: Age ATD-specific confirmation rates for cancer, $\theta_{c}$, and non-cancer, $\theta_{n c}$, in the Life Span Study Pathology Report 4.

| Cancer site | Age ATD, $\theta_{c}$ |  |  |  | Age ATD, $\theta_{n c}$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $<50$ | 50-59 | 60-69 | $70+$ | $<50$ | 50-59 | 60-69 | $70+$ |
| Leukemia | 0.809 | 0.800 | 1.000 | 0.857 | 0.998 | 0.998 | 0.999 | 1.000 |
| Nonleukemia | 0.936 | 0.944 | 0.920 | 0.927 | 0.927 | 0.907 | 0.897 | 0.893 |

and (deviance GOF $D=\sum r_{D}^{2}$ ) Freeman-Tukey, $r_{F r}$, residuals

$$
\begin{equation*}
r_{F T}=\sqrt{y_{i}}+\sqrt{y_{i}+1}-\sqrt{4 \hat{\mu}_{i}+1} \tag{17}
\end{equation*}
$$

1 statistic $G=\sum r_{F T}^{2}$ are more appropriate for assessing GOF.
A model is said to fit a given set of data if $\chi^{2}, D$, or $G$ do not exceed tabled values of $\chi_{(\alpha, n-s-p)}^{2}$ where $n$ is the total number of cells, $s$ is the total number of cells in the stratification and $p$ is the number of parameters in the model. (See Appendix A for a description of numerical methods employed in this study to determine GOF).

### 3.4 Projection of Lifetime Mortality Risks

### 3.4.1 Risk Coefficients and Projectir,n Models

Lifetime mortality risks of cancer for non-exposed and exposed populations were calculated using the program SURVRAD (Peterson et al., 1992). Age- and sex-specific AR and ERR coefficients for radiation-induced cancer were obtained from the dose-response analyses described earlier. Risk projections were made with four models in which 100,000 males and females were exposed to 1 Sv at age $25,45,65$ or to $0.02128 \mathrm{~Sv} /$ year continuously from

Table 8: T65DR-specific confirmation rates for cancer, $\theta_{c}$, and non-cancer, $\theta_{n c}$, in the Life Span Study Pathology Report 4.

|  | $\theta_{c}$ |  |  | $\theta_{n c}$ |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Shielded kerma (Gy) | Leukemia | Nonleukemia | Leukemia | Nonleukemia |  |  |
| $<0.01$ | 1.000 | 0.941 | 1.000 | 0.895 |  |  |
| $0.01-0.49$ | 0.769 | 0.914 | 0.990 | 0.916 |  |  |
| $0.50-0.99$ | 1.000 | 0.917 | 0.996 | 0.881 |  |  |
| $1.00-1.99$ | 1.000 | 0.980 | 1.000 | 0.875 |  |  |
| $2.00+$ | 0.714 | 0.926 | 0.988 | 0.867 |  |  |

age 18 to 65 , for a total career dose equivalent of $1 \mathrm{~Sv}^{11}$. The unconditional probability of radiation-induced cancer mortality, $\pi(\infty ; d)$, over a lifetime for the constant AR model was in the form

$$
\begin{equation*}
\pi(\infty ; d)=\int_{0}^{\infty} \int_{t-p}^{t-1} H(a) \Phi_{A R}(a) S(t ; d) d a d t \tag{18}
\end{equation*}
$$

where $\infty$ is by convention 100 years of age, $t-p$ prevents integration below the minimal latency period for the first (or only) age at exposure $a, t-l$ prevents integration beyond the plateau period for the last age at exposure (Checkoway et al., 1989), $H(a)$ is the annual dose equivalent in $\mathrm{Sv}, \Phi_{A R}(a)$ is the sex-, age ATB- and age ATD-specific absolute risk coefficient (deaths $/ 10^{4} \mathrm{PYSv}$ ) from $\S 3.3$, and $S(t ; d)$ is the all-cause survivorship function for each one-year interval of the complete life table. The number of radiation-induced cancer deaths per 100,000 exposed individuals is $\pi(\infty ; d) \times 10^{5}$. The unconditional probability of radiation-induced cancer mortality based on the constant transported $\mathrm{RR}(\mathrm{AR})$ model was calculated with the formula

$$
\begin{equation*}
\pi(\infty ; d)=\int_{0}^{\infty} \int_{t-p}^{t-1} H(a) \Phi_{E R R, U S}(a) h_{c}(t ; 0) S(t ; d) d a d t \tag{19}
\end{equation*}
$$

where the integrand $\Phi_{\text {ERR,US }}(a)$ is the ERR risk coefficient for the U.S. population determined by applying baseline cancer mortality rates over the relevant 35 -year ( $1950-85$ ) follow-up period ${ }^{12}$ in the LSS and $h_{c}(t ; 0)$ is the baseline cancer rate for spontaneously occurring cancer at age $t$.

Unconditional probabilities for the RR risk projection model were based on applying ERR coefficients obtained in this study directiy to baseline (spontaneously occurring) cancer rates and life tables for the U.S. population. This was functionally composed as

$$
\begin{equation*}
\pi(\infty ; d)=\int_{0}^{\infty} \int_{t-p}^{t-1} H(a) \Phi_{E R R}(a) h_{c}(t ; 0) S(t ; d) d a d t \tag{20}
\end{equation*}
$$

Finally, for the non-constant RERF and BEIR-V models, we used sex-, age ATB- and time-since-exposure (TSE)-specific ERR coefficients obtained in this study in the form

$$
\begin{equation*}
\pi(\infty ; d)=\int_{0}^{\infty} \int_{t-p}^{t-1} H(a) \Phi_{E R R}(a, t) h_{c}(t ; 0) S(t ; d) d a d t \tag{21}
\end{equation*}
$$

[^8]where $\Phi_{E R R}(a, t)$ is the ERR risk cuefficient at age $t$ for exposure at age $a$.
In the four projection models given above, risk was lagged for 2 years and held constant for 40 years for leukemia and was lagged 10 years and held constant until the end of life ( 100 years) for solid cancers. Appendix B outlines the underlying construct of lifetime risk projection and provides detailed explanations of each parameter used in the SURVRAD algorithm.

### 3.4.2 Baseline Rates and Life Tables

Age-specific mortality rates, $h_{\mathrm{c}}(t ; 0)$, for cancer in the 1985 U.S. population were obtained from data files used by the MONSON program (Monson, 1977). Rates for chronic lymphocytic leukemia (CLL) in whites and non-whites were extracted from the most recent Surveillance, Epidemiology and End Results monograph (NCI, 1986) and divided by the age-specific composite (total) leukemia rates to estimate the per cent contribution of CLL to overall leukemia for each sex and age group in 19:3. The resulting proportion of CLL in each quinquennium was then subtracted from the composite leukemia rates. Complete life tables for the 1990 U.S. population were based on data obtained from the Office of the Actuary of the Social Security Administration (Faber and Wade, 1983).

### 3.4.3 Death Certificate Correction Factors (DCCF)

A common misconception in epidemiology is that baseline cancer mortality rates in vital statistics registries represent precisely the risk for each quinquennium. Percy et al. give clear evidence to support the contention that this assumption does not hold (Percy et al., 1981; Percy et al., 1990). Thus, in an effort to correct for death certificate misclassification in the national cancer rates, we introduce the Death Cerlificate Correction Factor (DCCF), defined by

$$
\begin{equation*}
D C C F_{\mathrm{e}}=\frac{\theta_{\mathrm{c}}}{\phi_{\mathrm{c}}} \tag{22}
\end{equation*}
$$

where $\theta_{c}$ is the cancer confirmation rate and $\phi_{e}$ is the cancer detection rate (sensitivity) defined in Tables 1 and 2. The confirmation and detection rates are given for each site in the latest Percy et al. (1990) paper and were used to modify the baseline cancer rates used in lifetime risk projection described above. The DCCF has the unique property of increasing rates that are underreported aid decreasing rates that are overreported.

For the reader who is interested, see $\S 8$ "Notation" and $\$ 9$ "Abbreviations."

## 4 RESULTS

### 4.1 Per Cent Distribution of True Cancer Deaths

For most age ATB and ATD categories, a short minimal latency period was observed for leukemia because a majority of deaths occurred less than 10 years following exposure (data not shown). For older age ATB and ATD categories the shifting of leukemia deaths misclassified as non-cancer deaths increases because 1) the age-specific mortality rate of all deaths less leukemia outweighs the age-specific leukemia rates at all ages and because 2) the number of true cancer deaths is equal to the sum of the product of the observed cancer deaths and the probability that the observed cancer deaths are correctly classified and the product of the observed non-cancer deaths and one minus the the probability that the observed non-cancer deaths were correctly classified. This relationship will hold uniformly with increasing age ATD as long as the confirmation rates are not stratified by age ATD. For solid cancers, most deaths occurred at older age ATB and ATD levels and a visible minimal latency period was apparent (data not shown), because most deaths occurred greater than 10 years post-exposure.

### 4.2 Poisson Regression

### 4.2.1 Models with Age ATB, Age ATD and Gender

Regression models containing covariates for age ATB, age ATD and gender were used for leukemia and nonleukemia cancer beccuse of the guaranteed convergence at a global maximum, low scores and low values of the $\chi^{2}, D$ and $G$ goodness-of-fit (GOF) statistics. The modeling results in this section were, in general, in good agreement with with those reported in Table 6 of RERF Report 11, Part 2 (Shimizu et al., 1988). The only difference between the regression results of this study and those reported in Table 6 of RERF Report 11 (Part 2), was that in this study organ dose equivalents were calculated before performing regressions runs, whereas in Report 11, shielded kerma was used for dose.

Leukemia Tables C.1-C.11 of Appendix C list the ERR and AR coefficients for leukemia for various methods of adjustment for diagnostic misclassification without adjustment for DS86 random error. The GOF statistics for all of the models indicated that the model results were consistent with the observed data.

When no adjustment for diagnostic misclassification was made (Table C.1), the regression coefficient for dose ( $\% / \mathrm{Sv}$ ) was 42.04 and the $\chi^{2}, D$, and $G$ statistics (d.f. $=3022$ ) were 1262,632 and 238 ; however, when DS86-specific confirmation rates were employed (Table C.11) to estimate the number of cell-specific true cancer deaths, the regression coefficient for dose decreased by $18.4 \%(34.32)$ and $\chi^{2}$ and $D$ dropped to 888 and 508 , but the $G$ increased slightly to 268 . This reduction in GOF statistics indicates that the application DS86-specific confirmation rates for follow-up periods $1950-85$ resulted in a model that fitted better than the model in which no adjustments were made.

When marrow dose equivalents were adjusted for DS86 random error (Tables C.12C.22), ERR and AR coefficients increased in all age ATB and ATD categories. When diagnostic misclassification was not adjusted (Table C.12), the regression coefficient for dose was $6.0 \%$ higher ( 45.64 ) and $\chi^{2}, D$, and $G$ statistics were lower ( 1338,635 and 252) in comparison with the same model when DS86 random error was not adjusted. When DS86-specific confirmation rates were employed (Table C.22), the regression coefficient decreased by $18.1 \%$ and $\chi^{2}$ and $D$ dropped to 921 and 507 , but the $G$ increased slightly

Nonleukemia The ERR and AR for nonleukemia results when no adjustment for DS86 random error was made are listed in Tables C. 23 -C. 33 of Appendix C. The GOF statistics for all of the models indicated that the model results were concordant with the observed data. The regression coefficient for dose and $\chi^{2}, D$, and $G$ statistics $(\mathrm{d} . \mathrm{f} .=3022)$ when no adjustments for diagnostic misclassification were made (Table C.23) were $5.38,4636,2159$ and 1909 , respectively. However, when DS86-specific confirmation rates were employed (Table C.33) to estimate the number of cell-specific true cancer deaths, the regression coefficient changed to $3.55\left(-34 \%\right.$ reduction) and the $\chi^{2}, D$ and $G$ dropped to 2608,1585 , and 1816 , which indicated that GOF increased when DS86-specific contirmation rates were applied.

When a correction for DS86 random error (Table C.34) was made for the colon (large intestine) dose equivalent, the regression coefficient for dose increased by $2 \%$ (5.49) and GOF $\chi^{2}, D$, and $G$ statistics were were 4619,2159 and 1905 , when no adjustment for diagnostic misclassification was made. When DS86-specific confirmation rates were employed (Table C.44), the dose regression coefficient decreased by $30.4 \%$ (3.82) and $\chi^{2}, D$ and $G$ dropped to 2610,1582 and 1815 , which were essentially the same as the GOF statistics for the model in Table C. 33 , that is where DS86-specific confirmations were used, but no random error adjustments were made.

### 4.2.2 BEIR-V Models

The non-fully-parametric BEIR-V models included no more than 6 coefficients representing age ATB, time since exposure, and gender and therefore converged at a global maximum rather quickly with reliable goodness-of-fit statistics.

Leukemia Tables C. 45 and C. 46 list the ERR coefficients for leukemia. Regression coefficients (not in tables) and GOF statistics (d.f. $=2404$ ) when no adjustments were made for diagnostic misclassification or DS86 random error were similar to those in the BEIR-V report $\left(\alpha_{1}=0.28, \alpha_{2}=0.14, \beta_{1}=4.88, \beta_{2}=2.40, \beta_{3}=2.37, \beta_{4}=1.63, \chi^{2}=634, D=397\right.$, and $G=194)$. When adjusting for diagnostic misclassification using DS86-specific confirmation rates applied over the years $1950-85$, the linear dose coefficient increased substantially, however the remaining coefficients decreased ( $\alpha_{1}=0.72, \alpha_{2}=0.13, \beta_{1}=4.03, \beta_{2}=1.77, \beta_{3}=1.84$, $\beta_{4}=1.27, \chi^{2}=491, D=322$, and $G=223$ ).

When marrow dose equivalents were corrected for DS86 random error, the linear dose coefficient increased by $64.8 \%$ and the linear-quadratic term increased by $97.9 \%$ and the GOF statistics did not improve ( $\alpha_{1}=0.46, \alpha_{2}=0.28, \beta_{1}=4.42, \beta_{2}=1.96, \beta_{3}=1.83, \beta_{4}=1.20$, $\chi^{2}=718, D=412$, and $G=224$ ). The correction for both diagnostic misclassification using DS86-specific confirmation rates and DS86 random error resulted in a $113.4 \%$ increase in the linear dose coefficient, but the remaining coefficients were decreased and the GOF statistics decreased slightly ( $\alpha_{1}=0.97, \alpha_{2}=0.21, \beta_{1}=3.78, \beta_{2}=1.54, \beta_{3}=1.56, \beta_{4}=1.05$, $\chi^{2}=543, D=333$, and $G=218$ ).

Digestive System Tables C. 47 and C. 48 list the ERR coefficients for digestive system cancers. When no adjustments were made for diagnostic misclassification or DS86 random error the regression coefficients (not in tables) and GOF statistics (d.f. $=1910$ ) were identical to those in the BEIR-V report ( $\alpha_{1}=0.8068, \beta_{1}=0.5558, \beta_{2}=-0.1976, \chi^{2}=2159, D=1192$,
and $G=1039$ ). When adjusting for diagnostic misclassification using sex-city-specific confirmation rates applied over the years 1950-75, the linear dose coefficient decreased $9.9 \%$, however the other log-linear coefficients increased and the GOF statistics decreased moderately ( $\alpha_{1}=0.7267, \beta_{1}=0.604, \beta_{2}=-0.1861, \chi^{2}=1591, D=920$, and $G=932$ ).

When the stomach dose equivalents (stomach transmission factor were used for the digestive site) were corrected for DS86 random error, the linear dose coefficient decreased by $8.1 \%$ and the GOF statistics increased slightly $\left(\alpha_{1}=0.7356, \beta_{1}=0.6698, \beta_{2}=-0.1762\right.$, $\chi^{2}=2272, D=1245$, and $G=1052$ ). The correction for both diagnostic misclassification using sex-city-specific confirmation rates, and DS86 zandom error resulted in $15.6 \%$ decrease in the linear dose coefficient and a reduction of the GOF statistics $\left(\alpha_{1}=0.6204, \beta_{1}=0.7422\right.$, $\beta_{2}=-0.1631, \chi^{2}=1678, D=962$, and $G=944$ ).

### 4.3 Lifetime Risk Projection

### 4.3.1 Lifetime Risks Without Adjustments

Tables 9-11 list the site- and sex-specific lifetime risks (\%/Sv) based on the absolute, transported relative and relative projection models. The trends of excess risks of leukemia as a function age at exposure were similar for the absolute and transported relative models in Tables 9 and 10. Appendix D provides tables of lifetime risks for the 18-65 age at exposure profile for all results in this section.

In Table 11, using a Dose-rate Reduction Effectiveness Factor (DRREF) of two and no correction for DS86 random error or diagnostic misclassification in the non-constant relative projection model, lifetime risks (\%/Sv) of nonleukemia among males exposed acutely to 1 Sv at age $25,45,65$ or continuously from ages 18 to 65 were $2.10 \%, 2.78 \%$, $1.20 \%$ and $1.91 \%$. For females, nonleukemia lifetime risks for the same exposure profiles were $3.49 \%, 4.32 \%, 1.97 \%$ and $3.23 \%$. Excess leukemia risks for 1 Sv at 25,45 and 65 and over the years 18 to 65 were $0.35 \%, 0.46 \%, 2.46 \%$ and $0.87 \%$ for males and $0.26 \%, 0.41 \%$, $1.96 \%$ and $0.73 \%$ for females. By way of comparison, excess nonleukemia risks based on the constant projection models were $2.84 \%$ for males and $4.75 \%$ for females; risks of leukemia among males was $0.75 \%$ and among females was $0.64 \%$. Thus, lifetime risk estimates based on constant models did not underestimate risks projected by non-constant models.

The results in Tables 9-11 are in very close agreement with lifetime risks used in an ICRP analysis (Land and Sinclair, 1991) and in most cases only differed by several cancers per 100,000 . Small differences were noted with the transported relative and constant relative models which were attributable to 1) use of different baseline rates [our baseline rates were for the 1885 epoch, Land and Sinclair's were for the years 1973-77] and 2) a small variation in the estimation of hazard function for the transported relative risk model (Land, 1989). The negligible differences in absolute risks between the present study and those of Land and Sinclair supports the tentative use of projected all-cause vital statistics (Faber and Wade, 1983) in this study, for the study of birth-cohort effects on lifetime risk projection (Peterson et al., 1992) and projection of lifetime risks for the Hanford cohort (Peterson et al., 1993).

Since Sposto et al. (1992) did not project lifetime risks for various exposure profiles, we used their regression coefficients and the SURVRAD progrem to generate lifetime risks. In Table 11, one notices that excess nonleukemia risks based on the Sposto et al. data for exposure over a career ( 18 to 65 y ) were $156.4 \%$ greater for males and $83.3 \%$ greater for females when compared with nonleukemia results of our analysis.

Excess risks for ages at exposure 25,4565 and 18 -65 for the BEIR-V model are listed in Table 12. Results were in good agreement with lifetime risks reported by the BEIR-V

Table 9: Site- and sex-specific excess risks (\%/Sv) for the absolute projection model without adjustment for diagnostic misclassification and DS86 random error (DRREF=2).

| Site |  |  |  |
| :--- | :--- | :--- | :---: |
|  | Age at exposure ${ }^{\text {a }}$ | Excess risk (\%/Sv) |  |
|  | 25 | 0.84 | Males |
|  | 45 | 0.42 | 0.34 |
|  | 65 | 0.44 | 0.45 |
|  | $18-65$ | 0.67 | 0.40 |
| Nonleukemia | 25 | 2.24 | 2.95 |
|  | 45 | 2.86 | 3.78 |
|  | 65 | 0.79 | 1.44 |
|  | $18-65$ | 1.93 | 2.77 |

${ }^{a}$ For exposure at ages 25,45 and 65 , the population of 100,000 was acutely exposed to 1 Sv ; the exposure from age 18 to 65 involved chronic exposure to $0.02128 \mathrm{~Sv} / \mathrm{y}$ for 47 y (total $=1 \mathrm{~Sv}$ ).

Table 10: Site- and sex-specific excess risks (\%/Sv) for the transported relative projection model without adjustment for diagnostic misclassification and DS86 random error ( $\mathrm{DRREF}=2$ ).

| $=$ |  |  |  |
| :--- | :--- | :---: | :---: |
| Site |  | Excess risk $(\% /$ Sv $)$ |  |
| Leukemia | Age at exposure ${ }^{a}$ | Males | Females |
|  | 25 | 0.06 | 0.04 |
|  | 45 | 0.34 | 0.24 |
|  | 65 | 0.44 | 0.48 |
|  | $18-65$ | 0.35 | 0.27 |
| Nonleukemia | 25 | 1.89 | 2.29 |
|  | 45 | 2.79 | 3.66 |
|  | 65 | 0.52 | 0.91 |
|  | $18-65$ | 1.74 | 2.38 |

${ }^{\circ}$ For exposure at ages 25,45 and 65 , the population of 100,000 was acutely exposed to 1 Sv ; the exposure from age 18 to 65 involved chronic exposure to $0.02128 \mathrm{~Sv} / \mathrm{y}$ for 47 y (total $=1 \mathrm{~Sv}$ ).

Table 11: Site- and sex-specific excess risks ( $\%$ 'Sv) for the relative projection model without adjustment for diagnostic misclassification and DS86 random error (DRREF=2).

|  |  |  |  |
| :--- | :--- | :---: | :---: |
| Site | Age at exposure | Excess risk $(\% / \mathrm{Sv})$ |  |
|  | Males | Females |  |
|  | 25 | 0.35 | 0.26 |
|  | 45 | 0.46 | 0.41 |
|  | 65 | 2.46 | 1.96 |
| Nonleukemia | $18-65$ | 0.87 | 0.73 |
|  | 25 | 2.10 | 3.49 |
|  | 45 | 2.78 | 4.32 |
|  | 65 | 1.20 | 1.97 |
| Nonleukemia | $18-65$ | 1.91 | 3.23 |
| (Sposto. et al.) | 45 | 2.61 | 5.02 |
|  | 65 | 5.72 | 6.49 |
|  | $18-65$ | 2.39 | 2.46 |
|  |  | 4.90 | 5.92 |

> a For exposure at ages 25,45 and 65 , the population of 100,000 was acutely exposed to 1 Sv ; the exposure from age 18 to 65 involved chronic exposure to $0.02128 \mathrm{~Sv} / \mathrm{y}$ for 47 y (total $=1 \mathrm{~Sv}$ ).
committee (NRC, 1990). The leukemia risks for all ages at exposure were similar to those of the relative projection model in Table 11, in that they increased as age at exposure increased and were the greatest at age 65 ( 1.46 and $1.14 \% / \mathrm{Sv}$ for males and females). Solid cancers, such as the digestive system also had excess risks that closely resembled risks for the relative model listed in Table 11.

### 4.3.2 Bias in Absolute Projection Models

Lifetime risks for males and females based on the absolute projection model that were negatively biased are listed in Tables 13 and 14. The most negative bias due to diagnostic misclassification was indicated by the liver site (range -68 to -521\%). While diagnostic misclassification decreased with increasing age at exposure to negative values less than $-50 \%$, bias due to DS86 random error remained above $-30 \%$ and was relatively stable over varying levels of age at exposure. Another interesting trend that was noted was that exposure over a career (ages 18 to 65) usually led to a total bias that was greater (more positive) than $-50 \%$. In addition, when DS86-specific confirmation rates for a particular site were available, their use usually resulted in a bias for diagnostic misclassification that was lower than the other covariates on which confirmation rates were stratified.

### 4.3.3 Bias in Transported Relative Projection Models

Tables 15 and 16 list bias of excess risk for the transported relative model that were negative. A similar picture emerged with the transported model when comparing results with the purely absolute model in Tables 13 and 14. Overall, there was a tendency for

Table 12: Site- and sex-specific excess risks (\%/Sv) for the BEIR-V relative projection model without adjustment for diagnostic misclassification and DS86 random error ( $\mathrm{DRREF}=2$ ).

| Site |  |  | Age at exposure |  | Excess risk $(\% /$ Sv) |  |
| :--- | :--- | :---: | :---: | :---: | :---: | :---: |
|  | Males | Females |  |  |  |  |
|  | 25 | 0.35 | 0.27 |  |  |  |
|  | 45 | 0.99 | 0.76 |  |  |  |
|  | 65 | 1.46 | 1.14 |  |  |  |
|  | $18-65$ | 0.53 | 0.43 |  |  |  |
| Digestive | 25 | 2.06 | 3.36 |  |  |  |
|  | 45 | 0.36 | 0.59 |  |  |  |
|  | 65 | 0.30 | 0.50 |  |  |  |
|  | $18-65$ | 0.77 | 1.30 |  |  |  |

${ }^{a}$ For exposure at ages 25,45 and 65 , the population of 100,000 was acutely exposed to 1 Sv ; the exposure from age 18 to 65 involved chronic exposure to $0.02128 \mathrm{~Sv} / \mathrm{y}$ for 47 y (total $=1 \mathrm{~Sv}$ ).

Table 13: Negative bias of excess risk (\%/Sv) among males for the absolute projection model (DRREF=2).

| Site | Age at exposure | $\begin{gathered} \text { Risk } \\ (\% / \mathrm{Sv}) \end{gathered}$ | Bias |  |  | Strata of $\theta_{c}$ and $\theta_{n c}^{a}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Diag. misc. | DS86 random error | Total |  |
| Leukemia | 25 | 0.96 | 0.0 | -14.8 | -14.8 | N/A |
|  | 45 | 0.72 | -48.1 | -13.6 | -61.7 | DS86(1950-85) |
|  | 65 | 0.69 | -37.1 | -12.8 | -50.0 | DS86(1950-85) |
|  | 18-65 | 0.85 | -11.6 | -13.6 | -25.2 | DS86(1950-85) |
| Nonleukemia | 25 | 3.36 | -18.4 | -26.6 | -45.0 | DS86(1950-85) |
|  | 45 | 4.66 | -32.1 | -23.6 | -55.7 | DS88(1950-85) |
|  | 65 | 1.97 | -99.8 | -24.3 | -124.1 | DS86(1950-85) |
|  | 18-65 | 3.23 | -35.5 | -23.6 | -59.1 | DS86(1950-85) |

[^9]Table 14: Negative bias of excess risk (\%/Sv) among females for the absolute projection model (DRREF=2).

| Site | Age at exposure | $\begin{aligned} & \text { Risk } \\ & (\% / \mathrm{Sv}) \end{aligned}$ | Bias |  |  | Strata of $\theta_{c}$ and $\theta_{n e}^{a}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Diag. misc. | $\begin{gathered} \text { DS86 } \\ \text { random } \\ \text { error } \end{gathered}$ | Total |  |
| Leukemia | 25 | 0.40 | -0.8 | -12.8 | -13.6 | DS86(1950-85) |
|  | 45 | 0.61 | -59.3 | -10.9 | -70.2 | DS86(1950-85) |
|  | 65 | 0.69 | -42.3 | -8.8 | -51.1 | DS86(1950-85) |
|  | 18-65 | 0.59 | -34.2 | -10.3 | -44.5 | DS86(1950-85) |
| Nonleukemia | 25 | 3.66 | -5.2 | -18.2 | -23.4 | DS86(1950-85) |
|  | 45 | 5.77 | -31.6 | -15.7 | -47.3 | DS86(1950-85) |
|  | 65 | 3.33 | -99.3 | -16.1 | -115.4 | DS86(1950-85) |
|  | 18-65 | 4.28 | -33.7 | -15.6 | -49.3 | DS86(1950-85) |

${ }^{0}$ itrata of confirmation rates for which correction of diagnostic misclassification resulted in negative bias. (1950-85) denotes that confirmation rates were applied only to deaths which occurred during 1950-85.
diagnostic misclassification and DS86 random error to be the same with respect to the absolute projection model.

### 4.3.4 Bias in Relative Projection Models

There were fewer sites and exposure categories for which bias was negative under the relative projection model (Tables 17 and 18). A particularly interesting finding was that in most cases the bias was more positive and less erratic than bias for the absolute and transported relative models. Among males (Table 17), diagnostic misclassification bias for leukemia and nonleukemia for a career exposure was $-23.9 \%$ and $-11.3 \%$ and for DS86 random error was $-14.0 \%$ and $-23.7 \%$. Females (Table 18) had a bias of $-42.8 \%$ and $-6.3 \%$ for diagnostic misclassification of leukernia and nonleukemia. Bias due to DS86 random error for female leukemia and nonleukemia was $-11.4 \%$ and $-14.9 \%$ for exposure over a career.

In comparison, the bias due to diagnostic misclassification in males and females for lifetime risks based on the Sposto et al. anayisis for exposure over a career ( 18 to 65 y ) was $-10.0 \%$ and $2.5 \%$. The adjustrnent of cancer misclassification in U.S. cancer rates used for risk projection resulted in a bias of $11 \%$ for leukemia and $-2 \%$ for nonleukemia.

The total bias for leukemia and nonleukemia among males exposed over a career was $-27.1 \%$ and $-37.1 \%$ and resulted in changes of excess risk (\%/Sv) from 0.87 to 1.1 and 1.91 to 2.68 . Females had a total bias of $-43.4 \%$ and $-23.3 \%$ for leukemia and nonleukemia which led to changes in excess risk (\%/Sv) of 0.73 to 1.04 and 3.23 to 4.02 .

Figures 8 and 9 illustrate schematically, for males and females, the conditional probabilities, $\pi(t ; d)$, (see Eq. 68) of radiation-induced nonleukemia based on non-constant relative projections for this investigation and results based on projections using the Sposto et al. (1992) regression coefficients. Figure 8 shows that, for males exposed to 1 Sv at age 25, the difference between $\pi(t ; d)$ when a $22 \%$ correction for diagnostic misclassification was made and $\pi(t ; d)$ when no correction was made for the Sposto et al. data is smaller than

Table 15: Negative bias of excess risk (\%/Sv) among males for the transported relative projection model (DRREF=2).

| Site | Age at exposure | $\begin{gathered} \text { Risk } \\ (\% / \mathrm{Sv}) \end{gathered}$ | Bias |  |  | Strata of $\theta_{c}$ and $\theta_{\text {ne }}^{a}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Diag. misc. | $\begin{aligned} & \text { DS86 } \\ & \text { random } \\ & \text { error } \\ & \hline \end{aligned}$ | Total |  |
| Leukemia | 25 | 0.08 | -17.5 | -15.3 | -32.8 | DS86(1950-85) |
|  | 45 | 0.60 | -54.6 | -13.4 | -68.0 | DS86(1950-85) |
|  | 65 | 0.65 | -32.1 | -12.3 | -44.4 | DS86(1950-85) |
|  | 18.65 | 0.54 | -36.1 | -12.8 | -48.9 | DS86(1950-85) |
| Nonleukemia | 25 | 2.80 | -17.1 | -26.2 | -43.3 | DS86(1950-85) |
|  | 45 | 4.50 | -31.5 | -22.3 | -53.8 | DS86(1950-85) |
|  | 65 | 1.36 | -108.9 | -26.15 | -135.1 | DS86(1950-85) |
|  | 18-65 | 2.90 | -35.7 | -22.8 | $-58.5$ | DS86(1950-85) |

${ }^{\text {a }}$ Strata of confirmation rates for which correction of diagnostic misclassification resulted in negative bias. (1950-85) denotes that confirmation rates were applied only to deaths which occurred during 1950-85.

Table 16: Negative bias of excess risk (\%/Sv) among females for the transported relative projection model (DRREF $=2$ ).

| Site | A.ge at exposure | $\begin{aligned} & \text { Risk } \\ & (\% / S v) \end{aligned}$ | Bias |  |  | Strata of $\theta_{c}$ and $\theta_{n \text { e }}^{a}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Diag. misc. | $\begin{gathered} \text { DS86 } \\ \text { random } \\ \text { error } \end{gathered}$ | Total |  |
| Leukemia | 25 | 0.06 | -42.7 | -13.7 | -56.4 | DS86(1950-85) |
|  | 45 | 0.46 | -73.7 | -10.9 | -84.6 | DS86(1950-85) |
|  | 65 | 0.71 | -34.7 | -8.7 | -43.4 | DS86(1950-85) |
|  | 18-65 | 0.43 | -48.3 | -9.6 | -57.9 | DS86(1950-85) |
| Nonleukemia | 25 | 2.91 | $-7.7$ | -18.4 | -26.1 | DS86(1950-85) |
|  | 45 | 5.47 | -30.1 | -14.8 | -44.9 | DS86(1950-85) |
|  | 65 | 2.11 | -98.8 | -17.1 | -115.9 | DS86(1950-85) |
|  | 18-65 | 3.69 | -34.1 | 15.1 | -49.2 | DS86(1950-85) |

${ }^{a}$ Strata of confirmation rates ior which correction of diagnostic misclassification resulted in negative bias. (1950-85) denotes that confirmation rates were applied only to deaths which occurred during 1950-85.
the difference when an adjustment was made with the DS86-specific confirmation rates. The same was true for exposure at ages 45 and 65 and for the continuous exposure ( 18 to 65). At 65 years of age, an acute exposure to 1 Sv seemed to cause $\pi(t ; d)$ to increase rapidly with increasing attained age. This finding may be in accord with a suggestion by Moolgavkar and Knudson (1981) that relative risk is highest at older ages at exposure because the number of premalignant clones in the body increases with attained age. One also notes the striking similarity in the shapes of the curves for the Sposto et al. data and this study. The height of the curves for $\pi(t ; d)$ based on the Sposto et al. data was higher than $\pi(t ; d)$ for this study because a neutron RBE of unity was used (this study used a neutron RBE of 10). For fernales (Figure 9), the $22 \%$ correction for diagnostic misclassification made in the Sposto et al. analysis always resulted in a corrected $\pi(t ; d)$ that was lower than the $\pi(t ; d)$ when no correction was made because the regression coefficient for females in the uncorrected model was reduced from 0.356 to 0.315 after correction. In addition, the patterns of $\pi(t ; d)$ for all exposure profiles in Figure 9 indicate that neutron RBE had a lower impact on risk of nonleukemia among females.

### 4.3.5 Bias in BEIR-V Relative Projection Models

Tables 19 and 20 list the sites and exposure profiles for which bias from diagnostic misclassification and DS86 random errors in the BEIR-V models were negative. For males (Table 19), bias due to DS86 random errors were more positive than in the relative models with values typically above $-10 \%$. Diagnostic misclassification bias for leukemia for males was more negative when compared to the relative models (Tables 17) and was more positive for fernales when compared with the relative model (Table 18). While the the bias due to DS86 random error for leukemia ( $1 \mathrm{~Sv} 18-65$ ) among males and females were $-3.7 \%$ and $-3.9 \%$, the same bias was $-14.0 \%$ and $-11.4 \%$ in the relative models. One also notices in Tables 19 and 20 that, for the digestive site, there were only two exposure profiles ( 1 Sv acute at ages 45 and 65 ) for which correction of diagnostic misclassification was negatively biased; however, the magnitude of the bias is negligible.


Figure 8: Conditional probabilities of radiation-induced nonleukemia among males for this study and results based on the Sposto et al. regression coefficients. ( $x$-no adjustment in Sposto et al.; - $22 \%$ adjustment in Sposto et al. ; no adjustment in this study; - DS86-specific adjustment in this study).


Figure 9: Conditional probabilities of radiation-induced nonleukemia among females for this study and results based on the Sposto et al. regression coefficients. ( X -no adjustment in Sposto et al.; - $-22 \%$ adjustment in Sposto et al.; $\mathbf{\Delta}$ - no adjustment in this study; - DS86-specific adjustment in this study).

Table 17: Negative bias of excess risk (\%/Sv) among males for the relative projection model (DRREF $=2$ ).

| Site | Age at exposure | $\begin{aligned} & \text { Risk } \\ & (\% / \mathrm{Sv}) \end{aligned}$ | Biss |  |  |  | Strata of $\theta_{c}$ and $\theta_{n}^{a}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Diag. misc. | DS86 random error | DCCF | Total |  |
| Leukemia | 25 | 0.39 | -6.0 | $-17.4$ | 10.8 | -12.6 | Sex(1950-85) |
|  | 45 | 0.79 | -69.7 | -14.2 | 10.8 | -73.1 | DS86(1950-85) |
|  | 65 | 3.08 | -23.3 | -13.3 | 10.5 | -26.1 | DS86(1950-85) |
|  | 18-65 | 1.10 | -23.9 | -14.0 | 10.8 | -27.1 | DS86(1950-85) |
| Nonleukemia | 25 | 2.90 | -5.0 | -27.1 | -2.1 | -34.2 | DS86(1950-85) |
|  | 45 | 3.76 | -7.3 | -23.5 | -2.1 | -32.9 | DS86(1950-85) |
|  | 65 | 2.12 | -38.9 | -24.6 | -2.1 | -65.6 | DS86(1250-85) |
|  | 18-65 | 2.68 | -11.3 | -23.7 | -2.1 | -37.1 | DS86(1950-85) |
| Nonleukemia (Sposto et al.) | 25 | 2.64 | -1.0 | . | , | -1.0 | N/A-EM algorithm |
|  | 45 | 5.95 | -4.0 | - | - | -4.0 | N/A-EM algorithm |
|  | 65 | 2.36 | 1.3 | - | - | 1.3 | N/A-EM algorithm |
|  | 18-65 | 5.39 | -10.0 | - | $-$ | -10.0 | N/A-EM algorithm |

[^10]Table 18: Negative bias of excess risk ( $\% / \mathrm{Sv}$ ) among females for the relative projection model (DRREF $=2$ ).

| Site | Bias |  |  |  |  |  | Strata of $\theta_{c}$ and $\theta_{n c}^{a}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Age at exposure | $\begin{gathered} \text { Risk } \\ (\% / \mathrm{Sv}) \\ \hline \end{gathered}$ | Diag. <br> misc. | DS86 random error | DCCF | Total |  |
| Leukemia | 25 | 0.31 | -12.1 | -15.1 | 10.9 | -16.3 | DS86(1950-85) |
|  | 45 | 0.75 | -83.4 | -11.6 | 10.8 | -84.2 | DS86(1950-85) |
|  | 65 | 2.75 | -40.9 | -10.9 | 10.6 | -41.? | DS86(1950-85) |
|  | 18-65 | 1.04 | -42.8 | -11.4 | 10.8 | -43.4 | DS86(1950-85) |
| Nonleukemia | 25 | 4.33 | -1.5 | -19.6 | -2.1 | -23.2 | DS86(1950-85) |
|  | 45 | 5.22 | -3.9 | -13.9 | -2.0 | -19.8 | DS86(1950-85) |
|  | 65 | 2.94 | -26.2 | -15.9 | -2.1 | -44.2 | DS86(1950-85) |
|  | 18-65 | 4.02 | -6.3 | -14.9 | -2.1 | -23.3 | DS86(1950-75) |
| Nonleukemia (Sposto et al.) | 25 | 2.62 | 13.3 | - | - | 13.3 | N/A-EM algorithm |
|  | 45 | 6.28 | 3.2 | - | - | 3.2 | N/A-EM algorithm |
|  | 65 | 2.32 | 5.2 | - | - | 5.2 | N/A-EM algorithm |
|  | 18-65 | 5.77 | 2.5 | - | - | 2.5 | N/A-EM algorithm |

${ }^{a}$ Strata of confirmation rates for which correction of diagnostic misclassification resulted in negative bias. N/A-EM algorithm denotes that the diagnostic misclassification was estimated from the Sposto et al. (1992) Poisson regression coefficients obtained with the EM algorithm. (1950-75) denotes that confirmation rates were applied only to deaths that occurred during 1950-75. Likewise, (1950-85) denotes that confirmation rates were applied only to deaths which occurred during 1950-85.

Table 19: Negative bias of excess risk (\%/Sv) among males for the BEIR-V projection model (DRREF=2).

| Site | Age at exposure | $\begin{aligned} & \text { Risk } \\ & (\% / \mathrm{Sv}) \end{aligned}$ | Bias |  |  |  | Strata of $\theta_{c}$ and $\theta_{n c}^{a}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Diag. misc. | $\begin{gathered} \text { DS86 } \\ \text { random } \\ \text { error } \end{gathered}$ | DCCF | Total |  |
| Leukernia | 25 | 0.44 | -32.2 | -8.4 | 10.8 | -29.8 | DS86(1950-85) |
|  | 45 | 1.20 | -28.8 | -7.3 | 10.8 | -25.3 | DS86(1950.85) |
|  | 65 | 1.64 | -20.4 | -4.3 | 10.7 | -14.0 | DS86(1950-85) |
|  | 18-65 | 0.79 | -59.4 | $-3.7$ | 10.8 | -52.3 | DS86(1950-85) |
| Digestive | 25 | 1.95 | 0.0 | 6.6 | -1.5 | 5.1 | N/A |
|  | 45 | 0.44 | -1.4 | -17.8 | -1.6 | -22.5 | Crude(1950-75) |
|  | 65 | 0.37 | -1.0 | -19.5 | -1.6 | -20.4 | Sex-city(1950-75) |
|  | 18-65 | 0.77 | 0.0 | 1.4 | -1.6 | -0.2 | N/A |

[^11]Table 20: Negative bias of excess risk (\%/Sv) among females for the BEIR-V projection model (DRREF=2).

| Site | Age at exposure | $\begin{gathered} \text { Risk } \\ (\% / \mathrm{Sv}) \end{gathered}$ | Bias |  |  |  | Strata of <br> $\theta_{c}$ and $\theta_{n e}^{a}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Diag. misc. | DS86 tandom error | DCCF | Total |  |
| Leukemia | 25 | 0.34 | -31.8 | -8.3 | 10.9 | -29.2 | DS86(1950-85) |
|  | 45 | 1.0 | -29.9 | -7.7 | 10.8 | -26.8 | DS86(1950-85) |
|  | 65 | 1.3 | -21.2 | -0.5 | 10.8 | -10.9 | DS86(1950-85) |
|  | 18-65 | 0.6 | -60.5 | -3.9 | 10.8 | -53.6 | DS86(1950-85) |
| Digestive | 25 | 3.55 | 0.0 | -4.18 | -1.53 | -5.71 | N/A |
|  | 45 | 0.80 | -6.2 | -26.5 | -1.6 | -28.5 | Sex(1950-75) |
|  | 65 | 0.67 | -6.2 | -26.6 | -1.6 | -28.6 | Sex(1950-75) |
|  | 18-65 | 1.46 | 0.0 | -10.5 | -1.57 | -12.1 | N/A |

[^12]
## 5 DISCUSSION

### 5.1 Confirmation Rates

The use of cancer and non-cancer confirmation rates or the predictive value positive to estimate the true number of cancer deaths in the LSS confirmed the impression that correction of non-differential misclassification does not alway lead to a bias that is toward the null (Green, 1983). Bias due to random error in the DS86, however, was always negative, the correction of which produced increased ERR and AR coefficients.

A major influence on the validity of correction methods for diagnostic misclassification in the LSS is the large number of non-cancer deaths that have occurred (and are occurring) at the higher age ATD groups and middle-aged age ATB groups. If precautions are not taken when shifting the small number of misclassified cancer deaths from the large number of non-cancer deaths into the presumed correct cells, then invalid results may be obtained. We did not use the EM algorithm (Dempster et al., 1977) for imputing the true cancer deaths in cells for which no autopsy data existed, rather we employed the sufficient statistics that are used before the first iteration of the EM algorithm and applied the results to all cells after a minimal latency period of 2 years for leukemia and 10 years for solid cancers.

Table 21: Weighted per cent distribution of non-cancer deaths among both sexes in both cities in the Life Span Study (1950-85).

|  | Age ATD |  |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |  |  |
| $<10$ | 0.88 | 0.66 | 0.76 | 0.33 | 0.00 | 0.00 | 0.00 |  |  |
| $10-19$ | 0.24 | 1.63 | 1.13 | 1.45 | 0.73 | 0.00 | 0.00 |  |  |
| $20-29$ | 0.00 | 0.45 | 1.05 | 1.03 | 1.56 | 0.77 | 0.00 |  |  |
| $30-39$ | 0.00 | 0.00 | 0.42 | 1.65 | 2.80 | 4.86 | 2.30 |  |  |
| $40-49$ | 0.00 | 0.00 | 0.00 | 0.95 | 4.80 | 7.59 | 15.00 |  |  |
| $50+$ | 0.00 | 0.00 | 0.00 | 0.00 | 1.50 | 7.63 | 37.81 |  |  |

Table 22: Weighted per cent distribution of non-cancer deaths among males in both cities in the Life Span Study (1950-85).

|  | Age ATD |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |  |
| $<10$ | 1.19 | 0.88 | 1.15 | 0.54 | 0.00 | 0.00 | 0.00 |  |
| $10-19$ | 0.28 | 2.00 | 1.49 | 2.04 | 0.98 | 0.00 | 0.00 |  |
| $20-29$ | 0.00 | 0.35 | 0.67 | 0.88 | 1.30 | 0.90 | 0.00 |  |
| $30-39$ | 0.00 | 0.00 | 0.34 | 1.66 | 3.14 | 4.35 | 1.87 |  |
| $40-49$ | 0.00 | 0.00 | 0.00 | 1.04 | 5.70 | 8.55 | 13.63 |  |
| $50+$ | 0.00 | 0.00 | 0.00 | 0.00 | 1.97 | 9.64 | 33.46 |  |

### 5.2 Regression Methods

The maximum likelihood results we obtained indicated that, i, some cases, and in some cross-classifications of the data, there were indeed locations or, the likelihood surface where

Table 23: Weighted per cent distribution of non-cancer deaths among females in both cities in the Life Span Study (1950-85).

|  | Age ATD |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |  |
| $<10$ | 0.62 | 0.47 | 0.43 | 0.16 | 0.00 | 0.00 | 0.00 |  |
| $10-19$ | 0.21 | 1.33 | 0.83 | 0.96 | 0.52 | 0.00 | 0.00 |  |
| $20-29$ | 0.00 | 0.53 | 1.37 | 1.17 | 1.78 | 0.65 | 0.00 |  |
| $30-39$ | 0.00 | 0.00 | 0.49 | 1.64 | 2.51 | 5.30 | 2.66 |  |
| $40-49$ | 0.00 | 0.00 | 0.00 | 0.87 | 4.03 | 6.78 | 16.17 |  |
| $50+$ | 0.00 | 0.00 | 0.00 | 0.00 | 1.09 | 5.93 | 41.50 |  |

incongruities exist. For example, the stomach model had to be fit for each sex since models that contained a parameter for gender either 1) did not converge after 100 iterations; 2) had highly non-significant Wald statistics; or 3) had $\log$-linear regression coefficients that were $<-10,000$ ! In the case of the liver, sometimes AMFIT warned us that the results may not be the maximum likelihood values. Such perturbations can be attributable to local maxima that are proximal to areas located near starting points on the likelihood sufface or a general lack of a signal-to-noise ratio in certain cross-classifications of the data. The choice of regression models must also be taken into consideration because AMFIT uses partial-likelihood models that are stratified, non-fully parametric mixtures of linear and $\log$-linear parameters. Therefore, interpretation of results when using such quasi-likelihood models for fitting data with little or no signal-to-noise, e.g., liver, should be treated with caution.

A very interesting finding was that the Freeman-Tukey goodness-of-fit statistic was much more stable than the Pearson $\chi^{2}$ or $D$ statistics. There were many situations where $\chi^{2}$ and $D$ decreased after adjusting for diagnostic misclassification and the Freeman-Tukey GOF $(G)$ either remained the same or increased. We can infer from this apparent pattern in GOF statistics that the $G$ may be a more reliable measure of goodness-of-fit and that $G$ may indicate when an appropriate adjustment is made when imputing missing data. It would be interesting to see how $\chi^{2}, D$ and $G$ would behave when the EM algorithm is used for adjusting for diagnostic misclassification.

Table 24: Weighted per cent distribution of non-cancer deaths among hiroshima males in the Life Span Study (1950-85).

|  |  |  |  |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATD |  |  |  |  |  |  |  |  |  |  |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | 60.69 | $70+$ |  |  |  |
| $<10$ | 0.96 | 0.74 | 0.94 | 0.36 | 0.00 | 0.00 | 0.00 |  |  |  |
| $10-19$ | 0.00 | 1.84 | 1.25 | 1.76 | 1.12 | 0.00 | 0.00 |  |  |  |
| $20-29$ | 0.00 | 0.28 | 0.70 | 0.90 | 1.31 | 1.00 | 0.00 |  |  |  |
| $30-39$ | 0.00 | 0.00 | 0.42 | 1.83 | 3.04 | 4.74 | 1.98 |  |  |  |
| $40-49$ | 0.00 | 0.00 | 0.00 | 1.06 | 5.57 | 8.62 | 14.45 |  |  |  |
| $50+$ | 0.00 | 0.00 | 0.00 | 0.00 | 1.90 | 9.04 | 34.19 |  |  |  |

Table 25: Weighted per cent distribution of non-cancer deaths among Nagasaki males in the Life Span Study (1950-85).

|  | Age ATD |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |  |
| $<10$ | 1.76 | 1.21 | 1.63 | 0.96 | 0.00 | 0.00 | 0.00 |  |
| $10-19$ | 0.96 | 2.36 | 2.05 | 2.68 | 0.67 | 0.00 | 0.00 |  |
| $20-29$ | 0.00 | 0.51 | 0.61 | 0.83 | 1.28 | 0.67 | 0.00 |  |
| $30-39$ | 0.00 | 0.00 | 0.16 | 1.25 | 3.39 | 3.42 | 1.63 |  |
| $40-49$ | 0.00 | 0.00 | 0.00 | 0.99 | 6.01 | 8.37 | 11.67 |  |
| $50+$ | 0.00 | 0.00 | 0.00 | 0.00 | 2.14 | 11.06 | 31.74 |  |

Table 26: Weighted per cent distribution of non-cancer deaths among Hiroshima females in the Life Span Study (1950-85).

|  | Age ATD |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |  |
| $<10$ | 0.41 | 0.35 | 0.35 | 0.12 | 0.00 | 0.00 | 0.00 |  |
| $10-19$ | 0.02 | 1.28 | 0.78 | 0.90 | 0.58 | 0.00 | 0.00 |  |
| $20-29$ | 0.00 | 0.50 | 1.24 | 1.09 | 1.94 | 0.60 | 0.00 |  |
| $30-39$ | 0.00 | 0.00 | 0.41 | 1.53 | 2.47 | 5.74 | 2.81 |  |
| $40-49$ | 0.00 | 0.00 | 0.00 | 0.82 | 3.99 | 6.75 | 16.78 |  |
| $50+$ | 0.00 | 0.00 | 0.00 | 0.00 | 1.09 | 5.61 | 41.86 |  |

### 5.3 DS86 Random Error

With regard to differential misclassification of exposure, where it was assumed that that sensitivities and specificities were unequal across exposure strata, the correction of DS86 random errors was successful and in many situations produced increased excess risks. This is in agreement with analyses performed by Pierce and Vaeth (1991) and Pierce et al. (1991). In most situations, the bias due to DS86 random errors was on average $-15 \%$ to $-30 \%$, and depended on the sex, site, or regression (with or without age ATD) or projection model that was used. The BEIR-V models provided a bias for DS86 random error that was in some cases positive, and was most likely due to truncation of dose equivalent to 4 Sv.

Table 27: Weighted per cent distribution of non-cancer deaths among Nagasaki females in the Life Span Study (1950-85).

|  | Age ATD |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |  |
| $<10$ | 1.10 | 0.73 | 0.63 | 0.26 | 0.00 | 0.00 | 0.00 |  |
| $10-19$ | 0.63 | 1.44 | 0.94 | 1.07 | 0.39 | 0.00 | 0.00 |  |
| $20-29$ | 0.00 | 0.60 | 1.68 | 1.34 | 1.44 | 0.76 | 0.00 |  |
| $30-39$ | 0.00 | 0.00 | 0.68 | 1.89 | 2.59 | 4.30 | 2.33 |  |
| $40-49$ | 0.00 | 0.00 | 0.00 | 0.97 | 4.14 | 6.86 | 14.80 |  |
| $50+$ | 0.00 | 0.00 | 0.00 | 0.00 | 1.10 | 6.65 | 40.68 |  |

### 5.4 Autopsy Program and Diagnostic Misclassiciation

The use of autopsy data to correct for diagnostic misclassification in RERF models proved useful and typically resulted in negative bias in the excess risks. On the other hand, correction of diagnostic misclassification in the BEIR-V models that excluded records for which mean age at death was greater than 75 may have been moot and not beneficial.

A plausible inference concerning exposure misclassification and the assumption that the true exposure-specific sensitivities are known, is that LSS subjects who have or who are undergoing either radiodiagnoses or radiotherapy would gravitate, if not adjusted for in an analysis, to the false negative exposure category because of underestimation of dose. Future studies should focus on, or at least take into consideration, medical exposures of LSS subjects when fitting dose-response.

It is certain that much of the information obtained from the Autopsy Program will not change. The vital status of autopsied decedents for whom no tissue or biological specimens exist may not change and may remain fixed forever. As the RERF tumor registries in Hiroshima and Nagasaki increase in size, the utility of cancer incidence data for determing the risk of developing radiation-induced neoplasia will begin to overshadow mortality data and lend itself well for verifying the decedents' true cause of death.

### 5.5 Bias in Lifetime Mortality Risks

There was a wide variation of bias for the various combinations of sites, gender and exposure profiles. Bias in the absolute and transported relative models was erratic and did not seem to follow any particular pattern. Land and Sinclair (1991) found that, when comparing lifetime risks of radiation across countries, international correlations under the absolute and transported relative models were lower than those provided by the relative model. Storer et al. (1988) also found that the relative risk model was more suitable for extrapolating risk from radiobiological studies in various mouse strains to man. In view of our findings in relationship to variability of bias across projection models, it is likely that similar findings could be obtained in future radiobiologic and international epidemiologic studies.

Cases for which bias was negative are shown for all models, sites, and sexes in Tables 13-20. For leukemia and nonleukemia, dose-related (T65DR) confirmatiou rates for cancer and non-cancer were available and always resulted in the most negative bias when compared with lifetime risks for which other or no adjustments were made. With regard to risk projection and studies of information bias, we recommend the relative model because its use, when compared with other models, results in fewer instances where total bias approaches $-50 \%$. The use of autopsy data to correct for diagnostic misclassification in BEIR-V sites that exclude records for which mean age at death is greater than 75 may be moot. in fewer instances where total bias approaches $-50 \%$.

Several authors suggest that when misclassification of outcome status is differentiul, that is can be corrected with dose-specific sensitivities and specificities, the results of correction will typically result in bias that is negative and toward the null, but can go in either direction (Fleiss, 1981; Kleinbaum et al., 1982; Flegal et al., 1986). This was not unexpected since it was shown in the Autopsy Program that the probability of autopsy increased with increasing radiation dose, therefore, cancer misclassification is greater in the exposed than it is in the zero-dose or not-in-city category (Yamamoto et al., 1978).

Green (1983) published a report of an extensive evaluation of the use of predictive value positive (confirmation rates) to adjust relative risk biased by misclassification of ontcome status. While equations and examples were given for the adjustment of RR in
a number of situations, the entire analysis was based on the predictive value positive of the non-exposed group. Under the constraint of using $P V^{+}$for only the nonexposed group, it would have been impossible to employ, with the exception of leukemia and nonleukemia, the confirmation rates from the Autopsy Study used in this study because for most sites confirmation rates were not available for the zero dose groups. There have been other studies on correction of diagnostic misclassification reported but, in the main, they address two-way tables used in log-linear analyses rather than maximum likelihood applications with Poisson regression modeling (Greenland and Robins, 1985; Savitz and Baron, 1989; Hsieh and Walter, 1988; Duffy et al., 1989; Greenland, 1989; Chen, 1989; Elton, 1989). Tables 21-27 give the weighted per cent distribution of non-cancer deaths on a crude, sex- and sex- and city-specific basis. Variation in the number of non-cancer deaths from table to table (Tables 21-27) suggest that, along with the excess radiation-induced cases, these data could strongly influence the bias due to misclassification. A thorough analysis of regression coefficients for most sites revealed that when the adjustment of diagnostic misclassification resulted in a negative bias (lifetime risks increased), it was wholly attributable to an increase of regression coefficients. Nevertheless, models for which the correction of misclassification resulted in negative bias always had goodness-offit statistics that were lower (better) than models for which no adjustment was made.

There is only a spattering of information on Poisson regression and diagnostic misclassification in the literature. In one particular study, the investigators developed likelihood equations based on binomial misclassification probabilities and international rates of cervical cancer mortality rates that followed the Poisson assumption (Whittemore and Gong, 1991). They developed four models to account for combinations of the presence of age and country covariates and error rates that were either independent of country (crude) or dependent on country. Their choice of a model was based solely on the log-likelihood ratio statistic, and interestingly, resulted in selection of the model with the most negative bias. The only otkar study was the one by Sposto et al. (1992) on the effect of diagnostic misclassification on the cancer dose-response curve in the LSS. Logistic regression was used to estimate cancer and non-cancer misclassification probabilities, along with the EM algorithm to impute true cancer deaths in cells for which no autopsy data existed. Poisson regression was used with a continuous model including covariates for age ATB, attained age, sex and stratified on city, sex, age ATB, and follow-up period. in order to compare our results with theirs, we used the Sposto et al. Poisson regression coefficients for nonleukemia to estimate lifetime risks with the non-constant relative projection model for acute exposures to 1 Sv at ages 25,45 and 65 and exposure to a total of 1 Sv from age 18 to 65 over a 47 year career.

Although a $12 \%$ increase ( $-12 \%$ bias) in ERR was reported for 50 -year old Hiroshima males exposed at age 25, the bias of lifetime risk (Tables 17 and 18) for exposure to 1 Sv at age 25 was only $-1 \%, 1 \mathrm{~Sv}$ at age 45 was $-4 \%, 1 \mathrm{~Sv}$ at age 65 was $1.3 \%$ and continuous exposure from age 18 to 65 ( 1 Sv total) was $-10 \%$. For females, bias of lifetime risk based on the Sposto et al. regression coefficients for nonleukemia were $13.3 \%$ for 1 Sv at 25 , $3.2 \%$ for 1 Sv at $45,5.2 \%$ for 1 Sv at 65 , and $2.5 \%$ for exposure over a career. Bias for the female lifetime risks was always positive for the Sposto et al data because the loglinear regression coefficient for gender ( 0 -males, 1 -females) changed from 0.356 when no adjustment for diagnostic misclassification was made to 0.315 (positive bias) when a $22 \%$ correction was made. The bias of lifetime risks of nonleukemia among males under the relative model (Tables 17) for exposure over a career ( $-11.3 \%$ ) was in accord with lifetime risk based on the Sposto et al. data ( $-10 \%$ ). However, for females, bias of lifetime risk of nonleukemia for exposure nver a career (Table 18) was dissimilar.

The implication of these findings is that investigators may focus on modeling to the extent that the relevance of modeling to worker protection (via lifetime risk projection) may become obfuscated and not portray the picture that is sought by policy makers. Since information bias was dependent on gender, site, method of correction, projection model and exposure profile, the full effect of diagnostic misclassification and DS86 random errors on risk for Western working populations, that is, the generalizability, ic best seen when lifetime risk projections are made following adjustments for information bias.

Variation of the misclassification bias in the two studies reflect the different methods which were used for estimating the true number of deaths in the cancer mortality data. Whereas Sposto et al. used logistic regression to estimate cancer misclassification probabilities and then used a full implementation of the EM algorithm to impute data in cells for which autopsy information did and did not exist, we used cancer and non-cancer confirmation rates to impute the true number of cancer deaths in all cells after a minimal latency period of 10 years. Differences existed in the models that were used: while Sposto et al. used a continuous model, the present study employed grouped models. Breslow and Day (1987) compared relative risk estimates from continuous and grouped Poisson regression models and concluded that there was no dramatic difference between results obtained with the two methods. However, a common assumption about using grouped methods is that the results will be less affected by distortion due to measurement error (Gilbert, 1982).

For all cases, lifetime risks based on Sposto et al. regression coefficients were higher than those for nonleukemia in the present study because Sposto et al. used a neutron RBE of unity when applying large intestine body self-shielding transmission factors to shielded kerma to obtain organ dose estimates. In consideration of our findings, and those of Sposto et al., it is likely that the two studies represented limited analyses of a larger problem related to information bias and the validity of generalizing LSS results to working Western working populations that are mostly chronically exposed to low doses of ionizing radiation.

## 6 SUMMARY

The numerical methods employed in the present study were extensive. Poisson regression results are provided for a variety of corrections made for diagnostic misclassification and DS86 random dosimetry ertor. Since there were so many combinations of correction methods, the results were listed in tabular notation because the use of a graphic format would result in figures that would been too difficult to comprehend. Readers who are interested in comparison figures can construct graphics from the tabular data in the text or the appendices.

The major findings of this investigation were:
(1). As age at death increased a greater proportion of true cancer deaths were attributable to non-cancer deaths because the true number of cancer deaths is equal to the sum of the product of the observed cancer deaths and the probability that the observed cancer deaths are correctly classified and the product of the observed non-cancer deaths and one minus the the probability that the observed non-cancer deaths were correctly classified (see Eq. 5 in $\$ 3.2 .2$ ).
(2). Poisson regression resulted in fitted maximum likelihood models that were in concordance with the observed data. When the goodness-of-fit of regression models containing time-dependent covariates is reasonable, non-constant lifetime risk projection should be used.
(3). Excess relative risk coefficients for the RERF and BEIR-V models were in good agreement with those published in RERF Report 11 (Part 2) and the BEIR-V report. Small differences existed between regression results for RERF models that contained parameters for age at-time-of-bombing (ATB), age at-time-of-death (ATD), and gender because organ dose estimates were used rather than shielded kerma. Thus, the lifetime risks based on these models were slightly higher than those that would obtain from the use of coefficients in RERF Report 11.
(4). Statistical modeling with the BEIR-V models provided regression coefficients that were almost exactly identical to those in the BEIR-V report. For leukemia, the linearquadratic contribution of dose to excess mortality was slightly lower than that in the BEIRV report. Lifetime risks based on the BEIR-V models were similar to those published in the BEIR-V Report (NRC, 1990). Bias due to DS86 random error for the digestive site was smaller than bias in the RERF non-constant nonleukemia projection models, which was most likely due to truncation of dose equivalent to 4 Sv . The correction of diagnostic misclassification in excess risks for the FEIR-V digestive cancer site had little effect on bias $(-2 \%)$ because records with an age at death beyond 75 , when cancer misclassification rises markedly, were excluded.
(5). Using a Dose-Rate Reduction Effectiveness Factor (DRREF) of two and no correction for DS86 random error or diagnostic misclassification in the non-constant relative projection model, lifetime risks (\%/Sv) of nonleukemia among males exposed acutely to 1 Sv at age $25,45,65$ or continuously from ages 18 to 65 were $2.10 \%, 2.78 \%, 1.20 \%$ and $1.91 \%$. For females, nonleukemia lifetime risks for the same exposure profles were $3.49 \%$, $4.32 \%, 1.97 \%$ and $3.23 \%$. Excess leukemia risks for 1 Sv at 25,45 and 65 and over the years 18 to 65 were $0.35 \%, 0.46 \%, 2.46 \%$ and $0.87 \%$ for males and $0.26 \%, 0.41 \%, 1.96 \%$ and $0.73 \%$ for females. These data were in good agreement with the results of Land and Sinclair (1991). By way of comparison, for exposure from ages 18 to 65 , excess nonleukemia risks based on the constant relative projection model were $2.84 \%$ for males and $4.75 \%$ for females. The risks of leukemia among males was $0.75 \%$ and among females was $0.64 \%$. Therefore, lifetime risk estimates based on constant models did not underestimate risks
projected by non-constant models.
(6). The correction of differential diagnostic misclassification with leukemia and nonleukemia (and non-cancer) confirmation rates that were stratified on T65DR dose (DS86 shielded kerma was converted to T65DR shielded in order to select T65DR-specific confirmation rates) resulted in bias that was negative. Confirmation rates for leukemia and nonleukemia that were stratified on age ATD did not provide bias that was more negative than that obtained with DS86-specific confirmation rates. Correction of diagnostic misclassification using confirmation rates that were crude or stratified on either gender or city and gender resulted in bias that was negative or positive. The bias of excess risk of nonleukemia due to diagnostic misclassification for 1 Sv at age $25,45,65$ or continuously from ages 18 to 65 under the non-constant relative projection model was $-5.0 \%(2.13 \% / \mathrm{Sv}$ vs. $2.24 \% / \mathrm{Sv}),-7.3 \%(2.78 \% / \mathrm{Sv}$ vs. $2.99 \% / \mathrm{Sv}),-38.9 \%(1.20 \% / \mathrm{Sv}$ vs. $1.67 \% / \mathrm{Sv})$ and $-11.3 \%$ $(1.91 \% / \mathrm{Sv}$ vs. $2.13 \% / \mathrm{Sv})$ for males and $-1.5 \%(3.49 \% / \mathrm{Sv}$ vs. $3.54 \% / \mathrm{Sv}),-3.9 \%(4.32 \% / \mathrm{Sv}$ vs. $4.49 \% / \mathrm{Sv}),-26.2 \%(1.97 \% / \mathrm{Sv}$ vs. $2.18 \% / \mathrm{Sv})$ and $-6.3 \%(3.23 \% / \mathrm{Sv}$ vs. $3.43 \% / \mathrm{Sv})$ for females. For leukemia excess risks under the same dose profiles, the bias due to diagnostic misclassification was $-6.0 \%(0.36 \% / \mathrm{Sv}$ vs. $0.37 \% / \mathrm{Sv}),-69.7 \%(0.46 \% / \mathrm{Sv}$ vs. $0.77 \% / \mathrm{Sv})$, $-23.3 \%(2.46 \% / \mathrm{Sv}$ vs. $3.04 \% / \mathrm{Sv})$ and $-23.9 \%(0.87 \% / \mathrm{Sv}$ vs. $1.09 \% / \mathrm{Sv})$ for males and $-12.1 \%(0.26 \% / \mathrm{Sv}$ vs. $0.30 \% / \mathrm{Sv}),-83.4 \%(0.41 \% / \mathrm{Sv}$ vs. $0.75 \% / \mathrm{Sv}),-40.9 \%(1.96 \% / \mathrm{Sv}$ vs. $2.77 \% / \mathrm{Sv})$, and $-42.8 \%(0.73 \% / \mathrm{Sv}$ vs. $1.05 \% / \mathrm{Sv})$ for females. When the nonleukemia Poisson regression coefficients from Sposto et al. (1992) were used to project lifetime risks under the non-constant relative model, the bias due to diagnostic misclassification for 1 Sv acute at 25,45 , or 65 and over a career ( 18 to 65 ) was $-1.0 \%(2.61 \% / \mathrm{Sv}$ vs. $2.64 \% / \mathrm{Sv}$ ), $-4.0 \%(5.72 \% / \mathrm{Sv}$ vs. $5.95 \% / \mathrm{Sv}), 1.3 \%(2.39 \% / \mathrm{Sv}$ vs. $2.36 \% / \mathrm{Sv})$, $九 \mathrm{n}$. $-10.0 \%(4.90 \% / \mathrm{Sv}$ vs. $5.39 \% / \mathrm{Sv}$ ) for males and $13.3 \%(3.02 \% / \mathrm{Sv}$ vs. $2.62 \% / \mathrm{Sv}), 3.2 \%(6.49 \% / \mathrm{Sv}$ vs. $6.28 \% / \mathrm{Sv})$, $5.2 \%(2.46 \% / \mathrm{Sv}$ vs. $2.32 \% / \mathrm{Sv})$ and $2.5 \%(5.92 \% / \mathrm{Sv}$ vs. $5.77 \% / \mathrm{Sv})$ for females.
(7). The use of reduction factors to corre it for DS86 random error in survivor doses indicated that lifetime risks were negatively biased $15 \%-30 \%$. Bias of excess risk (nonconstant relative projection ad correction for diagnustic misclassification) of nonleukemia due to DS86 random errors for 1 Sv at age $25,45,65$ or continuously from ages 18 to 65 was $-27.1 \%(2.24 \% / \mathrm{Sv}$ vs. $2.28 \% / \mathrm{Sv}),-23.5 \%(2.99 \% / \mathrm{Sv}$ vs. $3.69 \% / \mathrm{Sv}),-24.6 \%(1.67 \% / \mathrm{Sv}$ vs. $2.08 \% / \mathrm{Sv}$ ) and $-23.7 \%(2.13 \% / \mathrm{Sv}$ vs. $2.63 \% / \mathrm{Sv})$ for males and $-19.6 \%(3.54 \% / \mathrm{Sv}$ vs. $4.24 \% / \mathrm{Sv}),-13.9 \%(4.49 \% / \mathrm{Sv}$ vs. $5.12 \% / \mathrm{Sv}),-15.9 \%(2.48 \% / \mathrm{Sv}$ vs. $2.88 \% / \mathrm{Sv})$ and $-14.9 \%(3.43 \% / \mathrm{Sv}$ vs. $3.94 \% / \mathrm{Sv})$ for females. For leukemia excess risks under the same dose profiles, the bias due to DS86 random error was $-17.4 \%(0.37 \% / \mathrm{Sv}$ vs. $0.44 \% / \mathrm{Sv}$ ), $-14.2 \%(0.77 \% / \mathrm{Sv}$ vs. $0.88 \% / \mathrm{Sv}),-13.3 \%(3.04 \% / \mathrm{Sv}$ vs. $3.44 \% / \mathrm{Sv})$ and $-14.0 \%(1.09 \% / \mathrm{Sv}$ vs. $1.24 \% / \mathrm{Sv})$ for males and $-15.1 \%(0.30 \% / \mathrm{Sv}$ vs. $0.34 \% / \mathrm{Sv}),-11.6 \%(0.75 \% / \mathrm{Sv}$ vs. $0.84 \% / \mathrm{Sv}),-10.9 \%(2.77 \% / \mathrm{Sv}$ vs. $3.07 \% / \mathrm{Sv})$, and $-11.4 \%(1.05 \% / \mathrm{Sv}$ vs. $1.17 \% / \mathrm{Sv})$ for females.
(8). The correction of morteity misclassification in SEER baseline rates used in lifetime risk projection (non-constant relative model) increased exiess risks by $2.1 \%$ for nonleukemia and decreased risk by $10.8 \%$ for leukernia.
(9). The total bias of excess risk of nonleukemia for exposure from age 18 to 65 under the non-constant relative projection model was $-37.1 \%$ for males and $-23.3 \%$ for females. For leukemia excess riske under the relative projection model, the total bias was $-27.1 \%$ for males and $-43.4 \%$ for females. Thus, nonleukemia risks increased $37.1 \%$ for males ( $1.91 \% / \mathrm{Sv}$ to $2.68 \% / \mathrm{Sv}$ ) and $23.3 \%$ for females $(3.23 \% / \mathrm{Sv}$ to $4.02 \% / \mathrm{Sv})$ and leukemia risks increased $27.1 \%(0.87 \% / \mathrm{Sv}$ to $1.10 \% / \mathrm{Sv})$ for males and $43.4 \%(0.73 \% / \mathrm{Sv}$ to $1.04 \% / \mathrm{Sv}$ ).
(10). In most cases, bias due to diagnostic misclassification for lifetime risk projections
usirg the relative model was more positive and less erratic than bias for the absolute and transported relative models. With regard to risk projection and future studies of information bias, we recommend the relative model because its use, when compared with other models, resulted in biases with lower variation across gender, sites and exposure profiles.

It is patently clear that the effects of diagnostic nisclassification and DS86 random errors are dependent on gender, site, correction methods, exposure profiles and projection models. The effects of increased internal validity on the generalizability of Japanese radiation risk information to U.S. nuclear workers are only revealed when lifetime risks are projected after adjustments are made for random and systematic errors. Future studies in which LSS data are generalized to U.S. nuclear workers may be biased if lifetime risks are not adjusted for random and systematic errors.

Epidemiologic theories of bias were applied and expounded throughout the course of this investigation. Adherents of our results should not let their enthusiasm exceed their knowledge of bias, so that our assumptions become regarded as fixed verities, rather than empirical bypotheses. The major purpose for undertaking this study was to confirm the impression that there are certain advantages of projecting lifetime risk, in addition to performing Poisson regression, when studying information bias in the LSS. Since we did not employ logistic regression to estimate cancer misclassification probabilities and did not fully implement the EM algorithm to impute missing data where there was no autopsy information, this study should be regarded as an investigation into the most fundamental assumptions. As a result, new phenomena in the LSS should not force a reevaluation of this study's findings.

## 7 REFERENCES

Abramowitz, M. and Stegun, I.A. Handbook of Mathematical Functions. New York: Dover; 1965.

Algorithm AS32. J. Royal Stat. Soc. 19(3); 1970.
Angevine, D.M., Jablon, S., Matsumoto, Y.S. ABCC. JNIH Pathology Studies, Hiroshama and Nagasaki. Repori 1. October 1950-September 1962. Atomic Bomb Casualty Commission (ABCC) Technical Report 14-63. Hiroshima: ABCC; 1963.
Beebe, G.W., Usugawa, 1968. The Major ABCC Samples. Atomic Bomb Casualty Commission (ABCㄱ) Technical Report 12-68. Hiroshima: ABCC; 1968.
Beebe, G.W., Jamamoto, T., Niatsumoto, Y.S., Gould, S.E. ABCC-JNIH Pathology Studies, Firoshima and Nagasaki. Report 2. October 1950 . December 1965. Atomic Bumb Casualty Ccmmission (ABCC) Technical Report 8-67. Hiroshima: ABCC; 1967.
Breslow, N.E., Day, N.E. Fitting Models to Continuous Data, Chap. 5. In: Statistical Methecis in Cancer Research. Volume II: The design and analysis of cohort studies. Chapter 4. IARC Scientific Publication No. 82. Lyon: IARC; 1987.
Brown, C.C., Chu, K.C. Additive and multiplicative models and multistage carcinogenesis theory. Risk Anal. $9(1): 99-105 ; 1989$.
Bunger, B.B., Cook, J.R., Barrick, M.K. Life table methodology for evaluating radiation risk: An application baszd on occupational exposures. Health Physics. 40:439-455; 1981.
Chen, T.T. A review of methods for misclassified categorical data in epidemiology. Stat. Med. 8:1095-1106; 1989.
Checkoway, H., Pearce, N., Crawford-Brown, D.J. Special (ions of Occupational Epidemiology Data, Chap. 10. In: Research Methods in Occupational Epidemiology. New Yoik: Oxford; 1989.

Chiang, C.L. Some Stochastic Models of Growth - Pure Death Process, Chap. 8. In: An Introduction to Stochastic Processes and Their Applications. Huntington(NY): Krieger; 1968.

Chiang, C.L. The Life Table and Its Construction - The Complete Life Table, Chap. 6. In: The Life Table and Its Application. Malabar(FL): Krieger; 1984.
Cochran, W.G. The Estimation of Sample Size, Chap. 4. In: Sampling Techniques. New York: Wiley; 1977.

Cressie, N., Read, T.R.C. Multinomial goodness-of-fit tests. J. Royal Stat. Soc. (B), 46:440464; 1984.

Dempster, A.P., Laird, N.M., Rubin, D.B. Maximum likelihood from incomplete data via the EM algorithm. J. Royal Stat. Soc. B. 39:1-38; 1977.

Dorn, H.F., Horn, J.L. The reliability of certificates of deaths from cancer. Am. J. of Hygiene. 34:12-23; 1941.

Dorn, H.F., Cutler, S.J. Morbidity from cancer in the United States. Comparison of death certificates and case reports. Pub. Hlth. Mono. 1958. 56:117-124; 1958.

Duffy, S.W., Rohan, T.E., Day, N.E. Misclassification in more than one factor in a casecontrol study: A combination of Mantel-Haenszel and maximum likelihood approaches. Stat. Med. 8(12):1529-1536; 1989.

Elandt-Johnson, R.C., Johnson, N.L. Theory of Competing Causes: Probabilistic Approach, Chap 9. In: Survival Models and Data Analysis. New York: Wiley; 1980.

Elton, R.A., Duffy, 5.W. Correcting for the effect of misclassification bias in a case-control study using data from two different questionnaires. Biometrics. 39(3):659-663; 1983.

Espeland, M.A., Platt, O.S., Gallagher, D. Joint estimation of incidence and diagnostic error rates from irregular longitudinal data. J. Am. Stat. Assoc. 84(408):972-979; 1989.

Faber, J.H., Wade, A.H. Life Tables for the United States: 1900-2050. Actuarial Study No. 89, Office of the Actuary, U.S. Social Security Administration. SSA Pub. No. 11-11536 (1983). Bethesda: SSA; 1983.

Flegal K.M., Brownie, C., Haas J.D. The effects of exposure misclassification on estimates of relative risk. Am. J. Epid. 41(12):1167-1174; 1988.

Fleiss, J.L. The Effects of Misclassification Errors, Chap. 11. In: Statistical Methods for Rates and Proportions. New York: Wiley; 1981.

Freeman, M.R., Tukey, J.W. Transformation related to the angular and square root transform. Ann. Math. Staiist. 21:607-611; 1950.

Frome, E.L., Kutner, M.H., Beauchamp, J. J. Regression analysis of Poisson-distributed data. J. Am. Stat. Assoc.. 68; 935-940; 1973.

Frome, E.L. The analysis of rates using Poisson regression methods. Biometrics. 39; 665-674; 1983.

Frome, E.L. PREG: A computer program for Poisson regression anaiysis. Oak Ridge Associated Universities Report ORAU-178 (1981). Oak Ridge National Laboratory. Oak Ridge(TN): ORAU; 1981.

Gail, M. Measuring the benefit of reduced exposure to environmental carcinogens. J. Chron. Dis. 28:135-147; 1975.

Gilbert, E.S. Some effects of random dose measurement errors on analysis of atomic bomb survivor data. Radiation Effects Research Foundation (RERF) Technical Report 12-82. Hiroshima: RERF; 1982.

Gilbert, E.S. Late Somatic Effects. In: Health Effects Model for Nuclear Power Plant Accident Consequence Analysis. Part 1, Introduction, Integration, and Summary. U.S. Nuclear Regulatory Commission, Report NUREG/CR-4214, Rev. 1, Part I. Washington, D.C., NRC; 1989a.

Giibert, E.S. Late Somatic Effects, Chap 3. In: Health Effects Model for Nuclear Power Plant Accident Consequence Analysis. Part 2, Scientific Basis for Health Effects Models. U.S. Nuclear Regulatory Commission, Report NUREG/CR-4214, Rev. 1, Part II. Washington, D.C.: NRC; 1989b.

Gilbert, E.S. Late Somatic Effects. In: Modifications of Models Resulting From Recent Reports on Health Effects of Ionizing Radiation. Part II: Scientific Bases for Health Effects Models. Health Effects Model for Nuclsar Power Plant Accident Consequence Analysis, NUREG/CR-4214, Rev.1, Part II, Addel.dum 1. Washington, D.C.: NRC; 1991.

Green, M.S. Use of predictive value to adjust relative risk estimates biased by misclassification of outcome status. Am. J. Epid. 117(1):98-105; 1983.
Greenland, S. On correcting for misclassification in twin studies and other matched-pair studies. Stat. Med. 8(7):825-829; 1989.
Greenland, S., Robins, J. M. Confounding and misclassification. Am. J. Epid. 122(3):495-506; 1985.

Greenwood, M. The Natural Duration of Cancer, Reports on Public Health and Medical Subjects, Her Majesty's Stationary Office, London, Vol. 33:1-26; 1926.
Hsieh, C.C., Walter, S.D. The effect of non-differential exposure misclassification on estimates of the attributable and prevented fraction. Stat. Med. 7(10):98-105; 1983.
1CRP, International Commission on Radiological Protection, 1977. Recommendations of the ICRP. ICRP Publication 26. Annals of the ICRP, Vol. 1(3). Oxford: Pergamon; 1977.
ICRP, International Commission on Radiological Protection, 1991. 1990 Recommendations of the International Commission on Radiological Protection. ICRP Publication 60. Annals of the ICRP, Vol. 21(1-3). Oxford: Pergamon; 1991.
Ishida, M., Beebe, G.W. Research plan for joint NIH-ABCC study of life span of A-bomb survivors. Atomic Bomb Casualty Commission (ABCC) Technical Report 04-59. Hiroshima: ABCC; 1959.
Jablon, S., Angevine, D.M., Matsumoto, Y.S., Ishida, M. On the significance of cause of death as recorded on death certificates in Hiroshima and Nagasaki, Japan. Study of Cancer and Other Chronic Diseases. In: National Cancer Institute Monograph 19 (1966). Ed, Haenzel, W. National Cancer Institute, Public Health Service, U.S. Department of Health, Education, and Welfare. Bethesda: HEW; 1966.
Jablon, S. Atomic bomb radiation dose estimetion at ABCC. Atomic Bomb Casualty Commission (ABCC) Technical Report 23-71. Hiroshima: ABCC; 1971.
James, G., Patton, R.E., Heslin, A.S. Accuracy of cause-of-death statements on death certificates. Pub. Hith. Rep. 1955. 70;39-51; 1955.
Kahn, H., Sempos, C.T. Follow-up Studies: Person- Years, Chap. 8. In: Statistical Methods in Epidemiology. New York: Oxford; 1989.

Kennedy, W.J., Gentle, J.E. Unconstrained Optimization and Non-linear Regression, Ch.p. 10. In: Statistical Computing. New York: Marcel Dekker; 1980.

Kleinbaum, D.G., Kupper, L.L., Morgenstern, H. Information Bias, Chap. 12. In: Epidemiologic Research: Principles and Quantitative Methods. New York: Van Nostrand Reinhold; 1982.

Kleinbaum, D.G., Kupper, L.L., Muller, K.E. Maximum Likelihood Methods: Theory and Ap-
plications, Chap. 21. In: Applied Regression Analysis and Other Multivariable Methods. Boston: PWS-Kent; 1988.

Kodell, R.L., Gaylor, D.W. On the additive and multiplicative models of relative risk. Biom. J. 31:359-370; 1989.

Kodell, R.L., Krewski, D., Zielinski, J.M. Additive and multiplicative relative risk in the two-stage cloaal expansion model of carcinogenesis. Risk Anal. 11:483-490; 1991.

Kramer, M.S. Diagnostic Tests. Chap. 16. In: Clinical Epidemiology and Biostatistics. Berlin: Springer-Verlag; 1988.

Land, C.E. Projection of risk from one population to another. In: Risk Estimates for Radiation Carcinogenesis, (Renz, G.K., Ed). Proceedings of the International Workshop of the Institut für Strahlenschutz. Koln: Institut für Strahlenschutz: 1989.

Land, C.E., Sinclair, W.K. The relative contributions of different organ sites to the total cancer mortality associated with low-dose radiation exposure. Annals of the ICRP. 22, 31-37; 1991.

Lee, E.T. Functions of Survival Time, Chap. 2. In: Statistical Methods for Survival Data Analysis. Belmont(CA): Wadsworth; 1980.

Lilienfeld, A.M., Lilienfeld, D.E. Morbidity Statistics, Chap. 6. In: Foundations of Epidemiology (2nd Ed.). New York: Oxford; 1980.

McCullagh, P., Nelder, J.A. Model Checking, Chap. 12. In: Generalized Linear Models. London: Chapman and Hall; 1989.

Miller, A.B., Howe, G.R., Sherman, G.J., Lindsay, J.P., Yaffe, M.J., Dinner, P.J., Risch, H.A., Preston, D.L. Mortality from breast cancer after irradiation during fluoroscopic examinations in patients being treated for tuberculosis. N. Eng. Journ. Med. 321:12851289; 1989.

Milton, R.C., Shohoji, T. Tentative 1965 radiation dose estimation for atomic bomb survivors, Hiroshima and Nagasaki. Atomic Bomb Casualty Commission (ABCC) Technical Report 1-68. Hiroshima: ABCC; 1968.

Monson, R.R.. Analysis of relative survival and proportionate mortality. Comp. Biomed. Research. 7;325-332; 1974.

Moolgavkar, S.H., Knudson, A.G. Mutation and cancer: a model for human carcinogenesis. JNCI. 66:1037-1052; 1981.

Moriyama, L.M., Baum, W.S., Haenszel, W.M. et al. Inquiry into diagnostic evidence supporting medical certifications of death. Am. J. Pub. Hlth. 48:1376-1387; 1958.

Muirhead, C.R., Darby, S.C. Modelling the relative and absolute risks of radiation-induced cancers. J. Royal Stat. Soc. A. 150:83-118; 1987.

NCI, National Cancer Institute, 1989. Cancer Statistics Review 1979-1986. May 1989 Bethesda: HHS; 1989.

NRC, National Research Council, 1980. Committee on the Biological Effecis of Ionizing Radiations. The effects on populations of exposure to low-levels of ionizing radiation. Washington, D.C.: NAS; 1980.

NRC, National Research Council, 1990. Committee on the Brological Effects of Ionizing Radiations (BEIR-V). Health Effects of Exposure to Low Levels of lonizing Radiation. Waslingtou, D.C.: NAS; 1990.

Percy, C.L., Stanek, E.,III, Gloecker, L.A. Accuracy of cancer death certificates and its effect on cancer mortality statistics. Am. J. Pub. Hlih. 71:242-250; 1981.
Percy, C.L., Miller, B.A., Gloeckler Ries, L.A. Effects of changes in cancer classification and the accuracy of cancer death certificates on trends in cancer mortality. Ann. N.Y. Acad. Sci. 409:87-97; 1990.

Peterson, L.E., Schull, W.J., Buffler, P.A., Davis, B.R. Birth cohort and period effects in absolute and relative risk projection of radiation-induced leukemia. Abstracts of papers for the 40th Annual Meeting of the Radietion Research Society, March, 1992.
Peterson, L.E., Schull, W.J., Davis, B.R., Buffler, P.A. SURVRAD: Computer Program for Estimating Lifetime Mortality Risk of Radiation-Induced Cancer. Epidemiology Research Unit Technical Report, School of Public Health, University of Texas. University of Texas - Health Science Center, Houston, 1992.

Pierce, D.A., Preston, D.L., Ishimaru, T. A method of analysis of cancer incidence in Japanese atomic bomb survivors with application to acute leukemia. Radiation Effects Research Foundation (RERF) Technical Report 15-83. Hiroshima: RERF; 1983.
Pierce, D.A., Preston, D.L. Hazard function modeling for dose response analysis of cancer incidence in the A-bomb survivors. In: Atomic Bomb Survivor Data: Utilization and Analysis. pp. 51-66. SIAM Institute for Mathematics and Society. Philadelphia: SIAM; 1984.

Pierce, D.A., Preston, D.L. Analysis of cancer mortality in the A-bomb survivor cohort. In: Proc. 45th Seasion Int. Statist. Inst. (pp. 1-13). Amsterdam: Int. Stat. Inst.; 1985.
Pierce, D.A., Preston. D.L. Developments in cohort analysis with application to radiationinduced cancer. In: 46th Session Int. Stat. Inst. (pp. 557-570). Amsterdam: Int. Stat. Inst.; 1988.

Pierce, D.A. and Vaeth, M. The shape of the cancer mortality dose-response curve for the atomic bomb survivors. Radial. Res. 126: 36-42; 1991.
Pierce, D.A., Stram, D.O., Vaeth, M. Allowing for random errors in radiation exposure estimates for the atomic bomb survivor data. Radiat. Res. 123:275-284; 1990. [RERF TR 2-89].

Pierce, D.A., Vaeth, M. Cancer risk estimation from the A-bomb survivors: extrapolation to low doses, use of relative risk models, and other uncertainties. Radiation Effects Research Foundation, Commentary and Review Series, CR 2-89. Hiroshima: RERF; 1989.
Pierce, D.A., Preston, D.L., Stram, D.O., Vaeth, M. Allowing for dose-estimation errors for the A-bomb survivor data. J. Radiat. Res. (Japan). Supp. 108-121; 1991.

Preston, D.L., Kato, H., Kopecky, K.J., Fujita, S. Life Span Study Report 10. Part 1, Cancer Mortality among A-bomb Survivors in Hiroshima and Nagasaki, 1950-82. Radiation Effects Research Foundation (RERF) Technical Repori 1-86. Hiroshima: RERF; 1986.

Preston, D.L., Pierce, D.A. AMFIT: A program for parameter estimation in additive and
multiplicative rate models with grouped survival data. EPICURE User's Guide and Command Language Handbooks. Hirosoft International Corporation. Seattle: Hirosoft; 1993.

Puffer, RR., Griffith, G.W. Patterns of urban mortality. Pan American Health Organization, WHO; 1967.

Rall, J.F., Beebe, G.W., Hoel, D.G., Jablon, S., Land, C.E., Nygaard, O.F., Upton, A.C., Yalow, R.S. Report of the National Institutes of Health Ad Hoc Working Group to Develop Radioepidemiological Tables, DHHS Publication No.(NIH) 85-2748. Washington, D.C.: NIH; 1985.

Rao, C.R. Large sample tests of statistical hypothesis concerning several parameters with applications to problems of estimation. Proc. Camb. Phal. Soc. 44:50-57; 1947.
Rayner, J.C.W., Best, D.J. Asymptotically Optimal Tests, Chap. 3. In: Smooth Tests of Goodness of Fit. New York: Oxford; 1989.

Ries, L.A.G., Hankey, B.F., Miller, B.A., Hartman, A.M., Eiwards, B.K. Cancer Statistics Review, 1973-88. National Cancer Institute. NIH Publication No. 91-2789. Bethesda: $\mathrm{NIH}(\mathrm{NCl}) ; 1991$.
Ron, E., Hoel, D.G., Carter, R.L., Mabuchi, K. The influence of death-certificate errors on cancer mortality trends. Radiation Effects Research Foundation, Commentary and Review Series, CR-2-91. Hiroshima: RERF; 1991.

Samter, T.G., Hrubec, Z., et al. Epidemiologic Evaluation of ABCC and Nagasaki University School of Medicine Autopsies and ABCC Surgical Specimens, Nagasaki 1950-1959. Atomic Bomb Casualty Commission (ABCC) Technical Report 15-60. Hiroshima: ABCC; 1960.

Santuer, T.J., Duffy, D.E. Univariate Discrete Responses, Chap. 2. In: The Statistical Analysis of Discrete Data. New York: Springer-Verlag; 1989.
Savitz, D.A., Baron, A.E. Estimating and correcting for confounder misclassification. Am. J. Epid. 129(5):1062-71; 1989.

Shimizu, Y., Kato, H., Schull, W. J., Preston, D.L., Fujita, S., Pierce, D.A. Life Span Study Report 11. Part 1, Comparisons of Risk Coefficients for Site-specific Cancer Mortality Based on the DS86 and T65D Shielded Kerma and Organ Doses. Radiation Effects Research Foundation (RERF) Technical Report 12-87. Hiroshima: RERF; 1987.

Shimizu, Y., Kato, H., Schull, W.J. Life Span Study Report 11. Part 2, Cancer Mortality in the Years 1950-85 Based on the Recently Revised Doses (DS86). Radiation Effects Research Foundation (BERF) Technical Report 5-88. Hiroshima: RERF; 1988.

Shore, R.E., Hildreth, N., Woodard, E., Dvoretsky, P., Hempelmann, L., Pasternack, B. Breast cancer among women given $x$-ray therapy for acute postpartum mastitis. JNCI. 77:689-696; 1986.

Smith, T.L., Putman, J.E., Gehan, E.A. A computer program for estimating survival functions from the life table. Comp. Prgms. in Biomed. i:58-64; 1970.
Sposto, R.. Stram, D.O., Awa, A.A. An estimate of the magnitude of random errors in the DS86 Dosimetry System from data on chromosome aberrations and severe epilation. Radiat. Res. 128:157-169 (1991). [RERF TR 7-90].

Sposto, R., Preston, D.L, Shimizu, Y., Mabuchi, K. The effect of diagnostic misclassification on non-cancer and cancer mortality dose-response in A-bomb survivors. Biometrics. 48:605-617; 1992. [RERF TR 4-91].
Stather, J.W., Muirhead, C.R., Edwards, A.A., Harrison, J.D., Lloyd, D.C., Wood, N. Health Effects Models Developed from the 1988 UNSCEAR Report. National Radiological Protection Board, Report R226. Oxfordshire: NRPB; 1988.

Steer, A., Moriyama, I.M., Shimizu, K. ABCC-JNIH Pathology Studies, Hiroshima and Nagasaki. Report 3. January 1951 - December 1970. Atomic Bomb Casualty Commission (ABCC) Technical Report 16-73. Hiroshima: ABCC; 1973.
Stone, R.S., Anderson, P.S. Epidemiologic Evaluation of ABCC Autopsies, Hiroshima 195059. Atomic Bomb Casualty Commission (ABCC) Technical Report 22-60. Hiroshima: ABCC; 1960.

Stone, R.S., Anderson, P.S. A Comparison of Death Cerlificate and Autopsy Diagnoses, Hiroshima. Atomic Bomb Casualty Commission (ABCC) Technical Report 19-60. Hiroshima: ABCC; 1960.
Storer, J.B., Mitchell, T.J., Fry, R.J.M. Extrapolation of the relative risk of radiogenic neoplasms across mouse strains and to man. Radiat. Res. 114:331-353; 1988.
Thiessen, J.W. and Kaul, D.C. The dosimetry system 1986 (DS86) and the Tentative Dosimetry System 1965 (T65D): How do they compare, what is left to do? J. Radiat. Res. (Japan). Supp. 1-10; 1991.
Thomas, D.C., Darby, S., Fagnani, F., Hubert, P., Vaeth, M., Weiss, K. Definition and estimation of lifetime detriment from radiation exposure: Principles and methods. Healh Physics. 63:259-272; 1992.

UN, United Nations, 1988. Sources, Effects, and Risks of Ionizing Radiation. United Nations Scientific Committee on the Effects of Atomic Radiation 1988 Report to the General Assembly, with Annexes. United Nations sales publication No. E.88.IX.7. New York: UN; 1988.
Vaeth, M., Pierce, D.A. Caiculating excess lifetime risk in relative risk models. Radiation Effects Research Foundation, Commentary and Review Series, CR 3-89. Hiroshima: RERF; 1989.

Wald, A. Tests of statistical hypotheses concerning several parameters when the number of parameters is large. Trans. of Am. Math. Soc. 54:426-482; 1943.
White, E. The effect of misclassification of disease status in follow-up studies: implications for selecting disease classification criteria. Am. J. Epid. 124(5):816-825; 1986.
Whittemore, A.S., Gong, G. Poisson regression with misclassified counts: application to cervical cancer mortality rates. Appl. Stat. 40:81-93; 1991.
World Health Organization, 1959. Manual of the International Statistical Classification of Diseases, Injuries and Causes of Death, 1955 Ed. Geneva: WHO; 1959.

Yamamoto, T., Moriyama, 1.M., Asano, M., Guralnick, L. RERF Pathology Studies, Hiroshima and Nagasaki, Report 4. Autopsy Program and the Life Span Study, January 1961 - December 1975. Radiation Effects Research Foundation (RERF) Technical Report

18-78. Hiroshima: ABCC; 1978.
Zeldis, L.J., Matsumoto, Y.S. Research plan for joint ABCC-NIH pathology studies in Hiroshima and Nagasaki. Atomic Bomb Casualty Commission (ABCC) Technical Report 12-62. Hiroshima: ABCC; 1962.

## 8 NOTATION

$\theta$
$\theta$
$d_{c} \quad$ Cancer deaths observed on death certificates.
$d_{n c} \quad$ Non-cancer deaths observed on death certificates.
$d_{T} \quad$ Total observed deaths equal to the sum of cancer and non-cancer deaths.
$P V^{+} \quad$ Predictive value positive, equal to the cancer confirmation rate.
Defined as the probability that individuals with cancer X as the underlying cause of death on their death certificate truly died of cancer X .
$\mathrm{PV}^{-} \quad$ Predictive value negative, equal to the non-cancer confirmation rate. Defined as the probability that individuals without cancer X as the underlying cause of death on their death certificate truly did not die of cancer X .
True effect measure of risk for the target population.
The estimator of $\theta$ based on a sample from the target population called the study population.

The parameter estimated by $\hat{\theta}$ for the larger actual population that is obtained from the study population. When adjustments are only made for random error, risk estimates are equal to $\hat{\theta}$, but when corrections are made for both random and systematic errors, risks are equal to $\theta$.

True positive cancer deaths. True cancer deaths certified as cancer deaths.
False positive non-cancer deaths. Non-cancer deaths certified as cancer deaths.
False negative cancer deaths. Cancek deaths certified as non-cancer deaths.
True negative non-cancer deaths. Non-cancer deaths certified as non-cancer deaths.

True cancer deaths estimated by sufficient statistics (Eq. 5).
True non-cancer deaths estimated by sufficient statistics (Eq. 6).

Sensitivity, equal to the cancer detection rate. Defined as the probability of correctly assigning an underlying cause of death as cancer X for individuals who truly died of cancer X.

Specificity, equal to the non-cancer detertion rate. Defined as the probability of correctly assigning an underlying cause of death as cause X for individuals who truly died of cause X .

| $\theta$ c | Cancer confirmation rate, equal to $\mathrm{PV}^{+}$. |
| :---: | :---: |
| $\theta_{n \mathrm{n}}$ | Non-cancer confirmation rate, equal to $P V^{-}$. |
| $\pi_{e}$ | True cancer rate. |
| $\pi_{n c}$ | True non-cancer rate. |
| $R(z)_{n+\gamma, \text { eity }}$ | City-specific reduction factor for DS86 random error multiplied by organ dose. |
| $z_{n+\gamma, \text { eity }}$ | City-specific estimated person-year weighted subpopulation dose from neutrons and $\gamma$-rays. |
| $\operatorname{Avg}(x \mid z)$ | Average survivor true dose. |
| $\operatorname{Avg}(z \mid x)$ | Average survivor estimated dose. |
| $D_{i j, n, 66, \text { eity }}^{*}$ | Survivor neutron organ dose equivalent based on the T65DR dosimetry system. |
| $D_{i j, n, 86, \text { eity }}^{*}$ | Survivor neutron organ dose equivalent based on the DS86 dosimetry system. |
| $\Omega_{n, 65, \text { city }}$ | Average city-specific house ransmission factor for neutrons in the T65DR dosimetry system. |
| $\Omega_{n, 86, \text { city }}$ | Average city-specific house transmission factor for neutrons in the DS86 dosimetry system. |
| $D_{i j, \gamma, 65, \text { eity }}^{*}$ | Survivor $\gamma$ organ dose equivalent based on the T65DR dosimetry system. |
| $D_{i j, ~, ~, 86, \text { city }}^{*}$ | Survivor $\gamma$ organ dose equivalent based on the DS86 dosimetry system. |
| $\Omega_{\gamma, 65, \text { city }}$ | Average city-specific house transmission factor for $\gamma$-rays the T65DR dosimetry system. |
| $\Omega_{\gamma, 86, \text { city }}$ | Average city-specific house transmission factor for $\gamma$-rays the DS86 dosimetry system. |
| $D_{i j}$ | Radiation dose in $i j$ th subpopulation. |
| $D_{i j}^{*}$ | Radiation dose in $i j$ th subpopulation adjusted for DS86 random error. |
| $k_{n}$ | Neutron shielded kerma. |


| $R B E_{n}$ | Relative biological effectiveness factor for neutrons. <br> Set equal to 10 for RERF models and 20 for BEIR-V models. |
| :---: | :---: |
| $\Omega_{n, \text { city, ATB }}$ | City- and age ATB-specific body self-shielding transmission factor for neutrons. Multiplied by $k_{n}$ and $R B E_{n}$ to obtain organ dose equivalent from neutrons. |
| $k_{\gamma}$ | $\gamma$-ray shielded kerma. |
| RBE $\mathrm{F}_{7}$ | Relative biological effectiveness factor for $\gamma$-rays. Set equal to unity. |
| $\Omega_{\gamma, \text { eity, ATB }}$ | City- and age ATB-specific body self-shielding transmission factor for $\gamma$-rays. Multiplied by $k_{\gamma}$ and $R B E_{\gamma}$ to obtain organ dose equivalent from $\gamma$-rays. |
| $P Y_{i j}$ | Person-years of follow-up in subpopulation ij. |
| $\lambda_{10}$ | Cancer mortality rate in subpopulation $i$ for the zero-dose group. |
| $\lambda_{i j}$ | Cancer mortality rate in subpopulation $i$ for the $j$ th dose group. |
| $\alpha_{s}$ | Unknown nuisance parameter for background cancer rate in stratum $s$. |
| $\beta_{0}$ | Multiplicative constant term in regression model. |
| $\beta_{1}$ | Linear parameter for the contribution of dose to excess relative risk. |
| $z_{j}$ | Covariate for age ATB, age ATB, or gender in regression model. |
| $\beta_{j}$ | Regression coefficient for covariate $z_{j}$. |
| $\boldsymbol{\beta}^{T}$ | Transpose of row vector of regression coefficients $\beta_{1}, \ldots, \beta_{j}$. |
| $z$ | Row vector of covariates $z_{1}, \ldots, z_{j}$. |
| $\left(\boldsymbol{\beta}^{T} ; \boldsymbol{z}\right)$ | Linear predictor of effects for covariates $z_{1}, \ldots, z_{j}$. |
| $e^{\boldsymbol{\beta}^{7} \cdot \boldsymbol{z}}$ | Log-linear link function for linear predictor ( $\boldsymbol{\beta}^{T} ; \boldsymbol{z}$ ) |
| $f(d)$ | Dose function in BEIR-V model. |
| $g(\beta)$ | Link function for BEIR-V model. |
| $\hat{\mu}_{\text {i }}$ | Fstimated number of deaths, equal to $\lambda_{i j} \times P Y_{i j}$. |
| $y_{i}$ | Observed deaths in each subpopulation. |


| $r_{P}$ | Pearson $\chi^{2}$ residual. |
| :---: | :---: |
| $\chi^{2}$ | Chi-square goodness-of-fit statistic. Measure of model dispersion. |
| $r_{D}$ | Deviance residual. |
| D | Deviance goodness-of-fit statistic. |
| $r_{F T}$ | Freeman-Tukey residual. |
| G | Freernan-Tukey goodness-of-fit statistic. |
| $f(d)$ | Dose function for BEIR-V model. |
| $g(\beta)$ | Link function for BEIR-V model. |
| $Z_{\alpha}$ | Standard normal deviate to adjust test statistics for a Type I error. |
| $H(a)$ | Annual dose equivalent in sieverts (Sv). |
| $\Phi_{R R}(a)$ | Fitted relative risk for exposure at age $a$ in the constant RERF models. |
| $\Phi_{R R}(a, t)$ | Fitted relative risk at age at $t$ for exposure at age $a$ in the non-constant RERF and BEIR-V models. |
| $\Phi_{\text {RR,US(a) }}$ | Fitted relative risk for exposure at age a in the non-constant transported RERF model. |
| $\Phi_{\text {RR,US }}(a, t)$ | Fitted relative risk at age at $t$ for exposure at age $a$ in the non-constant transported RERF model. |
| $\Phi_{E R R}(a)$ | Excess relative risk (\%/Sv) for exposure at age $a$ in the constant RERF models. Equal to $\Phi_{R R}(a)$ minus unity. |
| $\Phi_{E R R}(a, t)$ | Excess relative risk (\%/Sv) at age $t$ for exposure at age $a$ for the RERF non-constant and BEIR-V models. Equal to $\Phi_{R R}(a, t)$ minus unity. |
| $\Phi_{E R R, U S}(a)$ | Excess relative risk (\%/Sv) for exposure at age $a$ in the constant transported RERF model. Equal to $\Phi_{\pi R, U S}(a)$ minus unity. |
| $\Phi_{E R R, U S}(a, t)$ | Excess relative risk (\%/Sv) at age $t$ for exposure at age $a$ for the non-constant transported RERF model. Equal to $\Phi_{R R, U S}(a, t)$ minus unity. |
| $\Phi_{A R}(a)$ | Absolute risk (deaths per person-year per sievert (PYSv)) for age at exposure $a$ in the constant absolute RERF models. |


| $\Phi_{A R}(a, t)$ | Absolute risk at age $t$ for exposure at age $a$ for the non-constant absolute RERF models. |
| :---: | :---: |
| $h(t ; 0)$ | Hazard function for all causes of death at age $t$ in the nonexposed population. |
| $h_{c}(t ; 0)$ | Hazard function for cancer at age $t$ in the nonexposed population. |
| $h_{c}(a ; t ; d)$ | Hazard function for radiation-induced cancer at age $t$ for exposure at age a in the exposed population. |
| $h_{c}(\infty ; t ; d)$ | Cumulative hazard function for radiation-induced cancer at age $t$ for multiple radiation exposures at ages $a_{1}, a_{2}, a_{3}, \ldots, a_{n}$ in the exposed population (Elandt-Johnson and Johnson method). |
| $q_{c}(\infty ; t ; d)$ | Attributable probability of radiation-induced cancer death at age $t$ for multiple radiation exposures at ages $a_{1}, a_{2}$, $a_{3}, \ldots, a_{n}$ in the exposed population (Bunger et al. method). |
| $q(t ; d)$ | Probability of death from all causes and radiation-induced cancer at age $t$ in the exposed population. |
| $p(t ; d)$ | Probability of surviving death from all causes and radiation-induced cancer at age $t$ in the exposed population. |
| $d(t ; d)$ | Number of deaths from all causes and radiation-induced cancer at age $t$ in the exposed population. |
| $N(t ; d)$ | Number alive at age $t$ in the exposed population. |
| $l(t ; d)$ | Person-years of life at age $t$ in the exposed population. |
| $S(t ; d)$ | Probability of surviving beyond age $t$ in the exposed population. |
| $q(t ; 0)$ | Probability of dying from all causes at age $t$ in the nonexposed population. |
| $p(t ; 0)$ | Probability of surviving death from all causes at age t in the nonexposed population. |
| $d(t ; 0)$ | Number of deaths from all causes at age $t$ in the nonexposed population. |
| $N(t ; 0)$ | Number alive at age $t$ in the nonexposed population. |
| $l(t ; 0)$ | Person-years at age $t$ in the nonexposed population. |


| $S(t ; 0)$ | Probability of surviving beyond age $t$ in the nonexposed <br> population. |
| :--- | :--- |
| $\pi(t ; d)$ | Conditional probability of death due to radiation-induced <br> cancer at age $t$ in the exposer population. |
| $\pi(t ; 0)$ | Conditional probability of death due to cancer at age $t$ in <br> the nonexposed population. |
| $\pi(\infty ; d)$ | Unconditional probability of radiation-induced death over a lifetime. <br> The number of radiation-induced cancer deaths over the lifetime <br> of the exposed population is $\pi(\infty ; d) \times 10^{5}$, since the double-decrement <br> life table staris with $N(a+L ; d)=100,000$. |
| $\pi(\infty ; 0)$ | Unconditional probability of death over a lifetime in the nonexposed <br> population. The number of deaths over the lifetime of the nonexposed <br> population is $\pi(\infty ; 0) \times 10^{5}$, since the single-decrement <br> life table starts with $N(a+L ; 0)=100,000$. |
| $Q(t ; d)$ | Unconditional probability of death due to radiation-induced cancer <br> at age $t$ in the exposed population. |
| $Q(t ; 0)$ | Unconditional probability of death due to cancer at age $t$ in the <br> nonexposed population. |
| $Y L P D$ | Years of life lost per premature radiation-induced cancer |
| death at age $t$ in the exposed population. |  |

## 9 ABBREVIATIONS

| \%/Sv | Per cent increase of risk at the 1 Sv level. This is the unit of risk for excess relative risk coefficients and lifetime risk coefficients. The risk coefficients, $\Phi_{E R R}(a), \Phi_{E R R}(a, t), \Phi_{E R R, U S}(a)$ and $\Phi_{E R R}, U S(a, t)$ are in units of $\% / \mathrm{Sv}$. If a regression coefficient is 0.5 , then the mortality rate is $50 \%$ higher in the exposed population, or, $1+0.5=1.5$ times greater than the baseline mortality rate in the nonexposed population. The value 0.5 is the excess relative risk (ERR) and the value 1.5 is the relative risk (RR). For lifetime risks, if the number of radiation-induced cancer deaths is 2,500 per 100,000 individuals, each given 1 Sv , then the excess risk is $2.5 \% / \mathrm{Sv}\left(2,500 / 10^{5} \times 100\right)$. |
| :---: | :---: |
| $10^{4} \mathrm{PYSv}$ | Person-year-dose denominator of absolute risk coefficients. The coefficients $\Phi_{A R}(a)$ and $\Phi_{A R}(a, t)$ are in units of deaths $/ 10^{4} \mathrm{PYSv}$. |
| ABCC | Atumic Bomb Casualty Commission. |
| AHS | Adult Health Study of Hiroshima and Nagasaki A-bomb survivors. |
| AMFIT | Computer program designed to fit Poisson regression models in the LSS. |
| AR | Absolute risk in deaths per $10^{4}$ person-years per Sv ( $10^{4} \mathrm{PYSv}$ ). |
| ATB | Age at-time-of-bombing. |
| ATD | Age at-time-of-death. |
| BEIR-V | NRC Committee on Biological Effects of Low Levels of Ionizing Radiation. |
| CA | Chromosome Aberrations. |
| CV | Coefficient of variation. |
| DRREF | Dose rate reduction effectiveness factor. |
| EM | Expectation-maximization algorithm used for imputing missing data. |
| ERR | Excess relative risk in \%/Sv. Equal to relative risk less unity. |
| GOF | Goodness-of-fit. |
| GSD | Geometric standard deviation. If $\sigma$ is given on the arithmetic scale, then GSD is $\exp (\sigma)$. However, if GSD is given, $\sigma$ can be determined as the natural logarithm of GSD. |
| Gy | Gray. Systems Internationale (SI) unit for radiation absorbed dose in units |



## 10 APPENDIX A. Dose-Response Modeling

### 10.1 General Approach

The two modeling approaches used in this study were the BEIR-V method (NRC,1990) and the one reported in RERF Report 11 (Shimizu et al., 1988). While the RERF method employed models that included hazards that were either constant or non-constant following exposure, the BEIR-V models were based exclusively on non-constant hazards that changed following exposure. In all of the models, the relative risk (RR), excess relative risk (ERR), and absolute risk (AR) of radiation-induced cancer are estimated at the 1 Gy level. The endpoint is a regression equation relating ERR to radiation dose equivalent and several covariates.

Maximum likelihood (ML) estimates of the regression coefficients (and standard errors) with the subgroup-specific, Poisson distributed, mortality rates (number of deaths/PY) as the dependent variable were based on commonly known procedures (Pierce et al., 1983; Pierce and Preston, 1984; Pierce and Preston, 1985; Pierce and Preston, 1988; Kleinbaum et al., 1988; Frome and Kutner, 1973; Frome, 1981; Frome, 1983). The following sections explain succinctly the methods of estimating organ dose equivalents, adjusting doses for random errors, and the coding methods used in each regression model.

### 10.2 Correction of Shielded Kerma for Random Uncertainty

The person-year weighted organ dose equivalents for each subpopulation were adjusted for Dosimetry System-86 (DS86) random errors by use of reduction factors (Peirce and Vaeth, 1991) written

$$
\begin{equation*}
R(z)_{n+\gamma, \text { eity }}=\left[z_{n+\gamma, \text { city }}-\operatorname{Avg}(x \mid z)_{n+\gamma, \text { city }}\right] / z_{n+\gamma, \text { eity }} \tag{23}
\end{equation*}
$$

where $z_{n+\gamma}$, eity is the city-specific estimated person-year weighted subpopulation dose (neutron and gamma shielded kerma in Gy), and $\operatorname{Avg}(x \mid z)_{n+\gamma, \text { city }}$ is the city-specific average true subpopulation dose at estimated dose level $z_{n+\gamma, \text { city }}$. The relationship to estimate $R(z)_{n+\gamma, \text { city }}$ for a random error of $45 \%(\sigma=0.45)$ in Hiroshima was

$$
\begin{equation*}
R(z)_{n+\gamma, \text { eity }}=0.07765+0.11770 \ln \left(z_{n+\gamma, \text { city }}\right)+0.01026 \ln ^{2}\left(z_{n+\gamma, \text { eity }}\right) \tag{24}
\end{equation*}
$$

and for Nagasaki was

$$
\begin{equation*}
R(z)_{n+\gamma, \text { city }}=0.03604+0.09612 \ln \left(z_{n+\gamma, \text { city }}\right)+0.01725 \ln ^{2}\left(z_{n+\gamma, \text { city }}\right) \tag{25}
\end{equation*}
$$

The city-specific reduction factors were applied to gamma and neutron portions of shielded kerma described in the next section.

### 10.3 Neutron Relative Biological Effectiveness and Estimation of Organ Dose Equivalents from Corrected Shielded Kerma

Estimation of the organ dose equivalents, $D_{i j}^{*}$, in sieverts ( Sv ) used in this analysis began by first applying the city-specific reduction factors, $R(z)_{n+\gamma, \text { esty }}$, to the neutron and gamma components of shielded kerma and adjusting for the Relative Biological Effectiveness factors (RBEs). This was in the form

$$
\begin{equation*}
D_{i j}^{*}=\left(1-R(z)_{n+\gamma, \text { city }}\right)\left\{\left(k_{4} \Omega_{n, \text { city, ATB }} R B E_{n}+k_{\gamma} \Omega_{\gamma, \text { city, ATB }} R B E_{\gamma}\right) 10^{3}\right\} \tag{26}
\end{equation*}
$$

where $R(z)_{n+y, \text { city }}$ is the reduction factor to adjust for DS86 random error ( $\sigma=0.45$ ), $k_{n}$ and $k_{\gamma}$ are the neutron and gamma components of shielded kerma in mGy, $\Omega_{n, \text { eity, ATB }}$ and $\Omega_{\gamma, \text { eity, } A T B}$ are the sex-specific body self-shielding transmission factors based on phantoms representing infants $(0-2$ y age ATB), children ( $3-11$ y age ATB) and adults ( $>12$ y age ATB), RBE $E_{n}$ is 10 for neutrons in the RERF models and 20 for neutrons in the BEIR-V motels, and $R B E_{y}$ is unity for gamma rays.

### 10.4 RERF Models

### 10.4.1 Relative and Excess Relative Risks

The first modeling approach to be used in this investigation followed that employed in RERF Report 11 (Shimizu et al., 1988). Write the mortality rate, $\lambda_{i j}$, in the ith stratum of city, sex and age ATB categories and $j$ th exposure category as

$$
\begin{equation*}
\lambda_{i j}=\lambda_{i 0} \Phi_{R R}(a) \tag{27}
\end{equation*}
$$

where $\lambda_{i 0}$ is the mortality rate ( $D_{c}$ /person-years $\times 10,000$ ) in the 0 dose category of the th stratum of city, sex and ATB cross-classification and $\Phi_{R R}(a)$ is the relative risk coefficient for exposure at age ATB a. Since the relative risk is related to the excess relative risk as

$$
\begin{equation*}
\Phi_{R R}(a)=\left[1+\Phi_{E R R}(a)\right] \tag{28}
\end{equation*}
$$

we can obtain maximum likelihood (ML) estimates of $\Phi_{E R R}(a)$ by first fitting a model of the form

$$
\begin{equation*}
\lambda_{i j}=\alpha_{s,} e^{\beta_{0}}\left[1+\left\{\beta_{1} D_{i j}^{*} e^{\left(\beta^{T} ; z\right)}\right\}\right] \tag{29}
\end{equation*}
$$

where $\alpha_{s}$ is an unknown nuisance parameter for the stratification of background rates ( $\lambda_{i 0}$ ) on sex ( 2 levels), city ( 2 levels), age ATB ( 13 levels) and follow-up period ( 7 levels) resulting in $i=364$ strata, $\exp \left(\beta_{0}\right)$ is a constant term, $\beta_{1}$ is the contribution of dose equivalent to excess relative risk, $D_{i j}^{*}$ is the organ dose equivalent and $z$ is a row vector of covariates representing age ATB, age ATD or gender.

If the algorithm to fit the $\alpha_{s}$ parameters in Eq. 29 were to use a $364 \times 364\left(\boldsymbol{Z}^{T} \boldsymbol{W} \boldsymbol{Z}\right)^{-1}$ weighted dispersion matrix with $132,496\left(364^{2}\right)$ elements the memory requirement would be 529,984 bytes ( 4 -bytes $\times 132,496$ ) - and this approaches the MS-DOS ${ }^{13}$ RAM limit of 640,000 bytes. A computer program, called AMFIT, can fit Eq. 29 and avoid the large memory requirement by use of a Gauss-Seidel algorithm to estimate the $364 \alpha_{s}$ terms recursively (Preston and Pierce, 1993). AMFIT uses a Newton-Raphson iteration (Kennedy and Gentle, 1980) to maximize the log-likelihood equations, and also adjusts the standard errors of the $\beta$ terms by the standard errors of the $\alpha$ terms.

### 10.4.2 Non-constant Excess Relative Risk Models for Leukemia, Nonleukemia, Stomach and Breast Sites

The ML estimates to determine $\Phi_{E R R}(a, t)$ for leukemia, nonleukemia, stomach and breast sites were obtained by regressing the mortality rate in the exposed subgroup $\lambda_{i j}$ with the relationship
$\lambda_{i j}=\alpha_{0,} e^{\beta_{0}}\left[1+\left\{\beta_{1} D_{i j}^{*} e^{\beta_{2} z_{2}+\beta_{3} z_{3}+\beta_{4} z_{4}+\beta_{5} z_{8}+\beta_{6} z_{6}+\beta_{7} z_{7}+\beta_{6} z_{8}+\beta_{9} z_{9}+\beta_{10} z_{10}+\beta_{11} z_{11}+\beta_{12} z_{13}+\beta_{13} z_{18}}\right\}\right]$

[^13]where $\alpha_{s}$ is for stratification of background rates ( $\lambda_{i 0}$ ) on sex (2 levels), city ( 2 levels), age ATB ( 13 levels) and follow-up period ( 7 levels), $\exp \left(\beta_{0}\right)$ is a constant term, $\beta_{1}$ is the linear coefficient for dose equivalent in $S \mathrm{v}, D_{i j}^{*}$ is the organ dose equivalent, $z_{2}$ is coded with a 1 when the age ATB is $10-19$ and 0 otherwise, $z_{3}$ is coded 1 when the age ATB is $20-29$ and 0 otherwise, $z_{4}$ is coded 1 when the age ATB is $30-39$ and 0 otherwise, $z_{5}$ is coded 1 when the age ATB is $40-49$, and 0 otherwise, $z_{6}$ is coded with a 1 when the age ATB is $50+$ and 0 otherwise, $z_{7}$ is coded 1 when the age ATD is $20-29$ and 0 otherwise, $z_{8}$ is coded 1 when the age ATD is $30-39$ and 0 otherwise, $z_{9}$ is coded 1 when the age ATD is $40-49$, and $z_{10}$ is coded 1 when the age ATD is $50-59, z_{11}$ is coded with a 1 when the age ATD is 60-69 and 0 otherwise, $z_{12}$ is cored 1 when the age ATD is $70+$ and 0 otherwise and $z_{13}$ is coded 1 for males and 0 for temales. The 0-9 age ATB and 0-19 age ATD stratum for females is the corner-point w tere $z_{2}-z_{13}$ are all dummy coded with zeros. Organ dose equivalents for leukemia, nonleukemia, stomach and breast sites were based on bone marrow, large intestine, stomach and breast body self-shielding transmission factors, respectively, with a neutron RBE of 10 . The stomach site did not contain a parameter for gender and was fitted separately for each sex.

### 10.4.3 Constant Excess Relative Risk Models for Lung, Bladder, Liver, Colon, and Ovary Sites

Since there are fewer deaths for the lung, bladder, liver, colon and ovary sites, it was necessary to use a model with fewer covariates so that the scores (Rao, 1947), Wald tests (Wald, 1943), and Pearson, deviance, and Freeman-Tukey goodness-of-fit tests would remain acceptable (Nelder a.d McCullagh, 1989; Freeman and Tukey, 1950; Santner and Duffy, 1989). The ML estimates to determine $\Phi_{E R R}(a)$ for these sites were obtained by regressing the mortality rate in the exposed subgroup $\lambda_{i j}$ with the relationship

$$
\begin{equation*}
\lambda_{i j}=\alpha_{s,} e^{\beta_{0}}\left[1+\left\{\beta_{1} D_{i j}^{*} e^{\beta_{2} z_{2}+\beta_{3} z_{3}+\beta_{4} z_{4}+\beta_{3} z_{0}}\right\}\right] \tag{31}
\end{equation*}
$$

where $\alpha$, is for stratification of background rates ( $\lambda_{i 0}$ ) on sex (2 levels), city ( 2 levels), age ATB ( 13 levels) and follow-up period ( 7 levels), $\exp \left(\beta_{0}\right)$ is a constant term, $\beta_{1}$ is the linear coefficient for dose equivalent in $\mathrm{Sv}, D_{i j}^{*}$ is the organ dose equivalent, $z_{2}$ is coded with a 1 when the age ATB is $20-29$ and 0 otherwise, $z_{3}$ is coded 1 when the age ATB is $30-39$ and 0 otherwise, $z_{4}$ is coded 1 when the age ATB is $40+$ and 0 otherwise, $z_{5}$ is coded 1 for males and 0 for females if the model contains a gender parameter. The $0-19$ age ATB exposure category for females is the corner-point where $z_{2}-z_{5}$ are all dummy coded with zeros. Organ dose equivalents for the lung, bladder, liver, colon and ovary sites were based on lung, urinary bladder, liver, large intestine and ovary body self-shielding transmission factors, respectively, with a neutron RBE of 10 .

### 10.4.4 Determining Excess Relative Risk from Regression Coefficients

Once the ML estimates of parameters were obtained, the non-constant, $\Phi_{E R R}(a, t)$, or constant, $\Phi_{E R R}(a)$, excess relative risk at the 1 Sv level for a given age ATB group and gender were calculated by multiplying the exponent of the sum of the respective age ATB and sex coefficents by the linear coefficent for the dose equivalent $\beta_{1}$. As an example, the linear predictor $\left(\beta^{T} ; \boldsymbol{z}\right)$ for a given age ATB group and gender in a constant excess n rative risk model was obtained by cross-multiplying the transposed column vector of coefficients
of covariates and row vector of predictor values of the form

$$
\left(\boldsymbol{\beta}^{T} ; \boldsymbol{z}\right)=\left(\begin{array}{c}
\hat{\beta}_{2}  \tag{32}\\
\hat{\beta}_{3} \\
\hat{\boldsymbol{\beta}}_{4} \\
\hat{\beta}_{5}
\end{array}\right)\left(\begin{array}{llll}
z_{2} & z_{3} & z_{4} & z_{5}
\end{array}\right)
$$

For example, the linear predictor $\left(\boldsymbol{\beta}^{\boldsymbol{T}} ; \boldsymbol{z}\right)$ for the male age ATB group $30-39$ is defined in the form

$$
\left(\boldsymbol{\beta}^{T} ; \boldsymbol{z}\right)=\left(\begin{array}{l}
\hat{\beta}_{2}  \tag{33}\\
\hat{\beta}_{3} \\
\hat{\beta}_{4} \\
\hat{\beta}_{5}
\end{array}\right)\left(\begin{array}{llll}
0 & 1 & 0 & 1
\end{array}\right)
$$

which when substituted into Eq. 30 yields gives

$$
\begin{equation*}
\Phi_{E R R}(a)=\hat{\beta}_{1} e^{\left(\boldsymbol{\beta}^{T}: z\right)} \tag{34}
\end{equation*}
$$

We notice that $D_{i j}^{*}$ is not included in Eq. 34 bece.use the unit of dose during regression was Sv . As an example, if the relationship between blood pressure and age is such that each year of life increases blood pressure by one mm of Hg , then when regressing blood pressure on age, the regression coefficient for age, $\beta_{\text {age }}$, would be equal to one because of the one-to-one relationship. Therefore, in Eq. 34, the linear regression coefficient for dose, $\beta_{1}$, represents the per cent change in risk per one Sv in units of $\% / \mathrm{Sv}$ and $D_{i j}^{*}$ is no longer needed when estimating $\Phi_{E R R}(a)$.

For constant hazard models this was done for the three age ATB groups (20-29, 30-39, and $40+$ ) and two genders (females and males) for the lung, bladder, liver, colon and ovary sites. Similar matrix operations were done for the non-constant excess relative risk models for leukemia, nonleukemia, stomach and breast cancer mortality.

### 10.4.5 Determining Absolute Risks from Regression Coefficients

The number of excess deaths per $10^{4}$ person-years at the 1 Sv level for the constant AR model were estimated by use of the formula

$$
\begin{equation*}
\Phi_{A R}(a)=\left(\sum_{i} \sum_{k}\left(P Y_{i j} \lambda_{i 0} \Phi_{E R R}(a) D_{i j}^{*}\right) / \sum_{i} \sum_{j}\left(P Y_{i j} D_{i j}^{*}\right)\right) \times 10^{4} \tag{35}
\end{equation*}
$$

where $P Y$ is the person-years of follow-up and $\Phi_{E R R}(a)$ is the constant ERR from a regression model containing Age ATB and sex parameters. When non-constant regression models contained age ATB, age ATD and sex parameters to estimate $\Phi_{E R R}(a ; t)$, the AR coefficients were determined as

$$
\begin{equation*}
\Phi_{A R}(a, t)=\left(\sum_{i} \sum_{k}\left(P Y_{i j} \lambda_{i 0} \Phi_{E R R}(a, t) D_{i j}^{*}\right) / \sum_{i} \sum_{j}\left(P Y_{i j} D_{i j}^{*}\right)\right) \times 10^{4} \tag{36}
\end{equation*}
$$

Coefficients (excess deaths $/ 10^{4} \mathrm{PYSv}$ ) for $\Phi_{A R}(a, t)$ were estimated for the leukemia, nonleukemia, stomach and breast sites with neutron RBEs of 10. For the lung, bladder, liver, colon and ovary sites, only $\Phi_{A R}(a)$ were estimated because the regression models did not include an age ATD term. Absolute risks were not estimated from regression coefficients of the BEIR-V models.

### 10.5 BEIR-V Models

The linear additive relative risk for each exposed subpopulation of LSS survivors using the BEIR-V linear model is

$$
\begin{equation*}
\lambda_{i j}=\alpha_{s i} e^{\beta_{0}}\left[1+\left\{\beta_{1} D_{i j}^{*} e^{\left(\boldsymbol{\beta}^{T}: \boldsymbol{z}\right)}\right\}\right] \tag{37}
\end{equation*}
$$

where $\alpha_{s}$ is for stratification of background rates ( $\lambda_{i 0}$ ) on sex (2 levels), city (2 levels), age ATB ( 13 levels) and follow-up period ( 7 levels), $\exp \left(\beta_{0}\right)$ is a constant term, $\beta_{1}$ is the contribution of the dose term to the excess relative risk and $z$ is a row vector of covariates for sex, mean age at exposure (E) and time since exposure (TSE). The linear-quadratic model is

$$
\begin{equation*}
\lambda_{i j}=\alpha_{s,} e^{\beta_{0}}\left[1+\left\{\left(\beta_{1} D_{i j}^{*}+\beta_{2} D_{i j}^{* 2}\right) e^{\left(\beta^{T} ; z\right)}\right\}\right] \tag{38}
\end{equation*}
$$

where $\beta_{1}$ and $\beta_{2}$ represent the dose and dose-squared contribution to excess relative risk. The $\Phi_{R R}(a, t)$ for the same exposed subpopulation in Eqs. 37 and 38 can be rewritten

$$
\begin{equation*}
\Phi_{R R}(a, t)=1+\Phi_{E R R}(a, t) \tag{39}
\end{equation*}
$$

The BEIR-V committee defined $\Phi_{E R R}(a, t)$ in the above equation as

$$
\begin{equation*}
\Phi_{E R R}(a, t)=f(d) g(\beta) \tag{40}
\end{equation*}
$$

where $f(d)$ is a function of either the linear $\left(\beta_{1} D_{i j}^{*}\right)$ or inear-quadratic $\left(\beta_{1} D_{i j}^{*}+\beta_{2} D_{i j}^{* 2}\right)$ contribution of radiation dose and $g(\beta)$ is a link function for sex, age ATB, and time-since-exposure (TSE). The above models were used for fitting excess relative risk lenkemia, respiratory cancers, breast cancer, digestive cancers and "other" cancers not included in the ICD rubric of malignent neoplasms.

### 10.5.1 BEIR-V Leukemia model

For modeling leukemia we choose to evaluate only the RR for the $L$ and $L Q$ models as a function of age ATB $<20$ years and age ATB $>20$ since there are so many structural zeros (empty cells in the cross-classified data). There is no need to adjust for a latency period for leukemia because the first follow-up in the LSS occurred five years afters the bombings. Cases for which the bone marrow dose equivalent (neutron $\mathrm{RBE}=20$ ) exceeded 4 Sv and TSE>75y were excluded. The Poisson regression model used for modeling the BEIR-V leukemia $\Phi_{R R}(a, t)$ was

$$
\begin{equation*}
\lambda_{i j}=\alpha_{s i} e^{\beta_{0}}\left[1+\left\{\left(\beta_{1} D_{i j}^{*}+\beta_{2} D_{i j}^{*}\right) e^{\beta_{3} z_{1}+\beta_{4} z_{3}+\beta_{8} z_{0}+\hat{\beta}_{6} z_{4}}\right\}\right] \tag{41}
\end{equation*}
$$

where $\alpha_{s}$ is for stratification of background rates ( $\lambda_{i 0}$ ) on sex (2 levels), city (2 levels), age ATB ( 13 levels) and follow-up period ( 7 levels), $\exp \left(\beta_{0}\right)$ is a constant term, $\beta_{1}$ is the linear term for dose equivalent, $\beta_{2}$ is the quadratic term fcr dose equivalent, $D_{i j}^{*}$ is the marrow dose equivalent, $z_{1}$ is an indicator variable coded as a one when $\operatorname{TSE} \leq 15$ and age $\mathrm{ATB} \leq 20, z_{2}$ is an indicator variable coded as one when $15<\mathrm{TSE} \leq 25$ and age ATB $\leq 20$, $z_{3}$ is an indicator variable coded as one when TSE $\leq 25$ and age ATB>20, and $z_{4}$ is an indicator variable when $25<\mathrm{TSE} \leq 30$ and age ATB $>20$.

### 10.5.2 -EIR-V Breast Model

For the breast the RR was modeled only for the L model at the 1 Sv level. Cases for which the breast dose equivalent (neutron $\mathrm{RBE}=20$ ) exceeded 4 Sv and $\mathrm{TSE}>75$ were excluded.

The TSE was normalized to a TSE of 20 , and cases with TSE<10 were excluded from the analysis. The Poisson regression model used for modeling the BEIR-V breast $\Phi_{R R}(a, t)$ was

$$
\begin{equation*}
\lambda_{i j}=\alpha_{s,} e^{\beta_{0}}\left[1+\left\{\beta_{1} D_{i j}^{*} e^{\beta_{2} z_{1}+\beta_{3} z_{2}+\beta_{4} z_{3}+\beta_{5} z_{4}}\right\}\right] \tag{42}
\end{equation*}
$$

where $\alpha_{s}$ is for stratification of background rates ( $\lambda_{i 0}$ ) on sex ( 2 levels), city ( 2 levels), age ATB ( 13 levels) and follow-up period ( 7 levels), $\exp \left(\beta_{0}\right)$ is a constant term, $\beta_{1}$ is the linear coefficient for dose equivalent, $D_{i j}^{*}$ is the breast dose equivalent, $z_{1}$ is a column vecior of ones, $z_{2}$ is a covariate set to $\ln (T S E / 20)$ when age ATB<i5, $z_{3}$ is a covariate set to $\ln ^{2}(T S E / 20)$ when age ATB $<15$, and $z_{4}$ is a covariate set to age ATB- 15 when age ATB15.

### 10.5.3 BEIR-V Respiratory Model

The lung model took into account a sex effect and an age ATD effect. Sex was dummy coded into male and female groups. TSE was normalized to a TSE of 20. RR was estimated for the L model at the 1 Sv level. Cases for which the lung dose equivalent (neutron RBE $=20$ ) exceeded 4 Sv and TSE $>75$ were excluded. Cases for which TSE $<10$ years were excluded. The Poisson regression model used for modeling the BEIR-V respiratory $\Phi_{R R}(a, t)$ was

$$
\begin{equation*}
\lambda_{i j}=\alpha_{s,} e^{\beta_{0}}\left[1+\left\{\beta_{1} D_{i j}^{*} e^{\beta_{2} z_{1}+\beta_{3} z_{2}}\right\}\right] \tag{43}
\end{equation*}
$$

where $\alpha_{s}$ is for stratification of background rates ( $\lambda_{i 0}$ ) on sex (2 levels), city (2 levels), age ATB ( 13 levels) and follow-up period ( 7 levels), $\exp \left(\beta_{0}\right)$ is a constant term, $\beta_{1}$ is the linear coefficient for dose equivalent, $D_{i j}^{*}$ is the lung dose equivalent, $z_{1}$ is a covariate set to $\ln (T S E / 20)$ and $z_{2}$ is a covariate for gender set to one for females and zero for males, independent of age ATB.

### 10.5.4 BEIR-V Digestive Model

Modeling mortality from digestive cancer included a sex effect and age ATB effect. Sex was coded into male and female groups. Age ATB was coded into 3 separate groups representing age ATB $<25,25<$ age ATB $<35$, and age ATB $>35$ since the BEIR-V committee reported age $A T B$ to be quite significant. Cases for which the stomach dose equivalent (neutron RBE $=20$ ) exceeded 4 Sv and TSE $>75$ were excluded. Cases for which TSE $<10$ years were excluded.

The Poisson regression model used for modeling the BEIR-V digestive $\Phi_{R R}(a, t)$ was

$$
\begin{equation*}
\lambda_{i j}=\alpha_{s i} e^{\beta_{0}}\left[1+\left\{\beta_{1} D_{i j}^{*} e^{\beta_{2} z_{1}+\beta_{3} \sigma_{E}}\right\}\right] \tag{44}
\end{equation*}
$$

where $\alpha$, is for stratification of background rates $\left(\lambda_{i 0}\right)$ on sex (2 levels), city (2 levels), age ATB ( 13 levels) and follow-up period ( 7 levels), $\exp \left(\beta_{0}\right)$ is a constant term, $\beta_{1}$ is the linear coefficient for dose equivalent, $D_{i j}^{*}$ is the stomach dose equivalent, $z_{1}$ is a covariate for gender set to 1 for females and zero for males, and $\sigma_{E}$ is a covariate for age ATB set to zero if age ATB $\leq 25$, $(\mathrm{E}-25)$ when age ATB is $>25$ and $\leq 35$, and 10 when age ATB $>35$.

### 10.5.5 BEIR-V Other Cancers Model

Radiation-induced mortality in the remaining sites will only account for age ATB effects. Cases for which the stomach dose equivalent (neutron $\mathrm{RBE}=20$ ) exceeded 4 Sv and TSE $>75$ were excluded. Cases for which TSE $<10$ years were excluded to account for an
assumed minimum latency period. The Poisson regression model used for modeling the BEIR-V breast $\Phi_{R R}(a, t)$ was

$$
\begin{equation*}
\lambda_{i j}=\alpha_{s i} e^{\beta_{0}}\left[1+\left\{\beta_{1} D_{i j}^{*} e^{\beta_{2} x_{1}}\right\}\right] \tag{45}
\end{equation*}
$$

where $\alpha_{s}$ is for stratification of background rates $\left(\lambda_{i 0}\right)$ on sex ( 2 levels), city ( 2 levels), age ATB ( 13 levels) and follow-up period ( 7 levels), $\exp \left(\beta_{0}\right)$ is a constant term, $\beta_{1}$ is the linear coefficient for dose equivalent, $D_{i j}^{*}$ is the stomach dose equivalent and $z_{1}$ is a covariate for age ATB set to one if age ATB $\leq 10$ and E-10 if age ATB $>10$.

### 10.6 Regression Diagnostics and Goodness-of-Fit (GOF)

The goodness-of-fit (GOF) of each model was estimated to determine the degree of concordance of the model under consideration with the data (Rayner and Best, 1989). Aggregate statistics to determine concordance were based on the squared difference between the observed, $y_{i}$, and fitted values, $\hat{\mu}_{i}$, of the number of deaths in each subpopulation. Cressie and Read (1984) introduced the power divergence family of GOF test statistics employed in this study. When $\hat{\mu}_{i} \geq 5$ for all $i$ then Pearson $\chi^{2}$ residuals

$$
\begin{equation*}
r_{P}=\left(y_{i}-\hat{\mu}_{i}\right)^{2} / \hat{\mu}_{i} \tag{46}
\end{equation*}
$$

and $\chi^{2}$ GOF statistic $\sum r_{P}^{2}$ are adequate measures of dispersion. If all $\hat{\mu}_{i} \leq 1$ or $\hat{\mu}_{i} \rightarrow 0$, then deviance residuals

$$
\begin{equation*}
r_{D}=2\left[y_{i} \log \frac{y_{i}}{\mu_{i}}\right]^{1 / 2} \tag{47}
\end{equation*}
$$

and (deviance GOF $D=\sum r_{D}$ ) Freeman-Tukey, $r_{F T}$, residuals

$$
\begin{equation*}
r_{F T}=\sqrt{y_{i}}+\sqrt{y_{i}+1}-\sqrt{4 \mu_{i}+1} \tag{48}
\end{equation*}
$$

and statistic $G=\sum r_{F T}^{2}$ are more appropriate for assessing GOF.
Numerically, the residuals are

$$
\begin{equation*}
r_{P}=\left(\left(\max \left(y_{i}, 10^{-12}\right)-\max \left(\hat{\mu}_{1}, 10^{-12}\right) / \sqrt{\max \left(\hat{\mu}_{i}, 10^{-12}\right)}\right)\right. \tag{49}
\end{equation*}
$$

where $\max \left(y_{i}, 10^{-12}\right)$ is the larger of the two values $y_{i}$ and $10^{-12}$ and $\max \left(\hat{\mu}_{i}, 10^{-12}\right)$ is the larger of the two values $\hat{\mu}_{i}$ and $10^{-12}$. The deviance is in the form

$$
\begin{equation*}
r_{D}=\max \left(y_{i}, 10^{-12}\right)\left[\ln \left(\max \left(y_{i}, 10^{-12}\right)-\ln \left(\max \left(\hat{\mu}_{i}, 10^{-12}\right)\right)\right]\right. \tag{50}
\end{equation*}
$$

where $\max \left(y_{i}, 10^{-12}\right)$ and $\max \left(\hat{\mu}_{i}, 10^{-12}\right)$ are defined above. Lastly, the Freeman-Tukey residuals were determined as

$$
\begin{equation*}
r_{F T}=\left(\sqrt{\max \left(y_{i}, 10^{-12}\right)}+\sqrt{\left(\max \left(y_{i}, 10^{-12}\right)+1\right.}-\sqrt{4 \max \left(\hat{\mu}_{i}, 10^{-12}\right)-1}\right. \tag{51}
\end{equation*}
$$

Under the null hypothesis, $\chi^{2}=\sum r_{P}^{2}=D=\sum r_{D}^{2}=G=\sum r_{F T}^{2} \sim \chi_{n-s-p}^{2}$. Values of $\chi^{2}, D$, and $G$ that are less than $n-s-p$ represent models that "fit" the data and will typically result in tail probabilities $\geq 0.25$; a perfect fit will yield a tail probability of unity (see Algorithm AS32 in the references).

## 11 APPENDIX B. Lifetime Risk Projection

### 11.1 Introduction

The lifetime mortality risk of multiple exposures to radiation is quantified by applying the risks from each age at exposure to the total force of mortality experienced over a lifetime. In one sense, we are applying radiation risk coefficients obtained from the follow-up of a bona fide exposed cohort to the survival of a theoretically exposed population whose mortality increases proportionally with baseline cancer rates (relative projection model) or independently of baseline cancer rates (absolute projection model). The following sections will explain succinctly the complexities involved in calculating the lifetime risks of radiation-induced cancer mortality.

### 11.2 Hazard Functions for Radiation-Induced Cancer

First define $a$ as the age at exposure for an exposed population. The hagard of radiationinduced cancer at age $t$ from exposure at age $a$ for the constant relative model is

$$
\begin{equation*}
h_{c}(a ; t ; d)=H(a) \Phi_{E R R}(a) h_{c}(t ; 0) \tag{52}
\end{equation*}
$$

where $H(a)$ is the annual dose (Sv) at age is, $\Phi_{E R R}(a)$ is the excess relative risk at age $a$ and $h_{c}(t ; 0)$ is the hazard rate for spontzieously occurring cancer at age $t$. The hazard function for radiation-induced cancer at age $t$ from multiple exposures at various ages is written

$$
\begin{equation*}
h_{e}(\infty ; t ; d)=\int_{t-p}^{t-1} H(a) \Phi_{E R R}(a) h_{c}(t ; 0) d a \tag{53}
\end{equation*}
$$

where the integrands are defined in Eq. 52. The upper limit of integration $t-l$ prevents integration at ages beyond the plateau period and the lower limit prevents integration below the minimal latency period (Checkoway et al., 1989). When using risk coefficients that are only age at-time-of-bombing (ATB) specific, $\Phi_{E R R}(a)$ for exposure at age a remains constant for all subsequent age intervals. However, when using risk coefficients that are age ATB and time-since-exposure (TSE) specific, then $\Phi_{E R R}(a)$ changes and is termed $\Phi_{E R R}(a, t)$ to indicate the hazard at age $t$ from exposure at age $a$. The terms $l$ and $p$ in the limits of integration of Eq. 53 represent the beginning (minimum latency) and end of the plateau period for exposure at age $a$.

Risk projection for each age interval under the constant absolute model is similar to that of the constant relative model, however the absolute risk (deaths/person-year-Sv), $\Phi_{A R}(\alpha)$, for exposure at age $a$ is applied to the dose equivalent, $H(a)$, received at age $a$ in the absence of baseline cancer mortality rates. Thus, Eq. 52 becomes

$$
\begin{equation*}
h_{e}(a ; t ; d)=H(a) \Phi_{A R}(a) \tag{54}
\end{equation*}
$$

and Eq. 53 becomes

$$
\begin{equation*}
h_{c}(\infty ; t ; d)=\int_{t-p}^{t-1} H(a) \Phi_{A R}(a) d a \tag{55}
\end{equation*}
$$

### 11.3 Double-Decrement Life Table (Radiation-Induced Cancers)

We recall that for a double-decrement life table (Elandt-Johnson and Johnson, 1980) the conditional death probability, $q(t ; d)$ in age interval $(t, t+1)$ due to the combination of death from all causes in the absence of exposure and deaths due to radiation-incuced cancer is

$$
\begin{equation*}
q(t ; d)=\frac{2\left(h(t ; 0)+h_{c}(\infty ; t ; d)\right)}{2+\left(h(t ; 0)+h_{c}(\infty ; t ; d)\right)} \tag{56}
\end{equation*}
$$

where $h(t ; 0)$ is the age-specific central death rate due to all causes in the absence of exposure and $h_{e}(\infty ; t ; d)$ is the total age-specific central death rate for cancer due to radiation exposure (Eqs. 53 and 55 ). The conditional probability that an individual will not die in the interval $(t, t+1)$ is

$$
\begin{equation*}
p(t ; d)=1-q(t ; d) \tag{57}
\end{equation*}
$$

and the number of expected deaths from radiation-induced cancer and all causes in the absence of exposure is

$$
\begin{equation*}
d(t ; d)=q(t ; d) N(t ; d) \tag{58}
\end{equation*}
$$

The expected number of survivors, $N(t ; d)$ in interval $(t, t+1)$ out of a population of $N(a+L ; d)$ is found recursively as

$$
\begin{equation*}
N(t ; d)=N(t-1 ; d)-d(t-1 ; d) \tag{59}
\end{equation*}
$$

and the number of person-years in each interval $(t, t+1)$ is approximated by

$$
\begin{equation*}
l(t ; d)=N(t ; d)-\frac{1}{2} d(t ; d) \tag{60}
\end{equation*}
$$

The survivorship function (Chiang, 1968; Chiang, 1984; Smith et al., 1970; Lee, 1980) or curnulative probability of surviving beyond each interval is estimated with the equation

$$
\begin{equation*}
S(t ; d)=\prod_{y=0}^{t-1}(1-q(t ; d))=\prod_{y=0}^{t-1} p(t ; d) \tag{61}
\end{equation*}
$$

which is used later for estimating the lifetime risks of radiation-induced cancer in an exposed working population.

### 11.4 Single-Decrement Life Table (Baseline cancers)

Whereas the double-decrement life table provides estimates of radiation-indu ad cancer mortality, the single-decrement life table is applied to obtain estimates of baseline (spontaneous) cancer mortality risks over a career or lifetime. The probability and number of baseline cancers for the relevant projection periods are calculated the same way as the number of radiation-induced cancers was determined. In this instance, Eq. 53 is rearranged to

$$
\begin{equation*}
q(t ; 0)=\frac{2 h(t ; 0)}{2+(h(t ; 0)} \tag{62}
\end{equation*}
$$

In the absence of radiation exposure, the conditional probability that an individual will not die in the interval $(t, t+1)$ is

$$
\begin{equation*}
p(t ; 0)=1-q(t ; 0) \tag{63}
\end{equation*}
$$

and the number of expected deaths from all causes in the absence of exposure is

$$
\begin{equation*}
d(t ; 0)=q(t ; 0) N(t ; 0) . \tag{64}
\end{equation*}
$$

The expected number of survivors, $N(t ; 0)$, in interval $(t, t+1)$ out of a population of $N(a+L ; 0)$ nonexposed workers is found recursively as

$$
\begin{equation*}
N(t ; 0)=N(t-1 ; 0)-d(t-1 ; 0) \tag{65}
\end{equation*}
$$

and the number of person-years in each interval $(t, t+1)$ is approximated by

$$
\begin{equation*}
l(t ; 0)=N(t ; 0)-\frac{1}{2} d(t ; 0) \tag{66}
\end{equation*}
$$

The cumulative probability of surviving beyond each interval (survivorship function) is estimated with the equation

$$
\begin{equation*}
S(t ; 0)=\prod_{y=0}^{t-1}(1-q(t ; 0))=\prod_{y=0}^{t-1} p(t ; 0) \tag{67}
\end{equation*}
$$

The above parameters are endpoints that are used for determining the lifetime risks of baseline cancers in a nonexposed population. The next two sections describe the method for obtaining lifetime risks.

### 11.5 Lifetime Risks Based on Method of Elandt-Johnson and Johnson

The conditional probability of death due to radiation-induced cancer is estimated using the formula

$$
\begin{equation*}
\pi(t ; d)=h_{c}(\infty ; t ; d) S(t ; d) \tag{68}
\end{equation*}
$$

where $h_{c}(\infty ; t ; d)$ is the hazard function defined in Eqs. 53 and 55 and $S(t ; d)$ is the survivorship function from the double-decrement life table (Eq. 61) for the exposed population. The unconditional probability of death due to radiation-induced cancer at age $t$ is

$$
\begin{equation*}
Q(t ; d)=\int_{0}^{t} \pi(x ; d) d x=\int_{0}^{t} h_{c}(\infty ; x ; d) S(x ; d) d x \tag{69}
\end{equation*}
$$

Over a lifetime, the unconditional probability of radiation-induced cancer mortality ior an exposed population over a lifetime is

$$
\begin{equation*}
\pi(\infty ; d)=\int_{0}^{\infty} \pi(x ; d) d x=\int_{0}^{\infty} h_{c}(\infty ; x ; d) S(x ; d) d x \tag{70}
\end{equation*}
$$

where $\infty$ is by convention 100 years of age. The number of radiation-induced cancer deaths (per $10^{5}$ exposed individuals) is $\pi(\infty ; d) \times 10^{5}$.

The unconditional death probability for the constant RR risk projection model was based on applying ERR coefficients obtained in this study directly to baseline (spontaneously occurring) cancer rates and life tables for the U.S. population. This was functionally composed by substituting the integrands of Eq. 53 into Eq. 70 as

$$
\begin{equation*}
\pi(\infty ; d)=\int_{0}^{\infty} \int_{t-p}^{t-1} H(a) \Phi_{E R R}(a) h_{c}(t ; 0) S(t ; d) d a d t \tag{71}
\end{equation*}
$$

where $\infty$ is by convention 100 years of age, $t-p$ prevents integration below the minimal latency period for the first (or only) age at exposure $a, t-l$ prevents integration beyond the plateau period for the last age at exposure (Checkoway et al., 1989), $H(a)$ is the annual dose equivalent in $\mathrm{Sv}, \Phi_{E R R}(a)$ is the sex- and age ATB-specific excess risk coefficient (\%/Sv) from $£ 3.3, h_{c}(t ; 0)$ is the hazard rate of spontaneously occuring cancer in the interval $(t, t+1)$ and $S(t ; d)$ is the all-cause survivorship function for each one-year interval of the complete life table. The number of radiation-induced cancer deaths per 100,000 exposed individuals is $\pi(\infty ; d) \times 10^{5}$.

The unconditional death probability for the non-constant RERF and BEIR-V relative models were based on sex-, age ATB- and either age ATD- or TSE-specific ERR coefficients obtained in this study in the form

$$
\begin{equation*}
\pi(\infty ; d)=\int_{0}^{\infty} \int_{t-p}^{t-1} H(a) \Phi_{E R R}(a, t) h_{c}(t ; 0) S(t ; d) d a d t \tag{72}
\end{equation*}
$$

where $\Phi_{\text {ERR }}(a, t)$ is the ERR risk coefficient at age $t$ for exposure at age $a$.
The unconditional probability, $\pi(\infty ; d)$, of radiation-induced cancer mortality over a lifetime for the constant AR model is obtained by substitution of integrands of Eq. 55 into Eq. 70 in the form

$$
\begin{equation*}
\pi(\infty ; d)=\int_{0}^{\infty} \int_{t-p}^{t-1} H(a) \Phi_{A R}(a) S(t ; d) d a d t \tag{73}
\end{equation*}
$$

where $\Phi_{A R}(a)$ is the sex-, age ATB- and/or age ATD-specific absolute risk coefficient (deaths $/ 10^{4} \mathrm{PYSv}$ ) from $\S 3.3$, and $S(t ; d)$ is the all-cause survivorship function for each one-year interval of the complete life table. Non-constant absolute unconditional probabilities were estimated with $\Phi_{A R}(a, t)$ using the equation

$$
\begin{equation*}
\pi(\infty ; d)=\int_{0}^{\infty} \int_{t-p}^{t-1} H(a) \Phi_{A R}(a, t) S(t ; d) d a d t \tag{74}
\end{equation*}
$$

Unconditional probabilities based on the constant transported RR(AR) model were calculated with the formula

$$
\begin{equation*}
\pi(\infty ; d)=\int_{0}^{\infty} \int_{t-p}^{t-1} H(a) \Phi_{E R R, U S}(a) h_{e}(t ; 0) S(t ; d) d a d t \tag{75}
\end{equation*}
$$

where the integrand $\Phi_{E R R, U S}(a)$ is based on the relationship

$$
\begin{equation*}
\Phi_{E R R, U S}(a)=\frac{\int_{a+5}^{a+40} \Phi_{A R}(a) h_{\mathrm{e}}(t ; 0) S(t ; d) d t}{\int_{a+5}^{a+40} h_{\mathrm{c}}(t ; 0) S(t ; 0) d t} \tag{76}
\end{equation*}
$$

over the relevant 35 -year ( $1950-85$ ) follow-up period in the LSS from $a+5$ to $a+40$ where $h_{e}(t ; 0)$ is the baseline cancer rate for spontaneously occurring cancer at age $t$ and $S(t ; d)$ and $S(t ; 0)$ are the survivorship functions for the radiation exposed (Eq. 62) and the nonexposed populations (Eq. 67), respectively. For unconditional probabilities based on the
non-constant transported relative model of $\Phi_{E R R, U S}(a, t)$, lifetime risks were calcula ved with the formula

$$
\begin{equation*}
\pi(\infty ; d)=\int_{0}^{\infty} \int_{t-p}^{t-1} H(a) \Phi_{E R R, U S}(a, t) h_{c}(t ; 0) S(t ; d) d a d t \tag{77}
\end{equation*}
$$

where the integrand $\Phi_{E R R, U S}(a, t)$ is based on the relationship

$$
\begin{equation*}
\Phi_{E R R, U S}(a, t)=\frac{\int_{a+5}^{a+40} \Phi_{A R}(a, t) h_{c}(z ; 0) S(x ; d) d x}{\int_{a+5}^{a+40} h_{c}(t ; 0) S(t ; 0) d t} \tag{78}
\end{equation*}
$$

It is noteworthy to point out that the Risk of Exposure-Induced L•ath (REID) iutroduced by Thomas et al. (1992) as

$$
\begin{equation*}
R E I D_{c}(e, D)=\int_{e}^{\infty}\left[\mu_{c}(a \mid e, D)-\mu_{c}(a)\right] S(a \mid e, D) d a \tag{79}
\end{equation*}
$$

is equivalent to $\pi(\infty ; d)$ because the hazard function $h_{e}(\infty ; x ; d)$ in Eqs. 53 and 55 does not include the baseline hazard function $h_{c}(t ; 0)$ for spontaneously occurring cancer. Thus, the hazard functions in Eqs. 6 and 7 of Thomas et al. would be stated in this report as

$$
\begin{equation*}
\mu_{e}(a, e, t, s, y, D)=\beta(a, e, t, s) g(D) \tag{80}
\end{equation*}
$$

for the additive projection model and

$$
\begin{equation*}
\mu_{c}(a, e, t, s, y, D)=\mu(a, y), \mu(a, e, t, s) g(D) \tag{81}
\end{equation*}
$$

for the multiplicative model. Results of the Elandt-Johnson and Johnson (1980) method of estimating lifetime risks have been found to be similar to those estimated by the Bunger et al. (1981) and Gail (1975) methods because the SURVRAD algorithm implements all three methods of estimation (Peterson et al., 1992). The only difference between the Elandt-Johnson and Johnson method and Bunger method is that the former is based on the integral product of a hazard function, $h_{c}(\infty ; t ; d)$, and $S(t ; d)$ and the latter is based on the integral product of the conditional probability, $q(t ; d)$, and $S(t ; d)$. Kahn and Sempos (1989) suggest that the use of hazard rates will not underestimate risks based on probabilities because the denominator of a rate is comprised of fewer individuals (personyears) since it is based on the midpoint of the interval - probabilities, on the other hasd, are based on denominator data at the beginning of the interval where the average personyears of follow-up is greater. Thus, the use of hazard rates in lifetime risk projection will result in estimates that are essentially slightly greater than risks based on probabilities.

The conditional probability of death due to spontaneously occurring cancer at age $t$ is estimated using the formula

$$
\begin{equation*}
\pi(t ; 0)=h_{c}(t ; 0) S(t ; 0) \tag{82}
\end{equation*}
$$

where $h_{c}(t ; 0)$ is the hazard function for spontaneous cancer and $S(t ; 0)$ is the survivorship function from the single-decrement life table (Eq. 67) for the nonexposed population. To determine the unconditional probability of death and lifetime risk of spontaneously (baseline) occurring cancer at age $t$ Eq. 69 is rewritten

$$
\begin{equation*}
Q(t ; 0)=\int_{0}^{t} \pi(x ; 0) d x=\int_{0}^{t} h_{c}(x ; 0) S(x ; 0) d x \tag{83}
\end{equation*}
$$

The unconditional probability of spontaneously occurring cancer in the nonexposed population over a lifetime is

$$
\begin{equation*}
\pi(\infty ; 0)=\int_{0}^{\infty} \pi(x ; 0) d x=\int_{0}^{\infty} h_{e}(x ; 0) S(x ; 0) d x \tag{84}
\end{equation*}
$$

and once again $\infty$ is by convention 100 years of age. The number of baseline cancer deaths in the nonexposed population (per $10^{5}$ individuals) is $\pi(\infty ; 0) \times 10^{5}$.

### 11.6 Years of Life Lost Per Premature Radiation-Induced Cancer Death

One of the most useful, if not most important, indices of radiation risk in un exposed population is the number of years of life lost per premature radiation-induced cancer death. The years of life lost by the exposed cohort per premature radiation-induced cancer death at age $t$ is

$$
\begin{equation*}
Y L P D=\frac{\int_{0}^{t} l(x ; 0)-l(x ; d) d x}{Q(t ; d) \times 10^{5}} \tag{85}
\end{equation*}
$$

where $l(x ; d)$ and $l(x ; 0)$ are the number of person-years in each age interval $(x, x+1)$ and $Q(t ; d)$ is the unconditional probability of radiation-induced cancer in the exposed population at age $t$.

### 11.7 Probability of Causation

Sometimes it is useful to determine the attributable risk caused by one or more radiation exposures. In principle, the attributable risk or probability of causation (PC) is defined as the fraction of radiation-induced cancer deaths out of the total cancer deaths in an exposed population. Using the lifetime risks of radiation-induced cancer explained earlier, the PC at age $t$ is calculated with the equation

$$
\begin{equation*}
P C=\frac{Q(t ; d) / Q(t ; 0)}{1+(Q(t ; d) / Q(t ; 0))} \tag{86}
\end{equation*}
$$

### 11.8 Error Propagation

A thorough evaluation of statistical uncertainty in numerical analysis will always involve the propagation of error. Estimates of the total uncertainty are determined several ways depending on the numerical methods used.

### 11.8.1 Constant and Non-constant Absolute and Relative Projection Models

The cause-specific hazard rates for radiation-induced cancer in the double-decrement life table have standard error

$$
\begin{equation*}
\sigma_{h_{e}(a ; t, d)}=h_{c}(a ; t ; d) \sqrt{\left(\frac{\sigma_{H(a)}}{H(a)}\right)^{2}+\left(\frac{\sigma_{\Phi_{E R R}(a)}}{\Phi_{E R R}(a)}\right)^{2}+\left(\frac{\sigma_{h_{c}(t ; 0)}}{h_{e}(t ; 0)}\right)^{2}} \tag{87}
\end{equation*}
$$

where $\sigma_{H(a)}$ is the standard frror of the annual dose equivalent (assumed to be 0.1 ), $\sigma_{\Phi_{E A R}(a)}$ is the standard erro; of the excess relative risk, and $\sigma_{h(t ; 0)}$ is written

$$
\begin{equation*}
J_{h_{c}(t ; 0)}=\sqrt{\frac{h_{c}(t ; 0)\left(1-h_{c}(t ; 0)\right)}{N(t ; 0)}} \tag{88}
\end{equation*}
$$

The standard error of the central mortality rate for the absolute model is

$$
\begin{equation*}
\sigma_{h_{e}(a ; t ; d)}=h_{c}(a ; t ; d) \sqrt{\left(\frac{\sigma_{H(a)}}{H(a)}\right)^{2}+\left(\frac{\sigma_{\Phi_{A R}(a)}}{\Phi_{A R}(a)}\right)^{2}} \tag{89}
\end{equation*}
$$

where $\sigma_{H(a)}$ is the standard error of the annual dose equivalent (assumed to be 0.1 ) and $\sigma_{A R(a)}$, the standard error of the absolute risk is defined by the equation

$$
\begin{equation*}
\sigma_{\Phi_{A R}(a)}=\sqrt{\frac{\left(\Phi_{A R}(a)\right)\left(1-\Phi_{A R}(a)\right)}{10^{4} P Y S v}} \tag{80}
\end{equation*}
$$

Next, using the standard error of $h_{c}(a ; t ; d)$, estimate the standard error of $h_{e}(\infty ; t ; d)$ with the relationship

$$
\begin{equation*}
\sigma_{h_{c}(\infty ; t ; d)}=\sqrt{\sum_{a+L}^{100}\left(\sigma_{h_{e}(a ; t ; d)}^{2}\right)} \tag{91}
\end{equation*}
$$

The survivorship function's standard error is obtained with Greenwood's (1926) formula

$$
\begin{equation*}
\operatorname{Var}[S(t ; d)]=S(t ; d)^{2} \sum_{a+L}^{t-1} \frac{q(t ; d)}{N(t ; d) p(t ; d)} \tag{92}
\end{equation*}
$$

where $a+L$ is the first age at exposure plus the minimal latency period and $t$ and $t-1$ are somewhere in the plateau period. The standard error of the survivorship function is the square root of $\operatorname{Var}[S(t ; d)]$. The standard error of the conditional death probability is written

$$
\begin{equation*}
\sigma_{\pi(t ; d)}=\pi(t ; d) \sqrt{\left(\frac{\sigma_{S(t ; d)}}{S(t ; d)}\right)^{2}+\left(\frac{\sigma_{h_{s}(\infty ; t ; d)}}{h_{c}(\infty ; t ; d)}\right)^{2}} \tag{93}
\end{equation*}
$$

and the standard error for the unconditional probability $c^{6}$ radiation-induced cancer risk is defined as

$$
\begin{equation*}
\sigma_{\pi(\infty ; d)}=\sqrt{\sum_{a+L}^{100} \sigma_{\pi(t ; d)}^{2}} \tag{94}
\end{equation*}
$$

For 'saseline cancers in the non-exposed population, we do the same as that for propagating error in the double-decrement life table but with different and far fewer steps. The standard erro: of $h_{e}(t ; 0)$ is given in Eq. 88 and the survivorship function has, according to Greenwood (1926), variance

$$
\begin{equation*}
\operatorname{Var}[S(t ; 0)]=S(t ; 0)^{2} \sum_{a+L}^{t-1} \frac{q(t ; 0)}{N(t ; 0) p(t ; 0)} \tag{95}
\end{equation*}
$$

where $a+L$ is based on the same first age at exposure of the exposed population plus the minimal latency period and $t$ and $t-1$ are somewhere in the plateau period. The standard
B-7
error of the survivorship function is the square root of $\operatorname{Var}[S(t ; 0)]$. The conditional death probability of baseline cancer at each interval is

$$
\begin{equation*}
\sigma_{\pi(t ; 0)}=\pi(t ; 0) \sqrt{\left(\frac{\sigma_{S(t ; 0)}}{S(t ; 0)}\right)^{2}+\left(\frac{\sigma_{h_{c}(t ; 0)}}{h_{c}(t ; 0)}\right)^{2}} \tag{96}
\end{equation*}
$$

and the standard error for the unconditioual probability of spontaneously occurring cancer is

$$
\begin{equation*}
\sigma_{\pi(\infty ; 0)}=\sqrt{\sum_{a+L}^{100} \sigma_{\pi(t ; 0)}^{2}} \tag{97}
\end{equation*}
$$

### 11.8.2 BEIR-V Relative Projection Model

Although the numerical methods for estimating central death rates of the BEIR-V relative risk projection model are similar to those used for the constant models, there are several additional steps that must be taken to determine the uncertainty. We pointed out earlier that the excess relative risk of the BEIR-V model, $\Phi_{E R R}(a, t)$, is the product of a dose function $f(d)$ and a link function $g(\beta)$. The standard error of the link function $g(\beta)$ is the natural logarithm of its geometric standaid deviation (GSD)

$$
\begin{equation*}
\sigma_{g(\beta)}=\sqrt{\sigma_{\beta_{1}}^{2}+\sigma_{\beta_{2}}^{2}+\sigma_{\beta_{3}}^{2}+\sigma_{\beta_{4}}^{2}} \tag{98}
\end{equation*}
$$

where $\sigma^{2}$ is the variance of the coefficients of the BEIR-V regression models. The standard error of the dose function $f(d)$ is functionally composed as

$$
\begin{equation*}
\sigma_{f(d)}=\sqrt{\sigma_{\alpha_{1}}^{2}+\sigma_{\alpha_{2}}^{2}} \tag{99}
\end{equation*}
$$

Since the excias relative risk $\Phi_{E R R}(a, t)$ is the product of the dose $f(d)$ and link function $g(\beta)$, its standard error is of the form

$$
\begin{equation*}
\sigma_{\Phi_{E R R}(a, t)}=\Phi_{E R R}(a, t) \sqrt{\left(\frac{\sigma_{g(\beta)}}{g(\beta)}\right)^{2}+\left(\frac{\sigma_{f(d)}}{f(d)}\right)^{2}} \tag{100}
\end{equation*}
$$

The hazard function for radiation-induced cancer at age $t$ in the double-decrement life table is the product of $\Phi_{E R R}(a, t)$ and the age-specific cancer mortality rate $h_{c}(t ; 0)$

$$
\begin{equation*}
h_{e}(a ; t ; d)=\Phi_{E R R}(a, t) h_{c}(t ; 0) \tag{101}
\end{equation*}
$$

and its standard error is

$$
\begin{equation*}
\sigma_{h_{c}(a ; ; ; d)}=h_{c}(a ; t ; d) \sqrt{\left(\frac{\sigma_{\Phi_{E R R}(a, t)}}{\Phi_{E R R}(a, t)}\right)^{2}+\left(\frac{\sigma_{h_{c}(t ; 0)}}{h_{c}(t ; 0)}\right)^{2}} \tag{102}
\end{equation*}
$$

Once the standard errors of the age-specific central death rates are known, we next estimate the standard error for the total central death rate as the square root of the sum of their variances given in the form

$$
\begin{equation*}
\sigma_{h_{e}(\infty ; t ; d)}=\sqrt{\sum_{a+L}^{100}\left(\sigma_{h_{c}(a ; t ; d)}^{2}\right)} \tag{103}
\end{equation*}
$$

This standard error is then used in the right-hand side of Eq. 93.

Table 28: Error components of lifetime risk.

| Sex | Race | $\sigma_{\pi(0, d)}$ | $\sigma_{D S 86}$ | $\sigma_{\text {SEER }}$ | $\sigma \sigma_{0 p}$ |
| :--- | :--- | :---: | :---: | :---: | :---: |
| Males | White | Eq. 94 | 0.45 | 0.2 | $\ln (1.2)$ |
|  | Nonwhite | $"$ | $"$ | 0.1 | $"$ |
| Females | White | $"$ | $"$ | 0.8 | $"$ |
|  | Nonwhite | $"$ | $"$ | 0.5 | $"$ |
|  |  |  |  |  |  |

### 11.8.3 Probability of Causation

Calculating the standard error of the PC is a rather simple task. We recall that the PC is the ratio of $(Q(t ; d) / Q(t ; 0))$ to $(1+((Q(t ; d) / Q(t ; 0))$ and therefore, the standard error of the PC is determined according to the formula

which reduces to

$$
\begin{equation*}
\sigma_{P C}=P C \sqrt{2\left(\left(\frac{\sigma_{Q(t ; d)}}{Q(t ; d)}\right)^{2}+\left(\frac{\sigma_{Q(t ; 0)}}{Q(t ; 0)}\right)^{2}\right)} \tag{104}
\end{equation*}
$$

### 11.9 Credibility Intervals of Lifetime Risk

Credibility intervals $(1-\alpha)$ for lifetime risks and PCs are based on the geometric standard deviation (GSD). Therefore, the arithmetic parameter, i.e., $\sigma$, will need to be exponentiated after it is adjusted for a Type I error, that is, multiplied by the standard normal deviate, $Z_{a}$ (Abramowitz and Stegun, 1965). The ( $1-\alpha$ ) credibility interval for radiationinduced lifetime risk is defined as

$$
\begin{equation*}
\pi(\infty ; d) / \exp \left(Z_{\alpha} \sigma_{T}\right)<\pi(\infty ; d)<\pi(\infty ; d) \exp \left(Z_{\alpha} \sigma_{T}\right) \tag{106}
\end{equation*}
$$

where $\pi(\infty ; d)$ is the unconditional death probability and $\sigma_{T}$ is the quadrature sum of errors for lifetime risks, DS86 standard error, sampling variation of the SEER mortality rates and differences between the U.S. and Japanese populations. The total error is of the form

$$
\begin{equation*}
\sigma_{T}=\sqrt{\sigma_{\pi(\infty, d)}^{2}+\sigma_{D S 86}^{2}+\sigma_{S E E R}^{2}+\sigma_{P o p}^{2}} \tag{107}
\end{equation*}
$$

where the component standard errors are given in Table 28. Similarly, the ( $1-\alpha$ ) credibility interval for the PC is

$$
\begin{equation*}
P C / \exp \left(Z_{\alpha} \sigma_{T}\right)<P C<P C \exp \left(Z_{\alpha} \sigma_{T}\right) \tag{108}
\end{equation*}
$$

with total error of the form

$$
\begin{equation*}
\sigma_{T}=\sqrt{\sigma_{P C}^{2}+\sigma_{D S 86}^{2}+\sigma_{S E E R}^{2}+\sigma_{P o p}^{2}} \tag{109}
\end{equation*}
$$

where the standard error of the $\mathrm{PC}, \sigma_{P C}$, is from Eq. 105 and the standard error $\sigma_{D S 86}$ is from Sposto et al. (1991), the standard error $\sigma_{\text {SEER }}$ is from the "Total U.S." row of Tables I-22 and I-23 of the Cancer Statistics Review (Ries et al., 1989) and $\sigma_{P o p}$ is from the standard error of "Population differences" row in the table of GSDs on page 214 of the BEIR-V report (NRC, 1990).

## 12 APPENDIX C. Relative and Absolute Risk Coefficients.

Tables C. 1 - C. 48 .

Table C.1. Excess relative and absolute risk coefficients for leukemia.

Male Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 37.770 | 6.294 | 2.865 | . 000 | . 000 | . 000 | . 000 |
| 10-19 | . 000 | 12.227 | 5.566 | 2.313 | . 585 | . 000 | . 000 |
| 20-29 | . 000 | 19.856 | 9.039 | 3.757 | . 951 | . 435 | . 000 |
| 30-39 | . 000 | . 000 | 43.664 | 18.147 | 4.593 | 2.103 | 1.527 |
| 40-49 | . 000 | . 000 | . 000 | 7.297 | 1.817 | . 846 | . 614 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 16.095 | 7.371 | 5.352 |

Male Absolute risk (deaths/10^4PYSV)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATP | -20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 7.514 | 2.853 | 136 | . 000 | . 000 | . 000 | 000 |
| 10-19 | . 000 | 4.503 | 1.360 | 1. 656 | . 155 | . 000 | 000 |
| 20-29 | . 000 | 13.323 | 8.046 | 3.168 | . 470 | . 520 | 000 |
| 30-39 | . 000 | . 000 | 8.859 | 4.831 | 4.076 | 7.629 | 6.977 |
| 40-49 | . 000 | . 000 | . 000 | 5.718 | 3.302 | . 606 | 2.894 |
| $50+$ | . 000 | . 000 | . 000 | 000 | 4. 626 | 6,975 | 6.124 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 42.040 | 7.005 | 3.189 | 1.325 | . 000 | . 000 | 000 |
| 10-19 | . 000 | 13.609 | 6.195 | 2.575 | . 652 | . 000 | . 000 |
| 20-29 | . 000 | 22.101 | 10.061 | 4.181 | 1.058 | . 485 | . 000 |
| 30-39 | . 000 | . 000 | . 000 | 20.199 | 5.112 | 2.341 | 1.700 |
| 40-49 | . 000 | . 000 | . 000 | 8.122 | 2.056 | . 941 | . 684 |
| 50 * | . 000 | . 000 | . 000 | . 000 | 17.915 | 8.204 | 5.957 |

Female Absolute risk (deaths/10^4PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 6.429 | .474 | .742 | .555 | .000 | .000 | .000 |
| $10-19$ | .000 | 1.688 | .503 | 1.109 | .175 | .000 | .000 |
| $20-29$ | .000 | 6.560 | 1.629 | 2.802 | .278 | .959 | .000 |
| $30-39$ | .000 | .000 | .000 | 2.274 | 1.212 | 2.570 | 1.603 |
| $40-49$ | .000 | .000 | .000 | 7.283 | 2.094 | 1.001 | .918 |
| $50+$ | .000 | .000 | .000 | 000 | 5.186 | 5.213 | 4.703 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | ---: | ---: | ---: |
| Chi-square | 1262.2200 | 3022 | 1.0000 |
| Deviance | 632.3860 | 3022 | 1.0000 |
| Freeman-Tukey | 238.5760 | 3022 | 1.0000 |

Table C.2. Excess relative and absolute risk coefficients for leukeria adjusted for confirmation rates not stratified, i.e., crude(1950-75).

Male Excess relative risk (o/Sv)

| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| <10 | 36.162 | 6.567 | 3.111 | . 000 | . 000 | 000 | 000 |
| 10-19 | 56.120 | 10.191 | 4.828 | 2.037 | . 529 | . 000 | . 000 |
| 20-29 | . 000 | 17.097 | 8.100 | 3.417 | . 887 | . 465 | . 000 |
| 30-39 | . 000 | . 000 | 34.570 | 14.585 | 3.785 | 1.986 | 1.551 |
| 40-49 | . 000 | . 000 | . 000 | 6.760 | 1.754 | . 920 | . 719 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 7.885 | 4.137 | 3.232 |

Male Absolute risk (deaths/10^4PYSv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 6.446 | 2.545 | . 135 | . 000 | . 000 | . 000 | 000 |
| 10-19 | . 545 | 3.730 | 1.158 | 1. 367 | . 141 | . 000 | . 000 |
| 20-29 | . 000 | 11.325 | 6.736 | 2. 672 | . 476 | . 553 | . 000 |
| 30-39 | . 000 | . 000 | 7.623 | 4.137 | 3.341 | 7.092 | 7.073 |
| 40-49 | . 000 | . 000 | . 000 | 5.021 | 2.902 | . 771 | 3.348 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 3.812 | 5.297 | 5.697 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 40.960 | 7.438 | 3.524 | 1.487 | 000 | . 000 | 000 |
| 10-19 | 63.567 | 11.543 | 5.469 | 2.307 | . 599 | . 000 | 000 |
| 20-29 | . 000 | 19.366 | 9.175 | 3.871 | 1.005 | . 527 | . 000 |
| 30-39 | . 000 | . 000 | 39.158 | 16.520 | 4.287 | 2.249 | 1.757 |
| 40-45 | . 000 | . 000 | . 000 | 7.657 | 1.987 | 1.042 | . 814 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 8.932 | 4.686 | 3.660 |

Female Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 5.525 | .446 | .768 | .607 | .000 | .000 | .000 |
| $10-19$ | .261 | 1.440 | .451 | .943 | .162 | .000 | .000 |
| $20-29$ | .000 | 5.568 | 1.429 | 2.352 | .268 | 1.038 | .000 |
| $30-39$ | .080 | .000 | .223 | 2.032 | 1.075 | 2.497 | 1.649 |
| $40-49$ | .000 | .000 | .000 | 6.295 | 1.876 | 1.077 | 1.109 |
| $50+$ | .000 | .000 | .000 | .000 | 3.957 | 3.905 | 4.601 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | ---: | ---: | ---: | ---: |
| Chi-square | 1077.8200 | 3022 | 1.0000 |
| Deviance | 534.8830 | 3022 | 1.0000 |
| Freeman-Tukey | 226.3790 | 3022 | 1.0000 |

Table C.3. Excess relative and absolute risk coefficients for leukemia adjusted for confirmation rates not stratified, i.e., crude (1950-85).

Male Excess relative risk (q/Sv)


| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 6.472 | 2. 469 | . 141 | . 024 | . 000 | . 000 | . 000 |
| 10-19 | . 573 | 3.805 | 1.138 | 1.417 | . 163 | . 000 | . 000 |
| 20-29 | . 000 | 11.577 | 6.657 | 2.801 | . 485 | . 535 | . 000 |
| 30-39 | . 000 | . 000 | 7.494 | 4.200 | 3.496 | 6.492 | 5.390 |
| 40-49 | . 000 | . 000 | . 000 | 4.842 | 2.826 | . 690 | 2.434 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 4.186 | 5.852 | 5.407 |
| Female | Excess | elativ | risk | Sv) |  |  |  |


| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 40.000 | 5.786 | 3.001 | 1. 371 | . 000 | . 000 | 000 |
| 10-19 | 68.791 | 11.671 | 5.161 | 2.359 | . 627 | . 000 | . 000 |
| 20-29 | . 000 | 19.685 | 8.705 | 3.978 | 1.057 | . 508 | . 000 |
| 30-39 | . 000 | . 000 | 36.266 | 16.575 | 4.406 | 2.115 | 1.347 |
| 40-49 | . 000 | . 000 | . 000 | 6.937 | 1.844 | . 885 | . 564 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 10.945 | 5.253 | 3.346 |

Female Absolute risk (deaths/10^4PYSV)


Age ATD

Age ATB |  | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | $70+$

| $<10$ | 5.507 | .422 | .626 | .505 | .000 | .000 | .000 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $10-19$ | .268 | 1.447 | .437 | .924 | .161 | .000 | .000 |
| $20-29$ | .000 | 5.591 | 1.397 | 2.394 | .280 | .891 | .000 |
| $30-39$ | .000 | .000 | .220 | 2.033 | 1.097 | 2.225 | 1.38 C |
| $40-49$ | .000 | .000 | .000 | 5.973 | 1.776 | .934 | .8 |
| $50+$ | .000 | .000 | .000 | .000 | 4.308 | 4.189 | 4.52 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | ---: | ---: | ---: |
| Chi-square | 1042.8300 | 3022 | 1.0000 |
| Deviance | 501.8550 | 3022 | 1.0000 |
| Freman-Tukey | 227.3180 | 3022 | 1.0000 |

Table C.4. Excess relative and absolute risk coefficients for leukemia adjusted for confirmation rates stratified on gender (1950-75).

Male Excess relative risk (o/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | 0+ |
| <10 | 39.293 | 7.008 | 3.330 | . 000 | . 000 | 000 | 000 |
| 10-19 | . 000 | 11.476 | 5.453 | 2.264 | . 574 | . 000 | . 000 |
| 20-29 | . 000 | 19.318 | 9.180 | 3.812 | . 967 | . 493 | . 000 |
| 30-39 | . 000 | . 000 | 39.847 | 16.544 | 4.196 | 2.141 | 1. 602 |
| 40-49 | . 000 | . 000 | . 000 | 7.886 | 2.000 | 1.020 | . 763 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 11.101 | 5.664 | 4. 238 |

Male Absolute risk (deaths/10^4PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 6.429 | 2.569 | .127 | .000 | .000 | .000 | .000 |
| $10-19$ | .000 | 3.743 | 1.152 | 1.426 | .152 | .000 | .000 |
| $20-29$ | .000 | 11.196 | 6.892 | 2.716 | .477 | .584 | .000 |
| $30-39$ | .000 | .000 | 7.425 | 4.001 | 3.263 | 7.353 | 7.248 |
| $40-49$ | .000 | .000 | .000 | 5.069 | 2.971 | .607 | 3.425 |
| $50+$ | .000 | .000 | .000 | .000 | 3.530 | 5.186 | 4.854 |

Female Excess relative risk (\%/Sv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 38.730 | 6.908 | 3.283 | 1.363 | .000 | .000 | .000 |
| $10-19$ | 63.422 | 11.311 | 5.375 | 2.232 | .566 | .000 | .000 |
| $20-29$ | .000 | 19.041 | 9.049 | 3.757 | .953 | .486 | .000 |
| $30-39$ | .000 | .000 | 39.276 | 16.307 | 4.136 | 2.110 | 1.579 |
| $40-49$ | .000 | .000 | .000 | 7.773 | 1.971 | 1.006 | .752 |
| $50+$ | .000 | .000 | .000 | .000 | 10.942 | 5.582 | 4.177 |

Female Absolute risk (deaths/10^4PYSV)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 5.493 | 429 | . 731 | . 567 | . 000 | 000 | 000 |
| 10-19 | . 261 | 1.434 | . 448 | . 922 | . 154 | . 000 | . 000 |
| 20-29 | . 000 | 5.562 | 1.426 | 2.317 | . 256 | . 963 | . 000 |
| 30-39 | . 000 | . 000 | . 223 | 2.030 | 1.054 | 2.379 | 1.509 |
| 40-49 | . 000 | . 000 | . 000 | 6.369 | 1. 868 | 1. 047 | 1.033 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 4.318 | 4.351 | 5.047 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | ---: | :--- | :--- | :--- |
| Chi-square | 1093.6500 | 3022 | 1.0000 |
| Deviance | 547.2830 | 3022 | 1.0000 |
| Freeman-Tukey | 222.2210 | 3022 | 1.0000 |

Table C.5. Excess relative and absolute risk coefficients for leukemia adjusted for confirmation rates stratified on gender (1950-85).

Male Excess relative risk (s/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 50-69 | $70+$ |
| $<10$ | 41.068 | 7.080 | 3.099 | . 000 | . 000 | 000 | . 000 |
| 10-19 | . 000 | 12.362 | 5.412 | 2.603 | . 695 | . 000 | . 000 |
| 20-29 | . 000 | 21. 269 | 9.311 | 4.478 | 1.196 | . 578 | . 000 |
| 30-39 | . 000 | . 000 | 35.101 | 16.882 | 4.510 | 2.178 | 1. 266 |
| 40-49 | . 000 | . 000 | . 000 | 7.685 | 2.053 | . 991 | . 577 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 13.082 | 6.317 | 3.674 |
| Male | Absolut | risk | deaths/ | ^4PYSV |  |  |  |

Age ATD

| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| <10 | 6.488 | 2.581 | . 119 | . 000 | . 000 | . 000 | 000 |
| 10-19 | . 000 | 3.846 | 1.136 | 1.521 | . 154 | . 000 | 000 |
| 20-29 | . 000 | 11.561 | 6.866 | 3.027 | . 486 | . 573 | 000 |
| 30-39 | . 000 | . 000 | 7.130 | 4.043 | 3.427 | 6.673 | 5.063 |
| 40-49 | . 000 | . 000 | . 000 | 4.989 | 3.037 | . 590 | 2.323 |
| $50+$ | .000 | . 000 | . 000 | . 000 | 3.716 | 5.560 | 4.029 |

Female Excess relative risk (\%/Sv)

|  | Age ATD |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 32.640 | 5.627 | 2.463 | 1.185 | . 000 | . 000 | . 000 |
| 10-19 | 56.994 | 9.825 | 4.301 | 2.069 | . 553 | . 000 | . 000 |
| 20-29 | . 000 | 16.904 | 7.400 | 3.559 | . 951 | . 459 | . 000 |
| 30-39 | . 000 | . 000 | 27.898 | 13.417 | 3.584 | 1.731 | 1.007 |
| 40-49 | . 000 | . 000 | . 000 | 6.108 | 1. 632 | . 788 | . 458 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 10.397 | 5.021 | 2.920 |

Female Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 5.388 | .403 | .560 | .466 | .000 | .000 | .000 |
| $10-19$ | .531 | 1.438 | .439 | .876 | .160 | .000 | .000 |
| $20-29$ | .000 | 5.532 | 1.389 | 2.314 | .298 | .849 | .000 |
| $30-39$ | .000 | .000 | .427 | 2.099 | 1.104 | 2.081 | 1.296 |
| $40-49$ | .000 | .000 | .000 | 5.838 | 1.741 | .963 | .909 |
| $50+$ | .000 | .000 | .000 | .000 | 4.800 | 4.793 | 5.795 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | ---: | :--- | :--- |
| Chi-square | 1049.3300 | 3022 | 1.0000 |
| Deviance | 509.2610 | 3022 | 1.0000 |
| Freeman-Tukey | 243.6890 | 3022 | 1.0000 |

Table C.6. Excess relative and absolute risk coefficients for leukemia adjusted for confirmation retes stratified on gender and city ( $1950-75$ ) .

Male Excess relative risk (o/Sv)
Age ATD

| Age ATB | $<20$ | $20-9$ | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| <10 | 41.459 | 7.140 | 3.472 | . 000 | . 000 | 000 | 000 |
| 10-19 | . 000 | 11.434 | 5.560 | 2.341 | . 646 | . 000 | 000 |
| 20-29 | . 000 | 18.091 | 8.797 | 3.704 | 1.022 | . 553 | 000 |
| 30-39 | . 000 | . 000 | 34.977 | 14.727 | 4.062 | 2.198 | 1.565 |
| 40-49 | . 000 | . 000 | . 000 | 8.026 | 2.214 | 1. 198 | . 853 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 9.735 | 5.268 | 3.750 |

Male Absolute risk (deaths/10^4PYSv)

| Age ATL |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <iU | 6.515 | 2. 583 | 131 | . 000 | . 000 | . 000 | . 000 |
| 10-19 | . 000 | 3.725 | 1.160 | 1.453 | . 170 | . 000 | . 000 |
| 20-29 | . 000 | 10.933 | 6.743 | 2.670 | . 500 | . 648 | . 000 |
| 30-39 | . 000 | . 000 | 7.217 | 3.838 | 3.196 | 7.513 | 7.110 |
| 40-49 | . 000 | . 000 | . 000 | 5.100 | 3.191 | . 693 | 3.778 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 3.360 | 4.984 | 4.461 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATk | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 31.690 | 5.458 | 2.654 | 1.117 | . 000 | . 000 | 000 |
| 10-19 | 50.749 | 8.740 | 4.250 | 1.789 | .494 | . 000 | . 000 |
| 20-29 | . 000 | 13.828 | 6.724 | 2.831 | . 781 | . 423 | . 000 |
| 30-39 | . 0000 | . 000 | 26.736 | 11.257 | 3.105 | 1.680 | 1.196 |
| 40-49 | . 00 | . 000 | . 000 | 6.135 | 1.692 | . 916 | . 652 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 7.441 | 4.027 | 2.866 |

Female Absolute risk (deaths/10^4PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| 10 | 5.251 | .471 | .664 | .485 | .000 | .000 | .000 |
| $10-19$ | .944 | 1.269 | .416 | .744 | .137 | .000 | .000 |
| $20-29$ | .000 | 5.128 | 1.235 | 1.838 | .223 | .846 | .000 |
| $30-39$ | .000 | .000 | .583 | 1.928 | .965 | 2.017 | 1.193 |
| $40-49$ | .000 | .000 | .000 | 5.311 | 1.677 | 1.034 | .955 |
| $50+$ | .000 | .000 | .000 | .000 | 3.825 | 3.878 | 5.164 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | ---: | ---: | ---: |
| Chi-square | 1083.3700 | 3022 | 1.0000 |
| Deviance | 533.4790 | 3022 | 1.0000 |
| Freman-Tukey | 238.3740 | 3022 | 1.0000 |

Table C.7. Excess relative and absolute risk coefficients for leukemia adjusted for confirmation rates stratified on gender and city (1950-85).

Male Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 42.681 | 6. 644 | 2.871 | . 000 | . 000 | . 000 | 000 |
| 10-19 | . 000 | 12.493 | 5.399 | 2.558 | . 731 | . 000 | . 000 |
| 20-29 | . 000 | 20.181 | 8.721 | 4.132 | 1.181 | . 588 | . 000 |
| 30-39 | . 000 | . 000 | 33.041 | 15.655 | 4.476 | 2.227 | 1.260 |
| 40-49 | . 000 | . 000 | . 000 | 7.635 | 2.183 | 1.086 | . 614 |
| $50+$ | 000 | 000 | 000 | . 000 | 12.985 | 6.461 | 3.654 |

Male Absolute risk (deaths/10^4RYSv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 6.563 | 2.483 | 113 | . 000 | . 000 | . 000 | 000 |
| 10-19 | . 000 | 3.852 | 1.128 | 1.498 | . 161 | . 000 | . 000 |
| 20-29 | . 000 | 11.386 | 6.663 | 2.883 | . 479 | . 580 | . 000 |
| 30-39 | . 000 | . 000 | 7.040 | 3.927 | 3.410 | 6.773 | 5.011 |
| 40-49 | . 000 | . 000 | . 000 | 4.958 | 3.168 | . 636 | 2.449 |
| 50 + | . 000 | . 000 | . 000 | . 000 | 3.693 | 5.618 | 4.010 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 6c-69 | $70+$ |
| <10 | 29.930 | 4.659 | 2.013 | . 954 | . 000 | . 000 | 000 |
| 10-19 | 56.276 | 8.761 | 3.786 | 1.794 | . 513 | . 000 | . 000 |
| 20-29 | . 000 | 14.152 | 6.116 | 2.898 | . 828 | . 412 | . 000 |
| 30-39 | . 000 | . 000 | 23.170 | 10.978 | 3.139 | 1. 562 | . 883 |
| 40-49 | . 000 | . 000 | . 000 | 5.354 | 1. 531 | . 762 | . 431 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 9.105 | 4.531 | 2.562 |

Female Absolute risk (deaths/10^4PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 5.192 | .424 | .476 | .358 | .000 | .000 | .000 |
| $10-19$ | .975 | 1.272 | .388 | .749 | .137 | .000 | .000 |
| $20-29$ | .000 | 5.166 | 1.177 | 1.870 | .252 | .693 | .000 |
| $30-39$ | .000 | .000 | .558 | 1.906 | .975 | 1.777 | 1.156 |
| $40-49$ | .000 | .000 | .000 | 4.898 | 1.552 | .877 | .786 |
| $50+$ | .000 | .000 | .000 | .000 | 4.218 | 4.199 | 5.125 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | ---: | ---: | ---: |
| Chi-square | 1042.8900 | 3022 | 1.0000 |
| Deviance | 499.8330 | 3022 | 1.0000 |
| Freeman-Tukey | 230.1240 | 3022 | 1.0000 |

Table c.8. Excess relative and absolute risk coefficients for leukemia adjusted for confirmation rates stratified on age ATD (1950-75).

Male Excess relative risk (\%/Sv)

|  | Age ATD |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 33.289 | 6.549 | 3.106 | . 000 | . 000 | . 000 | 000 |
| 10-1.9 | 40.978 | 8.061 | 3.823 | 1.817 | . 494 | . 000 | 000 |
| 20-29 | . 000 | 13.984 | 6.632 | 3.152 | . 856 | . 587 | . 000 |
| 30-39 | . 000 | . 000 | 22.404 | 10.647 | 2.893 | 1.982 | 1.618 |
| 40-49 | . 000 | . 000 | . 000 | 5.597 | 1.521 | 1.042 | . 851 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 6.919 | 4.741 | 3.870 |

Male Absolute risk (deaths/10^4PYSV)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 6.104 | 2.476 | 149 | . 000 | . 000 | . 000 | 000 |
| 10-19 | 1.369 | 3.402 | 1.072 | 1.257 | . 133 | . 000 | . 000 |
| 20-29 | . 000 | 10.763 | 6.031 | 2.612 | . 514 | . 684 | . 000 |
| 30-39 | . 000 | . 000 | 7.024 | 3.928 | 2.913 | 7.373 | 7.323 |
| 40-49 | . 000 | . 000 | . 000 | 4.798 | 2.680 | . 941 | 3.828 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 4.417 | 6.433 | 5.534 |
| Female | Excess relative risk (\%/Sv) |  |  |  |  |  |  |

Age ATD

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 37.190 | 7.316 | 3.470 | 1.649 | .000 | .000 | .000 |
| $10-19$ | 45.780 | 9.006 | 4.271 | 2.030 | .551 | .000 | .000 |
| $20-29$ | .000 | 15.623 | 7.409 | 3.521 | .957 | .656 | .000 |
| $30-39$ | .000 | .000 | 25.029 | 11.894 | 3.232 | 2.215 | 1.808 |
| $40-49$ | .000 | .000 | .000 | 6.253 | 1.699 | 1.164 | .950 |
| $50+$ | .000 | .000 | .000 | .000 | 7.729 | 5.296 | 4.323 |

Female Absolute risk (deaths/10^4PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 5.206 | .463 | .758 | .665 | .000 | .000 | .000 |
| 10.19 | .685 | 1.359 | .445 | .854 | .151 | .000 | .000 |
| $20-29$ | .000 | 5.173 | 1.373 | 2.206 | .269 | 1.270 | .000 |
| $30-39$ | .000 | .000 | .576 | 2.035 | 1.089 | 2.531 | 1.690 |
| $40-49$ | .000 | .000 | .000 | 5.734 | 1.739 | 1.355 | 1.244 |
| $50+$ | .000 | .000 | .000 | .000 | 4.200 | 4.785 | 4.217 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | ---: | ---: | ---: |
| Chi-square | 1046.1700 | 3022 | 1.0000 |
| Deviance | 520.0840 | 3022 | 1.0000 |
| Freeman-Tukey | 254.2180 | 3022 | 1.0000 |

Table C.9. Excess relative and absolute risk coefficients for leukemia adjusted for confirmation rates stratified on age ATD (1950-85).

Male Excess relative risk (\%/Sv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | < 20 | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 33.223 | 5.908 | 2.530 | 1.306 | .000 | .000 | .000 |
| $10-19$ | 47.458 | 8.439 | 3.614 | 1.866 | .519 | .000 | .000 |
| $20-29$ | .000 | 14.771 | 6.326 | 3.266 | .909 | .573 | .000 |
| $30-39$ | .000 | .000 | 20.848 | 10.764 | 2.996 | 1.889 | 1.380 |
| $40-49$ | .000 | .000 | .000 | 5.276 | 1.469 | .926 | .677 |
| $50+$ | .000 | .000 | .000 | .000 | 8.267 | 5.213 | 3.809 |

Male Absolute risk (deaths/10^4PYSV)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 0.112 | 2.336 | . 175 | . 071 | . 000 | . 000 | 000 |
| 10-19 | 1.477 | 3.478 | 1.028 | 1.312 | . 189 | . 000 | . 000 |
| 20-29 | . 000 | 11.057 | 5.854 | 2.690 | . 529 | . 756 | . 000 |
| 30-39 | . 000 | . 000 | 6.821 | 3.956 | 2.995 | 7.288 | 5.850 |
| 40-49 | . 000 | . 000 | . 000 | 4.624 | 2. 612 | . 848 | 2.888 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 4.772 | 6.824 | 5.243 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 36.990 | 6.577 | 2.817 | 1.454 | . 000 | . 000 | . 000 |
| 10-19 | 52.839 | 9.396 | 4.024 | 2.078 | . 578 | . 000 | . 000 |
| 20-29 | . 000 | 16.445 | 7.043 | 3.637 | 1.012 | . 638 | . 000 |
| 30-39 | . 000 | . 000 | 23.211 | 11.985 | 3.336 | 2.104 | 1.537 |
| 40-49 | . 000 | . 000 | . 000 | 5.875 | 1.635 | 1.031 | . 753 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 9.204 | 5.805 | 4.240 |

Female Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | R | 20 | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| 10 | 5.205 | .433 | .597 | .530 | .000 | .000 | .000 |  |
| $10-19$ | .720 | 1.383 | .428 | .852 | .166 | .000 | .000 |  |
| $20-29$ | .000 | 5.249 | 1.339 | 2.254 | .318 | 1.258 | .000 |  |
| $30-39$ | .000 | .000 | .562 | 2.041 | 1.116 | 2.553 | 1.448 |  |
| $40-49$ | .000 | .000 | .000 | 5.531 | 1.690 | 1.221 | .992 |  |
| $50+$ | .000 | .000 | .000 | .000 | 4.554 | 5.048 | 4.169 |  |


| Goodness of fit | Value | d.f. | Prob |
| :--- | ---: | ---: | ---: |
| Chi-square | 1020.8000 | 3022 | 1.0000 |
| Deviance | 493.6020 | 3022 | 1.0000 |
| Freeman-Tukey | 244.6780 | 3022 | 1.0000 |

Table C.10. Excess relative and absolute risk coefficients for leukemia adjusted for confirmation rates stratified on DS86 shielded kerma (1950-75).

Male Excess relative risk (\%/Sv)

| Age ATL |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 28.757 | 5.338 | 2.713 | . 000 | . 000 | . 000 | 000 |
| 10-19 | 60.045 | 11. 146 | 5. 664 | 2.235 | . 611 | . 000 | . 000 |
| 20-29 | . 000 | 14.540 | 7.388 | 2.915 | . 797 | . 550 | . 000 |
| 30-39 | . 000 | . 000 | 32.717 | 12.909 | 3.529 | 2.435 | 1.931 |
| 40-49 | . 000 | . 000 | . 000 | 8.052 | 2. 201 | 1.519 | 1.204 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 13.789 | 9.515 | 7.545 |
| Male | Absolu | risk | deaths | ^ 4 EYSV |  |  |  |

Age ATD

| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| <10 | 5. 662 | 2.275 | 149 | . 000 | 000 | 000 | 000 |
| 10-19 | . 080 | 3.976 | 1.483 | 1.537 | 161 | . 000 | . 000 |
| 20-29 | . 000 | 8.993 | 6.213 | 2.413 | . 450 | . 645 | . 000 |
| 30-39 | . 000 | . 000 | 6.492 | 3.976 | 3. 462 | 8.586 | 8.402 |
| 40-43 | . 000 | . 000 | . 000 | 5.738 | 3.459 | 1.291 | 5.312 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 5.509 | 8.902 | 10.731 |

Female Excess relative risk (\%/Sv)

|  | Age ATD |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-6? | $70+$ |
| $<10$ | 34.360 | 6.378 | 3.241 | 1.279 | . 000 | . 000 | 000 |
| 10-19 | 71.743 | 13.318 | 6.767 | 2.670 | . 730 | . 000 | . 000 |
| 20-29 | . 000 | 17.373 | 8.828 | 3.483 | . 952 | . 657 | . 000 |
| 30-39 | . 000 | . 000 | 39.091 | 15.424 | 4.216 | 2.909 | 2. 307 |
| 40-49 | . 000 | . 000 | . .000 | 9.620 | 2.630 | 1.815 | 1.439 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 16.476 | 11.369 | 9.015 |

Female Absolute risk (deaths/10^4PYSV)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 5.397 | 463 | . 764 | . 537 | . 000 | . 000 | . 000 |
| 10-19 | . 320 | 1.274 | . 515 | 1.036 | . 194 | . 000 | . 000 |
| 20-29 | . 000 | 5.752 | 1.310 | 2.149 | . 271 | 1.273 | . 000 |
| 30-39 | . 000 | . 000 | . 313 | 2.222 | 1.014 | 3.074 | 2.045 |
| 40-49 | . 000 | . 000 | . 000 | 7.524 | 2.437 | 1.976 | 1.900 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 5.996 | 6.056 | 8.473 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | ---: | :--- | :--- |
| Chi-square | 952.8740 | 3022 | 1.0000 |
| Deviance | 536.0480 | 3022 | 1.0000 |
| Freeman-Tukey | 255.8280 | 3022 | 1.0000 |

Table C.11. Excess relative and absolute risk coefficients for leukemia adjusted for confirmation rates stratified on DS86 shielded kerma (1950-85).

Male Excess relative risk (o/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 28.350 | 5.021 | 2.405 | 1.241 | . 000 | . 000 | 000 |
| 10-19 | 64.886 | 11.492 | 5.505 | 2.840 | . 980 | . 000 | 000 |
| 20-29 | . 000 | 14.949 | 7.161 | 3. 694 | 1.275 | . 890 | . 000 |
| 30-39 | . 000 | . 000 | 21.023 | 10.844 | 3.742 | 2. 613 | 2.577 |
| 40-49 | . 000 | . 000 | . 000 | 6.832 | 2.358 | 1. 646 | 1. 624 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 13.061 | 9.121 | 8.995 |
| Male |  |  |  |  |  |  |  |
| Age ATD |  |  |  |  |  |  |  |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 5.648 | 2. 193 | . 157 | . 028 | . 000 | . 000 | . 000 |
| 10-19 | . 082 | 4.028 | 1.453 | 1.840 | . 332 | . 000 | . 000 |
| 20-29 | . 000 | 9.126 | 6.057 | 2.863 | . 643 | . 944 | . 000 |
| 30-39 | . 000 | . 000 | 5. 688 | 3.746 | 3.619 | 9.281 | 9.743 |
| 40-49 | . 000 | . 000 | . 000 | 5.172 | 3. 644 | 1.375 | 6.915 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 5.422 | 8.587 | 11.753 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 34.320 | 6.078 | 2.912 | 1.502 | 000 | . 000 | 000 |
| 10-19 | 78.549 | 13.912 | 6.664 | 3.437 | 1.186 | . 000 | . 000 |
| 20-29 | . 000 | 18.097 | 8.669 | 4.472 | 1.543 | 1.078 | . 000 |
| 30-39 | . 000 | . 000 | 25.450 | 13.128 | 4.530 | 3.164 | 3.120 |
| 40-49 | . 000 | . 000 | . 000 | 8.271 | 2.854 | 1.993 | 1.965 |
| $50 \%$ | . 000 | . 000 | . 000 | . 000 | 15.811 | 11.042 | 10.889 |

Female Absolute risk (deaths/10^4PYSv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 5.398 | . 448 | . 613 | 574 | . 000 | . 000 | 000 |
| 10-19 | . 329 | 1.292 | . 510 | 1.179 | . 283 | . 000 | . 000 |
| 20-29 | . 000 | 5.816 | 1.294 | 2.511 | . 423 | 1.886 | . 000 |
| 30-39 | . 000 | . 000 | . 283 | 2.087 | 1.069 | 3.243 | 3. 338 |
| 40-49 | . 000 | . 000 | . 000 | 6.904 | 2. 602 | 2.123 | 3.030 |
| $50+$ | 000 | 000 | 000 | . 000 | 5.922 | 5.980 | 9.84 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 888.3530 | 3022 | 1.0000 |
| Deviance | 506.2570 | 3022 | 1.0000 |
| Freeman-Tukey | 267.5660 | 3022 | 1.0000 |

Table C.12. Excess relative and absolute risk coefficients for leukemia. Organ dose equivalent adjusted for DS86 random error.

Male Excess relative risk (o/sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | C1.963 | 7.069 | 3.465 | . 000 | . 000 | . 000 | 000 |
| 10-19 | . 000 | 14.993 | 7.349 | 2.871 | . 727 | . 000 | 000 |
| 20-29 | . 000 | 22.664 | 11.109 | 4.340 | 1. 099 | . 499 | . 000 |
| 30-39 | . 000 | . 000 | 52.867 | 20.651 | 5.232 | 2.374 | 1. 637 |
| 40-49 | . 000 | . 000 | . 000 | 8.222 | 2.083 | . 945 | . 652 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 18.950 | 8.600 | 5.929 |

Male Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | R | 20 | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| 10 | 0.231 | 3.142 | .161 | .000 | .000 | .000 | .000 |  |
| $10-19$ | .000 | 5.229 | 1.653 | 1.978 | .193 | .000 | .000 |  |
| $20-29$ | .000 | 14.873 | 9.430 | 3.633 | .534 | .603 | .000 |  |
| $30-39$ | .000 | .000 | 9.997 | 5.455 | 4.505 | 8.705 | 7.516 |  |
| $40-49$ | .000 | .000 | .000 | 6.403 | 3.715 | .678 | 3.096 |  |
| $50+$ | .000 | .000 | .000 | .000 | 5.226 | 7.975 | 6.875 |  |

Female Excess relative risk (\%/Sv)
Age ATD

| Age ATB | 20 |  | $20-29$ | $30-39$ | $40-49$ | $50-50$ | $60-69$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $<10$ | 45.640 | 7.689 | 3.769 | 1.472 | .000 | .000 | .000 |
| $10-19$ | .000 | 16.307 | 7.993 | 3.122 | .791 | .000 | .000 |
| $20-29$ | .000 | 24.650 | 12.083 | 4.720 | 1.196 | .543 | .000 |
| $30-39$ | .000 | .000 | .000 | 22.461 | 5.690 | 2.583 | 1.780 |
| $40-49$ | .000 | .000 | .000 | 8.942 | 2.265 | 1.028 | .709 |
| $50+$ | .000 | .000 | .000 | .000 | 20.610 | 9.354 | 6.448 |

Female Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | AT | 20 | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| 10 | 7.027 | .520 | .849 | .615 | .000 | .000 | .000 |  |
| $10-19$ | .000 | 1.914 | .578 | 1.320 | .207 | .000 | .000 |  |
| $20-29$ | .000 | 7.315 | 1.859 | 3.131 | .314 | 1.072 | .000 |  |
| $30-39$ | .000 | .000 | .000 | 2.516 | 1.331 | 2.828 | 1.683 |  |
| $40-49$ | .000 | .000 | .000 | 7.926 | 2.308 | 1.096 | .949 |  |
| $50+$ | .000 | .000 | .000 | .000 | 5.691 | 5.761 | 5.074 |  |


| Goodness of fit | Value | d.f. | Prob |
| :--- | ---: | :--- | :--- | :--- |
| Chi-square | 1337.7300 | 3022 | 1.0000 |
| Deviance | 635.1690 | 3022 | 1.0000 |
| Freeman-Tukey | 251.6360 | 3022 | 1.0000 |

Table C.13. Excess relative and absolute risk ooefficients for leukemia adjusted for confirmation rates not stratified, i.e., crude(1950-75). organ dose equivalent adjusted for DS86 random irror.

Male Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 40.233 | 7.416 | 3.769 | . 000 | . 200 | . 000 | 000 |
| 10-19 | 67.672 | 12.474 | 6.339 | 2.526 | . 656 | 000 | 000 |
| 20-29 | . 000 | 19.568 | 9.943 | 3.963 | 1.029 | . 538 | 000 |
| 30-39 | . 000 | . 000 | 41.624 | 15.588 | 4.309 | 2.252 | 1.670 |
| 40-49 | . 000 | . 000 | . 000 | 7. 673 | 1.993 | 1.042 | . 772 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 9.343 | 4.883 | 3.621 |
| Male Absolute risk (deaths/10^4PYSV) |  |  |  |  |  |  |  |
| Age ATD |  |  |  |  |  |  |  |
| Age ATB <20 $20-29 \quad 30-39 \quad 40-49 \quad 50-59 \quad 60-69$ |  |  |  |  |  |  |  |
| $\begin{aligned} & <10 \\ & 10-19 \\ & 20-29 \\ & 30-39 \\ & 40-49 \\ & 50+ \end{aligned}$ | $\begin{array}{r} 7.063 \\ .600 \\ .000 \\ .000 \\ .000 \\ .000 \end{array}$ | $\begin{array}{r} 2.810 \\ 4.341 \\ 12.662 \\ .000 \\ .000 \\ .000 \end{array}$ | .160 .000 <br> 1.410 1.636 <br> 7.901 3.074 <br> 8.609 4.677 <br> .000 5.645 <br> .000 .000 |  | $\begin{array}{r} .000 \\ .175 \\ .543 \\ 3.700 \\ 3.284 \\ 4.355 \end{array}$ | $\begin{array}{r} .000 \\ .000 \\ .646 \\ 8.122 \\ .872 \\ 6.133 \end{array}$ | $\begin{array}{r} .000 \\ .000 \\ .000 \\ 7.651 \\ 3.621 \\ 6.441 \end{array}$ |
|  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 44.400 | 8.184 | 4.159 | 1. 657 | 000 | . 000 | 000 |
| 10-19 | 74.682 | 13.767 | 6.995 | 2.788 | . 724 | . 000 | . 000 |
| 20-29 | . 000 | 21.595 | 10.973 | 4.373 | 1.136 | . 594 | . 000 |
| 30-39 | . 000 | . 000 | 45.936 | 18.306 | 4.755 | 2.485 | 1.843 |
| 40-49 | . 000 | . 000 | . 000 | 8.468 | 2.200 | 1.149 | . 852 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 10.311 | 5.388 | 3.996 |

Female Absolute risk (deaths/10^4PYSV)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 6.038 | . 489 | . 875 | . 679 | . 000 | . 000 | 000 |
| 10-19 | . 282 | 1.633 | . 519 | 1.124 | . 192 | . 000 | . 000 |
| 20-29 | . 000 | 6.208 | 1.632 | 2.631 | . 303 | 1.156 | . 000 |
| 30-39 | . 000 | . 000 | . 246 | 2.247 | 1.181 | 2.754 | 1.734 |
| 40-49 | . 000 | . 000 | . 000 | 6.868 | 2.075 | 1.189 | 1.157 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 4.393 | 4.375 | 5.003 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | ---: | :--- | :--- | :--- |
| Chi-square | 1139.4000 | 3022 | 1.0000 |
| Deviance | 537.3280 | 3022 | 1.0000 |
| Freeman-Tukey | 238.6100 | 3022 | 1.0000 |

Table C.14. Excess relative and absolute risk coefficients for leukemia adjusted for confirmation rates not stratified, i.e., crude (1950-85). organ dose equivalent adjusted for DS86 random error.

Male Excess relative risk (\%/Sv)

|  |  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | 20 |  | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 40.802 | 7.013 | 3.319 | 1.434 | .000 | .000 | .000 |  |
| $10-19$ | 76.404 | 13.132 | 6.215 | 2.686 | .715 | .000 | .000 |  |
| $20-29$ | .000 | 20.692 | 9.794 | 4.232 | 1.126 | .536 | .000 |  |
| $30-39$ | .000 | .000 | 40.074 | 17.318 | 4.608 | 2.192 | 1.310 |  |
| $40-49$ | .000 | .000 | .000 | 7.190 | 1.913 | .910 | .544 |  |
| $50+$ | .000 | .000 | .000 | .000 | 12.022 | 5.719 | 3.417 |  |

Male Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 7.092 | 2.723 | .166 | .027 | .000 | .000 | .000 |
| $10-19$ | .631 | 4.429 | 1.386 | 1.697 | .202 | .000 | .000 |
| $20-29$ | .000 | 12.951 | 7.808 | 3.223 | .554 | .622 | .000 |
| $30-39$ | .000 | .000 | 8.465 | 4.752 | 3.877 | 7.415 | 5.759 |
| $40-49$ | 000 | .000 | .000 | 5.435 | 3.189 | .774 | 2.580 |
| $50+$ | .000 | .000 | .000 | .000 | 4.785 | 6.779 | 6.089 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 43.320 | 7.446 | 3.524 | 1. 523 | . 000 | . 000 | 000 |
| 10-19 | 81.119 | 13.942 | 6.599 | 2.852 | . 759 | . 000 | . 000 |
| 20-29 | . 000 | 21.969 | 10.398 | 4.493 | 1.196 | . 569 | . 000 |
| 30-39 | . 000 | . 000 | 42.547 | 18.386 | 4.892 | 2.327 | 1.390 |
| 40-49 | . 000 | . 000 | . 000 | 7.634 | 2.031 | . 966 | . 577 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 12.764 | 6.072 | 3.628 |

Female Absolute risk (deaths/10^4PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 6.016 | .462 | .714 | .563 | .000 | .000 | .000 |
| $10-19$ | .290 | 1.642 | .504 | 1.100 | .191 | .000 | .000 |
| $20-29$ | .000 | 6.234 | 1.596 | 2.677 | .316 | .995 | .000 |
| $30-39$ | .000 | .000 | .243 | 2.248 | 1.205 | 2.447 | 1.440 |
| $40-49$ | .000 | .000 | .000 | 6.504 | 1.957 | 1.022 | .891 |
| $50+$ | .000 | .000 | .000 | .000 | 4.786 | 4.698 | 4.888 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | ---: | ---: | ---: |
| Chi-square | 1105.0700 | 3022 | 1.0000 |
| Deviance | 504.1510 | 302 | 1.0000 |
| Freeman-Tukey | 231.9940 | 3022 | 1.0000 |

Table C.1.5. Excess relative and absolute risk coefficients for leukemia adjusted for confirmation rates stratified on gender (1950-75).
organ dose equivalent adjusted for DS86 random error.

Male Excess relative risk ( $\mathrm{g} / \mathrm{Sv}$ )

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 43.722 | 7.900 | 4.030 | .000 | .000 | .000 | .000 |
| $10-19$ | .000 | 14.004 | 7.145 | 2.802 | .711 | .000 | .000 |
| $20-29$ | .000 | 22.084 | 11.267 | 4.419 | 1.121 | .568 | .000 |
| $30-39$ | .000 | .000 | 47.983 | 18.819 | 4.772 | 2.420 | 1.726 |
| $40-49$ | .000 | .000 | .000 | 8.934 | 2.266 | 1.149 | .819 |
| $50+$ | .000 | .000 | .000 | .000 | 13.038 | 6.612 | 4.715 |
| Male | Absolute $r$ Lisk (deaths/10^4PYSV) |  |  |  |  |  |  |


| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 7.043 | 2.833 | . 151 | . 000 | . 000 | . 000 | . 000 |
| 10-19 | . 000 | 4.345 | 1.396 | 1.702 | . 189 | . 000 | . 000 |
| 20-29 | . 000 | 12.503 | 8.067 | 3.121 | . 543 | . 680 | . 000 |
| 30-39 | . 000 | . 000 | 8.374 | 4.521 | 3.605 | 8.399 | 7.848 |
| 40-49 | . 000 | . 000 | . 000 | 5.691 | 3.356 | . 684 | 3.705 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 4.001 | 5.952 | 5.483 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | 704 |
| $<10$ | 42.050 | 7.598 | 3.876 | 1.520 | . 000 | . 000 | 000 |
| 10-19 | 74.542 | 13.469 | 6.871 | 2.695 | . 683 | . 000 | . 000 |
| 20-29 | . 000 | 21.239 | 10.836 | 4.250 | 1.078 | . 547 | . 000 |
| 30-39 | . 000 | . 000 | 46.148 | 18.099 | 4.590 | 2.328 | 1.660 |
| 40-49 | . 000 | . 000 | . 000 | 8.592 | 2.179 | 1.105 | . 788 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 12.539 | 6.359 | 4.535 |

Female Absolute risk (deaths/10^4PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 6.004 | .471 | .835 | .635 | .000 | .000 | .000 |
| $10-19$ | .283 | 1.627 | .516 | 1.098 | .182 | .000 | .000 |
| $20-29$ | .000 | 6.202 | 1.629 | 2.593 | .290 | 1.080 | .000 |
| $30-39$ | .000 | .000 | .246 | 2.246 | 1.159 | 2.621 | 1.590 |
| $40-49$ | .000 | .000 | .000 | 6.949 | 2.065 | 1.152 | 1.079 |
| $50+$ | .000 | .000 | .000 | .000 | 4.768 | 4.837 | 5.470 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | ---: | :--- | :--- |
| Chi-square | 1157.7300 | 3022 | 1.0000 |
| Deviance | 549.7620 | 3022 | 1.0000 |
| Freeman-Tukey | 222.2430 | 3022 | 1.0000 |

Table C.16. Excess relative and absolute risk coefficients for leukemia adjusted for confirmation rates stratified on gender (1950-85). organ dose equivalent adfusted for DS86 random error.

Male Excess relative risk (s/sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 45.773 | 7.970 | 3.735 | . 000 | . 000 | . 000 | 000 |
| 10-19 | . 000 | 15.067 | 7.060 | 3.227 | . 860 | . 000 | . 000 |
| 20-29 | . 000 | 24.329 | 11.401 | 5.211 | 1.389 | . 661 | . 000 |
| 30-39 | . 000 | . 000 | 42.169 | 19.273 | 5.138 | 2.444 | 1. 341 |
| 40-49 | . 000 | . 000 | . 000 | 8.712 | 2. 323 | 1.105 | . 606 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 15.606 | 7.424 | 4.074 |

Male Absolute risk (deaths $/ 10^{\wedge} 4$ PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 7.109 | 2.842 | .141 | .000 | .000 | .000 | .000 |
| $10-19$ | .000 | 4.459 | 1.374 | 1.812 | .191 | .000 | .000 |
| $20-29$ | .000 | 12.917 | 8.024 | 3.481 | .554 | .663 | .000 |
| $30-39$ | .000 | .000 | 8.039 | 4.575 | 3.791 | 7.586 | 5.404 |
| $40-49$ | .000 | .000 | .000 | 5.600 | 3.424 | .659 | 2.463 |
| $50+$ | .000 | .000 | .000 | .000 | 4.223 | 6.391 | 4.539 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 35.470 | 6.176 | 2.894 | 1.323 | . 000 | . 000 | 000 |
| 10-19 | 67.053 | 11.675 | 5.471 | 2.500 | . 667 | . 000 | 000 |
| 20-29 | . 000 | 18.853 | 8.835 | 4.038 | 1.076 | . 512 | . 000 |
| 30-39 | . 000 | . 000 | 32.678 | 14.935 | 3.982 | 1.894 | 1.039 |
| 40-49 | . 000 | . 000 | . 000 | 6.751 | 1.810 | . 856 | . 470 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 12.094 | 5.753 | 3.157 |

Female Absolute risk (deaths/10^4PYSV)

| Age ATD |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10 \quad 5.889$ | . 442 | 2. 641 | . 522 | . 000 | . 000 | . 000 |
| 10-19 . 577 | 1.633 | 3 . 509 | 1.044 | . 189 | . 000 | . 000 |
| 20-29 . 000 | 6.169 | 1.592 | 2.594 | . 336 | . 945 | . 000 |
| $30-39$. 000 | . 000 | - . 473 | 2.324 | 1.216 | 2.281 | 1.348 |
| 40-49 . 000 | . 000 | O . 000 | 6.370 | 1.921 | 1.049 | . 930 |
| $50+\quad .000$ | . 000 | . . 000 | . 000 | 5.339 | 5.358 | 6.267 |
| Goodness of fit |  | Value | d.f. | Prob |  |  |
| chi-square |  | 1113.3900 | 3022 | 1.0000 |  |  |
| Deviance |  | 511.5560 | 3022 | 1.0000 |  |  |
| Freeman-Tukey |  | 261.6730 | 3022 | 1.0000 |  |  |

Table C.17. Excess relative and absolute risk coefficients for leukemia adjusted for confirmation rates stratified on gender and city (1950-75). organ dose equivalent adjusted for DS86 random error.

Male Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 46.374 | 8.075 | 4.173 | . 000 | . 000 | . 000 | 000 |
| 10-19 | . 000 | 13.968 | 7.219 | 2.909 | . 802 | . 000 | . 000 |
| 20-29 | . 000 | 20.704 | 10.701 | 4.312 | 1.189 | . 637 | . 000 |
| 30-39 | . 000 | . 000 | 41.710 | 16.806 | 4.635 | 2.481 | 1.673 |
| 40-49 | . 000 | . 000 | . 000 | 9.141 | 2.521 | 1.350 | . 910 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 11.551 | 6.183 | 4.169 |

Male Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 7.142 | 2.851 | .154 | .000 | .000 | .000 | .000 |
| $10-19$ | .000 | 4.326 | 1.398 | 1.737 | .211 | .000 | .000 |
| $20-29$ | .000 | 12.225 | 7.865 | 3.080 | .571 | .754 | .000 |
| $30-39$ | .000 | .000 | 8.122 | 4.346 | 3.536 | 8.570 | 7.650 |
| $40-49$ | .000 | .000 | .000 | 5.741 | 3.617 | .781 | 4.068 |
| $50+$ | .000 | .000 | .000 | .000 | 3.824 | 5.739 | 5.035 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 34.310 | 5.974 | 3.088 | 1.244 | . 000 | . 000 | 000 |
| 10-19 | 59.349 | 10.334 | 5.341 | 2.152 | . 594 | . 000 | . 000 |
| 20-29 | . 000 | 15.318 | 7.917 | 3.190 | . 880 | . 471 | . 000 |
| 30-39 | . 000 | . 000 | 30.859 | 12.434 | 3.430 | 1.836 | 1.238 |
| 40-49 | . 000 | . 000 | . 000 | 6.763 | 1.865 | . 998 | . 673 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 8.546 | 4.574 | 3.085 |

Female Absolute risk (deaths/10^4PYSV)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 5.738 | 516 | . 753 | 543 | . 000 | . 000 | 000 |
| 10-1.9 | 1.014 | 1.443 | . 483 | . 888 | . 160 | . 000 | . 000 |
| 20-29 | . 000 | 5.716 | 1.411 | 2.057 | . 253 | . 941 | 000 |
| 30-39 | . 000 | . 000 | . 644 | 2.135 | 1.058 | 2.205 | 1.239 |
| 40-49 | . 000 | . 000 | . 000 | 5.800 | 1.854 | 1.136 | . 985 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 4.264 | 4.325 | 5.592 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | ---: | :--- | :--- | :--- |
| Chi-square | 1148.6400 | 3022 | 1.0000 |
| Deviance | 535.9970 | 3022 | 1.0000 |
| Freeman-Tukey | 242.9840 | 3022 | 1.0000 |

Table C.18. Excess relative and absolute risk coefficients for leukemia adjusted for confirmation rates stratified on gender and city (1950-85). organ dose equivalent adjusted for DS86 random error.

## Male Excess relative risk (b/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 47.718 | 7.481 | 3.429 | . 000 | 000 | . 000 | 000 |
| 10-19 | . 000 | 15.319 | 7.022 | 3.181 | . 908 | . 000 | 000 |
| 20-29 | . 000 | 23.181 | 10.626 | 4.813 | 1.373 | . 669 | . 000 |
| 30-39 | . 000 | . 000 | 39.619 | 17.945 | 5.121 | 2.495 | 1.320 |
| 40-49 | . 000 | . 000 | . 000 | 8.674 | 2. 475 | 1. 205 | . 638 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 15.679 | 7.639 | 4.040 |

Male Absolute risk (deaths/10^4PYSV)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 7.194 | 2.735 | 133 | 000 | . 000 | . 000 | 000 |
| 10-19 | . 000 | 4.474 | 1.361 | 1.788 | . 199 | . 000 | . 000 |
| 20--9 | . 000 | 12.747 | 7.774 | 3.321 | . 546 | . 667 | . 000 |
| $30-39$ | . 000 | . 000 | 7.930 | 4.451 | 3.781 | 7.680 | 5.299 |
| 40-49 | . 000 | . 000 | . 000 | 5.572 | 3.578 | . 708 | 2.568 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 4.210 | 6.472 | 4.505 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age AT'B | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 32.350 | 5.071 | 2.325 | 1.053 | . 000 | . 000 | 000 |
| 10-19 | 66.249 | 10.386 | 4.761 | 2.156 | . 615 | . 000 | . 000 |
| 20-29 | . 000 | 15.715 | 7.204 | 3.263 | . 931 | . 454 | . 000 |
| 30-39 | . 000 | . 000 | 26.859 | 12.166 | 3.472 | 1.691 | 895 |
| 40-49 | . 000 | . 000 | . 000 | 5.880 | 1.678 | . 818 | . 432 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 10.629 | 5.179 | 2.739 |

Female Absolute risk (deaths/10^4PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 5.671 | .463 | .539 | .397 | .000 | .000 | .000 |
| $10-19$ | 1.047 | 1.447 | .452 | .893 | .162 | .000 | .000 |
| $20-29$ | .000 | 5.761 | 1.345 | 2.090 | .284 | .763 | .000 |
| $30-39$ | .000 | .000 | .618 | 2.113 | 1.070 | 1.930 | 1.183 |
| $40-49$ | .000 | .000 | .000 | 5.335 | 1.707 | .948 | .792 |
| $50+$ | .000 | .000 | .000 | .000 | 4.725 | 4.695 | 5.515 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | ---: | :--- | :--- | :--- |
| Chi-square | 1108.8200 | 3022 | 1.0000 |
| Deviance | 502.1960 | 3022 | 1.0000 |
| Freeman-Tukey | 245.1860 | 3022 | 1.0000 |

Table C.19. Excess relative and absolute risk coefficients for leukemia adjusted for confirmation rates stratified on age ATD (1950-75). organ dose equivalent adjusted for DS86 random error.

Male Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 37.054 | 7.421 | 3.752 | . 000 | . 000 | . 000 | 000 |
| 10-19 | 49.125 | 9.839 | 4.975 | 2.253 | . 612 | . 000 | . 000 |
| 20-29 | . 000 | 16.011 | 8.095 | 3.667 | . 996 | . 684 | . 000 |
| 30-39 | . 000 | . 000 | 26.692 | 12.090 | 3.285 | 2. 255 | 1.769 |
| 40-49 | . 000 | . 000 | . 000 | 6.381 | 1.734 | 1.190 | . 934 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 8.064 | 5.537 | 4.342 |
| Male | Absolute risk (deaths/10^4PYSV) |  |  |  |  |  |  |

Age ATD

| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $<10$ | 6.691 | 2.741 | .176 | .000 | .000 | .000 | .000 |
| $10-19$ | 1.519 | 3.971 | 1.308 | 1.508 | .165 | .000 | .000 |
| $20-29$ | .000 | 12.061 | 7.074 | 3.016 | .587 | .804 | .000 |
| $30-39$ | .000 | .000 | 7.937 | 4.446 | 3.226 | 8.457 | 8.021 |
| $40-49$ | .000 | .000 | .000 | 5.418 | 3.044 | 1.073 | 4.225 |
| $50+$ | .000 | .000 | .000 | .000 | 5.025 | 7.401 | 6.259 |

Female Excess relative risk (\%/Sv)

|  | Age ATD |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 40.330 | 8.078 | 4.084 | 1. 850 | 000 | . 000 | 000 |
| 10-19 | 53.469 | 10.709 | 5.415 | 2.452 | . 666 | . 000 | . 000 |
| 20-29 | . 000 | 17.427 | 8.811 | 3.991 | 1.084 | . 745 | . 000 |
| 30-39 | . 000 | . 000 | 29.052 | 13.159 | 3.575 | 2.455 | 1.925 |
| 40-49 | . 000 | . 000 | . 000 | 6.946 | 1.887 | 1.296 | 1.016 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 8.777 | 6.026 | 4.726 |

Female Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |  |
| 10 | 5.689 | .509 | .863 | .748 | .000 | .000 | .000 |  |
| $10-19$ | .745 | 1.547 | .516 | 1.019 | .178 | .000 | .000 |  |
| $20-29$ | .000 | 5.770 | 1.569 | 2.474 | .306 | 1.437 | .000 |  |
| $30-39$ | .000 | .000 | .037 | 2.249 | 1.198 | 2.795 | 1.800 |  |
| $40-49$ | .000 | .000 | .000 | 6.282 | 1.930 | 1.508 | 1.324 |  |
| $50+$ | .000 | .000 | .000 | .000 | 4.642 | $r$ | .25 | 4.583 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | ---: | ---: | ---: |
| Chi-square | 1104.7200 | 3022 | 1.0000 |
| Deviance | 522.4780 | 3022 | 1.0000 |
| Freeman-Tukey | 252.4250 | 3022 | 1.0000 |

Table C.20. Excess relative and absolute risk coefficients for leukemia adjusted for confirmation rates stratified on age KTD (1950-85) organ dose equivalent adjusted for DS86 random error.

Male Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 36.962 | 6.672 | 3.034 | 1.496 | . 000 | 000 | 00 |
| 10-19 | 57.282 | 10.340 | 4.702 | 2.319 | . 646 | . 000 | 000 |
| 20-29 | . 000 | 16.971 | 7.717 | 3.806 | 1.060 | . 669 | 000 |
| 30-39 | . 000 | . 000 | 24.801 | 12.230 | 3.407 | 2.149 | 1.499 |
| 40-49 | . 000 | . 000 | . 000 | 6.013 | 1. 675 | 1.056 | 1.737 |
| 50 + | . 000 | . 000 | . 000 | . 000 | 9.688 | 6.110 | 4. 263 |

Male Absolute risk (deaths/10^4PYSV)

|  | Age ATD |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 6.700 | 2. 583 | . 206 | 081 | . 000 | . 000 | 000 |
| 10-19 | 1.636 | 4.063 | 1.257 | 1.576 | . 234 | . 000 | . 000 |
| 20-29 | . 000 | 12.408 | 6.872 | 3.108 | . 609 | . 888 | . 000 |
| 30-39 | . 000 | . 000 | 7.709 | 4.477 | 3.320 | 8.351 | 6.358 |
| 40-49 | . 000 | . 000 | . 000 | 5.217 | 2.967 | 8. 965 | 3.163 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 5.431 | 7.853 | 5.918 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 40.040 | 7.227 | 3.287 | 1. 621 | . 000 | . 000 | 000 |
| 10-19 | 62.052 | 11.201 | 5.094 | 2.512 | . 700 | . 000 | . 000 |
| 20-29 | . 000 | 18.384 | 8.360 | 4.123 | 1.149 | . 724 | . 000 |
| 30-39 | . 000 | . 000 | 26.866 | 13.248 | 3.691 | 2.328 | 1.624 |
| 49 | . 000 | . 000 | . 000 | 6.514 | 1.815 | 1.144 | 1.798 |
| \% | . 000 | . 000 | . 000 | . 000 | 10.495 | 6. 619 | 4.618 |

remale Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | < 20 | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 5.686 | .476 | .678 | .592 | .000 | .000 | .000 |
| $10-19$ | .782 | 1.574 | .498 | 1.015 | .196 | .000 | .000 |
| $20-29$ | .000 | 5.858 | 1.532 | 2.527 | .361 | 1.424 | .000 |
| $30-39$ | .000 | .000 | .623 | 2.254 | 1.227 | 2.817 | 1.525 |
| $40-49$ | .000 | .000 | .000 | 6.049 | 1.873 | 1.354 | 1.047 |
| $50+$ | .000 | .000 | .000 | .000 | 5.032 | 5.606 | 4.523 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | ---: | ---: | ---: | ---: |
| Chi-square | 1080.0500 | 3022 | 1.0000 |
| Deviance | 495.9100 | 3022 | 1.0000 |
| Freeman-Tukey | 243.6370 | 3022 | 1.0000 |

Table C.21. Excess relative and absolute risk coefficients for leukemia adjusted for confirmation rates stratified on DS86 shielded kerma (1950-75). organ dose equivalent adjusted for DS86 random error.

Male Excess relative risk (\%/Sv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | < 20 | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 32.038 | 6.031 | 3.270 | .000 | .000 | .000 | .000 |
| $10-19$ | 73.152 | 13.771 | 7.467 | 2.786 | .764 | .000 | .000 |
| $20-29$ | .000 | 16.793 | 9.106 | 3.397 | .931 | .645 | .000 |
| $30-39$ | .000 | .000 | 39.249 | 14.643 | 4.015 | 2.779 | 2.147 |
| $40-49$ | .000 | .000 | .000 | 9.253 | 2.537 | 1.756 | 1.357 |
| $50+$ | .000 | .000 | .000 | .000 | 15.974 | 11.056 | 8.541 |

Male Absolute risk (deaths/10^4PYSV)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 6.211 | 2.517 | . 177 | . 000 | . 000 | . 000 | . 000 |
| 10-19 | . 088 | 4. 632 | 1.799 | 1.841 | . 201 | . 000 | . 000 |
| 20-29 | . 000 | 10.081 | 7.300 | 2.788 | . 517 | . 762 | . 000 |
| 30-39 | . 000 | . 000 | 7.333 | 4.489 | 3.831 | 9.869 | 9.322 |
| 40-49 | . 000 | . 000 | . 000 | 6.488 | 3.947 | 1.486 | 5.987 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 6.210 | 10.105 | 12.109 |

Female Excess relative risk ( $\% / \mathrm{sv}$ )

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 37.350 | 7.031 | 3.813 | 1.422 | .000 | .000 | .000 |
| $10-19$ | 85.279 | 16.054 | 8.705 | 3.248 | .890 | .000 | .000 |
| $20-29$ | .000 | 19.577 | 10.616 | 3.960 | 1.086 | .752 | .000 |
| $30-39$ | .000 | .000 | 45.756 | 17.070 | 4.689 | 3.239 | 2.503 |
| $40-49$ | .000 | .000 | .000 | 10.787 | 2.957 | 2.047 | 1.581 |
| $50+$ | .000 | .050 | .000 | .000 | 18.622 | 12.888 | 9.958 |

Female Absolute risk (deaths/10^4PYSV)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 5.899 | . 509 | . 871 | . 599 | . 000 | . 000 | 000 |
| 10-19 | . 349 | 1.447 | . 593 | 1.237 | . 230 | . 000 | 000 |
| 20-29 | . 000 | 6.435 | 1.499 | 2.417 | . 307 | 1.449 | . 000 |
| 30-39 | . 000 | . 000 | . 345 | 2. 452 | 1.114 | 3.405 | 2.212 |
| 40-49 | . 000 | . 000 | . 000 | 8.254 | 2.723 | 2.219 | 2.069 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 6.556 | 6.647 | 9.205 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 991.1560 | 3022 | 1.0000 |
| Deviance | 537.1580 | 3022 | 1.0000 |
| Freman-Tukey | 269.2380 | 3022 | 1.0000 |

Table C.22. Excess relative and absolute risk coefficients for leukemia adjusted for confirmation rates stratified on DS86 shielded kerma ( $1950-85$ ). organ dose equivalent adjusted for DS86 random error.

Male Excess relative risk (s/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age hTB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 31.560 | 5. 674 | 2.903 | 1.419 | . 000 | . 000 | 000 |
| 10-19 | 78.799 | 14.167 | 7.249 | 3.543 | 1.227 | . 000 | 000 |
| 20-29 | . 000 | 17.252 | 8.828 | 4.314 | 1.495 | 1.055 | . 000 |
| 30-39 | . 000 | . 000 | 25.051 | 12.242 | 4.2 .41 | 2.995 | 2.903 |
| 40-49 | . 000 | . 000 | . 000 | 7.845 | 2.718 | 1.919 | 1.861 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 14.849 | 10.485 | 10.164 |

Male Absolute risk (deaths/10^4PYSv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 6.194 | 2.428 | . 187 | . 031 | . 000 | . 000 | 000 |
| 10-19 | . 091 | 4.686 | 1.764 | 2.198 | . 412 | . 000 | . 000 |
| 20-29 | . 000 | 10.228 | 7.123 | 3.307 | . 741 | 1.126 | . 000 |
| 30-39 | . 000 | . 000 | 6.437 | 4.227 | 3.998 | 10.703 | 10.866 |
| 40-49 | . 000 | . 000 | . 000 | 5.853 | 4.162 | 1.594 | 7.898 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 6.080 | 9.715 | 13.194 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 37360 | 6.717 | 3.437 | 1.680 | . 000 | . 000 | 000 |
| 10-19 | 280 | 16.770 | 8.581 | 4.194 | 1.453 | . 000 | . 000 |
| 20-29 | 000 | 20.422 | 10.450 | 5.107 | 1.769 | 1.249 | . 000 |
| 30-39 | . 000 | . 000 | 29.654 | 14.492 | 5.021 | 3.545 | 3.437 |
| 40-49 | . 000 | . 000 | . 000 | 9.287 | 3.217 | 2.272 | 2.203 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 17.577 | 12.411 | 12.032 |

Female Absolute risk (deaths/10^4PYSV)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 5.901 | . 494 | . 702 | . 643 | . 000 | . 000 | . 000 |
| 10-19 | . 358 | 1.467 | . 588 | 1. 400 | . 336 | . 000 | 000 |
| 20-29 | . 000 | 6.508 | 1.485 | 2.827 | . 480 | 2.174 | . 000 |
| 30-39 | . 000 | . 000 | . 313 | 2.301 | 1.174 | 3.612 | ¢ 568 |
| 40-49 | . 000 | . 000 | . 000 | 7.586 | 2.915 | 2.83 | +. 354 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 6.446 | f $\&$ ? | 1i, 4, 7 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 920.6160 | 3022 | 1.0000 |
| Deviance | 507.0910 | 3022 | 1.0000 |
| Freeman-Tukey | 277.7080 | 3022 | 1.0000 |

Table C.23. Excess relative and absolute risk coefficients for nonleu.rania.

Male Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 2.660 | 1.930 | . 627 | . 959 | . 000 | . 000 | 000 |
| 10-19 | . .000 | 1.292 | . 420 | . 642 | . 513 | . 000 | . 000 |
| 20-29 | . 000 | . 000 | . 419 | . 640 | . 512 | . 370 | . 000 |
| 30-39 | . 000 | . 000 | . 151 | . 230 | . 184 | . 133 | . 310 |
| 40-49 | . 000 | . 000 | . 000 | . 227 | . 182 | . 131 | . 306 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 087 | . 063 | .147 |
| Male | Absolut | risk | eaths/ | ^4PYSV |  |  |  |


| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | . 108 | 1.441 | 1.341 | 5.335 | . 000 | . 000 | 000 |
| 10-19 | . 000 | . 517 | . 912 | 7.701 | 15.690 | . 000 | . 000 |
| 20-29 | . 000 | . 000 | 1.426 | 7.394 | 15.981 | 26.584 | . 000 |
| 30-39 | . 000 | . 000 | . 738 | 1.923 | 6.320 | 12.224 | 47.012 |
| 40-49 | . 000 | . 000 | . 000 | 4.636 | 7.144 | 11.508 | 54.139 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 3.298 | 6.000 | 20.997 |

Female Excess relative risk (o / Sv)

| * Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 5.380 | 3.902 | 1.268 | 1.940 | . 000 | . 000 | . 000 |
| 10-19 | 3.602 | 2.612 | . 849 | 1.299 | 1.038 | . 000 | . 000 |
| 20-29 | . 000 | 2. 605 | . 847 | 1.295 | 1.035 | . 749 | . 000 |
| 30-39 | . 000 | . 000 | . 305 | . 466 | . 373 | . 269 | . 627 |
| 40-49 | . 000 | . 000 | . 000 | . 460 | . 367 | . 266 | . 619 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 177 | . 128 | 297 |

Female Absolute risk (deaths/10^4PYSv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 591 | 2. 681 | 4.985 | 10.782 | . 000 | . 000 | . 000 |
| 10-19 | 1.791 | 2.432 | 3.082 | 10.686 | 14.953 | . 000 | . 000 |
| 20-29 | . 000 | 3.290 | 3.595 | 14.294 | 21.768 | 24.194 | . 000 |
| 30-39 | . 000 | . 000 | 2.376 | 6.947 | 9.548 | 11.015 | 43.589 |
| 40-49 | . 000 | . 000 | . 000 | 7.283 | 9.457 | 12.345 | 52.494 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 4.351 | 6.837 | 25.983 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 4636.2000 | 3022 | .0000 |
| Deviance | 2159.1900 | 3022 | 1.0000 |
| Freeman-Tukey | 1909.5400 | 3022 | 1.0000 |

Table C.24. Excess relative and absolute risk coefficients for nonleukemia adjusted for confirmation rates not stratified, i.e., crude (1950-75).

Male Excess relative risk (o/Sv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | < 20 | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 1.026 | .319 | .331 | .716 | .000 | .000 | .000 |
| $10-19$ | .663 | .206 | .214 | .463 | .453 | .000 | .000 |
| $20-29$ | .000 | .170 | .176 | .381 | .373 | .396 | .000 |
| $30-39$ | .000 | .000 | .073 | .158 | .154 | .164 | .317 |
| $40-49$ | .000 | .000 | .000 | .114 | .111 | .118 | .228 |
| $50+$ | .000 | .000 | .000 | .000 | .033 | .035 | .067 |

Male Absolute risk (deaths/10^4PYSv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 2. 615 | . 980 | 1.003 | 4.085 | . 000 | 000 | 000 |
| 10-19 | 4.934 | 1.468 | 1.394 | 6.874 | 13.991 | . 000 | . 000 |
| 20-29 | . 000 | 3.100 | 2.078 | 6.602 | 12.937 | 28.428 | . 000 |
| 30-39 | . 000 | . 000 | 1.082 | 2. 657 | 6.580 | 15.406 | 47.850 |
| 40-49 | . 000 | . 000 | . 000 | 3.676 | 5.739 | 12.264 | 43.009 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 2.135 | 4.662 | 17.512 |

Female Excess relative risk (t/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 2.350 | . 732 | . 759 | 2. 640 | . 000 | . 000 | 000 |
| 10-19 | 1.518 | . 473 | . 491 | 1.060 | 1.036 | . 000 | 000 |
| 2.0-29 | . 000 | . 389 | . 404 | . 873 | . 853 | . 907 | . 000 |
| 30-39 | . 000 | . 000 | . 167 | . 361 | . 353 | . 376 | . 726 |
| 40-49 | . 000 | . 000 | . 000 | . 260 | . 254 | . 271 | . 523 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 075 | . 080 | 154 |

Female Absolute risk (deaths/10^4PYSv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 3.220 | 1.333 | 3.296 | 9.393 | . 000 | . 000 | . 000 |
| 10-19 | 8.230 | 2.204 | 2.477 | 9.065 | 14.959 | . 000 | . 000 |
| 20-29 | . 000 | 2. 648 | 2.876 | 9.757 | 17.998 | 28.879 | . 000 |
| 30-39 | . 000 | . 000 | 1.904 | 5.660 | 8.734 | 15.610 | 49.858 |
| 40-49 | . 000 | . 000 | . 000 | 5.658 | 8. 123 | 15.682 | 49.755 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 2.989 | 6.006 | 28.837 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 1998.6200 | 3022 | 1.0000 |
| Deviance | 1480.4500 | 3022 | 1.0000 |
| Freeman-Tukey | 1746.1600 | 3022 | 1.0000 |

Table C.25. Excess relative and absolute risk coefficients for nonleukeraia adjusted for confirmation rates not stratified, i.e., crude (1950-85).

Male Excess relative risk (\%/sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 1.1.35 | . 385 | . 372 | . 713 | . 000 | . 000 | 000 |
| 10-19 | . 628 | . 213 | . 206 | . 394 | . 374 | . 000 | . 000 |
| 20-29 | . 000 | . 193 | . 187 | . 358 | . 340 | . 318 | 200 |
| 30-39 | . 000 | . 000 | . 093 | . 179 | . 170 | . 159 | . 211 |
| 40-49 | . 000 | . 000 | . 000 | . 126 | . 119 | . 111 | . 148 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 052 | . 048 | 064 |

Male Absolute risk (deaths/10^4PYSV)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 2.861 | 1.171 | 1.597 | 5.601 | . 000 | . 000 | 000 |
| 10-19 | 4.671 | 1.511 | 1.342 | 5.897 | 10.848 | . 000 | . 000 |
| 20-29 | . 000 | 3.517 | 2.200 | 6.225 | 12.175 | 21.297 | . 000 |
| 30-39 | . 000 | . 000 | 1.380 | 3.004 | 7.233 | 15.373 | 33.535 |
| 40-49 | . 000 | . 000 | . 000 | 4.055 | 6.154 | 11.577 | 36.009 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 3.371 | 6.475 | 17.803 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 2.336 | . 792 | . 766 | 1.467 | . 000 | . 000 | 000 |
| 10-19 | 1. 292 | . 438 | . 424 | . 811 | . 771 | . 000 | . 000 |
| 20-29 | . 000 | . 398 | . 385 | . 737 | . 700 | . 654 | . 000 |
| 30-39 | . 000 | . 000 | . 192 | . 368 | . 350 | . 327 | . 434 |
| 40-49 | . 000 | . 000 | . 000 | . 258 | . 245 | . 229 | . 304 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 107 | . 100 | 132 |

Female Absolute risk (deaths/10^4PYSv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 3304 | 1.436 | 3.074 | 7.404 | . 000 | . 000 | 000 |
| 10-19 | 7.090 | 2.054 | 2.158 | 6.845 | 10.181 | . 000 | . 000 |
| 20-29 | . 000 | 2.703 | 2.749 | 8.376 | 14.173 | 20.270 | . 000 |
| 30-39 | . 000 | . 000 | 2.183 | 5.763 | 8.650 | 14.813 | 37.436 |
| 40-49 | . 000 | . 000 | . 000 | 5.621 | 7.844 | 13.377 | 41.644 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 4.236 | 7.493 | 28.656 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 1625.7000 | 3022 | 1.0000 |
| Deviance | 1141.8700 | 3022 | 1.0000 |
| Freeman-Tukey | 1604.5900 | 3022 | 1.0000 |

Table C.26. Excess relative and absolute risk coefficients for nonleukemia adjusted for confirmation rates stratified on gender (1950-75).

Male Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 1.054 | . 331 | . 330 | . 725 | . 000 | . 000 | 000 |
| 10-19 | . 680 | . 214 | . 213 | . 468 | . 453 | . 000 | . 000 |
| 20-29 | . 000 | . 176 | . 176 | . 386 | . 374 | . 394 | 000 |
| 30-39 | . 000 | . 000 | . 074 | . 162 | . 157 | . 165 | . 317 |
| 40-49 | . 000 | . 000 | . 000 | . 117 | . 113 | . 119 | . 229 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 033 | . 035 | . 067 |

Male Absolute risk (deaths/10^4PYSv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 2.750 | 1.044 | 1.017 | 4.133 | . 000 | . 000 | 000 |
| 10-19 | 5.192 | 1.559 | 1.430 | 7.069 | 14.011 | . 000 | 000 |
| 20-29 | . 000 | 3.298 | 2.133 | 6.925 | 13.212 | 28.283 | . 000 |
| 30-39 | . 000 | . 000 | 1.124 | 2.812 | 6.926 | 15.738 | 47.978 |
| 40-49 | . 000 | . 000 | . 000 | 3.913 | 6.061 | 12.856 | 43.685 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 2.222 | 4.818 | 17.920 |

Female Excess relative risk ( $8 / \mathrm{SV}$ )

## Age ATD

| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $<10$ | 2.414 | .758 | .756 | 1.661 | .000 | .000 | .000 |
| $10-19$ | 1.558 | .489 | .488 | 1.072 | 1.039 | .000 | .000 |
| $20-29$ | .000 | .404 | .402 | .884 | .856 | .903 | .000 |
| $30-39$ | .000 | .000 | .169 | .370 | .359 | .378 | .727 |
| $40-49$ | .000 | .000 | .000 | .268 | .259 | .273 | .525 |
| $50+$ | .000 | .000 | .000 | .000 | .076 | .080 | .153 |

Female Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| 10 | 3.096 | 1.308 | 3.238 | 9.495 | .000 | .000 | .000 |
| $10-19$ | 7.926 | 2.150 | 2.352 | 8.992 | 14.984 | .000 | .000 |
| $20-29$ | .000 | 2.587 | 2.728 | 9.485 | 17.894 | 28.749 | .000 |
| $30-39$ | .000 | .000 | 1.829 | 5.558 | 8.512 | 15.466 | 49.901 |
| $40-49$ | .000 | .000 | .000 | 5.552 | 7.908 | 15.135 | 49.234 |
| $50+$ | .000 | .000 | .000 | .000 | 2.863 | 5.715 | 27.234 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 2013.9200 | 3022 | 1.0000 |
| Deviance | 1482.2600 | 3022 | 1.0000 |
| Freeman-Tukey | 1750.8700 | 3022 | 1.0000 |

Table c.27. Excess relative and absolute risk coefficients for nonleukemia adjusted for confirmation rates stratified on gender (1950-85).

Male Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 1.162 | . 398 | . 370 | . 715 | . 000 | . 000 | 000 |
| 10-19 | . 639 | . 219 | . 204 | . 393 | . 374 | . 000 | . 000 |
| 20-29 | . 000 | . 200 | . 186 | . 360 | . 343 | . 318 | . 000 |
| 30-39 | . 000 | . 000 | . 094 | . 182 | . 173 | . 160 | . 212 |
| 40-49 | . 000 | . 000 | . 000 | . 127 | . 121 | . 112 | . 148 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 052 | . 048 | . 064 |
| Male | Absolut | risk | eaths / | ^4PYSV |  |  |  |

Age ATD

| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| <10 | 2.999 | 1.246 | 1.641 | 5.820 | . 000 | . 000 | 000 |
| 10-19 | 4.882 | 1.596 | 1.370 | 6.108 | 11.295 | . 000 | . 000 |
| 20-29 | . 000 | 3.741 | 2.262 | 6.485 | 12.744 | 22.162 | . 000 |
| 30-39 | . 000 | . 000 | 1.433 | 3.151 | 7.634 | 16.119 | 35.115 |
| 40-49 | . 000 | . 000 | . 000 | 4.244 | 6.465 | 12.072 | 37.139 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 3.518 | 6.712 | 18.382 |

Female Excess relative risk (\%/Sv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $.30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 2.396 | .820 | .763 | 1.475 | .000 | .000 | .000 |
| $10-19$ | 1.318 | .451 | .420 | .811 | .772 | .000 | .000 |
| $20-29$ | .000 | .413 | .385 | .743 | .707 | .655 | .000 |
| $30-39$ | .000 | .000 | .194 | .374 | .356 | .330 | .438 |
| $40-49$ | .000 | .000 | .000 | .262 | .249 | .231 | .306 |
| $50+$ | .000 | .000 | .000 | .000 | .108 | .100 | .133 |

Female Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| 10 | 3.076 | 1.407 | 2.943 | 7.160 | .000 | .000 | .000 |
| $10-19$ | 6.798 | 1.995 | 2.044 | 6.575 | 9.818 | .000 | .000 |
| $20-29$ | .000 | 2.645 | 2.616 | 8.105 | 13.756 | 19.508 | .000 |
| $30-39$ | .000 | .000 | 2.099 | 5.620 | 8.467 | 14.339 | 36.116 |
| $40-49$ | .000 | .000 | .000 | 5.434 | 7.611 | 12.868 | 39.856 |
| $50+$ | .000 | .000 | .000 | .000 | 4.078 | 7.165 | 27.178 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 1645.2500 | 3022 | 1.0000 |
| Deviance | 1145.5800 | 3022 | 1.0000 |
| Freeman-Tukey | 1594.7300 | 3022 | 1.0000 |

Table C.28. Excess relative and absolute risk coefficients for nonleukemia adjusted for confirmation rates stratified on gender and city (1950-75).

Male Excess relative risk (\%/Sv)

|  |  |  |  |  |  |  |  |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 1.065 | .336 | .323 | .726 | .000 | .000 | .000 |
| $10-19$ | .691 | .218 | .210 | .471 | .452 | .000 | .000 |
| $20-29$ | .000 | .179 | .173 | .387 | .372 | .397 | .000 |
| $30-39$ | .000 | .000 | .072 | .161 | .155 | .165 | .318 |
| $40-49$ | .000 | .000 | .000 | .116 | .112 | .119 | .229 |
| $50+$ | .000 | .000 | .000 | .000 | .032 | .034 | .066 |

Male Absolute risk (deaths/10^4PYSv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 2.784 | 1.063 | . 999 | 4.135 | . 000 | . 000 | 000 |
| 10-19 | 5.394 | 1.594 | 1.410 | 7.104 | 13.972 | . 000 | . 000 |
| 20-29 | . 000 | 3.367 | 2.093 | 6.942 | 13.147 | 28.448 | . 000 |
| 30-39 | . 000 | . 000 | 1.094 | 2.800 | 6.836 | 15.725 | 48.059 |
| 40-49 | . 000 | . 000 | . 000 | 3.884 | 5.973 | 12.808 | 43.697 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 2.178 | 4.774 | 17.776 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-.59 | 60-69 | $70+$ |
| $<10$ | 2.448 | . 773 | . 743 | 1.669 | . 000 | . 000 | . 000 |
| 10-19 | 1.588 | . 501 | . 482 | 1. 082 | 1.039 | . 000 | . 000 |
| 20-29 | . 000 | . 412 | . 397 | . 890 | . 855 | . 912 | . 000 |
| 30-39 | . 000 | . 000 | . 165 | . 370 | . 356 | . 379 | . 731 |
| 40-49 | . 000 | . 000 | . 000 | . 267 | . 257 | . 274 | . 527 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 074 | . 079 | 153 |

Female Absolute risk (deaths/10^4pYSV)

|  | ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age AT |  |  |  |  |  |  |  |  |
| $<10$ | 3.021 | 1.305 | 3.179 | 9.530 | .000 | .000 | .000 |  |
| $10-19$ | 7.519 | 2.173 | 2.329 | 9.064 | 14.991 | .000 | .000 |  |
| $20-29$ | .000 | 2.620 | 2.672 | 9.514 | 17.867 | 29.003 | .000 |  |
| $30-39$ | .000 | .000 | 1.769 | 5.533 | 8.407 | 15.508 | 50.152 |  |
| $40-49$ | .000 | .000 | .000 | 5.529 | 7.830 | 15.122 | 49.401 |  |
| $50+$ | .000 | .000 | .000 | .000 | 2.823 | 5.672 | 27.152 |  |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 2024.1000 | 3022 | 1.0000 |
| Deviance | 1483.0000 | 3022 | 1.0000 |
| Freman-Tukey | 1748.2500 | 3022 | 1.0000 |

Table C.29. Excess relative and absolute risk coefficients for nonleukemia adjusted for confirmation rates stratified on gender and city (1950-85).

Male Excess relative risk (\%/Sv)

|  |  |  |  |  |  |  |  |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 1.173 | .408 | .372 | .723 | .000 | .000 | .000 |
| $10-19$ | .639 | .222 | .203 | .394 | .372 | .000 | .000 |
| $20-29$ | .000 | .202 | .184 | .357 | .338 | .319 | .000 |
| $30-39$ | .000 | .000 | .093 | .181 | .171 | .161 | .212 |
| $40-49$ | .000 | .000 | .000 | .125 | .119 | .112 | .147 |
| $50+$ | .000 | .000 | .000 | .000 | .051 | .049 | .064 |

Male Absolute risk (deaths/10^4PYSv)
Age ATD

| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $<10$ | 3.034 | 1.280 | 1.656 | 5.907 | .000 | .000 | .000 |
| $10-19$ | 5.000 | 1.625 | 1.368 | 6.125 | 11.228 | .000 | .000 |
| $20-29$ | .000 | 3.776 | 2.231 | 6.429 | 12.553 | 22.267 | .000 |
| $30-39$ | .000 | .000 | 1.415 | 3.134 | 7.531 | 16.213 | 35.141 |
| $40-49$ | .000 | .000 | .000 | 4.186 | 6.337 | 12.070 | 37.149 |
| $50+$ | .000 | .000 | .000 | .000 | 3.455 | 6.724 | 18.244 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 2.431 | . 846 | . 771 | 1.498 | . 000 | . 000 | 000 |
| 10-19 | 1.325 | . 461 | . 420 | . 817 | . 772 | . 000 | . 000 |
| 20-29 | . 000 | . 418 | . 381 | . 741 | . 700 | . 662 | . 000 |
| 30-39 | . 000 | . 000 | . 193 | . 375 | . 354 | . 335 | . 440 |
| 40-49 | . 000 | . 000 | . 000 | . 260 | . 246 | . 232 | . 306 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 107 | . 101 | . 133 |
| Female | Absolute risk (deaths/10^4PYSV) |  |  |  |  |  |  |

Age ATD

| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $<10$ | 3.003 | 1.419 | 2.956 | 7.280 | 000 | . 000 | 000 |
| 10-19 | 6.370 | 2.012 | 2.050 | 6.613 | 9.806 | . 000 | . 000 |
| 20-29 | . 000 | 2.654 | 2.576 | 8.056 | 13.663 | 19.599 | . 000 |
| 30-39 | . 000 | . 000 | 2.063 | 5.594 | 8.369 | 14.510 | 36.187 |
| 40-49 | . 000 | . 000 | . 000 | 5.384 | 7.505 | 12.919 | 39.814 |
| $50+$ | . 000 | . 000 | . 000 | 000 | 4.033 | 7.198 | 27.148 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 1658.0000 | 3022 | 1.0000 |
| Deviance | 1148.4900 | 3022 | 1.0000 |
| Freeman-Tukey | 1594.3500 | 3022 | 1.0000 |

Table C.30. Excess relative and absolute risk coefficients for nonleukemia adjusted for confirmation rates stratified on age ATD (1950-75).

Male Excess relative risk (\%/Sv)

## Age ATD

| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $<10$ | 1.294 | .863 | .521 | .914 | .000 | .000 | .000 |
| $10-19$ | .793 | .529 | .319 | .560 | .486 | .000 | .000 |
| $20-29$ | .000 | .490 | .296 | .519 | .450 | .389 | .000 |
| $30-39$ | .000 | .000 | .113 | .198 | .171 | .148 | .313 |
| $40-49$ | .000 | .000 | .000 | .167 | .145 | .125 | .265 |
| $50+$ | .000 | .000 | .000 | .000 | .054 | .047 | .099 |

Male Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 1.200 | 1.403 | 1.307 | 5.110 | .000 | .000 | .000 |
| $10-19$ | 2.085 | 1.461 | 1.230 | 7.460 | 14.929 | .000 | .000 |
| $20-29$ | .000 | 3.123 | 1.953 | 7.491 | 15.356 | 27.871 | .000 |
| $30-39$ | .000 | .000 | .988 | 2.370 | 7.103 | 14.217 | 47.387 |
| $40-49$ | .000 | .000 | .000 | 4.355 | 7.081 | 13.031 | 49.802 |
| $50+$ | .000 | .000 | .000 | .000 | 2.954 | 5.820 | 21.703 |

Female Excess relative risk (\%/Sv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| 10 | 2.767 | 1.844 | 1.114 | 1.954 | .000 | .000 | .000 |
| $10-19$ | 1.696 | 1.130 | .682 | 1.198 | 1.039 | .000 | .000 |
| $20-29$ | .000 | 1.047 | .632 | 1.110 | .963 | .831 | .000 |
| $30-39$ | .000 | .000 | .241 | .422 | .366 | .316 | .669 |
| $40-49$ | .000 | .000 | .000 | .357 | .310 | .268 | .565 |
| $50+$ | .000 | .000 | .000 | .000 | .116 | .101 | .213 |

Female Absolute risk (deaths/10^4PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | < 20 | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 1.577 | 2.130 | 4.606 | 10.859 | .000 | .000 | .000 |
| $10-19$ | 3.876 | 2.623 | 2.984 | 10.268 | 14.976 | .000 | .000 |
| $20-29$ | .000 | 3.460 | 3.508 | 13.098 | 20.734 | 26.632 | .000 |
| $30-39$ | .000 | .000 | 2.303 | 6.815 | 10.147 | 13.476 | 46.261 |
| $40-49$ | .000 | .000 | .000 | 6.790 | 9.654 | 15.179 | 52.357 |
| $50+$ | .000 | .000 | .000 | .000 | 3.978 | 6.999 | 31.444 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 3098.7800 | 3022 | 1616 |
| Deviance | 1756.9900 | 3022 | 1.0000 |
| Freeman-Tukey | 1864.1500 | 3022 | 1.0000 |

Table C.31. Excess relative and absolute risk coefficients for nonleukemia adjust nd for confirmation rates stratified on age ATD (1950-85).

Male Excess relative risk (z/Sv)
Age ATD

| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| <10 | 1.325 | . 887 | . 523 | . 891 | . 000 | . 000 | 000 |
| 10-19 | . 799 | . 535 | . 315 | . 537 | . 455 | . 000 | 000 |
| 20-29 | .000 | . 508 | . 300 | . 510 | . 433 | . 349 | 000 |
| 30-39 | . 000 | . 000 | . 123 | . 210 | . 178 | . 143 | 256 |
| 40-49 | . 000 | . 000 | . 000 | . 175 | . 149 | 120 | 214 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 063 | . 051 | 091 |

Male Absolute risk (deaths/10^4PYSv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 1.224 | 1.439 | 1.591 | 5.972 | 000 | . 000 | 000 |
| 10-19 | 2.099 | 1.477 | 1.217 | 7.411 | 14.903 | . 000 | . 000 |
| 20-29 | . 000 | 3.235 | 1.980 | 7.377 | 15.761 | 26.575 | . 000 |
| 30-39 | . 000 | . 000 | 1.079 | 2.511 | 7.359 | 14.759 | 43.917 |
| 40-49 | . 000 | . 000 | . 000 | 4.567 | 7.260 | 12.482 | 49.575 |
| $50+$ | 000 | 000 | 000 | . 000 | 3.431 | 6.314 | 20.883 |

Female Excess relative risk (\%/Sv)
Age ATD

| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| <10 | 2.753 | 1.843 | 1.088 | 1.851 | . 000 | . 000 | 000 |
| 10-19 | 1.660 | 1.111 | . 656 | 1.116 | . 946 | . 000 | 000 |
| 20-29 | . 000 | 1.056 | . 623 | 1.061 | . 899 | . 724 | . 000 |
| 30-39 | . 000 | . 000 | . 256 | 436 | . 369 | . 297 | 532 |
| 40-49 | . 000 | . 000 | . 000 | . 365 | . 309 | . 249 | 445 |
| $50+$ | . 000 | 000 | . 000 | . 000 | . 132 | 106 | 190 |

Female Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 1.570 | 2.130 | 4.575 | 10.506 | .000 | .000 | .000 |
| $10-19$ | 3.805 | 2.585 | 2.877 | 9.835 | 14.424 | .000 | .000 |
| $20-29$ | .000 | 3.486 | 3.462 | 12.593 | 20.477 | 25.095 | .000 |
| $30-39$ | .000 | .000 | 2.444 | 7.019 | 10.221 | 13.988 | 45.530 |
| $40-49$ | .000 | .000 | .000 | 6.924 | 9.625 | 14.164 | 53.982 |
| $50+$ | .000 | .000 | .000 | .000 | 4.488 | 7.379 | 31.255 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 2929.7400 | 3022 | .8831 |
| Deviance | 1582.8300 | 3022 | 1.0000 |
| Freeman-Tukey | 1805.5000 | 3022 | 1.0000 |

Table C.32. Excess relative and absolute risk coefficients for nonleukemia adjusted for confirmation rates stratified on DS86 shielded kerma (1950-75)

Male Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | 70 |
| <10 | 1.793 | . 885 | 591 | . 929 | . 000 | 000 | 000 |
| 10-19 | 1.130 | . 558 | . 373 | . 585 | . 523 | . 000 | 000 |
| 20-29 | . 000 | . 538 | . 360 | . 565 | . 505 | . 400 | . 000 |
| 30-39 | . 000 | . 000 | . 166 | . 261 | . 233 | . 185 | . 316 |
| 40-49 | . 000 | . 000 | . 000 | . 247 | . 221 | . 175 | . 300 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 145 | . 115 | . 197 |

Male Absolute risk (deaths/10^4PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| 10 | 2.209 | 1.695 | 1.558 | 5.178 | .000 | .000 | .000 |
| $10-19$ | 4.071 | 2.087 | 1.670 | 8.096 | 15.973 | .000 | .000 |
| $20-29$ | .000 | 4.811 | 2.790 | 8.720 | 17.209 | 28.642 | .000 |
| $30-39$ | .000 | .000 | 1.703 | 3.477 | 9.692 | 17.718 | 47.819 |
| $40-49$ | .000 | .000 | .000 | 6.917 | 10.844 | 18.229 | 55.958 |
| $50+$ | .000 | .000 | .000 | .000 | 7.955 | 14.135 | 41.638 |

Female Excess relative risk (o/Sv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 3.559 | 1.757 | 1.174 | 1.843 | .000 | .000 | .000 |
| $10-19$ | 2.243 | 1.107 | .740 | 1.162 | 1.038 | .000 | .000 |
| $20-29$ | .000 | 1.068 | .714 | 1.121 | 1.002 | .795 | .000 |
| $30-39$ | .000 | .000 | .329 | .517 | .462 | .367 | .628 |
| $40-49$ | .000 | .000 | .000 | .491 | .439 | .348 | .596 |
| $50+$ | .000 | .000 | .000 | .000 | .288 | .228 | .391 |

Female Absolute risk (deaths/10^4PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 2.501 | 2.298 | 4.898 | 10.334 | .000 | .000 | .000 |
| $10-19$ | 6.310 | 3.133 | 3.470 | 10.195 | 14.967 | .000 | .000 |
| $20-29$ | .000 | 4.174 | 4.305 | 13.656 | 21.476 | 25.555 | .000 |
| $30-39$ | .000 | .000 | 3.368 | 8.537 | 12.632 | 15.551 | 43.516 |
| $40-49$ | .000 | .000 | .000 | 9.810 | 13.540 | 19.563 | 54.643 |
| $50+$ | .000 | .000 | .000 | .000 | 9.879 | 15.676 | 55.178 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 2787.9100 | 3022 | .9990 |
| Deviance | 1758.5000 | 3022 | 1.0000 |
| Freeman-Tukey | 1861.5000 | 3022 | 1.0000 |

Table C.33. Excess relative and absolute risk coefficients for nonleukemia adjusted for confirmation rates stratified on DS86 shielded kerma ( $1950-85$ )

Male Excess relative risk (\%/Sv)

|  |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 1.851 | .935 | .628 | .956 | .000 | .000 | .000 |
| $10-19$ | 1.110 | .561 | .377 | .573 | .514 | .000 | .000 |
| $20-29$ | .000 | .554 | .372 | .566 | .507 | .407 | .000 |
| $30-39$ | .000 | .000 | .181 | .275 | .247 | .198 | .322 |
| $40-49$ | .000 | .000 | .000 | .254 | .228 | .183 | .298 |
| 50 | .000 | .000 | .000 | .000 | .151 | .122 | .198 |

Male Absolute risk (deaths/10^4PYSV)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 2.268 | 1.783 | 2.104 | 6.855 | . 000 | . 000 | . 000 |
| 10-19 | 4.001 | 2.099 | 1.687 | 8.354 | 16.804 | . 000 | . 000 |
| 20-29 | . 000 | 4.941 | 2.881 | 8.738 | 18.473 | 31.026 | . 000 |
| 30-39 | . 000 | . 000 | 1.852 | 3.664 | 10.253 | 20.363 | 54.793 |
| 40-49 | . 000 | . 000 | . 000 | 7.113 | 11.192 | 19.039 | 67.384 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 8.291 | 14.911 | 43.730 |
| Female | Excess relative risk (\%/Sv) |  |  |  |  |  |  |

Age ATD

| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| <10 | 3.549 | 1.793 | 1.204 | 1.833 | 000 | . 000 | 000 |
| 10-19 | 2.128 | 1.075 | . 722 | 1.099 | . 986 | . 000 | . 000 |
| 20-29 | . 000 | 1.061 | . 713 | 1.085 | . 973 | . 781 | . 000 |
| 30-39 | . 000 | . 000 | . 346 | . 527 | . 473 | . 380 | . 618 |
| 40-49 | . 000 | . 000 | . 000 | . 488 | . 437 | . 351 | . 572 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 290 | . 233 | 379 |

Female Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 2.495 | 2.339 | 5.154 | 10.682 | .000 | .000 | .000 |
| $10-19$ | 6.030 | 3.056 | 3.396 | 10.063 | 14.901 | .000 | .000 |
| $20-29$ | .000 | 4.151 | 4.302 | 13.273 | 21.892 | 27.061 | .000 |
| $30-39$ | .000 | .000 | 3.537 | 8.695 | 12.912 | 17.702 | 51.996 |
| $40-49$ | .000 | .000 | .000 | 9.755 | 13.511 | 19.752 | 67.100 |
| $50+$ | .000 | .000 | .000 | .000 | 9.950 | 15.975 | 59.348 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 2608.3700 | 3022 | 1.0000 |
| Deviance | 1584.5700 | 3022 | 1.0000 |
| Freeman-Tukey | 1815.6600 | 3022 | 1.0000 |

Table C.34. Excess relative and absolute risk coefficients for nonleukemia. organ dose equivalent adjusted for DS86 random error.

Male Excess relative risk. (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 2.956 | 2.244 | . 787 | 1.167 | . 000 | . 000 | 000 |
| 10-19 | . 000 | 1.516 | . 532 | . 788 | . 667 | . 000 | . 000 |
| 20-29 | . 000 | . 000 | . 544 | . 807 | . 683 | . 489 | . 000 |
| 30-39 | . 000 | . 000 | . 190 | . 282 | . 239 | . 171 | . 364 |
| 40-49 | . 000 | . 000 | . 000 | . 295 | . 250 | . 179 | . 381 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 118 | . 085 | . 181 |

Male Absolute risk (deaths/10^4PYSv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | . 115 | 1. 641 | 1. 658 | 6.390 | 000 | . 000 | 000 |
| 10-19 | . 000 | . 620 | 1. 164 | 9.416 | 20.161 | . 000 | . 000 |
| 20-29 | . 000 | . 000 | 1.828 | 9.212 | 21.220 | 35.146 | . 000 |
| 30-39 | . 000 | . 000 | . 927 | 2.350 | 8.205 | 15.675 | 55.273 |
| 40-49 | . 000 | . 000 | . 000 | 5.986 | 9.788 | 15.582 | 67.153 |
| 50 + | . 000 | . 000 | . 000 | . 000 | 4.466 | 8.038 | 25.741 |

Female Excess relative risk (\%/Sv)

## Age ATD

| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| <10 | 5.491 | 6. 169 | 1.462 | 2.167 | . 000 | 000 | 000 |
| 10-19 | 3.708 | 2.815 | . 988 | 1.463 | 1.239 | . 000 | 000 |
| 20-29 | . 000 | 2.883 | 1.011 | 1.498 | 1.269 | . 908 | . 000 |
| 30-39 | . 000 | . 000 | . 353 | . 523 | . 443 | . 317 | 676 |
| 40-49 | . 000 | . 000 | . 000 | . 548 | . 464 | . 332 | 709 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 220 | . 157 | . 335 |

Female Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | .629 | 2.891 | 5.713 | 12.021 | .000 | .000 | .000 |
| $10-19$ | 1.903 | 2.656 | 3.609 | 12.101 | 17.748 | .000 | .000 |
| $20-29$ | .000 | 3.670 | 4.261 | 16.438 | 26.443 | 29.222 | .000 |
| $30-39$ | .000 | .000 | 2.746 | 7.794 | 11.332 | 12.967 | 47.324 |
| $40-49$ | .000 | .000 | .000 | 8.655 | 11.884 | 15.341 | 59.873 |
| $50+$ | .000 | .000 | .000 | .000 | 5.409 | 8.387 | 29.201 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 4619.1500 | 3022 | .0000 |
| Deviance | 2158.8800 | 3022 | 1.0000 |
| Freeman-Tukey | 1905.4100 | 3022 | 1.0000 |

Table C.35. Excess relative and absolute risk coefficients for nonleukemia adjusted for confirmation rates not stratified, i.e., crude(1950-75). organ dose equivalent adjusted for DS86 random error.

Male Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 1.217 | . 379 | . 446 | . 881 | . 000 | . 000 | 000 |
| 10-19 | . 786 | . 245 | . 288 | . 569 | . 586 | . 000 | . 000 |
| 20-29 | . 000 | . 207 | . 243 | . 481 | . 495 | . 525 | . 000 |
| 30-39 | . 000 | . 000 | . 096 | . 189 | . 195 | . 207 | . 384 |
| 40-49 | . 000 | . 000 | . 000 | . 141 | . 145 | . 154 | . 287 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 041 | . 044 | 082 |

Male Absolute risk (deaths/10^4PYSV)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 3.070 | 1.154 | 1.335 | 4.961 | . 000 | . 000 | . 000 |
| 10-19 | 5.831 | 1.739 | 1.871 | 8.422 | 17.923 | . 000 | . 000 |
| 20-29 | . 000 | 3.798 | 2.843 | 8.275 | 17.103 | 37.739 | . 000 |
| 30-39 | . 000 | . 000 | 1.408 | 3.193 | 8.335 | 19.434 | 58.202 |
| 40-49 | . 000 | . 000 | . 000 | 4.561 | 7.498 | 15.966 | 53.810 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 2.700 | 5.865 | 21.255 |
| Female | Excess | lativ | risk | (SV) |  |  |  |

## Age ATD

| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $<10$ | 2.562 | .798 | .939 | 1.855 | .000 | .000 | .000 |
| $10-19$ | 1.655 | .515 | .606 | 1.198 | 1.233 | .000 | .000 |
| $20-29$ | .000 | .435 | .512 | 1.012 | 1.041 | 1.105 | .000 |
| $30-39$ | .000 | .000 | .202 | .399 | .410 | .435 | .809 |
| $40-49$ | .000 | .000 | .000 | .297 | .306 | .325 | .603 |
| $50+$ | .000 | .000 | .000 | .000 | .087 | .093 | .172 |

Female Absolute risk (deaths/10^4PYSV)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 3.534 | 1.455 | 4.035 | 10.593 | . 000 | . 000 | 000 |
| 10-19 | 8.957 | 2.406 | 3.043 | 10.301 | 17.697 | . 000 | 000 |
| 20-29 | . 000 | 2.966 | 3.628 | 11.277 | 21.816 | 34.987 | . 000 |
| 30-39 | . 000 | . 000 | 2.294 | 6.252 | 10.134 | 18.102 | 55.873 |
| 40-49 | . 000 | . 000 | . 000 | 6.473 | 9.749 | 18.770 | 57.244 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 3.473 | 6.954 | 32.251 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 2004.1000 | 3022 | 1.0000 |
| Deviance | 1480.2700 | 3022 | 1.0000 |
| Freeman-Tukey | 1765.6700 | 3022 | 1.0000 |

Table C.36. Excess relative and absolute risk coefficients for nonleukemia adjusted for confirmation rates not stratified, i.e., crude (1950-85). organ dose equivalent adjusted for DS86 random error.

| Male | Excess relative risk (\%/Sv) |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Age ATD |  |  |  |  |  |  |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 1.310 | . 437 | . 473 | . 845 | . 000 | . 000 | . 000 |
| 10-19 | . 738 | . 246 | . 267 | . 476 | . 469 | . 000 | . 000 |
| 20-29 | . 000 | . 230 | . 248 | . 444 | . 437 | . 400 | . 000 |
| 30-39 | . 000 | . 000 | . 117 | . 208 | . 205 | . 188 | . 246 |
| 40-49 | . 000 | . 000 | . 000 | . 156 | . 154 | . 141 | . 184 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 065 | . 059 | . 077 |

Male Absolute risk (deaths/10^4PYSv)


Age ATD

| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $<10$ | 3.277 | 1.323 | 2.014 | 6.572 | .000 | .000 | .000 |
| $10-19$ | 5.484 | 1.750 | 1.738 | 7.114 | 13.513 | .000 | .000 |
| $20-29$ | .000 | 4.209 | 2.904 | 7.670 | 15.624 | 26.870 | .000 |
| $30-39$ | .000 | .000 | 1.713 | 3.511 | 8.789 | 18.257 | 39.354 |
| $40-49$ | .000 | .000 | .000 | 5.030 | 7.928 | 14.580 | 44.704 |
| $50+$ | .000 | .000 | .000 | .000 | 4.193 | 7.858 | 21.451 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 2.567 | . 856 | . 927 | 1. 655 | . 000 | . 000 | 000 |
| 10-19 | 1.447 | . 483 | . 522 | . 933 | . 919 | . 000 | . 000 |
| 20-29 | . 000 | . 450 | . 487 | . 869 | . 857 | . 783 | . 000 |
| 30-39 | . 000 | . 000 | . 229 | . 4.08 | . 403 | . 368 | . 483 |
| 40-49 | . 000 | . 000 | . 000 | . 306 | . 301 | . 276 | . 361 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 126 | . 116 | 152 |

Female Absolute risk (deaths/10^4PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | 704 |
| $<10$ | 3.538 | 1.555 | 3.692 | 8.324 | .000 | .000 | .000 |
| $10-19$ | 7.918 | 2.265 | 2.648 | 7.892 | 12.086 | .000 | .000 |
| $20-29$ | .000 | 3.060 | 3.460 | 9.833 | 17.229 | 24.248 | .000 |
| $30-39$ | .000 | .000 | 2.597 | 6.406 | 9.964 | 16.757 | 41.997 |
| $40-49$ | .000 | .000 | .000 | 6.651 | 9.605 | 16.034 | 49.226 |
| $50+$ | .000 | .000 | .000 | .000 | 5.013 | 8.665 | 32.788 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 1635.7500 | 3022 | 1.0000 |
| Deviance | 1142.3300 | 3022 | 1.0000 |
| Freeman-Tukey | 1589.2800 | 3022 | 1.0000 |

Table C.37. Excess relative and absolute risk coefficients for nonleukemia adjusted for confirmation rates stratified on gender (1950-75). organ dose equivalent adjusted for DS86 random error.

Male Excess relative risk (\%/Sv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 1.245 | .392 | .443 | .890 | .000 | .000 | .000 |
| $10-19$ | .804 | .253 | .286 | .575 | .586 | .000 | .000 |
| $20-29$ | .000 | .214 | .242 | .486 | .496 | .522 | .000 |
| $30-39$ | .000 | .000 | .096 | .194 | .197 | .208 | .384 |
| $40-49$ | .000 | .000 | .000 | .145 | .148 | .156 | .288 |
| $50+$ | .000 | .000 | .000 | .000 | .042 | .044 | .081 |

Male Absolute risk (deaths/10^4PYSV)

| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| <10 | 3.219 | 1.228 | 1.355 | 5.009 | . 000 | . 000 | 000 |
| 10-19 | 6.122 | 1.846 | 1.920 | 8.650 | 17.934 | . 000 | . 000 |
| 20-29 | . 000 | 4.037 | 2.921 | 8.670 | 17.451 | 37.530 | . 000 |
| 30-39 | . 000 | . .000 | 1.462 | 3.374 | 8.763 | 19.826 | 58.175 |
| 40-49 | . 000 | . 000 | . 000 | 4.851 | 7.915 | 16.729 | 54.641 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 2.803 | 6.047 | 21.696 |

Female Excess relative risk (\%/Sv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 2.625 | .825 | .934 | 1.876 | .000 | .000 | .000 |
| $10-19$ | 1.695 | .533 | .603 | 1.212 | 1.235 | .000 | .000 |
| $20-29$ | .000 | .451 | .511 | 1.025 | 1.045 | 1.100 | .000 |
| $30-39$ | .000 | .000 | .203 | .408 | .416 | .438 | .809 |
| $40-49$ | .000 | .000 | .000 | .306 | .312 | .328 | .606 |
| $50+$ | .000 | .000 | .000 | .000 | .088 | .092 | .170 |

Female Absolute risk (deaths/10^4PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 3.390 | 1.426 | 3.966 | 10.696 | .000 | .000 | .000 |
| $10-19$ | 8.615 | 2.346 | 2.893 | 10.218 | 17.725 | .000 | .000 |
| $20-29$ | .000 | 2.898 | 3.446 | 10.959 | 21.694 | 34.840 | .000 |
| $30-39$ | .000 | .000 | 2.205 | 6.134 | 9.871 | 17.935 | 55.906 |
| $40-49$ | .000 | .000 | .000 | 6.350 | 9.493 | 18.122 | 56.665 |
| $50+$ | .000 | .000 | .000 | .000 | 3.321 | 6.607 | 30.407 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 2019.6500 | 3022 | 1.0000 |
| Deviance | 1482.0900 | 3022 | 1.0000 |
| Freonan-Tukey | 1751.7700 | 3022 | 1.0000 |

Table C.38. Excess relative and absolute risk coefficients for nonleukemia adjusted for confirmation rates stratified on gender (1950-85) organ dose equivalent adjusted for DS86 randon error.

Male Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 1.339 | . 452 | . 470 | . 846 | . 000 | . 000 | 000 |
| 10-19 | . 752 | . 254 | . 264 | . 475 | . 470 | . 000 | . 000 |
| 20-29 | . 000 | . 239 | . 248 | . 447 | . 442 | . 401 | . 000 |
| 30-39 | . 000 | . 000 | . 118 | . 212 | . 209 | . 190 | . 248 |
| 40-49 | . 000 | . 000 | . 000 | . 158 | . 156 | . 142 | . 185 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 065 | . 059 | . 077 |

Male Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 3.431 | 1.407 | 2.070 | 6.819 | .000 | .000 | .000 |
| $10-19$ | 5.732 | 1.850 | 1.777 | 7.369 | 14.091 | .000 | .000 |
| $20-29$ | .000 | 4.486 | 2.993 | 7.997 | 16.393 | 28.022 | .000 |
| $30-39$ | .000 | .000 | 1.781 | 3.682 | 9.288 | 19.160 | 41.209 |
| $40-49$ | .000 | .000 | .000 | 5.266 | 8.345 | 15.229 | 46.519 |
| $50+$ | 000 | .000 | .000 | .000 | 4.375 | 8.141 | 22.115 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 2.627 | . 886 | . 922 | 1. 659 | . 000 | . 000 | 000 |
| 10-19 | 1.476 | . 498 | . 518 | . 932 | . 922 | . 000 | . 000 |
| 20-29 | . 000 | . 468 | . 487 | . 876 | . 867 | . 786 | . 000 |
| 30-39 | . 000 | . 000 | . 231 | . 415 | . 411 | . 372 | 487 |
| 40-49 | . 000 | . 000 | . 000 | . 310 | . 306 | . 278 | . 363 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 128 | . 116 | 151 |

Female Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATS | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 3.391 | 1.522 | 3.531 | 8.033 | .000 | .000 | .000 |
| $10-19$ | 7.588 | 2.201 | 2.510 | 7.575 | 11.664 | .000 | .000 |
| $20-29$ | .000 | 2.998 | 3.298 | 9.516 | 16.745 | 23.364 | .000 |
| $30-39$ | .000 | .000 | 2.498 | 6.240 | 9.756 | 16.219 | 40.484 |
| $40-49$ | .000 | .000 | .000 | 6.426 | 9.330 | 15.436 | 47.112 |
| $50+$ | .000 | .000 | .000 | .000 | 4.819 | 8.274 | 31.024 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 1655.3200 | 3022 | 1.0000 |
| Deviance | 1146.0100 | 3022 | 1.0000 |
| Freeman-Tukey | 1590.0000 | 3022 | 1.0000 |

Table C.39. Excess relative and absolute risk coefficients for nonleukemia adjusted for confirmation rates stratified on gender and city (1950-75). organ lose equivalent adjusted for DS86 random error.

Male Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | <20 | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 1.258 | . 398 | 435 | 890 | 000 | . 000 | 000 |
| 10-19 | . 817 | . 259 | . 282 | . 578 | . 584 | . 000 | 000 |
| 20-29 | . 000 | . 218 | . 238 | . 488 | . 493 | . 525 | . 000 |
| 30-39 | . 000 | . 000 | . 094 | . 193 | . 195 | . 207 | . 384 |
| 40-49 | . 000 | . 000 | . 000 | . 144 | . 146 | . 155 | . 288 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 041 | . 043 | 081 |

Male Absolute risk (deaths/10^4PYSv)
Age ATD

| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| <10 | 3.258 | 1.251 | 1.332 | 5.011 | . 000 | . 000 | 000 |
| 10-19 | 6.363 | 1.891 | 1.896 | 8.697 | 17.879 | . 000 | 000 |
| 20-29 | . 000 | 4.130 | 2.870 | 8.692 | 17.354 | 37.738 | . 000 |
| 30-39 | . 000 | . 000 | 1.424 | 3.358 | 8.639 | 19.795 | 58.249 |
| 40-49 | . 000 | . 000 | . 000 | 4.813 | 7.792 | 16.657 | 54.635 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 2.751 | 6.002 | 21.568 |
| Female | Excess | elative | Iisk ( | Sv) |  |  |  |

## Age ATD

| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| <10 | 2.663 | . 843 | . 921 | 1.885 | . 000 | . 000 | 000 |
| 10-19 | 1.729 | . 547 | . 598 | 1.224 | 1.236 | . 000 | 00 |
| 20-29 | . 000 | . 462 | . 504 | 1.033 | 1.043 | 1.111 | 00 |
| 30-39 | . 000 | . 000 | . 199 | . 408 | . 412 | . 439 | 81 |
| 40-49 | . 000 | . 000 | . 000 | . 305 | . 309 | . 329 | 60 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 086 | . 092 | . 17 |

Female Absolute risk (deaths/10^4PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 3.308 | 1.425 | 3.900 | 10.738 | .000 | .000 | .000 |
| $10-19$ | 8.187 | 2.374 | 2.869 | 10.308 | 17.737 | .000 | .000 |
| $20-29$ | .000 | 2.938 | 3.378 | 10.997 | 21.658 | 35.154 | .000 |
| $30-39$ | .000 | .000 | 2.135 | 6.106 | 9.743 | 17.981 | 56.191 |
| $40-49$ | .000 | .000 | .000 | 6.321 | 9.393 | 18.099 | 56.861 |
| $50+$ | .000 | .000 | .000 | .000 | 3.279 | 6.570 | 30.376 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 2029.3400 | 3022 | 1.0000 |
| Deviance | 1482.8200 | 3022 | 1.0000 |
| Freman-Tukey | 1755.7700 | 3022 | 1.0000 |

Table C.40. Excess relative and absolute risk coefficients for nonleukemia adjusted for confirmation rates stratified on gender and city (1950-85) orgar dose equivalent adfusted for DS86 random error.

Male Excess relative risk (\%/Sv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| -10 | 1.352 | .463 | .473 | .854 | .000 | .000 | .000 |
| $10-19$ | .754 | .258 | .263 | .476 | .467 | .000 | .000 |
| $20-29$ | .000 | .240 | .245 | .443 | .434 | .403 | .000 |
| $30-39$ | .000 | .000 | .116 | .210 | .206 | .191 | .248 |
| $40-49$ | .000 | .000 | .000 | .156 | .153 | .142 | .184 |
| $50+$ | .000 | .000 | .000 | .000 | .064 | .059 | .077 |

Male Absolute risk (deaths/10^4PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | L20 | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 3.473 | 1.447 | 2.087 | 6.917 | .000 | .000 | .000 |
| $10-19$ | 5.880 | 1.889 | 1.775 | 7.396 | 14.002 | .000 | .000 |
| $20-29$ | .000 | 4.536 | 2.950 | 7.929 | 16.126 | 28.176 | .000 |
| $30-39$ | .000 | .000 | 1.757 | 3.659 | 9.145 | 19.272 | 41.249 |
| $40-49$ | .000 | .000 | .000 | 5.190 | 8.163 | 15.225 | 46.157 |
| $50+$ | .000 | .000 | .000 | .000 | 4.291 | 8.162 | 21.972 |

Female Excess relative risk (\%/Sv)
Age ATD

| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| $<10$ | 2.669 | .915 | .933 | 1.686 | .000 | .000 | .000 |
| $10-19$ | 1.487 | .510 | .520 | .940 | .921 | .000 | .000 |
| $20-29$ | .000 | .474 | .484 | .874 | .857 | .795 | .000 |
| $30-39$ | .000 | .000 | .230 | .415 | .407 | .378 | .490 |
| $40-49$ | .000 | .000 | .000 | .308 | .302 | .280 | .363 |
| $50+$ | .000 | .000 | .000 | .000 | .126 | .117 | .152 |

Female Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | A | 20 | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 3.312 | 1.537 | 3.547 | 8.171 | .000 | .000 | .000 |  |
| $10-19$ | 7.133 | 2.223 | 2.519 | 7.629 | 11.654 | .000 | .000 |  |
| $20-29$ | .000 | 3.011 | 3.248 | 9.463 | 16.627 | 23.501 | .000 |  |
| $30-39$ | .000 | .000 | 2.454 | 6.211 | 9.634 | 16.428 | 40.607 |  |
| $40-49$ | .000 | .000 | .000 | 6.365 | 9.188 | 15.504 | 47.089 |  |
| $50+$ | .000 | .000 | .000 | .000 | 4.763 | 8.323 | 31.032 |  |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 1667.4400 | 3022 | 1.0000 |
| Deviance | 1148.8800 | 3022 | 1.0000 |
| Freeman-Tukey | 1592.1900 | 3022 | 1.0000 |

Table C.41. Excess relative and absolute risk coefficients for nonleukemia adjusted for confirmation rates stratified on age ATD (1950-75). organ dose equivalent adjusted for DS86 random error.

Male Excess relative risk (\%/Sv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $\ldots 10$ | 1.504 | 1.010 | .663 | 1.109 | .000 | .000 | .000 |
| $10-19$ | .934 | .627 | .412 | .689 | .631 | .000 | .000 |
| $20-29$ | .000 | .597 | .392 | .656 | .600 | .513 | .000 |
| $30-39$ | .000 | .000 | .143 | .240 | .219 | .188 | .372 |
| $40-49$ | .000 | .000 | .000 | .213 | .195 | .167 | .331 |
| $50+$ | .000 | .000 | .000 | .000 | .072 | .061 | .121 |

Male Absolute risk (deaths/10^4PYSv)
Age ATD

| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $<10$ | 1.380 | 1.622 | 1.642 | 6.116 | .000 | .000 | .000 |
| $10-19$ | 2.447 | 1.735 | 1.588 | 9.145 | 19.168 | .000 | .000 |
| $20-29$ | .000 | 3.823 | 2.850 | 9.385 | 20.366 | 36.864 | .000 |
| $30-39$ | .000 | .000 | 1.248 | 2.880 | 9.126 | 18.045 | 56.434 |
| $40-49$ | .000 | .000 | .000 | 5.534 | 9.502 | 17.288 | 61.912 |
| $50+$ | .000 | .000 | .000 | .000 | 3.887 | 7.569 | 26.484 |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 2.958 | 1.985 | 1.303 | 2.181 | . 000 | . 000 | 000 |
| 10-19 | 1.837 | 1.233 | . 809 | 1.355 | 1.240 | . 000 | . 000 |
| 20-29 | . 000 | 1.174 | . 770 | 1.290 | 1.181 | 1.009 | . 000 |
| 30-39 | . 000 | . 000 | . 282 | . 471 | . 432 | . 369 | . 731 |
| 40-49 | . 000 | . 000 | . 000 | . 419 | . 384 | . 328 | . 650 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 141 | . 120 | 239 |

Female Absolute risk (deaths/10^4PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 1.708 | 2.303 | 5.349 | 12.103 | .000 | .000 | .000 |
| $10-19$ | 4.215 | 2.874 | 3.544 | 11.668 | 17.769 | .000 | .000 |
| $20-29$ | .000 | 3.890 | 4.245 | 15.144 | 25.209 | 32.189 | .000 |
| $30-39$ | .000 | .000 | 2.690 | 7.612 | 11.936 | 15.722 | 50.911 |
| $40-49$ | .000 | .000 | .000 | 7.953 | 11.899 | 18.513 | 59.894 |
| $50+$ | .000 | .000 | .000 | .000 | 4.812 | 8.365 | 35.332 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 3104.6500 | 3022 | .1441 |
| Deviance | 1756.6600 | 3022 | 1.0000 |
| Freeman-Tukey | 1851.0100 | 3022 | 1.0000 |

Table C.42. Excess relative and absolute risk coefficients for nonleukemia adjusted for confirmation rates stratified on age ATD (1950-85).
organ dosa equivalent adjusted for DS86 random error.

| Male | Excess relative risk (\%/Sv) |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Age ATD |  |  |  |  |  |  |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 1.527 | 1.027 | . 657 | 1.070 | . 000 | . 000 | . 000 |
| 10-19 | . 936 | . 630 | . 403 | . 656 | . 585 | . 000 | . 000 |
| 20-29 | . 000 | . 615 | . 393 | . 640 | . 571 | . 451 | . 000 |
| 30-39 | . 000 | . 000 | . 154 | . 251 | . 224 | . 177 | . 300 |
| 40-49 | . 000 | . 000 | . 000 | . 224 | . 200 | . 158 | . 268 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 083 | . 066 | . 111 |

Male Absolute risk (deaths/10^4PYSV)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 1.397 | 1. 648 | 1.974 | 7.083 | . 000 | . 000 | . 000 |
| 10-19 | 2.451 | 1.742 | 1.556 | 9.034 | 18.960 | . 000 | . 000 |
| 20-29 | . 000 | 3.932 | 2.568 | 9.182 | 20.712 | 34.483 | . 000 |
| 30-39 | . 000 | . 000 | 1.341 | 3.013 | 9.301 | 18.288 | 51.634 |
| 40-49 | . 000 | . 000 | . 000 | 5.814 | 9.725 | 16.396 | 61.683 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 4.506 | 8.127 | 25.560 |
| Female | Excess | lative | risk | SV) |  |  |  |

Age ATD

| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $<10$ | 2.949 | 1.984 | 1.268 | 2.067 | .000 | .000 | .000 |
| $10-19$ | 1.808 | 1.217 | .778 | 1.267 | 1.129 | .000 | .000 |
| $20-29$ | .000 | 1.188 | .759 | 1.237 | 1.102 | .872 | .000 |
| $30-39$ | .000 | .000 | .298 | .485 | .432 | .342 | .579 |
| $40-49$ | .000 | .000 | .000 | .433 | .386 | .305 | .517 |
| $50+$ | .00 | .000 | .000 | .000 | .161 | .127 | .215 |

Female Absolute risk (deaths/10^4PYSv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 1.704 | 2.304 | 5.301 | 11.711 | .000 | .000 | .000 |
| $10-19$ | 4.157 | 2.841 | 3.418 | 11.214 | 17.122 | .000 | .000 |
| $20-29$ | .000 | 3.931 | 4.190 | 14.607 | 24.899 | 30.123 | .000 |
| $30-39$ | .000 | .000 | 2.840 | 7.823 | 11.954 | 16.112 | 50.008 |
| $40-49$ | .000 | .000 | .000 | 8.209 | 11.965 | 17.288 | 62.466 |
| $50+$ | .000 | .000 | .000 | .000 | 5.476 | 8.822 | 35.481 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 2936.8300 | 3022 | .8638 |
| Deviance | 1582.6900 | 3022 | 1.0000 |
| Freeman-Tukey | 1803.1600 | 3022 | 1.0000 |

Table C.43. Excess relative and absolute risk coefficients for nonleukemia adjusted for confirmation rates stratified on DS86 shielded kerma (1950-75) organ dose equivalent adjusted for DS86 random error.

| Male | Excess relative risk (\%/Sv) |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATD |  |  |  |  |  |  |  |
| Age ATB | $<20$ | 20-29 | $30=2$ | 40-49 | 50-59 | 60-69 | $70+$ |
| $<10$ | 2.071 | 1.035 | . 750 | 1.119 |  | . 000 | . 000 |
| $10-19$ | 1.330 | . .664 | . 481 | . 718 | . 674 | . 000 | . 000 |
| 20-29 | . 000 | . 656 | . 475 | . 709 | . 665 | . 522 | . 000 |
| 30-39 | . 000 | . 000 | . 211 | . 314 | . 295 | . 232 | . 373 |
| 40-49 | . 000 | . 000 | . 000 | . 314 | . 295 | . 231 | . 372 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 194 | . 152 | . 245 |

Male Absolute risk (deaths/10^4PYSV)

|  |  |  |  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | :---: | :---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |  |  |  |
| $<10$ | 2.523 | 1.959 | 1.951 | 6.155 | .000 | .000 | .000 |  |  |  |
| $10-19$ | 4.767 | 2.485 | 2.156 | 9.896 | 20.352 | .000 | .000 |  |  |  |
| $20-29$ | .000 | 5.888 | 3.646 | 10.851 | 22.547 | 37.419 | .000 |  |  |  |
| $30-39$ | .000 | .000 | 2.148 | 4.201 | 12.313 | 22.229 | 56.533 |  |  |  |
| $40-49$ | .000 | .000 | .000 | 8.742 | 14.396 | 23.924 | 69.122 |  |  |  |
| $50+$ | .000 | .000 | .000 | .000 | 10.597 | 18.615 | 51.620 |  |  |  |

Female Excess relative risk (\%/Sv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 3.815 | 1.905 | 1.381 | 2.060 | . 000 | . 000 | 000 |
| 10-19 | 2.449 | 1.224 | . 887 | 1.322 | 1.240 | . 000 | 000 |
| 20-29 | . 000 | 1.208 | . 875 | 1.305 | 1.224 | . 961 | 000 |
| 30-39 | . 000 | . 000 | . 388 | . 579 | . 543 | 426 | . 686 |
| 40-49 | . 000 | . 000 | . 000 | 579 | . 543 | 426 | 686 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 357 | . 280 | 451 |

Female Absolute risk (deaths/10^4PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| 10 | 2.714 | 2.502 | 5.715 | 11.534 | .000 | .000 | .000 |
| $10-19$ | 6.894 | 3.470 | 4.153 | 11.646 | 17.775 | .000 | .000 |
| $20-29$ | .000 | 4.726 | 5.235 | 15.816 | 26.025 | 30.790 | .000 |
| $30-39$ | .000 | .000 | 3.962 | 9.566 | 14.824 | 18.092 | 47.948 |
| $40-49$ | .000 | .000 | .000 | 11.529 | 16.653 | 23.810 | 62.652 |
| $50+$ | .000 | .000 | .000 | .000 | 12.195 | 19.125 | 63.540 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 2787.7400 | 3022 | .9990 |
| Deviance | 1756.8400 | 3022 | 1.0000 |
| Freeman-Tukey | 1866.0900 | 3022 | 1.0000 |

Table c.44. Excess relative and absolute risk coefficients for nonleukemia adjusted for confirmation rates stratified on DS86 shielded kerma (1950-85) organ dose equivalent adjusted for DS86 random error.

Male Excess relative risk (\%/SV)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 50-69 | $70+$ |
| $<10$ | 2.124 | 1.083 | . 785 | 1.139 | . 000 | 000 | 000 |
| 10-19 | 1.303 | . 664 | . 481 | . 699 | . 655 | . 000 | . 000 |
| 20-29 | . 000 | . 670 | . 485 | . 704 | . 661 | . 523 | . 000 |
| 30-39 | . 000 | . 000 | . 227 | . 329 | . 309 | . 245 | 382 |
| 40-49 | . 000 | . 000 | . 000 | . 321 | . 301 | . 239 | 373 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | . 199 | 158 | 246 |

Male Absolute risk (deaths/10^4PYSv)

| Age ATD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age ATB | $<20$ | 20-29 | 30-39 | 40-49 | 50-59 | 60-69 | $70+$ |
| <10 | 2.575 | 2.041 | 2.599 | 8.067 | . 000 | . 000 | 000 |
| 10-19 | 4.673 | 2.484 | 2.158 | 10.159 | 21.216 | . 000 | . 000 |
| 20-29 | . 000 | 6.003 | 3.722 | 10.790 | 23.950 | 39.927 | . 000 |
| 30-39 | . 000 | . 000 | 2.308 | 4.393 | 12.872 | 25.199 | 65.136 |
| 40-49 | . 000 | . 000 | . 000 | 8.938 | 14.716 | 24.687 | 83.759 |
| $50+$ | . 000 | . 000 | . 000 | . 000 | 10.867 | 19.270 | 54.352 |

Female Excess relative risk (\%/Sv)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age ATB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $50-69$ | $70+$ |
| 10 | 3.816 | 1.945 | 1.411 | 2.047 | .000 | .000 | .000 |
| $10-19$ | 2.341 | 1.193 | .865 | 1.255 | 1.277 | .000 | .000 |
| $20-29$ | .000 | 1.203 | .872 | 1.266 | 1.187 | .941 | .000 |
| $30-39$ | .000 | .000 | .408 | .592 | .555 | .440 | .687 |
| $40-49$ | .000 | .000 | .000 | .578 | .542 | .429 | .670 |
| $50+$ | .000 | .000 | .000 | .000 | .357 | .283 | .442 |

Female Absolute risk (deaths $/ 10^{\wedge} 4$ PYSV)

|  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Age AmB | $<20$ | $20-29$ | $30-39$ | $40-49$ | $50-59$ | $60-69$ | $70+$ |
| $<10$ | 2.714 | 2.546 | 5.991 | 11.911 | .000 | .000 | .000 |
| $10-19$ | 6.631 | 3.397 | 4.064 | 11.537 | 17.692 | .000 | .000 |
| $20-29$ | .000 | 4.711 | 5.224 | 15.403 | 26.489 | 32.489 | .000 |
| $30-39$ | .000 | .000 | 4.155 | 9.763 | 15.126 | 20.532 | 58.177 |
| $40-49$ | .000 | .000 | .000 | 11.514 | 16.626 | 23.991 | 78.208 |
| $50+$ | .000 | .000 | .000 | .000 | 12.210 | 19.325 | 68.991 |


| Goodness of fit | Value | d.f. | Prob |
| :--- | :--- | :--- | :--- | :--- |
| Chi-square | 2610.1000 | 3022 | 1.0000 |
| Deviance | 1582.5200 | 3022 | 1.0000 |
| Freeman-Tukey | 1815.2400 | 3022 | 1.0000 |

Table C.45. BEIR-V excess relative risk $(\% / \mathrm{Sv})$ of leukemia for various stratifications of confirmation rates.

|  |  | $E<=20$ | $E>20$ |  |
| :---: | :---: | :---: | :---: | :---: |
| Adjustment | $\mathrm{T}<15$ | $115<\pi<25$ | $T<=25$ | $25<T<=30$ |
| Mo adyustment | 5. 14 | 0.43 | 0.42 | 0.20 |
| Crude (1950-75) | 4. 56 | 0.41 | 0.13 | 0.18 |
| Crude (1950-85) | 4.25 | 0.39 | 0.33 | 0.17 |
| $\operatorname{Sex}(1950-75)$ | 4.84 | 0.43 | 0.39 | 0.19 |
| $\operatorname{sex}(1950-85)$ | 4.46 | 0.40 | 0.34 | 0.17 |
| Sex and city (1950-075) | 4. 43 | 0.39 | 0.33 | 0.17 |
| Sex and city (1950-075) | 3.82 | 0.34 | 0.29 | 0.15 |
| Age ATD (1950-75) | 3,37 | 0.37 | 0.31 | 0.16 |
| Age ATD (1950-85) | 3.06 | 0.34 | 0.29 | 0.15 |
| DS86(1950-75) | 2.91 | 0.30 | 0.33 | 0.18 |
| DS86(1950-85) | 5.19 | 0.54 | 0.58 | 0.33 |

E denotes age at exposure (age ATB)
T denotes time since exposure

Table C.46. BEIR-V excess relative risk ( $8 /$ SV) of leukemia for various stratifications of confirmation rates.
DS86 dose equivalents adjusted for random error
$E<=20 \quad E>20$

| Adjustment | $T<15$ | $15<T<=25$ | $T<=25$ | $25<T<=30$ |
| :--- | ---: | :---: | :---: | :---: | :---: |
| No adjustment | 10.56 | 0.90 | 0.80 | 0.42 |
| Crude(1950-75) | 9.36 | 0.85 | 0.67 | 0.38 |
| Crude(1950-85) | 8.62 | 0.80 | 0.62 | 0.35 |
| Sex (1950-75) | 9.94 | 0.89 | 0.75 | 0.40 |
| Sex(1950-85) | 8.99 | 0.81 | 0.64 | 0.36 |
| Sex and city (1950-075) | 9.12 | 0.80 | 0.63 | 0.35 |
| Sex and city (1950-075) | 8.66 | 0.77 | 0.60 | 0.33 |
| Age ATD(1950-75) | 6.88 | 0.75 | 0.60 | 0.33 |
| Age ATD(1950-85) | 6.50 | 0.72 | 0.57 | 0.32 |
| DS86(1950-75) | 6.36 | 0.67 | 0.69 | 0.42 |
| DS86(1950-85) | 8.85 | 0.95 | 0.96 | 0.58 |

E denotes age at exposure (age ATB)
T denotes time since exposure

Table C.47. BEIR $-V$ excess relative risk ( $8 / \mathrm{SV}$ ) of digestive cancer for various stratifications of confirmation races.

|  | MALES |  |  | FEMALES |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Adjustment | ER 2 25 | $25<\mathrm{E}<=35$ | E>35 | $E<=25$ | $25<\mathrm{E}<=35$ | E>35 |
| No adjustment | 0.81 | 0.30 | 0.11 | 1.41 | 0.52 | 0.19 |
| Crude (1950-75) | 0.62 | 0.26 | 0.11 | 1.11 | 0.48 | 0.20 |
| Crude (1950-85) | 0.48 | 0.22 | 0.10 | 0.96 | 0.44 | 0.20 |
| $\operatorname{Sex}(1950-75)$ | 0.60 | 0.26 | 0.11 | 1.13 | 0.48 | 0.21 |
| $\operatorname{sex}(1950-85)$ | 0.46 | 0.21 | 0.10 | 0.99 | 0.45 | 0.20 |
| Sex and city (1950-075) | 0.73 | 0.29 | 0.11 | 1. 33 | 0.52 | 0.21 |
| Sex and city (1950-075) | 0.66 | 0.26 | 0.11 | 1.27 | 0.51 | 0.20 |

E denotes age at exposure (age ATB)
T denotes time since exposure

Table C.48. BEIR-V excess relative $r i s k\left(\frac{8}{6} / \mathrm{SV}\right)$ of digescive cancer for various stratifications of confirmation rates bsह̂6 dose equivalent adjusted for random error.

|  | MALES |  |  | FEMALES |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Adjustment | E<<25 | $25<E<=35$ | E>35 | $\mathrm{E}<=25$ | $25<E<=35$ | E>35 |
| No adjustment | 0.74 | 0.30 | 0.13 | 1.44 | 0.60 | 0.25 |
| Crude (1950-75) | 0.59 | 0.28 | 0.14 | 1.13 | 0.54 | 0.26 |
| crude (1950-85) | 0.47 | 0.24 | 0.12 | 0.98 | 0.49 | 0.25 |
| Sex (1950-75) | 0.58 | 0.28 | 0.13 | 1.15 | 0.55 | 0.26 |
| Sex(1950-85) | 0.46 | 0.23 | 0.11 | 1.01 | 0.51 | 0.25 |
| Sex and city(1950-075) | 0.69 | 0.30 | 0.13 | 1.36 | 0.60 | 0.26 |
| Sex and city(1950-075) | 0.62 | 0.27 | 0.12 | 1.30 | 0.58 | 0.26 |

E denotes age at exposure (age ATB)
T denotes time since exposure

## 13 APPENDIX D. Lifetime Mortality Risks (\%/Sv).

Tables D. 1 - D. 56 .

Table D.1. Excess mortality risk ( $8 / \mathrm{Sv}$ ) of leukemia among males exposed to 0.02 Sv/y from ages 18 to 65 for the absolute projection model.


Table D.2. Excess mortality risk (\%/Sv) of leukemia among males exposed to $0.02 \mathrm{~Sv} / \mathrm{Y}$ from ages 18 to 65 for the transported relative projection model.

```
                Run date: 10/ 3/1903
                    Title: Lifetime risks
                        Sex: MALE
                        Race: "*'TTE
        Life table used: to 
    Risk coefficients
                Leukemia:
            Solid cancers
                    DRREF
        Minimal latency(yra): 2 Plateau(yrs): 40
        Minimal latency(yrs): 10 Plateau(yrs):101
        2.0
        18
    Age at last expos.: 65
    Total dose eq. (Sv): 1.000000
```



Table D.3. Excess mortality risk (\%/\$v) of leukemia among males exposed to 0.02 sv/y from ages 18 to 65 for the relative projection model.

```
            Run date: 10/ 3/1993
                                    Title: Lifetime risks
                                    Sex: MALE
                                    Race: WHITE
        Llfe table used: }199
    Risk coefficients: BASED ON AGE ATB- AND ATD-SPECIFIC (RBE=10) COEFFICIENTS
                Leukemia: Minimal latency(yrs): 2 Plateau(yrs): 40
            Solid cancers: Minimal latency(yrs): 10 Flateau(yrs): 101
                DRREF: 2.0
Age at first expos.: }1
    Age at last expos.: 65
Total dose eq. (Sv): 1.000000
```

| Strata for confirmation rates used in adjustment | Radiation induced cancer deaths per 10**5 | 90.08 | CI | Baseline cancer deaths per 10**5 | $\begin{gathered} \text { Years } \\ \text { of } \\ \text { life } \\ \text { lost } \end{gathered}$ | PCl | $0.0 \% \mathrm{CI})$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| None | 875 | 356. | 21521 | 750 | 34 | 541 | .19,1.00) |
| Crude (1950-75) | 291 | 118, | 7161 | 750 | 34 | 281 | .11, . 721 |
| Crude (1950-85) | 640 | 260, | 1574) | 750 | 36 | . 46 | .17,1.00) |
| Sex(1950-75) | 758 | 308. | 1863) | 750 | 35 | . 501 | . $18,1.001$ |
| Sex (1950-85) | 755 | 307. | $1856)$ | 750 | 36 | . 501 | .18,1.00) |
| Sex-city (1950-85) | 715 | 291. | 1758) | 750 | 35 | . 491 | .27,1.00) |
| Sex-city (1950-85) | 752 | 306 , | 1848) | 750 | 35 | . 501 | .18,1.00) |
| Age ATD (1950-75) | 624 ( | 254. | 1535) | 750 | 32 | . 451 | .16,1.00) |
| Age ATD (1950-85) | 623 1 | 253, | 1532) | 750 | 32 | . 45 i | .16,2.00) |
| DS86(1950-75) | 1010 ( | 410. | 2482) | 750 | 30 | . 571 | .19,1.00) |
| DS86 (1950-85) | 1085 | 441. | 2668) | 750 | 28 | . 591 | .20,1.001 |

Table D.4. Excess mortality risk ( $8 /$ SV) of leukemia among males exposed co 0.02 Sv/y from ages 18 to 65 for the transported relative projection model. Baseline rates adfusted with death certificate correction factors (DCCFs).

Run date: $10 / 3 / 1993$
Title: Lifetime Iisks
Sex: MALE
Race: WHITE
Life table used: 1990
Risk coefficients: BASED ON AGE ATB- AND ATD-SPECIFIC (RBE=10) COEFFICIENTS Leukemia: Minimal latency (yrs): 2 Plateau(yrs): 40
Solid cancers: Minimal latency (yrs): 10 Plateau(yrs): 101
DRREF: 2.0
Age at first expos.: 18
Age at last expos.: 65
Total dose eq. (Sv): $\quad 1.000000$


Table D.5. Excess mortality risk ( $/ / \mathrm{sv}$ ) of leukemia among males exposed to 0.02 sv/y from ages 18 to 65 for the relative projection model. Easeline rates adfusted with death certificate correction factors (DCCFs).


Table D.6. Excess mortality risk ( $2 /$ Sv) of leukemia among fumales exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the absolute projection model.

```
                            Run date: 10/ 3/1993
                        Title: Lifetime risks
                                Sex: FEMALE
                        Race: WHITE
        Life table used: }199
        Kisk coefficients: BASED ON AGE ATB-AND ATD-SPECIFIC (REE=10) COEFFICIENKS
            Leukemia: Minimal latency(yrs); 2 Plateau(yrs): 40
            Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
            DRREF: 2.0
Age at first expos.: 18
    Age at last expos.t 65
Total dose eq. (Sv): 1.000000
```



Table D.7. Excess mortality risk (8/Sv) of leukemia among females exposed to $0.02 \mathrm{SV} / \mathrm{y}$ from ages 18 to 65 for the transported relative projection model.

$$
\text { Run date: } 10 / 3 / 1993
$$

Title: Lifetime riskg
Sex: FEMALE
Race: WHITE
Life table used: 1990
Risk coefficient E: EASED ON AGE ATB-AND ATD-SPECIFIC (REE=10) COEFFICIENPS Leukemia: Minimal latency (yrs): 2 Plateau(yrs): 40
Solid cancers: Minimal latencyfyrs): 10 Plateau(yrs): 101
DRREF; 2.0
Age at first expos.; 18
Age at last expos.: 65
Total dose eq. (SV): 1.000000


Table D. B. Excess mortality risk ( $8 / 5 v$ ) of leukemia among females exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the relative projection model.

Run date: 10/3/1993
Title: Lifetime risks
Sex: FEMALE
Race: WHITE
Life table used: 1990
Risk coeffi nts: BASED ON AGE ATB-AND ATD-SPECIFIC (RBE=10) COEFFICIENTS
Solid cancers: Minima, latency(yrs): 2 Plateau(yrs): 40
DRREF: 2.0
Age at first expos. $\quad 18$
Age at last expos.: 65
Total dose eq. (Sv): 1.000000


Table D.9. Excess mortality risk ( $/ \mathrm{/Sv}$ ) of leukemia among females exposed to 0.02 sv/y from ages 18 to 65 for the transported relative projection model. Baseline rates adfusted with death certificate correction factors (DCCFs).

Run date: $10 / 3 / 1993$
Title: Lifetime risks
Sox: FEMALE
Race: WHITE
Life table used: 1990
Risk coefficiente: BASED ON AGE ATB- AND ATD-SPECIFIC (RBE=10) COEFFICIENTS
Leukemia: M1nımal latency(yrs): 2 Plateauiyrs): 40
Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
DRREF: 2.0
Age at first expos.: 18
Age at last expos.: 65
Total dose eq. (Sv): $\quad 1.000000$

| Strata for confirmation rates used in adjustment | $\begin{aligned} & \text { Radiation- } \\ & \text { Induced } \\ & \text { cancer } \\ & \text { deachs } \\ & \text { per } 10 * * 5 \end{aligned}$ | 90.08 | CI | Baseline cancer deaths per $10^{* * 5}$ | $\begin{gathered} \text { Years } \\ \text { of } \\ \text { life } \\ \text { lost } \end{gathered}$ | PC 19 | (90.08 C:) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| None | 267 | 113, | 627) | 581 | 26 | 311 | .12, .81) |
| Crude (1950-75) | 2281 | 97. | 535) | 581 | 28 | . 281 | . $11, .711$ |
| Crude (1950-85) | 242 1 | 103. | 568) | 581 | 26 | . 291 | .12, .75) |
| Sex (1950-75) | 243 | 103. | 572) | 581 | 26 | . 301 | .12, .75) |
| Sex (1950-85) | 256 | 109, | 600) | 581 | 26 | . 311 | .12, . 781 |
| Sex-city (1950-85) | 222 | 94. | 5221 | 581 | 26 | . 281 | .11, . 701 |
| Sex-city (1950-85) | 225 | 95. | $528)$ | 581 | 26 | . 28 | .11, . 701 |
| Age ATD (1950-75) | 246 | 105. | 578) | 581 | 27 | . 301 | .12,.761 |
| Age ATD (1950-85) | 247 | 105. | 580) | 581 | 26 | .301 | .12, .76) |
| DS86(1950-75) | 352 | 150. | 827) | 581 | 26 | .381 | .14.1.00) |
| DS86(1950-85) | 3961 | 168 , | 929) | 581 | 26 | . 411 | . 15.1 .001 |

Table D. 10. Excess mortalicy $r i s k$ ( $8 / S v$ ) of leukemia among females exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 10 to 65 for the relative projection model
Baseline rates adjusted with death ce:tificate correction factors (DCCFs)


Table D. 11 . Excess mortality risk ( $\mathrm{s} / \mathrm{Sv}$ ) of leukemia among males exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the absolute projection model. DS86 Dose eguivalenta adjueted for random error

```
                Run date: 9/30/1993
                    Title: Lifetime risks
                Sex: MALE
                    Race: WHITE
        Life cable used: }299
        Risk coefficients: BASED ON AGE ATB- AND ATD-SPECIFIC (RBE=10) COEFFICIENTS
            Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
            DRREF: 2.0
Age at first expos.: }1
    Age at last expos.: 65
Total dose eq. (Sv): 1.000000
```

| Strata for confirmation rates used in adjustment | Radiation <br> induced <br> cancer <br> deaths <br> per 10**5 | $90.0 \%$ | CI | Baseline cancer deaths per 10**5 | $\begin{gathered} \text { Years } \\ \text { of } \\ \text { life } \\ \text { logt } \end{gathered}$ | PC(90.0\% CI) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| None | 759 | 309, | 1867) | 750 | 50 | 50 |
| Crude (1950-75) | 6701 | 272 , | 1648) | 750 | 48 | . $471.27,1.00)$ |
| Crude (1950-85) | 651 | 264 , | 1600) | 750 | 50 | . $46(.17,1.00)$ |
| Sex (1950-75) | 661 ( | 269 , | 1626) | 750 | 49 | . $47(.17 .1 .00)$ |
| $\operatorname{Sex}(1950-85)$ | 632 1 | 257. | 1554) | 750 | 51 | . 461.17 .1 .00$)$ |
| Sex-city (1950-85) | 6551 | 266. | 1610) | 750 | 48 | $.47(.17,1.00)$ |
| Sex-city (1950-85) | 6291 | 256. | 1546) | 750 | 50 | $.461 .16,1.00)$ |
| Age ATD (1950-75) | 6761 | 275. | 1662) | 750 | 47 | $.47(.17,1.00)$ |
| Age ATD (1950-85) | 658 | 268 , | 1619) | 750 | 47 | $.47(.17,1.00)$ |
| DS86 (1950-75) | 810 | 329. | 1993) | 750 | 43 | $.52(.18,1.001$ |
| DS86(1950-85) | 8531 | 347. | 2097) | 750 | 41. | $.53(.18,1.00)$ |

Table D.12. Excess mortality risk (\%/Sv) of leukemia among males exposed to 0.02 Sv/y from ages 18 to 65 for the transported relative projection model. DS86 Dose equivalents adyuated for random error

Run date: $9 / 30 / 1993$
Title: Lifetime risks
Sex: MALE
Race: WHITE
Life table used: 1990
Risk coefficients : BASED ON AGE ATB - AND ATD-SPECIFIC (RBE= 10) COEFFICIENTS
Solid cancers
DRREF: 2
Age at first expos.: 18
Age at last expos.
Total dose eq. (Sv):
1.000000


Table D.13. Excess mortality risk (\%/Sv) of leukemia among males exposed to $0.02 \mathrm{sv} / \mathrm{y}$ from ages 18 to 65 for the relative projection model.
DS86 Dose equivalents adfusted for random error

```
Fun date: 9/30/1993
    Title: LIfetime Ilsks
    Sex: MALE
```

    Race: WHITE
    Life table used: 199
    Risk coefficients: BASED ON AGE ATB-AND ATD-SPECIFTC (REE=10) COEFFICIENTS
                Leukemia: Minimal latency(yrs): 2 Plateau(yrs): 4
    Solid cancers: Minimal latency(yrs); 10 Plateau(yrs): 101
        DRREF: 2.0
    Age at first expos.: 18
Age at last expos.; 65
Total dose eq. (Sv): 1.000000


Table D.14. Excess mortality risk ( $8 / 5 v$ ) of leukemia among males exposed co $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the transported relative projection model Baseline rates adjusted with death certificate correction factors (DCCFs). os86 Dose equivalents adjusted for random error

Run date: $9 / 30 / 1993$
Titie: Lifetime risks
Sex: MALE
Race: WHTTE

## Lile table used: 1990

Risk coefficients: BASED ON AGE ATB-AND ATD-SPECIFIC (RBE=10) COEFFICIENTS Minimal latency (yrs) : 2 Plateau(yrs): 40 Solid cancers: Minimal latency(yrs):

DRREF: 2.0
Age at first expos.
Age at last expos.
Total dose eq. (Sv)
1.001000


Table D.15. Excess mortality risk (\%/Sv) of leukemia among males exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the relative projection model
Baseline rates adfusted with death certificate correction factors (DCCFs)
DS86 Dose equivalents adjusted for random error



Table D.16. Excess mortality risk (8/Sv) of leukemia among females exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the absolute projection model. DS86 Dose equivalents adjusted for random error.

Fun date: $9 / 30 / 1993$
Title: Lifetime risks
Sex: FEMALE
Race: WHITE
Life table used: 1990
Risk coefficients: BASED ON AGE ATB-AND ATD-SPECIFIC (RBE $=10$ ) COEFFICIENTS
Leukemia: Minimal latency (yrs): 2 Plateau(yrs): 40
Solid cancers: Minimal latency (yrs): 10 Plateau(yrs): 101
DRREF: 2.0
Age at first expos.t 18
Age at last expos.; 65
Total dose eq. (SV): 1.000000


Table D.17. Excess mortality risk (\%/Sv) of leukemia among females exposed to $0.02 \mathrm{sv} / \mathrm{y}$ from ages 18 to 65 for the transported projection model Ds86 Dose equivalents adfusted for random error

Run date: $9 / 30 / 1993$
Th.tle: Lifetime risks
Sex: FEMALE
Race: WHITE
Life table used: 1990
Risk coefficients: BASED ON AGE ATB- AND ATD-SPECIFIC (RBE=10) COEFFICIENTS Leukemia: Minimal lacency(yrs): 2 Plateau(Vrs): 40
Solid cancers. Minimal latency(yrs): 10 Plateau(yrs): 101
DRREF: 2.0
Age at first expos.. 18
Age at last expos.: 65
Total dose eq. (Sv): 1.000000

| Strata for confirmation rates used in adjustment | Radiation induced cancar deaths per 10**5 | 90.08 | CI | Baseline cancer deaths per 10**5 | $\begin{gathered} \text { Years } \\ \text { of } \\ \text { life } \\ \text { lost } \end{gathered}$ | PC | 90.08 CI) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| None | 291 | 124. | 684) | 653 | 26 | 311 | .12, .78) |
| Crude (1950-75) | 255 | 108. | 599) | 653 | 27 | . 281 | .11, . 701 |
| Crude (1950-85) | 253 | 107. | 5941 | 653 | 26 | . 281 | .11, 701 |
| $\operatorname{Sex}(1950-75)$ | 267 | 113. | 626) | 653 | 26 | . 291 | .12, .731 |
| Sex (1950-85) | 2801 | 119. | 658) | 653 | 26 | .301 | .12, . 761 |
| Sex-city (1950-85) | 2431 | 103. | 572) | 653 | 26 | . 271 | .11, .68) |
| Sex-city (1950-85) | 246 1 | 104, | 577) | 653 | 26 | . 271 | . $11, .681$ |
| Age ATD $1950-75$ ) | 270 | 115. | 635) | 653 | 27 | . 291 | .12, .74) |
| Age ATD (1950-85) | 271 | 115 , | 6371 | 653 | 26 | . 291 | .12, .74) |
| DS86 (1950-75) | 386 | 164 , | 9061 | 653 | 26 | . 371 | .14, . 98 ) |
| DS86 (1950-85) | 4331 | 184. | 2018) | 653 | 26 | . 401 | .15,1.00) |

Table D.18. Excess mortality risk (s/SV) of leukemia among females exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the relative proyection model D586 Dose equivalents adfusted for random error.

| Run date: | $9 / 30 / 1993$ |
| ---: | :--- |
| Title: | Lifetime risks |
| Sex: | PFMALE |



Table D.19. Excess mortality risk ( $\% /$ Sv) of leukemia among females exposed to 0.02 Sv/y from ages 18 to 65 for the trangported relative projection model. Baseline rates adfusted with desth certiflcate correction factors (DCCFs) DS86 Dose equivalents adjusted for random error

| Run date: | $9 / 30 / 1993$ |
| ---: | :--- |
| Titie: | Lifetime risks |
| Sex: | FEMALE |
| Race: | HHTTE |

e used:
Risk coefficients: BASED ON AGE ATB-AND ATD-SPECTFIC (RBE=10) COEFFICIENTS
Leukemia: Minimal latency (yrs): 2 Sateau(yrs): 40
Solid cancers: Minimal latency (yrs): 10 Plateau(yrs): 101
DRREF: 2.0
Age at first expos.: 18
Age at last expos.: 65
Total dose eq. (Sv): 1.000000


Table D.20. Excess mortality risk (\%/Sv) of leukemia among females exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the relative projection model.
Baseline rates adyusted with death certiticate correction factors (DCCFs)
DS86 Dose equivalents adjusted for random error


Table D.21. Excess mortality risk ( $8 / \mathrm{SV}$ ) of nonleukemia among males exposed to 0.02 SV/y from ages 18 to 65 for the absolute projection model

| Rur date: TJTLE: <br> Sex: <br> Race: <br> Life table used: <br> Risk coefficients: <br> Leukemia: <br> Solid cancers: <br> DRREF: <br> Age at first expos. <br> ige at last expos. <br> Total dose eq. (SV): | $\begin{aligned} & 10 / 1 / 1993 \\ & \text { Lifetime } \\ & \text { MALE } \\ & \text { WHITE } \\ & 1990 \\ & \text { BASED ON } \\ & \text { Minimal } \\ & \text { Minimal } \\ & 2.0 \\ & 18 \\ & 65 \\ & \quad 1.0000 \end{aligned}$ | risks <br> AGE ATB- AND atency (yrs) : atency(yrs) : | $\begin{gathered} \text { ATD-SPECI } \\ 2 \\ 10 \end{gathered}$ | C (RBE=10) <br> ateau(yrs) <br> ateau(yrs) | $\begin{gathered} \text { COEFFI } \\ 40 \\ 101 \end{gathered}$ | ENTS |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Strata for confirmation rates used in adjustment | Radiationinduced cancer deaths per 10**5 | $90.08 \mathrm{CI}$ | Baseline cancer deaths per 10**5 | $\begin{gathered} \text { Years } \\ \text { of } \\ \text { life } \\ \text { lost } \end{gathered}$ | PC190.08 | CI) |
| None <br> Crude (1950-75) <br> Crude (1950-85) <br> Sex(1950-75) <br> Sex (1950-85) <br> Sex-city(1950-85) <br> Sex-city(1950-85) <br> Age ATD (1950-75) <br> Age ATD (1950-85) <br> DS86(1950-75) <br> DS86(1950-85) |  | 786, $4753)$ <br> 743, $4486)$ <br> 639, $3862)$ <br> 754, $4559)$ <br> 666, $4025)$ <br> 754, $4555)$ <br> 663, $4010)$ <br> 794, $4799)$ <br> 777, $4694)$ <br> 966, $5837)$ <br> 1066, $5441)$ | 24536 24536 24536 24536 24536 24536 24536 24536 24536 24536 24536 | 26 26 27 27 27 27 27 26 27 27 26 | .071 .03 <br> .071 .03 <br> .061 .02 <br> .07 .03 <br> .061 .03 <br> .071 .03 <br> .061 .03 <br> .071 .03 <br> .071 .03 <br> .101 .04 <br> .101  | $\begin{aligned} & .18) \\ & .17) \\ & .15) \\ & .17) \\ & .15) \\ & .17) \\ & .15) \\ & .18) \\ & .18 \\ & .22( \\ & .24) \end{aligned}$ |

Table D.22. Excess mortality risk (8/8v) of nonleukemia among males exposed to 0.02 Sv/y from ages 18 to 65 for the transported relative projection model.

Run date: $10 / 1 / 1993$
TITLE: Lifetime risks Sex: MALE
Rece: NHITE
Risk coefficients: BASED ON AGE ATB- AND ATD-SPECIFIC (RBE=10) COEFFICIENTS
$\begin{aligned} \text { Leukemia: Minimal latency (yrs): } & 2 & \text { Plateau(yrs): } 40 \\ \text { Solid cancers: Minimal lacency (yrs): } & 10 & \text { Plateau(yrs): } 101\end{aligned}$
DRREF: 2.0
Age at first expos.: 18
Age at last expos.: 65
Total dose eq. (sv): 1.000000


Table D.23. Excess mortality risk $\{\% / \mathrm{Sv}$. of nonleukemia among males exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the relative projection model

Run date: 10/1/1993
TITLE: Lafetime rigks
Sex: MALE
Race: WHITE
Life table used: 1990
Risk coefficients: BASED ON AGE ATB - AND ATD-SPECIFIC (RBE=10) COEFFICIENTS Leukemia: Minima! latency (yrs) : 2 Plateau(yrs): 40
Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
Age at flret DRREF: 2.0
Age at last expos.: 65
Total dose eq. (SV): 1.000000


Table D.24. Excess mortality risk ( $8 / \mathrm{Sv}$ ) of nonleukemia among males exposed to $0.02 \mathrm{sv} / \mathrm{y}$ from ages 18 to 65 for the transported relative projection model Baseline rates adjusted with death certificate correction factors (DCCFs)

```
                Run date: 10/1/1993
                    TITLE: Lifetime risks
                    Sex: MALE
                    Race: WHITE
    Life table used: }199
    Risk coefficients: BASED ON AGE ATE- AND ATD-SPECIFIC (RBE=10) COEFFICIENTS
                                    Leukemia.
    Solid cancers: M
                    DRREF: 2.0
Age at first expos.: 18
Age at last expos.: 65
Total dose eq. (SV): 1.000000
```



Table D.25. Fxcess mortality risk (\$/Sv) of nonleukemia among males exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the relative projection model.
Baseline rates adfusted with death certificate correction factors (DCCPs).
Run date: $10 / 1 / 1993$
TITLE: LIfetime risks
Sex: MALE
Race: WHITE
Life table used: 1990
Risk coefficients: BASED ON ACE ATB- AND ATD-SPECIFIC (RBE=10) COEFFICIENTS Leukemia: Minimal latency(yrs): 2 Plateau(yrs): 40 Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101

DRREF: 2.0
age at first expos.: 18
Age at last expos. : 65
Total dose eq. (Sv): 1.000000

| Strata for confirmation rates used in adjustment | Radiationinduced cancer deaths per $10^{* * 5}$ | 90.08 | CI | Baseline cancer deaths per 10**5 | $\begin{gathered} \text { Years } \\ \text { of } \\ \text { life } \\ \text { lost } \end{gathered}$ | PC1 | 0.08 CI) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| None | 1953 | 794. | 4801) | 25066 | 24 | . 07 | 03, .18) |
| Crude (1950-75) | 1677 | 682 , | 4123) | 25066 | 24 | . 06 | .03, .15) |
| Crude (1950-85) | 1317 | 536. | 32371 | 25066 | 25 | . 051 | .02, .12) |
| Sex(1950-75) | 1681 | 684. | 4132) | 25066 | 24 | . 061 | .03, .15) |
| $\operatorname{sex}(1950-85)$ | 1322 | 538. | 3250) | 25066 | 25 | . 051 | .02, .121 |
| Sex-city (1950-85) | 1681 | 684, | 41331 | 25066 | 24 | . 061 | .03, .15) |
| Sex-city ( $1950-85$ ) | 1320 | 537. | 3245) | 25066 | 25 | . 051 | .02, .12) |
| Age ATD (1950-75) | 1808 | 736, | 4445) | 25066 | 24 | . 071 | .03, .17) |
| Age ATD (1950-85) | 1594 | 648, | 39191 | 25066 | 25 | . 061 | .02, .151 |
| DS86(1950-75) | 2138 | 869. | 5255) | 25066 | 24 | . 081 | .03, .19) |
| DS86(1950-85) | 21731 | 884. | 5342) | 25066 | 24 | . 081 | 03, .201 |

Table D.26. Excess mortality risk ( $8 / \mathrm{SV}$ ) of nonleukemia among females exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the absolute projection model.

```
    Run date: 10/ 1/1993
                            TITLE: Lifetime rikks
                    Sex: FEMALE
                        Race: WHITE
        Life table used, 1990
        Risk coefficients: BASED ON AGE ATB- AND ATD-SPECIFIC (RBE=10) COEFFICIENTS
            Leukemia: Minimal latency(yrs): 2 Plateau(yrs): 40
            Solid cancers: Minimal latency(yrs): 10 Plateau(yrs); 101
            DRREF: }2.
Age at f1rst expos.; 18
    Age at last expos.: 65
Total dose eq. (Sv): 1.000000
```



Table D. 27 Excess mottality risk ( $8 / 8 v$ ) of nonleukemia among females exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the tranaported relative projection model

Run date: $10 / 1 / 1993$
TITIE: Lifetime risks
Sex: FEMALE
Face: WHITE
Life table used: 1990
RIsk coefficients: BASED ON AGE ATB-AND ATH-SPECIEIC (RBE=10) COEFFICIRNTS Leukemia: Minimal latency (yrs): 2 Plateau(yrs): 40
Solid cancers: Minimal Latency (yrs): 10 Plateau(yrs): 101
DRREF: 2.0

Age at first expos.: $\frac{18}{65}$
Age at last expos
Total dose eq. (Sv): 1.000000


Table D.28. Excess mortality risk (\%/Sv) of nonleukemia among females exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the relative plojection model

```
Run date: 10/ 1/1993
    TITLE: Lifetime risks
        Sex: FEMALE
        Race: WHITE
```

    Life table used: 1990
    Risk coefficients: BASED ON AGE ATB-AND ATD-SPECIFIC (RBE=10) COEFFICIENTS
            Leukemia: Minimal latency (yrs): 2 Plateau(yrs): 40
    Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
            DRREF: 2.0
    Age at first expos.: 18
Age at last expos-: 65
Total dose eq. (Sv): 1.000000


Table D.29. Excess mortality risk (i/8v) of nonleukemia among females exposed to 0.02 Sv/y from ages 18 to 65 for the transported relative projection model. Baseline rates adjusted with death certificate correction factors (DCCFs).

```
Run date: 10/ 1/1993
    TITLE: Lifetime risks
        Sex: FEMALE
    Race: WHITE
    Life table used:
Risk coefficients: BASED ON AGE ATB- AND ATD-SPECIFIC (RBE=10) COEFFICIENTS
    Leukemia: Minimal latency(yrs): 2 Plateau(yrs): 40
    Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
    DRREF: 2.0
Age at first expos.: 16
Age at. last expog.: }6
Total dose eq. (SV): 1.000000
```



Table D.30. Excess mortality risk (8/Sv) of nonleukemia among females exposed co 0,02 sv/y from ages 18 to 65 for the relative projection model.
Baseline rates adjusted with death certificate correction factors (DCCFs)

```
                Run date: 10/ 1/1993
```

                    TITLE: Hifetime risks
                        Sex: FEMALE
                            Race: WHITE
        Life table used: 1990
    Riak coefficients: BASED ON AGE ATB- AND ATD-SPECIFIC (RBEa10) COEFFICIENTS
                Leukemia: Minimal latency (yrs): 2 Plateau(yrs): 40
            Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
                    DRREF: 2.0
    Age at first expos.; 18
Age at last expos.: 65
Total dose eq. (Sv): 1.000000


Table D.31. Excess mortality risk (g/Sv) of nonleukemia among males exposed to $0.02 \mathrm{sv} / \mathrm{y}$ from ages 18 to 65 for the absolute projection model

Run date: $10 / 1 / 1993$
TITLE: Lifetime rigks
Sex: MALE
Race: WHITE
Life table used: 1990
Risk coefficients: BASED ON AGE ATB- AND ATD-SPECIFIC (RBE $=10$ ) COEFFICIENTS Leukemia: Minimal Latency(yrs): 2 Plateau(yis): 40
Solid cancers, Minimal latency $(y r s): 10$ Plateaufyts): 101
DRREF: 2.0
Age at first expos.: 18
Age at last expos.: 65
Total dose eq. (Sv): 1.000000


Table D.32. Excess mortality risk (事/Sv) of nonleukemia among males exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the transported relative projection model

Run date: $10 / 1 / 1993$
TITLE: Lifetime risks
Sex: MALE
Race: WHITE
Life table used: 1990
R1sk coefficients: BASED ON AGE ATB-AND ATD-SPECIFIC (RBE=10) COEFFICIENTS Solid Leukemia: Minimal latency (yrs): 2 Plateau(yrs): 40
solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
DRREF: 2.0
Age at first expos. 18
Age at last expos.: 65
Total dose eq. (SV): 1.000000


Table D.33. Excess mortality risk (8/Sv) of nonleukemia among males exposed to 0.02 Sv/y from ages 18 to 65 for the relative projection model


Table D.34. Excess mortality risk (o/Sv) of nonleukemia among males exposed to $0.02 \mathrm{sv} / \mathrm{y}$ from ages 18 to 65 for the transported relative projection model. Baseline rates adfusted with death certificate correction factors (DCCFs)

```
                Run date: 10/1/1993
```

                    TITLE: Lifetime risks
                    Sex: MALE
                    pace: WHTTE
        Life table used: 1990
        Risk coefficients: BASED ON AGE ATB- AND ATD-SPECIFIC ( \(\mathrm{RBE}=10\) ) COEFFICIENTS
                Leukemia: Minimal latency (yrs): 2 Plateau(yrs): 40
            Solid cancers: Minimal latancy(yrs): 10 Plateau(yrs): 101
                    DRREF: 2.0
    Age at first expos.: 28
Age at last expos.: 65
Total dose eq. (Sv) $\quad 1.000000$


Table D. 35. Excess mortality risk (象/Sv) of nonleukemia among males exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the relative projection model Baseline rates adjusted with death certificate correction factors (DCCFB).


| Strata for confirmation rates used in adjustment | Radiation <br> induced <br> cancer <br> deaths <br> per 10**5 | 90.08 | CI | Baseline cancer deaths per 10**5 | $\begin{gathered} \text { Years } \\ \text { of } \\ \text { life } \\ \text { last } \end{gathered}$ | PC | 0.08 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| None | 2434 | 990. | 5983) | 25066 | 24 | 09 |  |
| Crude (1950-75) | 2102 | 855. | 5166) | 25066 | 24 | . 081 | . 03. |
| Crude (1950-85) | 1602 ( | 651. | 3937) | 25066 | 25 | .061 | . 02. |
| Sex (1950-75) | 2104 | 856. | 5171) | 25066 | 24 | . 081 | . 03. |
| Sex (1950-85) | 1612 | 656. | 3963) | 25066 | 25 | . 061 | . 02. |
| Sex-city (1950-85) | 2102 ( | 855. | 5167) | 25066 | 24 | . 081 | . 03. |
| Sex-city (1950-85) | 1609 | 654. | 3955) | 25066 | 25 | . 061 | . 02. |
| Age ATD (1950-75) | 2254 ( | 917, | 55421 | 25066 | 24 | . 081 | . 03. |
| Age ATD (1950-85) | 1973 ( | 803, | 48501 | 25066 | 25 | . 071 | . 03 , |
| DS86(1950-75) | 2652 ( | 1079, | 6519) | 25066 | 24 | . 101 | . 04. |
| DS86(195 -85 ) | 2689 i | 1094, | 5610 ) | 25066 | 24 | .101 | . 04. |

Table D.36. Excess mortality risk (8/Sv) of nonleukemia among females expoged to 0.02 Sv/y from ages 18 to 65 for the absolute projection model.

Run date: $10 / 1 / 1993$
TITLE: Lifetime riaks
Sex: FEMALE
Race: WHITE
Life table used: 1990
Risk coefficients: BASED ON AGE ATB- AND ATD-SPECIFIC (RBE $=10$ ) COEFFICIEMTS
Leukemia: Minimal latency(yrs): 2 Plateau(yrs): 40
Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
Age at first expos. 18
Age at last expos.: 65
Total dose eq. (Sv): 1.000000


Table D.37. Excess mortality risk ( $8 / \mathrm{SV}$ ) of nonleukemia among females exposed to 0.02 sviy from ages 18 to 65 for the transported relative projection model.


```
    Run date: 10/1/1993
                TITLE: Lifetime risks
                Sex: FEMALIE
                Race: WHITE
    Life table used: }199
    Risk coefficients: BASED ON AGE ATB- AND ATD-SPECIFIC (RBE=10) COEFFICIENTS
        eukemia: Minimal latenc>(yrs): 2 Plateau(yra): 40
        Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
            DRREF: 2.0
Age at first expos.: 18
Age at last expos.: 65
Total dose eq. (Sv): 1.000200
```



Table D.38. Excess mortality risk (\%/Sy) of nonleukemia among females exposed to 0.02 Sv/y from ages 18 to 65 for the relative projection model.

| Run date: TITLE: <br> Sex: <br> Race: <br> Life table used: <br> Risk coefficients: <br> Leukemia: <br> Solid cancers: <br> DRREF: <br> Age at first expos. <br> Age at last expos.: <br> Total dose eq. (Sv): | ```10/1/1993 Lifetime risks FEMALE WHITE 1990 BASED ON AGE ATB- AND ATD-SPECIFIC (RBE=10) COEFFICIBNTS Minimal latency(yrs): 2 Plateau(yrs): 40 Minimal latency(yrs): 10 Flateau(yrs): 101 2.0``` |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Strata for <br> confirmation <br> rates used in <br> adjustment | Radiation induced cancer deaths per 10**5 | 90.09 | CI | Baseline cancer deaths per $10 * * 5$ | $\begin{gathered} \text { Years } \\ \text { of } \\ \text { life } \\ \text { lost } \end{gathered}$ | PC | . 08 | CI) |
| None <br> Crude (1950-75) <br> Crucie (1950-85) <br> Sex (1950-75) <br> Sex (1950-85) <br> Sex-city $(1950-85)$ <br> Sex-city(1.950-85) <br> Age ATD (1950-75) <br> Age ATD (1950-85) <br> DS86(1950-75) <br> DS86(1950-85) | 3689 3580 2545 3586 2562 3602 2572 3606 3103 3945 3906 | $\begin{aligned} & 1575, \\ & 1525, \\ & 1084, \\ & 1528, \\ & 1092, \\ & 1535, \\ & 1095, \\ & 1537, \\ & 1322, \\ & 1681, \\ & 1664, \end{aligned}$ | $\begin{aligned} & 8656 \\ & 8400 \\ & 5973 \\ & 8415 \\ & 6013 \\ & 8452 \\ & 6035 \\ & 8463 \\ & 7282 \\ & 9257 \\ & 9166 \end{aligned}$ | $\begin{aligned} & 21029 \\ & 21029 \\ & 21029 \\ & 21029 \\ & 21029 \\ & 21029 \\ & 21029 \\ & 21029 \\ & 21029 \\ & 21029 \\ & 21029 \end{aligned}$ | 31 30 33 30 33 30 33 31 32 31 31 | $\begin{aligned} & .15 \\ & .15 \\ & .11 \\ & .15 \\ & .11 \\ & .15 \\ & .15 \\ & 11 \\ & .15 \\ & .13 \\ & .16 \\ & .16 \end{aligned}$ | .06 , <br> .06 , <br> .05 , <br> .05 , <br> .05 , <br> .06 . <br> .05 , <br> .06 , <br> .05 . <br> .07, <br> .07 . | $\begin{aligned} & .35) \\ & .34) \\ & .25) \\ & .34) \\ & .25) \\ & .34) \\ & .26) \\ & .34) \\ & .30) \\ & .37) \\ & .37) \end{aligned}$ |

Table D.39. Excess mortality risk ( $\% / 5 \mathrm{~V}$ ) of nonleukemia among females exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 28 to 65 for the transported relative projection model. Baseline rates adjusted with death certificate correction factors (DCCFs)

```
Run date: 10/1/1993
    TITLE: Lifetime risks
            Sex: FEMALE
            Race: WHITE
    Life table used: }199
    Risk coefflcients: BASED ON AGE ATB- AND ATD-SPECIFIC (RBE=10) COEFFICIENTS
            Leukemia: Minimal latency(yrs): 2 Plateau(yIg): 40
    Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
            DRREF: 2.0
Age ac first expos.: 18
Age at last expos.: }6
Total dose eq. (SV): 1.000000
```



Table D. 40 . Excess mortality risk ( $8 / 5 v$ ) of nonleukemia among females exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the relative projection model
Baseline rates adjusted with death certificate correction factors (DCCFs). Run date: $10 / 1 / 1993$
TITHE: Lifetime İsks
Sex: FEMALE
Race: WHJTE
Life table used: 1990
Risk coefficients: BASED CN AGE ATB- AND ATD-SPECIFIC (RBE=10) COEFFICIENTS
Leukemia: Minimal latency (yrs): 2 plateau(yrs): 40
Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
DRREF: 2.0
Age at first expos. : 18
Age at last expos.: 65
Total dose eq. (Sv): 1.000000


Table D.41. Excess mortality risk (8;'sv) of digestive cancer among males exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the AEIR-V relative projection model.

```
                Run date: 10/ 3/1993
                    Tytle: Lifetime risks
                    Sex: MALE
                    Race: WhlTE
    Life table used: 1990
    Risk coefficients: BASED ON BEIR-V MODELS
        Leukemia: Minimal latency(yrs): 2 Plateau(yrs): 40
    Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
            DRREF: 2.0
Age at first expos.: 18
Age at last expos.: }6
Total dose eq. (Sv): 1.000000
```



Table D.42. Excess mortality risk (\%/Sv) of leukemia among males exposed to 0.02 Sv/y from ages 18 to 65 for the BEIR-V relative projection model

$$
\begin{aligned}
\text { Run date: } & 10 / 3 / 1993 \\
\text { Title: } & \text { Lifetime risks } \\
\text { Sex: } & \text { MALE }
\end{aligned}
$$

Race: WHITE

## Life table used: 1990

Risk coefficients: BASED ON BEIR-V MODELS
Solideukemia: Minimal latency(yrs):
Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
Age at first expos.: 18
Age at last expos. 65
Total dose eq. (Sv): 1.000000


Table D.43, Excess mortality risk (s/Sv) of digestive cancer among males exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the EEIR-V relative projection model Baseline rates adjusted with death certificate correction factors (DCCFs)


Table D.44, Excess mortality risk ( $8 / S V$ ) of leukemia among malez exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the $\operatorname{sEIR}-V$ relative projection model Baseline rates adjusted with death certificate correction factors (DCCFs)

## Run date: $10 / 3 / 1993$

Title: Lfecime risks
Sex: MALE
Race: WHITE
Life table used 1990
Risk coefficients: BASED ON BEIR-V MODELS
Leukemia: Minimal latency(yrs)
Minimal latency (yrs)
solid cancers: Minima latency (yrs) ; 10- Plateau(yte) 10
Age at first DRREF: 2.0
Age at last expos.: 65
Total dose eq. (Sv): 1.00000

| Strata for confirmation rates used in adjustment | Radiation- <br> induced cancer deaths <br> per 10**5 | 90.08 | CI | Baseline cancer deaths per 10**5 | $\begin{gathered} \text { Years } \\ \text { of } \\ \text { life } \\ \text { lost } \end{gathered}$ | PC( 90.0 CI) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| None | 478 | 194, | 1175) | 668 | 9 |  |
| Crude (1950-75) | 2461 | 100, | 605) | 668 | 29 | $27(.10, .70)$ |
| Crude (1950-85) | 4091 | 166 , | 1006) | 668 | 29 | $.381 .14,1.00)$ |
| Sex(1950-75) | 4391 | 178. | 1079) | 668 | 29 | $.40(-151.00)$ |
| Sex(1950-85) | 411 | 167 , | 1011) | 668 | 29 | $.381 .14,1.001$ |
| Sex-city (1950-85) | 3691 | 150. | 907) | 668 | 29 | . $36(.13, .95)$ |
| Sex-city (1950-85) | 3731 | 151. | 917) | 668 | 29 | . $36(.13, .96)$ |
| Age ATD (1950-75) | 3781 | 153, | 929) | 668 | 29 | $.36(.13, .97)$ |
| Age ATD (1950-85) | 3811 | 155. | 937) | 668 | 29 | $.361 .14, .98)$ |
| DS86 (1950-75) | 7541 | 307 , | 1855) | 668 | 28 | . $53(.18,1.00)$ |
| DS86 (1950-85) | 762 1 | 310. | 1874) | 668 | 28 | . $536.18,1.00$ ) |

Table D. 45. Excess mortality risk (8/Sv) of digestive cancer among females exposed to $0.02 \mathrm{SV} / \mathrm{Y}$ from ages 18 to 65 for the BEIR-V relative projection model.

Run date: 10 / 3/1993
Titie: Lifetime risks Sex: FEMALE Race: WHTTE
Life table used: 1990
Risk coefficients: BASED ON BEIR-V MODELS
Leukemia: Minimal latency(yrs):
Plateau(yrs): 40
Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
DRREF: 2.0
Age at first expos.: 18
Age at last expos.: 65
Total dose eq. 1 SV ; : 1.000000


Table D.46. Excess mortality risk (\%/SV) of leukemia among females exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 fcr the BEIR-V relative projection model.


```
    Run date: 10/ 3/1993
    Title: Lifetime risks
        Sex: FEMALE
            Race: WhITE
        Life table used: }199
    Risk coefficients: BASED ON BEIR-V MODELS
            Leukemia: Minimal latency(yrs): }
        Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): }10
```

            DRREF: 2.0
    Age at first expos.: 18
Age at last expos.: 65
Total dose eq. (Sv): 1.000000

| Strata for confirmation rates used in adjustment | $\begin{aligned} & \text { Radiation- } \\ & \text { induced } \\ & \text { cancer } \\ & \text { deaths } \\ & \text { per } 2.0^{* * 5} \end{aligned}$ | 90.08 | CI | Baseline cancer deaths per 10**5 | $\begin{gathered} \text { Yeara } \\ \text { of } \\ 11 \mathrm{fe} \\ \text { lost } \end{gathered}$ |  | .08 CI) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| None | 4331 | 184. | 1016) | 653 | 36 | 40 | 15,1.001 |
| Crude (1950-75) | 2291 | 97. | 537) | 653 | 34 | . 26 | . $10, .65$ ) |
| Crude (1950-85) | 3711 | 158, | 871) | 653 | 36 | . 36 | .14, .95) |
| $\operatorname{Sex}(1950-75)$ | 3971 | 169. | 933) | 653 | 36 | . 38 | .14,1.00) |
| Sex (1950-85) | 3731 | 159, | 875) | 653 | 36 | . 36 | .14. .95) |
| Sex-city (1950-85) | 3341 | 142, | 785) | 653 | 36 | . 34 | .13, .88) |
| Sex-city (1950-85) | 338 | 144. | 794) | 653 | 36 | . 34 | .13, .88) |
| Age ATD (1950-75) | 3431 | 146. | 8051 | 653 | 35 | . 34 | .13, .891 |
| Age ATD (1950-85) | 345 | 147. | 811) | 653 | 35 | . 35 | .13, . 901 |
| DS86 (1950-75) | 687 | 293. | 1613) | 653 | 34 | . 51 | . 18,1.00) |
| DS86(1950-85) | 6941 | 296. | $1630)$ | 653 | 34 | 52 | 18,1.00) |

Table D.47. Excess mortality risk (\%/Sv) of digestive cancer among females exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the BEIR-V relative projection model Baseline rates adjusted with death certificate correction factors (DCCFB)

$$
\begin{aligned}
\text { Run date: } & 10: 3 / 1993 \\
\text { Title: } & \text { Lifetime risk } \\
\text { Sex: } & \text { FEMALE } \\
\text { Race: } & \text { WHITE }
\end{aligned}
$$

Life table use3: 1990
Riak coefficients: BASED ON BEIR-V MODELS
Leukemia: Minimal latency (yrs): 2 Plateau(yrs): 40
Solid cancers: Minimal latency (yrs)
DRREF: 2.0
Age at first expos.: 18
Age at last expos.: 65
Total dose eq. (Sv): 1.000000
$\begin{array}{lll}2 & \text { Plateau(yrs): } & 40 \\ 10 & \text { Plateau(yrs) }: & 101\end{array}$
1.000000

| Strata for confirmation rates used in adjustment | Radiation induced cancer deaths per 10**5 | 90.08 | CI | $\begin{gathered} \text { Baseline } \\ \text { cancer } \\ \text { deaths } \\ \text { per } 10 * * 5 \end{gathered}$ | $\begin{gathered} \text { Years } \\ \text { of } \\ \text { life } \\ \text { lost } \end{gathered}$ | PC | 0.08 | CI) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| None | 1323 | 564 | 31061 | 6194 | 23 |  |  |  |
| Crude (1950-75) | 1165 | 496, | 2735) | 6194 | 23 | . 161 |  | . 391 |
| Crude (1950-85) | 1060 ( | 452, | 2489) | 6194 | 23 | . 151 | . 06 , | . 361 |
| $\operatorname{sex}(1950-75)$ | 1188 | 506, | 2789) | 6194 | 23 | . 161 | . 07 , | . 391 |
| Sex(1950-85) | 1088 | 463 , | 2554) | 6194 | 23 | . 151 | . 06 , | . 371 |
| Sex-city (1950-85) | 1305 | 556 , | 3063) | 6194 | 23 | .171 | . 07 , | .42) |
| Sex-city(1950-85) | 1259 | 536, | 2955) | 6194 | 23 | .171 | .07, | .41) |

Table D.48, Excess mortality risk (8:SV) of leukemia among females exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the BEIR-V relative projection model Baseline rates adjusted with death certificate correction factors (DCCFs)
Run date: 10/3/1993
Title: Lifetime risks
Sex: FEMALE

Race: WHITE
Life table used: 1990
Risk coefficients: BASED ON BEIR-V MODELS
Leukemia: Minimal latency(yrs): 2 Plateau(yrs): 40
Solid cancers: Minimal latency (yrs): 10 Plateau(yrs): 101
DRREF: 2,0
Age at first expos.: 18
Age at last expos.:
Total dose eq. (Sv): 1.000000


Table D. 49. Excess mortality risk (8/Sv) of digestive cancer among fiales expoged to $0.02 \mathrm{~Sv} / \mathrm{y}$ trom ages 18 to 65 for the relative projection model D586 Dose equivalents adjusted for random errot.

Run date: $10 / 3 / 1993$
Title: Lifetime risks
Sex: MALE
Race: WHITE
Life table used: 1990
Risk coefficients BASED ON BEIR -V MODELS
Leukemia: Minimal latency (yrs): 2 Plateau(yrs): 40
Solid cancers: Minimal latency(yrs); 10 Plateau(yrs): 101
DRREF: 2.0
Age at first expos. ; 18
Age at last expos.: 65
Total dose eq. (Sv): 1.000000

| Radiation- |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| induced |  |
| Strata for |  |
| confirmation | cancer |

Table D.50. Excess mortality risk (\%/Sv) of leukemia amonc males exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the relative projection model. DS86 Dose equivalents adjusted toz random ercor.

Run date: $10 / 3 / 1993$
Title: Lifetime risks
Sex MALE
Race: WHITE
Life table usea: 1990
R1sk coefficients: BASED ON BEIR-V MODELS
Leukemia: Minimal latency(yrs): 2 Plateau(yrs) : 40
Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
Age at first expos. 2.0
Age at last expos.: 65
Total dose eq. (Sv): 1.000000


Table D.51. Excess mortality risk (8/sv) of digestive cancer among males exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the relative projection model
Baseline rates adjusted with death certificate correction factors (DCCPA)
DS86 Dose equivalents adjusted for random error
Run dete: $10 / 3 / 1993$
Title: Lifetime risks
Sex: MALE
Race: WHITE
Life table used: 1990
Risk coefficients: BASED ON BEIR-V MONELS
Leukemia: Minimal latency (yrs):
Minimal latency (yre).
DRREF: 2.0
Age at first expos.: 18
Age at last expos.: 65
Total dose eq. (sv): 1.000000

| Strata for confirmation rates used in adjustment | Radiationinduced cancer deaths per 10**5 | 90.0 \% | CI | Baseline cancer deaths per 10**5 | $\begin{gathered} \text { Years } \\ \text { of } \\ \text { life } \\ \text { lost } \end{gathered}$ | PC 1 | $0.08 \mathrm{Cl})$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| No..e | 944 | 343 | 2075) | 6474 | 21 | 12 ! |  |
| Crude (1950-75) | 763 | 310 , | 1875) | 6474 | 21 | 111 |  |
| Crude (1950-85) | 634 | 258, | 15601 | 6414 | 21 | . 091 | .04, . 221 |
| Sex (1950-75) | 746 ! | 303. | 1835) | 6474 | 21 | . 101 | 04, . 26 ; |
| Sex(1950-85) | 615 ( | 250. | 1513) | $64^{*} 4$ | 21 | . 091 | 03, .22) |
| Sex-city (1950-85) | 829 ! | 337. | 2037) | 6474 | 21 | .111 | .05, .281 |
| Sex-city (1950-85) | 750 | 305, | 1843) | 6474 | 21 | . 101 | .04, . 261 |

Table D. 52. Excess mortality risk (8/Sv) of leukemia among males exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the relative projection model.
Baseline rates adjusted with death certificate correction factors (DCCFs)
DS86 Dose equivalents adjusted for random error.
Run date: 10 / $3 / 1993$
Title: Lifetime risks
Sex: MALE
Race: WHITE
Life table used: 1990
Risk coefficients: BASED ON BEIR-V MODELS
Leukemia: Minimal latency (yrs): 2 Plateau(yrs): 40
Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
Age at first expos.: $\quad 18$
Age at last expos.: 65
Total dose eq. (Sv): 1.000000


Table D.53. Excess mortalitv risk (\%/Sv) of digestive cancer among females exposed to $0.02 \mathrm{sv} / \mathrm{y}$ from ages 18 to 65 for the relative projection model.
DSB6 Dose equivalents adfusted for random error.
Run date: $10 / 3 / 1993$
$\begin{aligned} \text { Run date: } & 10 / 3 / 1993 \\ \text { Ticle: } & \text { Lifetime risks }\end{aligned}$
Sex: FEMALE
Race: WHITE
Life table used: 1990
Risk coefficients: BASED ON BEIR-V MODELS
Leukemia: Minimal latency (yrs): $2 \quad$ Plateau(yrs): 40
Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
hge at firat DRREF: 2.0
Age at last expos.: 65
Total dose eq. (Sv): $\quad 1.000000$

|  | Radiation- |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Strata for | induced |  |  | Baseline | Years |  |  |  |
| Cs .t. rmation | cancer |  |  | cancer | of |  |  |  |
| rates used in | deaths |  |  | deaths | 11fe |  |  |  |
| adjust, ent | per 10**5 | 90.07 | CI | per 10**5 | lost | PC 1 | 9.0\% | CI) |
| None | 1505 | 641. | 3531) | 6096 | 23 | . 201 | . 08 , | 48) |
| Crude (195 -75) | 1332 | 567. | $3126)$ | 6096 | 23 | . 181 | .07, | 44) |
| crude (1956.85) | 1213 | 517. | 28481 | 6096 | 23 | . 271 | .07, | .41) |
| Sex(1950-75) | 1358 | 579, | 3188) | 6096 | 23 | . 181 | .07, | .45) |
| Sex (1950-85) | 1244 | 530, | 2919) | 6096 | 23 | 171 | .07, | .42) |
| Sex-city (1950-85) | 1490 ( | 635, | 3497) | 6096 | 23 | . 201 | . 08 , | .48) |
| Sex-city (1950-85) | 1438 | 613. | 3376) | 6096 | 23 | . 191 | . 03. | .47) |

Table D.54. Excess mortality risk (\%/Sv) of leukemia among femiles exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the relative projection model. DS8s Dose equivalents adjusted for random error.


Table D.55. Excess mortality risk ( $8 / \mathrm{sv}$ ) of digestive cancer among females exposed to $0.02 \mathrm{SV} / \mathrm{y}$ from ages 18 to 65 for the relative projection model.
Baseline rates adyusted with death certificate correction factors (DCCFS)
DS86 Dose equivalents adjusted for random error

$$
\text { Run date: } 10 / 3 / 1993
$$

Run date: $\quad 10 / 3 / 1993$
Title: Lifetime risks
Sex: FEMALE
Race: WHITE
Life table used: 1990
Risk coefficients: BASED ON BEIR-V MODELS Leukemia: Minimal latency $y r$ s)
Solid cancers: Minimal latency (yrs):
2 Plateau(yrs): 40
10
DRREF: 2.0
Age at first expos. :
Age at last expos.
Total dose eq. (sv):


Table D.56. Excess mortality risk (8/Sv) of leukemia among females exposed to $0.02 \mathrm{~Sv} / \mathrm{y}$ from ages 18 to 65 for the relative projection model.
Baseline rates adjusted with death certificate correction factors (DCCFs)
DS86 Dose equivalents adjusted for random error.
Run date: $10 / 3 / 1993$
Title: Lifetime risks
Sex: FEMALE
Race: WKITE
Life table used: 1990
Risk coefficients: BASED ON BEIR-V MODELS
Leukemia: Minimal latency(yrs): 2 Plateau(yrs): 40
Solid cancers: Minimal latency(yrs): 10 Plateau(yrs): 101
Age at first expos.: 18
Age at last expos.: 65
Total dose eq. (Sv): 1.000000

| Strata for confirmation rates used in adjustment | Radiationinduced cancer deaths per 10**5 | 90.08 | CI | ```Baseline cancer deaths per 10**5``` | $\begin{gathered} \text { Years } \\ \text { of } \\ \text { life } \\ \text { lost } \end{gathered}$ | PC | 0.08 CI) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| None | 387 | 165. |  |  | 36 |  |  |
| Crude (1950-75) | 328 ( | 140 , | 7711 | 581 | 35 | .461 | 15,1.00) |
| crude (1950-85) | 3341 | 142 , | 785) | 581 | 35 35 | . 37 ( | 14, .96) |
| Sex (1950-75) | 3531 | 150 , | 829) | 581 | 35 | . 381 | .14,1.00) |
| Sex (1950-85) | 335 | 143. | 787) | 581 | 36 | . 371 | .14, .97) |
| Sex-city (1950-85) | 2971 | 126 , | 698) | 581 | 36 | . 34 ( | .13, .88) |
| Sex-city (1950-85) | 3011 | 128, | $706)$ | 581 | 36 | . 341 | . $13, .89$ ) |
| Age ATD (1950-75) | 3091 | 132, | 727) | 581 | 35 | . 351 | . $13, .91$ ) |
| Age ATD (1950-85) | 313 ( | 133. | 735) | 581 | 35 | . 351 | .13, . 921 |
| DS86 (1950-75) | 6281 | 267. | 1474) | 581 | 34 | . 521 | .17,1.00) |
| DS86(1950-85) | 6431 | 274. | 1510 ) | 581 | 34 | . 531 | .18,1.00) |




Federal Recycling Program

```
120555139531 1 IAN1:F
US NRC-OADM OWQtTCATTAN:S
```



```
TPS-ODR-NURZG
TWFN-GETON IC 20555
```


[^0]:    *School of Public Health, University of California, Berkeley, CA 94720

[^1]:    ${ }^{1}$ This study was a dissertation thesis in the School of Public Health at the The University of Texas-Health Science Center at Houston. The first author's current address is: Kelsey-Seybold Clinic, Mail Code SD23, Lyndon B. Johnson Space Center, Houston, Texas 77058.

[^2]:    ${ }^{1}$ Information bias is the distortion of risk estimates caused by random and systematic misclassification of a subject's exposure status or diagnosis of death or disease.
    ${ }^{2}$ AMFIT is trademark of Hirosoft International Corporation. See $\$ 3.3 .1$.
    ${ }^{3}$ SURVRAD is neither an abbreviation nor an acronym. See $\$ 3.4 .1$.

[^3]:    ${ }^{4}$ Confirmation and detection rates are defined in $\S 3.2 .1$.

[^4]:    ${ }^{5}$ ABCC Technical Report 12-62 was based on a draft report by L.J. Zeldis and Y.S. Matsumoto and, in part, on previous suggestions by T. Francis, Jr., S. Jablon, and F.E. Moore.

[^5]:    ${ }^{6}$ The expectation-maximization (EM) algorithm is a generic statistical method based on sufficient statistics to impute missing data. See Sposto et al. (1992) and Dempster et al. (1977).
    ${ }^{7}$ Radiation absorbed dose is the amount of energy deposited in tissue.
    ${ }^{8}$ Kinetic Energy Released in Matter, KERMA, is the total amount of kinetic energy released by charged particles created from the interaction of radiation in tissue.

[^6]:    ${ }^{9}$ Although other RERF regression models for the stomach, breast, luag, bladder and liver were fitted in this investigation, the results are not provided in the text because the tabular output tables were so voluminous. However, the coding schemes for all Poisson regression runs are provided in Appendix A. Results of all modeling sessions are available on request by writing to the address on the botiom of page $v$ (acknowledgement page).

[^7]:    ${ }^{10}$ Other BEIR-V models for the respiratory, female breast and "other" cancer sites were fitted but ase not described in the results or discussion. However, for the reader who is interested, Appendix A includes the coding format for all BEIR-V models.

[^8]:    ${ }^{11}$ The annual dose for the continuous exposure from age 18 to 65 was based on dividing the total dose of 1 Sv by 47 years, which resulted in $0.02128 \mathrm{~Sv} /$ year. In the complete (complete life tables have one-year age intervals; abridged life tables have, say, 5-year intervals or quinquennia) life table calculations, the first annual dose of 0.02128 Sv was applied to the 19 th one-year age interval becsuse an individual is 18 years old in the 19 th age interval. The last exposure occurred in the 65th age interval, in which an individual spends an entire year. It is assumed that each individual retires when they enter the 66th age interval at their 65 th birthday. These one-year differences between a given age and respective life table interval are easily explained by the fact that when an individual is born, the first year of life is spent in the first interval when the infant is zero years old. After the infant's first birthday, the infant enters the second one-year life table interval, but is still only one year old.
    ${ }^{12}$ Although the total follow-up time from 1 October 1950 to 31 December 1985 was equal to 35 years and 3 months (Table 1), it is assumed to be 35 years in this study.

[^9]:    aStrata of confirmation rates for which correction of diagnostic misclassification resulted in negative bias. N/A denotes that the use of confirmation rates in Tables 4-8 did not result in negative bias from diagnostic misclassification. (1950-85) denotes that confirmation rates were applied only to deaths which occurred during 1950-85.

[^10]:    ${ }^{\text {a }}$ Strata of confirmation rates for which correction of diagnostic misclassification resulted in negativ - bias. N/A-EM algorithm denotes that the diagnostic misclassification was estimated from the Sposto et al. (1992) Poisson regression coefficients obtained with the EM algorithm. (1950-85) denotes that confirmation rates were applied only to deaths which occurred during 1950-85.

[^11]:    ${ }^{\text {a }}$ Strata of confirmation rates for which correction of diagnostic misclassification resulted in negative bias. N/A denotes that the use of confirmation rates in Tables 4-8 did not result in negative bias from diagnostic misclassification. (1950-75) denotes that confirmation rates were applied only to deaths that occurred during 1950-75. Likewise, (1950-85) denotes that confirmation rates were applied only to deaths which occurred during 1950-85.

[^12]:    ${ }^{a}$ Strata of confirmation rates for which correction of diagnostic misclassification resulted in negative bias. N/A denotes that the use of confirmation rates in Tables 4-8 did not result in negative bias from diagnostic misclassification. (1950-75) denotes that confirmation rates were applied only to deaths that occurred during 1950-75. Likewise, (1950-85) denotes that confirmation rates were applied only to deaths which occurred during 1950-85.

[^13]:    ${ }^{13}$ MS-DOS is a registered trademark of the Microsoft Corporation.

